

ON THE PROBABILITY THAT AN INTEGER CHOSEN  
 ACCORDING TO THE BINOMIAL DISTRIBUTION  
 BE  $k$ -FREE

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**Introduction.** Let  $s$  and  $t$  be integers chosen from among the first  $n + 1$  non-negative integers according to a binomial distribution with parameter  $p$ ,  $0 < p < 1$ . Consider the probability that  $s$  and  $t$  be relatively prime. In [1] we showed that this probability tends to  $6/\pi^2$ , independent of  $p$ , as  $n \rightarrow \infty$ . Suppose now we choose a single integer  $s$  from the first  $n + 1$  non-negative integers according to a binomial distribution and ask what is the probability that  $s$  be square-free. In this paper we show that the techniques of [1] can also be used to show that this probability is  $6/\pi^2$  in the limit. In fact we show something more, viz., that the probability that  $s$  be  $k$ -free,  $k$  any integer greater than 1, is  $1/\zeta(k)$  where  $\zeta$  denotes the Riemann zeta-function. ( $s$  is  $k$ -free if and only if  $s$  is not divisible by the  $k$ -th power of any prime.) In section 1 we deal with the case  $k > 2$  and in section 2, with the case  $k = 2$ .

1. Let  $n$  be a non-negative integer and denote by  $N_n$  the set of integers  $0, 1, 2, \dots, n$ . Let  $P_n$  be a probability distribution on  $N_n$  and let  $Q_k$  denote the set of non-negative  $k$ -free integers. Set  $Q_k(n) = Q_k \cap N_n$ . For any positive integer  $d$ , let  $A_n(d) = \{j \in N_n : j \equiv 0 \pmod{d}\}$ . We then have the following.

**LEMMA 1.** *Let  $P_n$  be any probability measure on  $N_n$ . Then for  $n > 1$ ,*

$$P_n(Q_k(n)) = \sum_{1 \leq d \leq n^{1/k}} \mu(d) \{P_n(A_n(d^k)) - P_n(\{0\})\}.$$

**PROOF.** Let  $p_1 < p_2 < \dots < p_s$  be the primes less than or equal to  $n^{1/k}$ . Then, if  $\tilde{Q}_k(n)$  denotes the complement of  $Q_k(n)$  in  $N_n$ , we have

$$\tilde{Q}_k(n) = \bigcup_{i=1}^s A_n(p_i^k).$$

Therefore

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$$\begin{aligned}
 P_n(Q_k(n)) &= 1 - P_n(\tilde{Q}_k(n)) = 1 - P_n\left(\bigcup_{i=1}^s A_n(p_i^k)\right) \\
 &= 1 - \sum_{r=1}^s \sum_{(i_1, i_2, \dots, i_r)} (-1)^{r-1} P_n(A_n(p_{i_1}^k) \\
 &\quad \cap A_n(p_{i_2}^k) \cap \dots \cap A_n(p_{i_r}^k)),
 \end{aligned}$$

where the inner sum is taken over all  $r$ -tuples  $(i_1, i_2, \dots, i_r)$  such that  $1 \leq i_1 < i_2 < \dots < i_r \leq s$ . Now it is clear that if  $(d_1, d_2) = 1$ , then  $A_n(d_1) \cap A_n(d_2) = A_n(d_1 d_2)$ . Hence this last expression can be rewritten as

$$1 + \sum_{r=1}^s \sum_{(i_1, i_2, \dots, i_r)} (-1)^r P_n(A_n((p_{i_1} p_{i_2} \dots p_{i_r})^k)).$$

Now if  $(p_{i_1} p_{i_2} \dots p_{i_r})^k > n$ ,  $A_n((p_{i_1} p_{i_2} \dots p_{i_r})^k) = \{0\}$ . Hence this last expression is the same as

$$\sum_{1 \leq d \leq n^{1/k}} \mu(d) P_n(A_n(d^k)) + \sum_{r=1}^s \sum_{p_{i_1} p_{i_2} \dots p_{i_r} > n^{1/k}} \mu(p_{i_1} p_{i_2} \dots p_{i_r}) P_n(\{0\}).$$

Since

$$\begin{aligned}
 \sum_{d|p_1 p_2 \dots p_s} \mu(d) &= 0, \\
 \sum_{r=1}^s \sum_{p_{i_1} p_{i_2} \dots p_{i_r} > n^{1/k}} \mu(p_{i_1} p_{i_2} \dots p_{i_r}) &= - \sum_{1 \leq d \leq n^{1/k}} \mu(d).
 \end{aligned}$$

This observation completes the proof of the lemma.

If  $P_n$  is the uniform distribution on  $N_n$  ( $P_n(j) = (n+1)^{-1}$  for all  $j \in N_n$ ) then it is easy to check that  $|P_n(A_n(d)) - d^{-1}| < n^{-1}$  uniformly in  $d$ . Using this estimate along with Lemma 1 and the fact that  $\sum \mu(d) d^{-k} \rightarrow 1/\zeta(k)$ , it is not difficult to prove

$$\lim_{n \rightarrow \infty} P_n(Q_k(n)) = 1/\zeta(k)$$

for all  $k \geq 2$ .

From now on  $P_n$  will always be taken to be a binomial distribution relative to some fixed parameter  $p$  with  $0 < p < 1$ . Thus  $P_n(j) = \binom{n}{j} p^j (1-p)^{n-j}$ . For  $1 \leq d \leq n$  define  $\epsilon_n(d)$  by

$$\epsilon_n(d) = P_n(A_n(d)) - d^{-1} = \sum_{j=0(d)} \binom{n}{j} p^j(1-p)^{n-j} - d^{-1}.$$

LEMMA 2.  $|\epsilon_n(d)| \ll n^{-1/2}$  uniformly in  $d$ .

PROOF. See [1].

THEOREM 3. If  $P_n$  is a binomial distribution, then  $\lim_{n \rightarrow \infty} P_n(Q_k(n)) = 1/\zeta(k)$  for all  $k \geq 3$ .

PROOF. By Lemma 1 we have

$$\begin{aligned} P_n(Q_k(n)) &= \sum_{1 \leq d \leq n^{1/k}} \mu(d) \{P_n(A_n(d^k)) - P_n(\{0\})\} \\ &= \sum_{1 \leq d \leq n^{1/k}} \mu(d) \{d^{-k} + \epsilon_n(d^k) - (1-p)^n\} \\ &= \sum_{1 \leq d \leq n^{1/k}} \mu(d)d^{-k} + \sum_{1 \leq d \leq n^{1/k}} \mu(d)\epsilon_n(d^k) \\ &\quad - (1-p)^n \sum_{1 \leq d \leq n^{1/k}} \mu(d). \end{aligned}$$

The first sum tends to  $1/\zeta(k)$  while the last sum goes to zero as  $n \rightarrow \infty$ . For the middle sum we have by Lemma 2

$$\left| \sum_{1 \leq d \leq n^{1/k}} \mu(d)\epsilon_n(d^k) \right| \ll n^{1/k}n^{-1/2}.$$

Thus for  $k > 2$  this term goes to zero which proves the theorem.

2. In this section we show that  $\lim_{n \rightarrow \infty} P_n(Q_2(n)) = 6/\pi^2 (= 1/\zeta(2))$  where  $P_n$  is a binomial distribution. As in the proof of Theorem 3 it is sufficient to show that

$$\lim_{n \rightarrow \infty} \sum_{1 \leq d^2 \leq n} |\epsilon_n(d^2)| = 0.$$

We need the following lemmas. For proofs of the first two we refer to [1].

LEMMA 4.

$$\sum_{|k-pn| > pn^{3/4}} \binom{n}{k} p^k(1-p)^{n-k} \ll n^{-1}.$$

LEMMA 5. If  $d > p(n + n^{3/4})$ , then  $|\epsilon_n(d)| \ll d^{-1}$  uniformly in  $d$ .

LEMMA 6. Let  $K_n$  be the number of integers  $d$  which satisfy  $pn^{3/4} \leq d^2 \leq p(n - n^{3/4})$  and which have the property that for some integer  $k$ ,  $kd^2$  is in the interval  $(p(n - n^{3/4}), p(n + n^{3/4}))$ . Then  $K_n \ll n^{3/8}$ .

PROOF. Let  $u = pn$ ,  $v = pn^{3/4}$  and let  $s = [(u + v)/v]$ . Suppose  $kd^2 \in (u - v, u + v)$ . Then we must have  $2 \leq k \leq s$ . For each such  $k$  we ask how many possible  $d$ 's are there such that  $kd^2 \in (u - v, u + v)$ . Such  $d$ 's must lie in the interval

$$(((u - v)/k)^{1/2}, ((u + v)/k)^{1/2}).$$

Hence there are not more than  $z_k = ((u + v)^{1/2} - (u - v)^{1/2})k^{-1/2} + 1$  of them. Now it is easy to verify that  $(u + v)^{1/2} - (u - v)^{1/2} \leq (2v^2/u)^{1/2}$ . Therefore

$$K_n = \sum_{k=2}^s z_k < (2v^2/u)^{1/2} \sum_{k=2}^s k^{-1/2} + s - 1 \leq v(2s/u)^{1/2} + s - 1 \ll n^{3/8}.$$

We now state and prove our main result as

THEOREM 7. Let  $P_n$  be a binomial distribution. Then

$$\lim_{n \rightarrow \infty} P_n(Q_2(n)) = 6/\pi^2.$$

PROOF. As stated at the beginning of this section we need to show

$$(1) \quad \lim_{n \rightarrow \infty} \sum_{1 \leq d^2 \leq n} |\epsilon_n(d^2)| = 0.$$

Let  $n_1 = pn^{3/4}$ ,  $n_2 = p(n - n^{3/4})$  and  $n_3 = p(n + n^{3/4})$ . The sum in (1) can then be written as

$$(2) \quad \sum_{1 \leq d^2 \leq n} = \sum_{1 \leq d^2 \leq n_1} + \sum_{n_1 < d^2 \leq n_2} + \sum_{n_2 < d^2 \leq n_3} + \sum_{n_3 < d^2 \leq n}.$$

(We assume that  $n$  is large enough so that  $n_1 < n_2$  and  $n_3 < n$ .) We will examine each of these sums separately. By Lemma 2

$$\sum_{1 \leq d^2 \leq n_1} |\epsilon_n(d^2)| \ll n_1^{1/2} n^{-1/2} \leq n^{3/8} n^{-1/2} = n^{-1/4}$$

and hence the first term on the right-hand side of (2) goes to zero as  $n \rightarrow \infty$ . A similar argument works for the third sum on the right-hand side of (2).

By Lemma 5  $|\epsilon_n(d^2)| \ll d^{-2}$  for  $d^2 > p(n + n^{3/4})$ . Hence the fourth sum

$$\sum_{n_3 < d^2 \leq n} |\epsilon_n(d^2)| \ll \sum_{n_3 < d^2 \leq n} d^{-2} < \sum_{d=[n_3^{1/2}] }^{\infty} d^{-2}$$

and hence goes to zero because it is less than the tail of a convergent series.

The second sum on the right-hand side of (2) is somewhat more difficult to deal with. We break it into two parts

$$(3) \quad \sum_{n_1 < d^2 \leq n_2} = \sum'_{n_1 < d^2 \leq n_2} + \sum''_{n_1 < d^2 \leq n_2}$$

where the summation with the prime on it is taken over those  $d^2$  which have the property that for some integer  $k$ ,  $kd^2$  is in the interval  $(n_2, n_3)$  and the double primed summation is taken over the remaining  $d^2$ . By Lemmas 2 and 6 we have

$$\sum'_{n_1 < d^2 \leq n_2} |\epsilon_n(d^2)| \ll n^{3/8}n^{-1/2} = n^{-1/8}.$$

Hence the single primed sum goes to zero as  $n \rightarrow \infty$ . We now examine the double primed sum. Recall that

$$\epsilon_n(d^2) = \sum_{k=0(d^2)} \binom{n}{k} p^k(1-p)^{n-k} - d^{-2}.$$

For the  $d^2$  under consideration we have by Lemma 4

$$\begin{aligned} \sum_{k=0(d^2)} \binom{n}{k} p^k(1-p)^{n-k} &= \sum_{\substack{k=0(d^2) \\ |k-pn| > pn^{3/4}}} \binom{n}{k} p^k(1-p)^{n-k} \\ &\leq \sum_{|k-pn| > pn^{3/4}} \binom{n}{k} p^n(1-p)^{n-k} \ll n^{-1}. \end{aligned}$$

Hence for those  $d^2$ ,  $|\epsilon_n(d^2)| \ll d^{-2}$ . Thus for the double primed sum

$$\sum''_{n_1 < d^2 \leq n_2} |\epsilon_n(d^2)| \ll \sum_{n_1 < d^2 \leq n_2} d^{-2} < \sum_{d=[n_1^{1/2}] }^{\infty} d^{-2}$$

and hence goes to zero as  $n \rightarrow \infty$ . This completes the proof of Theorem 7.

## REFERENCE

1. J. E. Nymann and W. J. Leahey, *On the probability that integers chosen according to the binomial distribution are relatively prime*, Acta Arithmetica 31 (1976), 205-211.

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