

Variance of Lipschitz functions and an isoperimetric problem for a class of product measures

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The maximal variance of Lipschitz functions (with respect to the ℓ^1 -distance) of independent random vectors is found. This is then used to solve the isoperimetric problem, uniformly in the class of product probability measures with given variance.

Keywords: isoperimetry; Lipschitz function; variance inequality

1. Statements

Let $\xi = (\xi_1, \dots, \xi_n)$ be a vector of independent random variables with finite variance $\sigma_i^2 = \text{var } \xi_i$, $1 \leq i \leq n$. Denote by \mathcal{F}_1 the class of all functions on \mathbb{R}^n which are Lipschitz with respect to the ℓ^1 -distance

$$d_1(x, y) = \|x - y\|_1 = \sum_{k=1}^n |x_k - y_k|, \quad x, y \in \mathbb{R}^n.$$

By definition, $f \in \mathcal{F}_1$, if for all $x, y \in \mathbb{R}^n$, $|f(x) - f(y)| \leq d_1(x, y)$. Let $S_n = \xi_1 + \dots + \xi_n$.

Theorem 1. *In the class \mathcal{F}_1 , the maximal value of $\text{var } f(\xi)$ is attained at the function $f(x) = x_1 + \dots + x_n$. In other words, for any $f \in \mathcal{F}_1$,*

$$\text{var } f(\xi) \leq \text{var } S_n = \sum_{i=1}^n \sigma_i^2. \quad (1.1)$$

Fernique (1981, Theorem 3.2) proved an inequality similar to (1.1) for $f \in \mathcal{F}_1$ convex. However, in that case ξ is only assumed to be symmetrically distributed, i.e. for all $\epsilon_i = \pm 1$, the random vectors $(\epsilon_i \xi_1, \dots, \epsilon_n \xi_n)$ have the same distribution (of course, this assumption holds if the ξ_i are i.i.d. with a symmetric one-dimensional distribution). In contrast to Fernique's difficult proof, Theorem 1 can easily be obtained by induction.

Theorem 1 also has the following consequence: Denote by $M^n(\sigma)$ the family of all the

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product measures $\mu = \mu_1 \otimes \dots \otimes \mu_n$ on \mathbb{R}^n with given variance $\text{var}(\mu) = \sigma^2$, where

$$\text{var}(\mu) = \sum_{i=1}^n \int_{\mathbb{R}} \left| x - \int_{\mathbb{R}} t d\mu_i(t) \right|^2 d\mu_i(x).$$

Hence, with the above notation, $\text{var}(\mu) = \text{var} S_n$. Now, given a set $A \subset \mathbb{R}^n$ and $h > 0$, denote by

$$A^h = A + hB_1 = \{x \in \mathbb{R}^n : d_1(a, x) < h, \text{ for some } a \in A\}$$

the open h -neighbourhood of A (B_1 is the open ℓ^1 -unit ball in \mathbb{R}^n). From Theorem 1 we obtain a solution to the isoperimetric problem with respect to the ℓ^1 -distance uniformly in the class $M^n(\sigma)$ controlled by the parameter σ .

Theorem 2. For any $h > 0, \sigma > 0$ and $p \in (0, 1)$,

$$\inf_{\mu \in M^n(\sigma)} \inf_{\mu(A) \geq p} \mu(A^h) = \begin{cases} p, & \text{if } h \leq \frac{\sigma}{\sqrt{p(1-p)}}, \\ 1 - \frac{p\sigma^2}{ph^2 - \sigma^2}, & \text{if } h \geq \frac{\sigma}{\sqrt{p(1-p)}}. \end{cases} \tag{1.2}$$

The first infimum in (1.2) is taken over all the $\mu \in M^n(\sigma)$, the second is taken over all the Borel sets A of μ -measure greater or equal to p . In particular, from Theorem 2, we have:

Corollary 3. Given $\sigma > 0$ and $p \in (0, 1)$, one can guarantee that $\mu(A^h) > p$ regardless of the dimension $n \geq 1$, regardless of the measure $\mu \in M^n(\sigma)$, and regardless of the set $A \subset \mathbb{R}^n$ of μ -measure p , if and only if

$$h > h(p, \sigma) \equiv \frac{\sigma}{\sqrt{p(1-p)}}.$$

Otherwise, it is possible to have $\mu(A^h) = p$.

Equality in (1.2) is easy to obtain when $n = 1$. Indeed, denote by δ_x the unit mass at the point $x \in \mathbb{R}$. If $h \leq h(p, \sigma)$, take

$$\mu = p\delta_0 + (1-p)\delta_{h(p,\sigma)}, \quad A = \{0\}.$$

Then, $\text{var}(\mu) = \sigma^2, A^h = (-h, h)$, so $\mu(A^h) = p = \mu(A)$. If $h \geq h(p, \sigma)$, take

$$\mu = p\delta_0 + q\delta_x + r\delta_h, \quad A = \{0\},$$

with $r = p\sigma^2/(ph^2 - \sigma^2), q = 1 - p - r, x = rh/(p + q)$. Then it is again easy to verify that $\text{var}(\mu) = \sigma^2$, and that $\mu(A^h) = 1 - p\sigma^2/(ph^2 - \sigma^2)$.

Since equality in (1.2) is attained when $n = 1$, (1.2) will not change if the h -neighbourhood is defined with respect to the ℓ^2 -distance, or, more generally, with respect to the ℓ^α -distance in $\mathbb{R}^n, 1 \leq \alpha \leq +\infty$. Indeed, the ℓ^α -unit ball B_α is larger than B_1 , hence, $A + hB_1 \subset A + hB_\alpha$, and therefore $\mu(A + hB_1) \leq \mu(A + hB_\alpha)$. Hence, the same inequality holds when one takes the second infimum in (1.2). But all the balls B_α coincide when $n = 1$ (in which case equality in (1.2) is attained).

For individual measures μ (for example, for those having finite exponential moments) there exist estimates for $1 - \mu(A^h)$ which decrease exponentially when $h \rightarrow +\infty$ (see Talagrand 1994). For example, given $\sigma_i > 0, 1 \leq i \leq n$, let $\mu = \mu_1 \otimes \dots \otimes \mu_n \in M^n(\sigma)$, where $\mu_i = (\delta_{\sigma_i} + \delta_{-\sigma_i})/2, \sigma^2 = \sigma_1^2 + \dots + \sigma_n^2$. Then, as shown in Talagrand (1994, Proposition 2.1.1., Theorem 2.4.1.) (see also Ledoux 1994, p. 24, for an extension to non-identical marginals), if $h > 0, \mu(A) = p$, then

$$\mu(A + hB_1) \geq 1 - \frac{1}{p} \exp(-h^2/4\sigma^2). \tag{1.3}$$

When all the $\sigma_i = 1$, the extremal sets minimizing $\mu(A^h)$, while $\mu(A) = p$ is fixed, are known, having been obtained by Harper (1966). If one minimizes $\mu(A^h)$ over all convex sets A , the situation changes considerably, and we are then dealing with a much more powerful concentration principle discovered by Talagrand (1988; 1994). In particular, when all the $\sigma_i = 1$, one has

$$\mu(A + hB_2) \geq 1 - \frac{1}{p} \exp(-h^2/8).$$

In our case, since one is looking for a uniformly minimal value of $\mu(A + hB_1)$, it does not matter whether one considers convex sets or arbitrary sets, since the extremal $A = \{0\}$ is convex.

To complete this section, we give an inequality which is actually equivalent to the second part of (1.2). For non-empty sets $A, B \subset \mathbb{R}^n$, let $d_1(A, B) = \inf \{d_1(a, b) : a \in A, b \in B\}$.

Corollary 4. For any $\mu \in M^n(\sigma)$, and any non-empty Borel sets $A, B \subset \mathbb{R}^n$,

$$d_1(A, B) \leq \sigma \sqrt{\frac{1}{\mu(A)} + \frac{1}{\mu(B)}}. \tag{1.4}$$

Let $\mu(A) > 0, \mu(B) > 0$ be such that $\mu(A) + \mu(B) \leq 1$. Then, choosing $B = \{h\}$ with h equal to the right-hand side of (1.4), it is easily seen that equality in (1.4) is attained at the same measure μ and for the same set $A = \{0\}$ as the second inequality in (1.2).

2. Proofs

A statement slightly more general than Theorem 1 will actually be proved. Assume we have n measurable spaces (X_k, Σ_k) and n measurable functions $h_k = h_k(x_k, y_k)$ defined on $X_k \times X_k, 1 \leq k \leq n$, and which vanish on the diagonal $x_k = y_k$. let ξ_k be independent random variables with values in $X_k, 1 \leq k \leq n$, such that

$$2\sigma_k^2 = E h_k^2(\xi_k, \eta_k) < +\infty,$$

where η_k is an independent copy of ξ_k . Put $\xi = (\xi_1, \dots, \xi_n)$.

Lemma 5. Let f be a measurable function defined on $X_1 \times \dots \times X_n$ such that

$$|f(x) - f(y)| \leq \sum_{k=1}^n h_k(x_k, y_k), \quad (2.1)$$

for $x = (x_1, \dots, x_n), y = (y_1, \dots, y_n) \in X_1 \times \dots \times X_n$. Then

$$\text{var} f(\xi) \leq \sum_{k=1}^n \sigma_k^2. \quad (2.2)$$

Proof. This lemma is proved by induction on the dimension n . For $n = 1$, and since $2\text{var} f(\xi) = \int \int (f(\xi) - f(\eta))^2 d\mu(\xi) d\mu(\eta)$, (2.2) is immediate. Assume now that (2.2) is true for n . Denote by μ_{n+1} the distribution of ξ_{n+1} , and by P_n the distribution of the random vector (ξ_1, \dots, ξ_n) , thus $P_{n+1} = P_n \otimes \mu_{n+1}$ is the distribution of $(\xi_1, \dots, \xi_{n+1})$. Let $f : X_1 \times \dots \times X_{n+1} \rightarrow \mathbb{R}$ satisfy (2.1). Now, fix x_{n+1} . Since the function $g(x_1, \dots, x_n) = f(x_1, \dots, x_n, x_{n+1})$ satisfies (2.1), making use of the induction hypotheses and writing (2.2) for g , we obtain:

$$\int g^2 dP_n \leq \left(\int g dP_n \right)^2 + \sum_{k=1}^n \sigma_k^2. \quad (2.3)$$

The function $m(x_{n+1}) = \int g dP_n$ is well defined, measurable and as a function of one variable,

$$|m(x_{n+1}) - m(y_{n+1})| \leq h_{n+1}(x_{n+1}, y_{n+1}).$$

Thus m satisfies (2.1), hence

$$\int m^2 d\mu_{n+1} \leq \left(\int m d\mu_{n+1} \right)^2 + \sigma_{n+1}^2. \quad (2.4)$$

Integrating (2.3) over X_{n+1} (with respect to μ_{n+1}), and taking into account (2.4), gives (2.2) for f . Lemma 5 and thus Theorem 1 are proved. \square

Proof of Theorem 2. Let $A \subset \mathbb{R}^n$ be such that $\mu(A) \leq p$. Since the function $f(x) = \inf_{a \in A} d_1(a, x)$ belongs to \mathcal{F}_1 , we have, by Theorem 1, that $\text{var} f \leq \sigma^2$. In addition, $f \geq 0$ and $\mu(f = 0) \geq p$. Note also that $A^h = \{x \in \mathbb{R}^n : f(x) < h\}$. To get (1.2), it just remains to appeal to the following result:

Lemma 6. For any $h > 0, \sigma > 0$ and $p \in (0, 1)$,

$$\sup P(\xi \geq h) = \begin{cases} 1 - p, & \text{if } h \leq \frac{\sigma}{\sqrt{p(1-p)}}, \\ \frac{p\sigma^2}{ph^2 - \sigma^2}, & \text{if } h \geq \frac{\sigma}{\sqrt{p(1-p)}}, \end{cases} \quad (2.5)$$

where the supremum is taken over all non-negative random variables ξ on a probability space (Ω, \mathcal{B}, P) such that $P(\xi = 0) \geq p$ and $\text{var} \xi \leq \sigma^2$.

Proof. Denote by $\mathcal{L}(\xi)$ the distribution of ξ . The cases of equality in (2.5) were, in fact, already settled in Section 1. If

$$h \leq h(p, \sigma), \quad \mathcal{L}(\xi) = p\delta_0 + (1 - p)\delta_{h(p, \sigma)},$$

then $\text{var}(\xi) = \sigma^2, P(\xi = 0) = p, P(\xi \geq h) = 1 - p$. If

$$h > h(p, \sigma), \quad \mathcal{L}(\xi) = p\delta_0 + q\delta_x + r\delta_h,$$

where $r = p\sigma^2 / (ph^2 - \sigma^2), q = 1 - p - r, x = rh / (p + q)$, then, as easily verified, we have $\text{var}(\xi) = \sigma^2, P(\xi = 0) = p$, and $P(\xi \geq h) = r = p\sigma^2 / (ph^2 - \sigma^2)$. So, one need only show that whenever $h \geq h(p, \sigma)$,

$$P(\xi \geq h) \leq \frac{p\sigma^2}{ph^2 - \sigma^2}. \tag{2.6}$$

To prove this, we first show, following a suggestion by M. Talagrand, that in (2.6) it suffices to consider only those ξ whose distribution is of the type $\mathcal{L}(\xi) = p_0\delta_0 + p_1\delta_x + p_2\delta_h$, for some $0 \leq x < h$. Then, keeping p_2 constant, we maximize the functional $J = p_2$ over all $p_0 \geq p$ and $x \in [0, h)$ such that $\text{var}(\xi) \leq \sigma^2$.

Note first that in (2.6), ξ can be replaced by $\eta = \min(\xi, h)$ since $P(\eta \geq h) = P(\xi \geq h)$, while $\text{var}(\eta) \leq \text{var}(\xi) \leq \sigma^2$ (η is a Lipschitz function of $\xi: \eta = f(\xi)$, where $f(t) = \min(t, h)$). Then, let ξ take values in $[0, h]$, have distribution ν , and assume that $\nu(0, h) > 0$. Then, ν can (uniquely) be written as

$$\nu = p_0\delta_0 + p_1\lambda + p_2\delta_h,$$

where the distribution λ is concentrated in $(0, h)$. Let ξ_1 be a random variable whose distribution is λ , and let $x = E(\xi_1)$. Then

$$\begin{aligned} \text{var}(\xi) &= p_1 E(\xi_1^2) + p_2 h^2 - (p_1 x + p_2 h)^2 \\ &= p_1 \text{var}(\xi_1) + p_1(1 - p_1)x^2 - 2p_1 p_2 h x - p_2 h^2. \end{aligned}$$

Therefore, given the mean value x , $\text{var}(\xi)$ is minimal if and only if $\lambda = \delta_x$, when $\text{var}(\xi_1) = 0$. Thus, ξ can be replaced in (2.6) by a random variable η which takes three values, 0, x and h .

So let us assume that $\mathcal{L}(\xi) = p_0\delta_0 + p_1\delta_x + p_2\delta_h, 0 \leq x < h, p_0 \geq p, p_1, p_2 \geq 0, p_0 + p_1 + p_2 = 1$. Again, $J = p_2$ is constant. By simple algebra, and for fixed p_0, p_1, p_2 , the minimal value of

$$\text{var}(\xi) = (p_1 x^2 + p_2 h^2) - (p_1 x + p_2 h)^2,$$

as a function of x in $(0, h)$, is attained at

$$x = \frac{p_2}{p_0 + p_2} h.$$

For this value of x , we find

$$\begin{aligned}
 \text{var}(\xi) &= \left(p_1 \frac{p_2^2}{(p_0 + p_2)^2} h^2 + p_2 h^2 \right) - \left(p_1 \frac{p_2}{p_0 + p_2} h + p_2 h \right)^2 \\
 &= \left[\frac{p_2}{(p_0 + p_2)^2} (p_1 p_2 + (p_0 + p_2)^2) - \frac{p_2^2}{(p_0 + p_2)^2} (p_1 + (p_0 + p_2))^2 \right] h^2 \\
 &= \frac{p_2}{(p_0 + p_2)^2} [p_1 p_2 + (p_0 + p_2)^2 - p_2] h^2 \\
 &= \frac{p_0 p_2}{p_0 + p_2} h^2.
 \end{aligned}$$

Now, we have to maximize $J = p_2$ under the condition

$$\text{var}(\xi) = \frac{p_0 p_2}{p_0 + p_2} h^2 \leq \sigma^2. \quad (2.7)$$

From (2.7), when p_0 decreases, $\text{var}(\xi)$ also decreases, while $J = p_2 = 1 - p_0 - p_1$ increases (p_1 is fixed). Hence, to conclude, it is enough to consider only the case $p_0 = p$. The possible maximal value $p_2 = 1 - p$ satisfies (2.7) if and only if $p(1 - p) \leq \sigma^2/h^2$, that is, if and only if $h \leq h(p, \sigma)$. Otherwise, if $h > h(p, \sigma)$, or even if $h = h(p, \sigma)$, the maximal value of $J = p_2$ is, according to (2.7), the value which satisfies

$$\frac{p p_2}{p + p_2} h^2 = \sigma^2.$$

The only solution to this equation is given by

$$p_2 = \frac{p \sigma^2}{p h^2 - \sigma^2}.$$

Lemma 6 follows. □

Proof of Corollary 4. Let $p = \mu(A)$, $q = \mu(B)$. If $p = 0$ or $q = 0$, there is nothing to prove. If $p + q > 1$, then $A \cap B \neq \emptyset$, so $d_1(A, B) = 0$. Thus, we need consider only the case $p + q \leq 1$. Let $p, q > 0$, $p + q \leq 1$, and assume $A \cap B = \emptyset$. Note that

$$h \equiv \sigma \sqrt{\frac{1}{p} + \frac{1}{q}} \geq h(p, \sigma).$$

Therefore, by (1.2),

$$1 - \mu(A^h) \leq \frac{p \sigma^2}{p h^2 - \sigma^2} = q,$$

and again by (1.2), for all $h_1 > h$, $1 - \mu(A^{h_1}) < q$. Hence, $B \cap (A^{h_1} \setminus A) \neq \emptyset$, and therefore, $d_1(A, B) \leq h_1$. Letting $h_1 \rightarrow h$ completes the proof. □

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