

37. A Version of the Central Limit Theorem for Martingales

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(Communicated by Kôzaku YOSIDA, M. J. A., May 12, 1979)

§ 1. Introduction. In the present note let $\{X_n, \mathcal{F}_n\}$ be a zero-mean square-integrable martingale on a probability space (Ω, \mathcal{F}, P) and let $Y_1 = X_1$, $Y_n = X_n - X_{n-1}$ ($n \geq 2$). Then our purpose is to prove the following

Theorem. *Suppose that there exist a sequence $\{A_n\}$ of positive numbers for which $\lim_{n \rightarrow +\infty} A_n = +\infty$ and a random variable $Z(\omega)$ such that*

(L-I) *for any given $\varepsilon > 0$, $\lim_{n \rightarrow +\infty} A_n^{-2} \sum_{k=1}^n E\{Y_k^2 I(|Y_k| \geq \varepsilon A_n)\} = 0$,^{*)}*

(L-II) *$\lim_{n \rightarrow +\infty} A_n^{-2} \sum_{k=1}^n Y_k^2 = Z$, in probability.*

Then for any set $F \in \sigma(\cup_{n=1}^{\infty} \mathcal{F}_n)$ and any real number x ($x \neq 0$)

$$\lim_{n \rightarrow +\infty} P\{F, X_n(\omega)/A_n \leq x\} = (2\pi)^{-1/2} \int_F \left\{ \int_{-\infty}^{x/\sqrt{Z}} \exp(-u^2/2) du \right\} dP,$$

where $\sigma(\cup_{n=1}^{\infty} \mathcal{F}_n)$ denotes the σ -algebra generated by the algebra $\cup_{n=1}^{\infty} \mathcal{F}_n$ and x/\sqrt{Z} is $+\infty$ (or $-\infty$) if x is positive (or negative).

In the important special case when Y_n 's are independent and \mathcal{F}_n is the σ -algebra generated by $\{X_k, k \leq n\}$ the condition (L-I) for $A_n^2 = EX_n^2$ is called Lindeberg's condition for the central limit theorem and in this case (L-I) implies (L-II) with $Z(\omega) = 1$. But in general (L-I) does not imply (L-II) and even if the conditions (L-I) and (L-II) are satisfied the limit Z is not necessarily a constant. When $Z(\omega)$ is a constant, the central limit theorems are proved by many authors (cf. [1]).

As an application of Theorem we can prove the central limit theorem for $\{X_n\}$. In fact we prove the following

Corollary. *Under the conditions (L-I) and (L-II) if $P\{Z(\omega) \neq 0\} > 0$, then we have for any real number x*

$$\lim_{n \rightarrow +\infty} P\{X_n(\omega)/A_n \leq x\sqrt{Z(\omega)} | Z(\omega) \neq 0\} = (2\pi)^{-1/2} \int_{-\infty}^x \exp(-u^2/2) du.$$

In § 2 we prove Theorem. By Lévy's continuity theorem it is enough to show that, for any fixed real number λ ,

$$(1.1) \quad \lim_{n \rightarrow +\infty} \int_F \exp(i\lambda X_n/A_n) dP = \int_F \exp(-\lambda^2 Z/2) dP.$$

The right hand side of the above formula is the Fourier-Stieltjes transform of the function $(2\pi)^{-1/2} \int_F \left\{ \int_{-\infty}^{x/\sqrt{Z}} \exp(-u^2/2) du \right\} dP$, $-\infty < x$

^{*)} $I(A)$ denotes the indicator of the set A .

$< +\infty$.

§ 2. Proof of Theorem. By the condition (L-I) there exists a sequence $\{\varepsilon_n\}$ of positive numbers such that

$$(2.1) \quad \begin{cases} 2\varepsilon_m \leq 1 \text{ for all } m, \varepsilon_n \rightarrow 0, \sum_{k=1}^n P(|Y_k| \geq \varepsilon_n A_n) \rightarrow 0 \text{ and} \\ A_n^{-2} \sum_{k=1}^n E\{Y_k^2 I(|Y_k| \geq \varepsilon_n A_n)\} \rightarrow 0, \text{ as } n \rightarrow +\infty. \end{cases}$$

Using this sequence $\{\varepsilon_n\}$ let us put for $k \geq 1$ and $n \geq 1$

$$Y_{k,n} = Y_k I(|Y_k| < \varepsilon_n A_n) - E\{Y_k I(|Y_k| < \varepsilon_n A_n) | \mathcal{F}_{k-1}\}.$$

Then for each n , $\{Y_{k,n}, \mathcal{F}_k, k \geq 1\}$ is a martingale difference sequence.

Lemma 1. We have

- (i) $\lim_{n \rightarrow +\infty} A_n^{-1} \sum_{k=1}^n |Y_k - Y_{k,n}| = 0$, in *pr.*,
- (ii) $\lim_{n \rightarrow +\infty} A_n^{-2} \sum_{k=1}^n Y_{k,n}^2 = Z$, in *pr.*

Proof. Since $|E\{Y_k I(|Y_k| < \varepsilon_n A_n) | \mathcal{F}_{k-1}\}| = |E\{Y_k I(|Y_k| \geq \varepsilon_n A_n) | \mathcal{F}_{k-1}\}|$, we have by (2.1)

$$\begin{aligned} A_n^{-1} \sum_{k=1}^n E|Y_k - Y_{k,n}| &\leq 2A_n^{-1} \sum_{k=1}^n E|Y_k I(|Y_k| \geq \varepsilon_n A_n)| \\ &\leq 2A_n^{-1} \sum_{k=1}^n \{P(|Y_k| \geq \varepsilon_n A_n) EY_k^2 I(|Y_k| \geq \varepsilon_n A_n)\}^{1/2} \\ &\leq 2 \left\{ \sum_{k=1}^n P(|Y_k| \geq \varepsilon_n A_n) \right\}^{1/2} \left\{ A_n^{-2} \sum_{k=1}^n EY_k^2 I(|Y_k| \geq \varepsilon_n A_n) \right\}^{1/2} \rightarrow 0, \\ &\hspace{15em} \text{as } n \rightarrow +\infty, \end{aligned}$$

and we can prove the first part. Next we have

$$A_n^{-2} \sum_{k=1}^n |Y_k^2 - Y_{k,n}^2| \leq \max_{1 \leq k \leq n} (|Y_k| + |Y_{k,n}|) A_n^{-2} \sum_{k=1}^n |Y_k - Y_{k,n}|.$$

Therefore, we can prove (ii), by (2.1), (i) and (L-II).

Now for any fixed positive number M and $n \geq 1$ let us put

$$S_n(\omega, M) = S_n(\omega) = \begin{cases} +\infty, & \text{if } \sum_{k=1}^{\infty} Y_{k,n}^2(\omega) \leq MA_n^2, \\ \min\{m; \sum_{k=1}^m Y_{k,n}^2(\omega) > MA_n^2\}, & \text{otherwise.} \end{cases}$$

Then for each n , $S_n(\omega)$ is a stopping time with respect to $\{\mathcal{F}_k\}$ and

$$(2.2) \quad S_n(\omega) \geq M/4\varepsilon_n^2$$

because $|Y_{k,n}(\omega)| \leq 2\varepsilon_n A_n$. Next we put, for $k=1, 2, \dots, n$ and $n=1, 2, \dots$,

$$(2.3) \quad \hat{Y}_{k,n} = Y_{k,n} I(S_n \geq k) \text{ and } \mathcal{F}_{k,n} = \mathcal{F}_{\min(k, S_n)}.$$

Then for each n , $\{\hat{Y}_{k,n}, \mathcal{F}_{k,n}, k \geq 1\}$ is a martingale difference sequence (cf. [2, p. 300]) and by (2.1) and (2.3), we have

$$(2.4) \quad A_n^{-2} \sum_{k=1}^n \hat{Y}_{k,n}^2 \leq M + 4\varepsilon_n^2 \leq M + 1.$$

In the following let λ denote any fixed real number and

$$P_{k,n}(\omega, \lambda) = P_{k,n}(\omega) = \prod_{j=1}^k (1 + i\lambda \hat{Y}_{j,n}(\omega) A_n^{-1}).$$

Lemma 2. We have, for any set $F \in \sigma(\bigcup_{n=1}^{\infty} \mathcal{F}_n)$,

$$\lim_{n \rightarrow +\infty} \int_F P_{n,n}(\omega) dP = P(F).$$

Proof. Since $|P_{k,n}|^2 \leq \exp(\lambda^2 A_n^{-2} \sum_{j=1}^k \hat{Y}_{j,n}^2)$, (2.4) implies that

$$(2.5) \quad |P_{k,n}|^2 \leq \exp\{\lambda^2(M+1)\}, \text{ for } 1 \leq k \leq n \text{ and } n \geq 1.$$

On the other hand from the theory of measure it follows that if $F \in \sigma(\bigcup_{n=1}^{\infty} \mathcal{F}_n)$, then for any given $\varepsilon > 0$ there exists a set G such that $P\{F \Delta G\} < \varepsilon$ and $G \in \bigcup_{n=1}^{\infty} \mathcal{F}_n$. Therefore, by (2.5) and the above fact it suffices to prove the lemma for any set $F \in \mathcal{F}_m$ where m is any fixed positive integer. Hereafter we assume that $F \in \mathcal{F}_m$. Then by (2.2) it is seen that

(2.6) $F \in \mathcal{F}_{k,n}$, for all (k, n) such that $m \leq k$ and $m \leq M/4\varepsilon_n^2$. Hence by (2.1), (2.6) and (2.5), we have for sufficiently large n

$$\begin{aligned} \int_F P_{n,n}(\omega) dP &= \int_F \left\{ 1 + \sum_{k=1}^n i\lambda \hat{Y}_{k,n}(\omega) A_n^{-1} P_{k-1,n}(\omega) \right\} dP \\ &= P(F) + \sum_{k=1}^m \int_F i\lambda \hat{Y}_{k,n}(\omega) A_n^{-1} P_{k-1,n}(\omega) dP \\ &= P(F) + O(m |\varepsilon_n \lambda| (1 + |2\varepsilon_n \lambda|)^{m-1}) \\ &= P(F) + o(1), \quad \text{as } n \rightarrow +\infty. \end{aligned}$$

Lemma 3. We have, for any set $F \in \sigma(\bigcup_{n=1}^{\infty} \mathcal{F}_n)$,

$$\lim_{n \rightarrow +\infty} \int_F \exp \left(i\lambda \sum_{k=1}^n \hat{Y}_{k,n} / A_n \right) dP = \int_F \exp(-\lambda^2 Z_M / 2) dP,$$

where $Z_M(\omega) = \min\{Z(\omega), M\}$.

Proof. From (2.3) and (2.1) it is seen that

$$\begin{aligned} A_n^{-2} \sum_{k=1}^n \hat{Y}_{k,n}^2 &= A_n^{-2} \sum_{k=1}^n Y_{k,n}^2, & \text{if } S_n(\omega, M) \geq n, \\ M < A_n^{-2} \sum_{k=1}^n \hat{Y}_{k,n}^2 &\leq M + 4\varepsilon_n^2, & \text{if } S_n(\omega, M) < n. \end{aligned}$$

Therefore, (ii) in Lemma 1 and (2.1) imply that

$$\lim_{n \rightarrow +\infty} A_n^{-2} \sum_{k=1}^n \hat{Y}_{k,n}^2 = Z_M, \quad \text{in pr.}$$

Hence, by (2.5)

$$\lim_{n \rightarrow +\infty} E \left| P_{n,n} \left\{ \exp \left(-\lambda^2 \sum_{k=1}^n \hat{Y}_{k,n}^2 / 2A_n^2 \right) - \exp(-\lambda^2 Z_M / 2) \right\} \right| = 0.$$

On the other hand since Lemma 2 and (2.5) imply that

$$\lim_{n \rightarrow +\infty} \int_F P_{n,n} \exp(-\lambda^2 Z_M / 2) dP = \int_F \exp(-\lambda^2 Z_M / 2) dP,$$

we have

$$(2.7) \quad \lim_{n \rightarrow +\infty} \int_F P_{n,n} \exp \left(-\lambda^2 \sum_{k=1}^n \hat{Y}_{k,n}^2 / 2A_n^2 \right) dP = \int_F \exp(-\lambda^2 Z_M / 2) dP.$$

Further it is easily seen that if $|x| < 1/2$, then

$$\exp(x) = (1+x) \exp\{(x^2/2) + \theta(x)\} \quad \text{and} \quad |\theta(x)| \leq |x|^3.$$

Therefore, by (2.1) and (2.4)

$$\left| \sum_{k=1}^n \theta(i\lambda \hat{Y}_{k,n} / A_n) \right| \leq 2\varepsilon_n |\lambda|^3 (M+1) = o(1), \quad \text{as } n \rightarrow +\infty,$$

and since $|P_{n,n} \exp(-\lambda^2 \sum_{k=1}^n \hat{Y}_{k,n}^2 / 2A_n^2)| \leq 1$, we have

$$\begin{aligned} \exp \left(i\lambda \sum_{k=1}^n \hat{Y}_{k,n} A_n^{-1} \right) &= P_{n,n} \exp \left\{ -\lambda^2 \sum_{k=1}^n \hat{Y}_{k,n}^2 / 2A_n^2 + \sum_{k=1}^n \theta(i\lambda \hat{Y}_{k,n} / A_n) \right\} \\ &= P_{n,n} \exp \left(-\lambda^2 \sum_{k=1}^n \hat{Y}_{k,n}^2 / 2A_n^2 \right) + o(1), \\ &\quad \text{uniformly on } \Omega, \text{ as } n \rightarrow +\infty. \end{aligned}$$

Thus by the above relation and (2.7) we can prove the lemma.

Lemma 4. *We have, for any set $F \in \sigma(\bigcup_{n=1}^{\infty} \mathcal{F}_n)$,*

$$\lim_{n \rightarrow +\infty} \int_F \exp\left(i\lambda \sum_{k=1}^n Y_k/A_n\right) dP = \int_F \exp(-\lambda^2 Z/2) dP.$$

Proof. By (i) in Lemma 1 it is enough to show that

$$(2.8) \quad \lim_{n \rightarrow +\infty} \int_F \exp\left(i\lambda \sum_{k=1}^n Y_{k,n}/A_n\right) dP = \int_F \exp(-\lambda^2 Z/2) dP.$$

If we put $E_M = \{Z(\omega) > M\}$ and $E_{n,M} = \{A_n^{-2} \sum_{k=1}^n Y_{k,n}^2(\omega) > M\}$, then (ii) in Lemma 1 implies that $P(E_{n,M}) \rightarrow P(E_M)$ as $n \rightarrow +\infty$, at the continuity points M of $P(E_M)$. Therefore, for any given $\varepsilon > 0$ we can take M and n_0 such that

$$P(E_M) < \varepsilon \quad \text{and} \quad P(E_{n,M}) < \varepsilon \quad \text{if } n \geq n_0.$$

Since $\omega \in E_{n,M}^c$ and $k \leq n$ imply that $\hat{Y}_{k,n}(\omega) = Y_{k,n}(\omega)$, we have for $n \geq n_0$

$$\begin{cases} P\left\{E \mid \exp\left(i\lambda \sum_{k=1}^n \hat{Y}_{k,n}/A_n\right) - \exp\left(i\lambda \sum_{k=1}^n Y_{k,n}/A_n\right) \mid < 2\varepsilon, \right. \\ \left. P\left\{E \mid \exp(-\lambda^2 Z_M/2) - \exp(-\lambda^2 Z/2) \mid < \varepsilon. \right. \end{cases}$$

Thus by Lemma 3 and above relations we can prove (2.8).

By Lemma 4 (1.1) holds and Theorem is proved.

§ 3. Proof of Corollary. For simplicity of writing the formulas we prove Corollary only for positive x . Let ε ($0 < \varepsilon < 1$) be any given number and put for $k=0, \pm 1, \pm 2, \dots$ and h ($0 < 2h < \varepsilon$)

$$a(k) = \exp(kh) \quad \text{and} \quad E_k = \{a(k) \leq \sqrt{Z}(\omega) < a(k+1)\}.$$

Then clearly E_k 's are disjoint sets in $\sigma(\bigcup_{n=1}^{\infty} \mathcal{F}_n)$ and $\bigcup_{k=-\infty}^{\infty} E_k = \{Z(\omega) \neq 0\}$. Therefore we have, by Theorem,

$$\varliminf_{n \rightarrow +\infty} P\{X_n/A_n \leq x\sqrt{Z}, Z \neq 0\}$$

$$\begin{aligned} &\leq \sum_{|k| < m_0} \lim_{n \rightarrow +\infty} P\{X_n/A_n \leq xa(k+1), E_k\} + \sum_{|k| \geq m_0} P(E_k) \\ &\leq \sum_{k=-\infty}^{\infty} (2\pi)^{-1/2} \int_{E_k} \left\{ \int_{-\infty}^{xa(k+1)/\sqrt{Z}} \exp(-u^2/2) du \right\} dP + \sum_{|k| \geq m_0} P(E_k), \end{aligned}$$

and in the same way

$$\varliminf_{n \rightarrow +\infty} P\{X_n/A_n \leq x\sqrt{Z}, Z \neq 0\}$$

$$\geq \sum_{k=-\infty}^{\infty} (2\pi)^{-1/2} \int_{E_k} \left\{ \int_{-\infty}^{xa(k)/\sqrt{Z}} \exp(-u^2/2) du \right\} dP - \sum_{|k| \geq m_0} P(E_k).$$

Since $x \exp(-x^2/2) < 1$ for $x > 0$, we have

$$\begin{aligned} &\sum_k \int_{E_k} \left\{ \int_{xa(k)/\sqrt{Z}}^{xa(k+1)/\sqrt{Z}} \exp(-u^2/2) du \right\} dP \\ &\leq \sum_k \int_{E_k} (e^h - 1)(xa(k)/\sqrt{Z}) \exp\{-x^2 a^2(k)/2Z\} dP \\ &\leq \sum_k 2h P(E_k) < \varepsilon P(Z \neq 0) < \varepsilon. \end{aligned}$$

Since $xa(k)/\sqrt{Z} \leq x < xa(k+1)/\sqrt{Z}$ on E_k and we can take m_0 so large that $\sum_{|k| \geq m_0} P(E_k) < \varepsilon$, Corollary is proved.

References

- [1] B. M. Brown: Martingale central limit theorems. *Ann. Math. Stats.*, **42**, 59-66 (1971).
- [2] J. L. Doob: *Stochastic Processes*. John Wiley and Sons, New York (1953).