

## 100. A Note on Sums and Maxima of Independent, Identically Distributed Random Variables

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**§ 1. Introduction.** Let  $X_1, X_2, \dots$  be i.i.d. (independent, identically distributed) random variables and put  $S_n = X_1 + \dots + X_n$ , and  $M_n = \max(X_1, \dots, X_n)$ ,  $n = 1, 2, \dots$ . Chow-Teugels [2] studied the joint limiting distributions of  $(S_n, M_n)$  as  $n \rightarrow \infty$  after suitable normalizations. In this note we will consider this problem using the theory of point processes and generalize the result of [2] to a functional limit theorem for the sums and the maxima of triangular arrays of i.i.d. random variables.

**§ 2. Main theorem.** Let  $\{\xi_{nk}\}_{k=1}^\infty$  be i.i.d. random variables with distribution function  $F_n(x)$ ,  $n = 1, 2, \dots$ . Throughout this paper we assume that for suitably chosen constants  $A_n$ ,  $n = 1, 2, \dots$  and a nondegenerate distribution function  $F(x)$  we have

$$(2.1) \quad \lim_{n \rightarrow \infty} P[\sum_{j=1}^n \xi_{nj} - A_n \leq x] = F(x) \text{ at all continuity points of } F(x).$$

The characteristic function  $\phi(\theta)$  of  $dF(x)$  has the following representation.

$$(2.2) \quad \phi(\theta) = \exp \left[ \gamma \theta - \frac{1}{2} \sigma^2 \theta^2 + \int \{e^{i\theta x} - 1 - i\theta x I(|x| \leq \delta)\} \mu(dx) \right]$$

where  $\gamma \in \mathbf{R}$ ,  $\sigma^2 \geq 0$ ,  $\int \min(1, x^2) \mu(dx) < \infty$  and  $\delta > 0$  is chosen so that  $\mu\{\pm \delta\} = 0$ .

It is well known that (2.1) implies a functional limit theorem; the process

$$(2.3) \quad \hat{\xi}_n(t) = \sum_{j \leq nt} \xi_{nj} - A_n[nt]/n$$

converges in law to the Lévy process  $\xi(t)$  with characteristic (2.2) over the Skorohod function space  $D([0, \infty); \mathbf{R})$  endowed with the  $J_1$ -topology (see [5] for the definition). We also assume that there exist constants  $B_n > 0$ ,  $C_n$ ,  $n = 1, 2, \dots$  and nondegenerate distribution function  $G(x)$  such that

$$(2.4) \quad \lim_{n \rightarrow \infty} P[B_n \max_{k \leq n} \xi_{nk} - C_n \leq x] = G(x) \text{ at all continuity points of } G(x)$$

(see Lemma 4.1.).

(As we will see later, if  $\mu(0, \infty) > 0$  then this condition is automatic from (2.1) with  $B_n = 1$ ,  $C_n = 0$ .) It is also well known that (2.4) implies a functional limit theorem: Define

$$(2.5) \quad m_n(t) = \begin{cases} B_n \max_{k \leq nt} \xi_{nk} - C_n, & t \geq 1/n \\ m_n(1/n), & 0 < t < 1/n. \end{cases}$$

Then we have that  $\{m_n(t)\}_{t>0}$  converges in law over  $D((0, \infty); \mathbf{R})$  to a nondecreasing process  $m(t)$  with marginals as follows.

$$(2.6) \quad \begin{aligned} P[m(t_j) \leq x_j, j=1, \dots, n] \\ = G(x_1)^{t_1} G(x_2)^{t_2-t_1} \dots G(x_n)^{t_n-t_{n-1}}, \end{aligned}$$

for  $0 < t_1 < \dots < t_n, x_1 < x_2 < \dots < x_n, n=1, 2, \dots$ .

$\{m(t)\}$  is called the extremal process associated with  $G(x)$ . We now consider the joint limiting process of  $\{(\xi_n(t), m_n(t))\}$ . Since  $m_n(t) = \sup_{s \leq t} \{\xi_n(s) - \xi_n(s-)\}$  ( $t > 1/n$ ), applying the continuity theorem (see e.g. [1] p. 31) we see that  $\{(\xi_n(t), m_n(t))\}$  converges in law to  $\{(\xi(t), \max_{s \leq t} (\xi(s) - \xi(s-)))\}$ . If  $\mu(0, \infty) > 0$  then  $\max_{s \leq t} \{\xi(s) - \xi(s-)\}$  is not trivial and therefore we have a complete answer to the joint convergence of  $(\xi_n(t), m_n(t))$  with  $B_n=1, C_n=0$ . Thus we need to consider only the case where

$$(2.7) \quad \mu(0, \infty) = 0.$$

Our main theorem is

**Theorem 1.** *Assume (2.1), (2.4) and (2.7). Let  $\xi_n, m_n, \xi$  and  $m$  be as before. Then  $\{(\xi_n(t), m_n(t))\}_{t>0}$  converges in law to  $\{(\xi(t), \tilde{m}(t))\}$  as  $n \rightarrow \infty$  in  $D((0, \infty); \mathbf{R}^2)$  endowed with the  $J_1$ -topology, where  $\{\xi(t)\}$  and  $\{\tilde{m}(t)\}$  are independent and are identical in law to  $\{\xi(t)\}$  and  $\{m(t)\}$ , respectively.*

**§ 3. Outline of the proof.** Let  $p$  be a Poisson point process on  $\{(t, x) | t > 0, x \in \mathbf{R} \setminus \{0\}\}$  with intensity measure  $\hat{N}_p(dt dx) = dt(dx/x^2)$ , and let  $N_p(ds dx)$  denote the counting measure of  $p$ . We also put  $\tilde{N}_p(ds dx) = N_p(ds dx) - \hat{N}_p(ds dx)$ . We refer to the textbook of Ikeda and Watanabe [4] for the details of definitions, notations and fundamental results of (Poisson) point processes.

**Proposition.** *Assume (2.1) and (2.4) but we drop the condition (2.7). Put*

$$(3.1) \quad f(x) = \begin{cases} \inf \{t > 0 | \mu[t, \infty) < 1/x\}, & x > 0 \\ \sup \{t < 0 | \mu(-\infty, t) < -1/x\}, & x < 0 \end{cases}$$

and

$$(3.2) \quad g(x) = G^{-1}(e^{-1/x}), \quad x > 0.$$

Then,  $\{(\xi_n(t), m_n(t))\}_{t>0}$  converges in law to

$$(3.3) \quad (\gamma t + \sigma B(t) + \zeta_\delta(t) + \eta_\delta(t), g(\max_{s \leq t} p(s)))$$

where  $\gamma, \sigma$  and  $\delta$  are the constants in (2.2),  $B(t)$  is a standard Brownian motion independent of the Poisson point process  $p$  and where

$$(3.4) \quad \zeta_\delta(t) = \int_0^{t+} \int_{\{|f(x)| \leq \delta\}} f(x) \tilde{N}_p(ds dx),$$

$$(3.5) \quad \eta_s(t) = \int_0^{t+s} \int_{\{|f(x)| > s\}} f(x) N_p(ds dx), \quad t > 0.$$

This proposition can be proved by using the idea of [3] and we omit the details. Theorem 1 is an easy consequence of this proposition; if  $\mu(0, \infty) = 0$  then we have  $f(x) = 0$  on  $(0, \infty)$ . Therefore, the (stochastic) integrals in (3.4) and (3.5) are in fact functionals of the restriction  $p^-$  of  $p$  to the lower half plane  $\{x < 0\}$  while  $\max_{s \leq t} p(s)$  depends only on the restriction  $p^+$  of  $p$  to the upper half plane. Since  $p^+$  and  $p^-$  are independent, we see that the first and the second components of (3.3) are mutually independent, which prove the theorem.

**§ 4. Supplement.** By a slight modification of the proof, Theorem 1 may be extended to all order statistics. Let  $M^k(n)$  denote the  $k$ -th largest among  $\{\xi_{n1}, \dots, \xi_{nn}\}$ . For  $k > n$ , we define  $M^k(n) = M^k(k)$  for convenience. (Thus  $M^k(n)$  is nondecreasing in  $n$ .)

**Theorem 2.** Assume (2.1), (2.4) and (2.7) and let  $g(x) = G^{-1}(e^{-1/x})$ ,  $x > 0$  as before. Define  $m_n^{(k)}(t) = B_n M^k([nt]) - C_n$ ,  $t > 0$ ,  $n, k \geq 1$ . Then, for  $N \geq 1$ ,  $\{\{\xi_n(t), m_n^{(1)}(t), \dots, m_n^{(N)}(t)\}_{t > 0}\}$  converges in law in  $D((0, \infty); R^{1+N})$  to  $\{\{\tilde{\xi}(t), g(J^{(1)}(t)), \dots, g(J^{(N)}(t))\}_{t > 0}\}$  as  $n \rightarrow \infty$ , where  $J^{(k)}(t)$  denotes the  $k$ -th largest among  $\{p(s); s \leq t, s \in D_p\}$ , and where  $\{\tilde{\xi}(t)\}$  is a process which is independent of  $p$  and is identical in law to  $\{\xi(t)\}$ . ( $p$  is the Poisson point process in § 3.)

We can also consider similar problems for the case where the sums and the maxima are based on different arrays of i.i.d. random variables. Let  $\{(\xi_{nk}, \zeta_{nk})\}_{k=1}^\infty$  be i.i.d. random vectors such that  $\{\xi_{nk}\}$  satisfies (2.1) and that (2.4) holds replacing  $\{\xi_{nk}\}$  by  $\{\zeta_{nk}\}$ . Let  $\xi_n(t)$  be as in (2.3) and define  $m_n(t) = B_n \max_{k \leq nt} \zeta_{nk} - C_n$ . If  $\mu(dx)$  vanishes identically, we have a result similar to Theorem 1;

**Theorem 3.** Assume that  $\mu(R \setminus \{0\}) = 0$  as well as above assumption. Then the assertion of Theorem 1 is still valid.

Finally, we give a necessary and sufficient condition for (2.4).

**Lemma 4.1.** (2.4) is equivalent to the following condition.

$$(4.1) \quad \lim_{n \rightarrow \infty} B_n F_n^{-1} \left( 1 - \frac{1}{nx} \right) - C_n = g(x) \quad \text{at all continuity points } x > 0,$$

where  $g(x) = G^{-1}(e^{-1/x})$  as before.

*Proof.* It is easy to see that (2.4) is equivalent to

$$(4.2) \quad \lim_{n \rightarrow \infty} n \{ F_n((C_n + x)/B_n) - 1 \} = \log G(x).$$

By considering the inverse functions of the both sides of (4.2) we have the assertion after a change of the variable ( $x \rightarrow -1/x$ ).

## References

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