

19. Tail Probabilities for Positive Random Variables Satisfying Some Moment Conditions

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1. Let X be a positive random variable such that the asymptotic inequality

$$(c(1-\varepsilon))^{2n} \Gamma(2n+1)^\beta \leq E[X^{2n}] \leq (d(1+\varepsilon))^{2n} \Gamma(2n+1)^\beta$$

(n : integer)

holds for all ε , $0 < \varepsilon < 1$, where $0 < c \leq d < +\infty$ and $0 < \beta < 1$. Then L. Davies [1] has proved the following inequalities as a corollary of his theorem:

$$\begin{aligned} \beta d^{-1/\beta} &\leq \liminf_{x \rightarrow +\infty} -\log P(X \geq x) / x^{1/\beta} \\ &\leq \limsup_{x \rightarrow \infty} -\log P(X \geq x) / x^{1/\beta} \\ &\leq \beta d^{-1/\beta} (r_u / r_l)^{1/\beta}, \end{aligned}$$

where $0 < r_l \leq 1 \leq r_u < +\infty$ are the two positive roots of $f(y) = 0$,

$$f(y) = \beta(c/d)^{1/\beta} y^{1/\beta} / (1-\beta) - y / (1-\beta) + 1.$$

We will extend his result to a class of positive random variables satisfying some moment conditions which includes his result. For this aim, we shall define “nearly regularly varying function with index α ” which is first introduced in [2].

2. Let $\sigma(x)$ be a positive measurable function defined on $[c_0 + \infty)$, ($c_0 > 0$). We say that $\sigma(x)$ is a “nearly regularly varying function with index α ” if and only if there exist two positive constants $r_1 \geq r_2$ and a slowly varying function $s(x)$ such that

$$r_2 x^\alpha s(x) \leq \sigma(x) \leq r_1 x^\alpha s(x).$$

In particular, we say that $\sigma(x)$ is a “nearly slowly varying function” if $\alpha = 0$.

As is well known (for example see [3]) $s(x)$ is represented as follows:

$$s(x) = b(x) \exp \int_x^{\infty} a(t) / t dt,$$

where $a(x)$ and $b(x)$ are measurable functions such that

$$\lim_{x \rightarrow \infty} b(x) = b > 0 \quad \text{and} \quad \lim_{x \rightarrow \infty} a(x) = 0.$$

3. **Theorem 1.** *Let X be a positive random variable. Assume that there exist two positive constants c_1 and h , and also a non-decreasing nearly regularly varying function $\sigma(x)$ with index α . $0 < \alpha < 1$, defined on $[1/h, +\infty)$ such that*

$$E[X^{2n}] \leq c_1^{2n} \prod_{k=1}^{2n} \sigma(k/h)$$

holds for any $n \geq n_0$. Then

$$P(X \geq A\sigma(x)) \leq \exp \{ -(\alpha - \delta(1/h))g(hx) \}$$

holds for any $x \geq 2n_0/h$, where

$$b_1 = \sup_{x \geq 1/h} r_1 b(x), \quad b_2 = \inf_{x \geq 1/h} r_2 b(x),$$

$$A = c_1 b_1 / b_2, \quad \delta(x) = \sup_{t \geq x} |a(t)|$$

and

$$g(y) = y - (1/2) \log y - \log \sqrt{2\pi} - 2 - 0(1/y).$$

Theorem 2. In addition to the conditions of Theorem 1, we assume that there exists a constant $c_2 (\leq c_1)$ such that

$$E[X^{2n}] \geq c_2^{2n} \prod_{k=1}^{2n} \sigma(k/h)$$

holds for any $n \geq n_0$. Then

$$\lim_{x \rightarrow \infty} \log P(X > A\sigma(x/h)) / x \geq -\kappa(\alpha + \delta(1/h)) \quad (\kappa > 0)$$

holds uniformly in $h \leq 1$. In particular, we can choose $\kappa = 1$ if $c_1 = c_2$, and $b_1 = b_2$.

4. First we prove Theorem 1. By Chebyshev's inequality it follows that

$$P(X \geq A\sigma(x_n)) \leq E[X^{2n}] / (A\sigma(x_n))^{2n}$$

$$\leq b_2^{2n} \sigma(x_n)^{-2n} \prod_{k=1}^{2n} (k/h)^\alpha \bar{s}(k/h),$$

where

$$\bar{s}(x) = \exp \int_x^\infty a(t) / t dt \quad \text{and} \quad x_n = 2n/h, \quad n \geq n_0.$$

With elementary calculus by making use of Stirling's formula for $n!$, we have

$$\prod_{k=1}^{2n} \bar{s}(k/h) = \exp \left\{ 2n \int_0^{2n/h} a(t) / t dt - \sum_{k=1}^{2n-1} k \int_{k/h}^{(k+1)/h} a(t) / t dt \right\}$$

and

$$\sum_{k=1}^{2n-1} k \int_{k/h}^{(k+1)/h} |a(t)| / t dt \leq \delta(1/h) \sum_{k=1}^{2n-1} k \log(1 + 1/k)$$

$$= \delta(1/h) (2n - (1/2) \log 2n - \log \sqrt{2\pi} - 0(1/n)).$$

Again by Stirling's formula, we have

$$b_2^{2n} \sigma(x_n)^{-2n} \prod_{k=1}^{2n} (k/h)^\alpha \bar{s}(k/h)$$

$$\leq \exp \{ -(\alpha - \delta(1/h))(2n - (1/2) \log 2n - \log \sqrt{2\pi} - 0(1/n)) \}.$$

Finally, it follows for $x_n \leq x \leq x_{n+1}$ that

$$P(X \geq A\sigma(x)) \leq P(X \geq A\sigma(x_n))$$

$$\leq \exp \{ -(\alpha - \delta(1/h))(hx_n - (1/2) \log hx_n - \log \sqrt{2\pi} - 0(1/n)) \}$$

$$= \exp \{ -(\alpha - \delta(1/h))g(hx) \}.$$

Now we prove Theorem 2. Setting

$$F(x) = P(X \leq x), \quad B = (1 + \varepsilon)A, \quad m = [\kappa n] \quad (\kappa > 1), \quad \text{and} \quad p = [\mu m] \quad (\mu > 0),$$

we have

$$\begin{aligned} P(X > A\sigma(x_n)) &\geq \int_{A\sigma(x_n)+0}^{B\sigma(x_n)} dF(x) \\ &\geq (B\sigma(x_n))^{-2m} \left(\int_0^{+\infty} - \int_0^{A\sigma(x_n)} - \int_{B\sigma(x_n)+0}^{+\infty} \right) x^{2m} dF(x). \end{aligned}$$

Let us estimate each term.

$$\begin{aligned} \text{I} &= \int_0^{+\infty} x^{2m} dF(x) = E[X^{2m}] \\ &\geq (c_2 b_2)^{2m} \exp \left\{ 2m \left(\alpha \log(2m/h) + \int_{\cdot}^{2m/h} a(t)/tdt \right) \right. \\ &\quad \left. - \alpha(2m - (1/2) \log 2m - \log \sqrt{2\pi} - 0(1/m)) - \sum_{k=1}^{2m-1} k \int_{k/h}^{(k+1)/h} a(t)/tdt \right\}, \\ \text{II} &= \int_0^{A\sigma(x_n)} x^{2m} dF(x) = \left(\int_0^{A\sigma(x_{n_0})} + \int_{A\sigma(x_{n_0}+0)}^{A\sigma(x_n)} \right) x^{2m} dF(x) = \text{II}_1 + \text{II}_2, \\ \text{II}_1 &\leq (A\sigma(x_{n_0}))^{2m}, \\ \text{II}_2 &\leq (A\sigma(x_{n_0+1}))^{2m} (1 - F(A\sigma(x_{n_0}))) + \sum_{k=n_0+1}^{n-1} (A\sigma(x_{k+1}))^{2m} (1 - F(A\sigma(x_k))) \\ &= \text{II}_{2,1} + \text{II}_{2,2}, \\ \text{II}_{2,1} &\leq (A b_1)^{2m} \exp \left\{ 2m \left(\alpha \log(2n_0+2)/h + \int_{\cdot}^{(2n_0+2)/h} a(t)/tdt \right) \right. \\ &\quad \left. - \left(\alpha - \delta(1/h) \right) \left(2n_0 - \frac{1}{2} \log 2n_0 - \log \sqrt{2\pi} - O\left(\frac{1}{n_0}\right) \right) \right\}, \\ \text{II}_{2,2} &\leq (A\sigma(x_n))^{2m} \exp \left\{ -\alpha \left(2n - 2 - \frac{1}{2} \log(2n-2) - \log \sqrt{2\pi} - O\left(\frac{1}{n-1}\right) \right) \right. \\ &\quad \left. - \sum_{k=1}^{2n-3} k \int_{k/h}^{(k+1)/h} a(t)/tdt \right\} \sum_{k=n_0+1}^{n-1} (\sigma(x_{k+1})/\sigma(x_n))^{2m} \\ &\quad \times \exp \left\{ \alpha \left(2n - 2k + 2 - \frac{1}{2} \log(n-1)/k + O\left(\frac{1}{k}\right) - O\left(\frac{1}{n-1}\right) \right) \right. \\ &\quad \left. + \sum_{j=2k}^{2n-3} j \int_{j/h}^{(j+1)/h} a(t)/tdt \right\} \\ &\leq (A b_1)^{2m} \exp \left\{ 2m \left(\alpha \log 2n/h + \int_{\cdot}^{2n/h} a(t)/tdt \right) \right. \\ &\quad \left. - \alpha \left(2n - 2 - \frac{1}{2} \log(2n-2) - \log \sqrt{2\pi} - O\left(\frac{1}{n-1}\right) \right) \right. \\ &\quad \left. - \sum_{j=1}^{2n-3} j \int_{j/h}^{(j+1)/h} a(t)/tdt \right\} \sum_{k=0}^{\infty} \exp \{ -2k(\alpha(k-1) - 2\kappa\delta(2n_0/h)) \}, \\ \text{III} &= \int_{B\sigma(x_n)+0}^{+\infty} x^{2m} dF(x) \leq (B\sigma(x_n))^{-2p} E[X^{2(m+p)}] \\ &\leq (1+\varepsilon)^{-2p} (c_1 b_1)^{2m} \exp \left\{ 2m \left(\alpha \log(2m+2p)/h + \int_{\cdot}^{(2m+2p)/h} a(t)/tdt \right) \right. \\ &\quad \left. + 2p \left(\alpha \log(m+p)/n + \int_{2n/h}^{(2m+2p)/h} a(t)/tdt \right) \right. \\ &\quad \left. - \alpha \left(2m + 2p - \frac{1}{2} \log(2m+2p) - \log \sqrt{2\pi} - O\left(\frac{1}{m+p}\right) \right) \right. \\ &\quad \left. - \sum_{k=1}^{2m+2p-1} k \int_{k/h}^{(k+1)/h} a(t)/tdt \right\}. \end{aligned}$$

Therefore, for suitably chosen κ and ε we have

$$\lim_{n \rightarrow +\infty} (\text{II} + \text{III})/I = O$$

uniformly in $h \leq 1$. This yields the proof of Theorem 2.

References

- [1] Davies, L.: Tail probabilities for positive random variables with entire characteristic functions of very regular growth. *Z. Angew. Math. Mech.*, **56**(3), 334–336 (1976).
- [2] Kôno, N.: On the modulus of continuity of sample functions of Gaussian processes. *J. Math. Kyoto Univ.*, **10**(3), 493–536 (1970).
- [3] Seneta, E.: Regularly varying functions. *Lecture Notes in Mathematics*, Springer-Verlag, 508 (1976).