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Multidimensional q -Normal and related distributions - Markov case

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Abstract

We define and study distributions in \mathbb{R}^d that we call q -Normal. For $q = 1$ they are really multidimensional Normal, for $q \in (-1, 1)$ they have densities, compact support and many properties that resemble properties of ordinary multidimensional Normal distribution. We also consider some generalizations of these distributions and indicate close relationship of these distributions to Askey-Wilson weight function i.e. weight with respect to which Askey-Wilson polynomials are orthogonal and prove some properties of this weight function. In particular we prove a generalization of Poisson-Mehler expansion formula.

Key words: Normal distribution, Poisson-Mehler expansion formula, q -Hermite, Al-Salam-Chihara Chebyshev, Askey-Wilson polynomials, Markov property.

AMS 2000 Subject Classification: Primary 62H10, 62E10; Secondary: 60E05, 60E99.

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1 Introduction

The aim of this paper is to define, analyze and possibly 'accustom' new distributions in \mathbb{R}^d . They are defined with a help of two one-dimensional distributions that first appeared recently, partially in noncommutative context and are defined through infinite products. That is why it is difficult to analyze them straightforwardly using ordinary calculus. One has to refer to some extent to notations and results of so called q -series theory.

However the distributions we are going to define and examine have *purely commutative, classical probabilistic meaning*. They appeared first in an excellent paper of Bożejko et al. [4] as a by product of analysis of some non-commutative model. Later they also appeared in purely classical context of so called one-dimensional random fields first analyzed by W. Bryc et al. in [1] and [3]. From these papers we can deduce much information on these distributions. In particular we are able to indicate sets of polynomials that are orthogonal with respect to measures defined by these distributions. Those are so called q -Hermite and Al-Salam-Chihara polynomials - a generalizations of well known sets of polynomials. Thus in particular we know all moments of the discussed one-dimensional distributions.

What is interesting about distributions discussed in this paper is that many of their properties resemble similar properties of normal distribution. As stated in the title we consider three families of distributions, however properties of one, called multidimensional q -Normal, are main subject of the paper. The properties of the remaining two are in fact only sketched.

All distributions considered in this paper have densities. The distributions in this paper are parametrized by several parameters. One of this parameters, called q , belongs to $(-1, 1]$ and for $q = 1$ the distributions considered in this paper become ordinary normal. Two out of three families of distributions defined in this paper have the property that all their marginals belong to the same class as the joint, hence one of the important properties of normal distribution. Conditional distributions considered in this paper have the property that conditional expectation of a polynomial is also a polynomial of the same order - one of the basic properties of normal distributions. Distributions considered in this paper satisfy Gebelein inequality -property discovered first in the normal distribution context. Furthermore as in the normal case lack of correlation between components of a random vectors considered in the paper lead to independence of these components. Finally conditional distribution $f_C(x|y, z)$ considered in this paper can be expanded in series of the form $f_C(x|y, z) = f_M(x) \sum_{i=0}^{\infty} h_i(x) g_i(y, z)$ where f_M is a marginal density, $\{h_i\}$ are orthogonal polynomials of f_M and $g_i(y, z)$ are also polynomials. In particular if $f_C(x|y, z) = f_C(x|q)$ that is when instead of conditional distribution of $X|Y, Z$ we consider only distribution of $X|Y$ then $g_i(y) = h_i(y)$. In this case such expansion formula it is a so called Poisson-Mehler formula, a generalization of a formula with h_i being ordinary Hermite polynomials and $f_M(x) = \exp(-x^2/2)/\sqrt{2\pi}$ that appeared first in the normal distribution context.

On the other hand one of the conditional distributions that can be obtained with the help of distributions considered in this paper is in fact a re-scaled and normalized (that is multiplied by a constant so its integral is equal to 1) Askey-Wilson weight function. Hence we are able to prove some properties of this Askey-Wilson density. In particular we will obtain mentioned above, generalization of Poisson-Mehler expansion formula for this density.

To define briefly and swiftly these one-dimensional distributions that will be later used to construct

multidimensional generalizations of normal distributions, let us define the following sets

$$S(q) = \begin{cases} [-2/\sqrt{1-q}, 2/\sqrt{1-q}] & \text{if } |q| < 1 \\ \{-1, 1\} & \text{if } q = -1 \end{cases} .$$

Let us set also $m + S(q) \stackrel{df}{=} \{x = m + y, y \in S(q)\}$ and $\mathbf{m} + \mathbf{S}(q) \stackrel{df}{=} (m_1 + S(q)) \times \dots \times (m_d + S(q))$ if $\mathbf{m} = (m_1, \dots, m_d)$. Sometimes to simplify notation we will use so called indicator functions

$$I_A(x) = \begin{cases} 1 & \text{if } x \in A \\ 0 & \text{if } x \notin A \end{cases} .$$

The two one-dimensional distributions (in fact families of distributions) are given by their densities. The first one has density:

$$f_N(x|q) = \frac{\sqrt{1-q}}{2\pi\sqrt{4-(1-q)x^2}} \prod_{k=0}^{\infty} ((1+q^k)^2 - (1-q)x^2q^k) \prod_{k=0}^{\infty} (1-q^{k+1}) I_{S(q)}(x) \quad (1.1)$$

defined for $|q| < 1, x \in \mathbb{R}$. We will set also

$$f_N(x|1) = \frac{1}{\sqrt{2\pi}} \exp(-x^2/2) . \quad (1.2)$$

For $q = -1$ considered distribution does not have density, is discrete with two equal mass points at $S(-1)$. Since this case leads to non-continuous distributions we will not analyze it in the sequel.

The fact that such definition is reasonable i.e. that distribution defined by $f_N(x|q)$ tends to normal $N(0, 1)$ as $q \rightarrow 1^-$ will be justified in the sequel. The distribution defined by $f_N(x|q), -1 < q \leq 1$ will be referred to as q -Normal distribution.

The second distribution has density:

$$f_{CN}(x|y, \rho, q) = \frac{\sqrt{1-q}}{2\pi\sqrt{4-(1-q)x^2}} \times \quad (1.3a)$$

$$\prod_{k=0}^{\infty} \frac{(1-\rho^2q^k)(1-q^{k+1})((1+q^k)^2 - (1-q)x^2q^k)}{(1-\rho^2q^{2k})^2 - (1-q)\rho q^k(1+\rho^2q^{2k})xy + (1-q)\rho^2(x^2+y^2)q^{2k}} I_{S(q)}(x) \quad (1.3b)$$

defined for $|q| < 1, |\rho| < 1, x \in \mathbb{R}, y \in S(q)$. It will be referred to as (y, ρ, q) -Conditional Normal, distribution. For $q = 1$ we set

$$f_{CN}(x|y, \rho, 1) = \frac{1}{\sqrt{2\pi(1-\rho^2)}} \exp\left(-\frac{(x-\rho y)^2}{2(1-\rho^2)}\right)$$

(in the sequel we will justify this fact). Notice that we have $f_{CN}(x|y, 0, q) = f_N(x|q)$ for all $y \in S(q)$.

The simplest example of multidimensional density that can be constructed from these two distributions is two dimensional density

$$g(x, y|\rho, q) = f_{CN}(x|y, \rho, q) f_N(y|q) ,$$

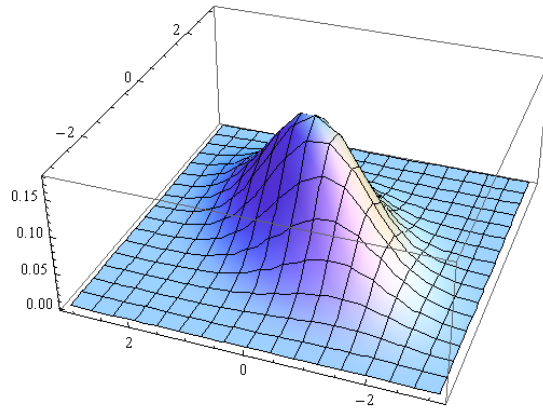


Figure 1: $\rho = .5, q = .8$

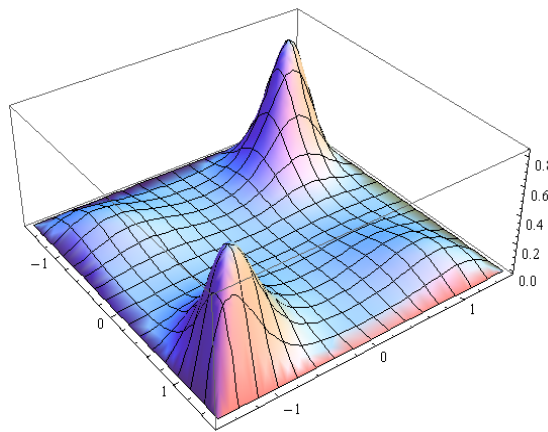


Figure 2: $\rho = .5, q = .8$

that will be referred to in the sequel as $N_2(0,0,1,1,\rho|q)$. Below we give some examples of plots of these densities. One can see from these pictures how large and versatile family of distributions is this family.

It has compact support equal to $S(q) \times S(q)$ and two parameters. One playing similar rôle to parameter ρ in two-dimensional Normal distribution. The other parameter q has a different rôle. In particular it is responsible for modality of the distribution and of course it defines its support.

As stated above, distribution defined by $f_N(x|q)$ appeared in 1997 in [4] in basically non-commutative context. It turns out to be important both for classical and noncommutative probabilists as well as for physicists. This distribution has been 'accustomed' i.e. equivalent form of the density and methods of simulation of i.i.d. sequences drawn from it are e.g. presented in [18]. Distribution f_{CN} , although known earlier in nonprobabilistic context, appeared (as an important probability distribution) in the paper of W. Bryc [1] in a classical context as a conditional distribution of certain Markov sequence. In the following section we will briefly recall basic properties of these distributions as well as of so called q -Hermite polynomials (a generalization of ordinary Hermite polynomials). To do this we have to refer to notation and some of the results of q -series

theory.

The paper is organized as follows. In section 2 after recall some of the results of q -series theory we present definition of multivariate q -Normal distribution. The following section presents main result. The last section contains lengthy proofs of the results from previous section.

2 Definition of multivariate q -Normal and some related distributions

2.1 Auxiliary results

We will use traditional notation of q -series theory i.e. $[0]_q = 0$; $[n]_q = 1 + q + \dots + q^{n-1} = \frac{1-q^n}{1-q}$, $[n]_q! = \prod_{i=1}^n [i]_q$, with $[0]_q! = 1$, $\begin{bmatrix} n \\ k \end{bmatrix}_q = \begin{cases} \frac{[n]_q!}{[n-k]_q! [k]_q!} & , n \geq k \geq 0 \\ 0 & , otherwise \end{cases}$. It will be useful

to use so called q -Pochhammer symbol for $n \geq 1$: $(a|q)_n = \prod_{i=0}^{n-1} (1 - aq^i)$, with $(a|q)_0 = 1$, $(a_1, a_2, \dots, a_k|q)_n = \prod_{i=1}^k (a_i|q)_n$. Often $(a|q)_n$ as well as $(a_1, a_2, \dots, a_k|q)_n$ will be abbreviated to $(a)_n$ and $(a_1, a_2, \dots, a_k)_n$, if it will not cause misunderstanding.

It is easy to notice that $(q)_n = (1 - q)^n [n]_q!$ and that

$$\begin{bmatrix} n \\ k \end{bmatrix}_q = \begin{cases} \frac{(q)_n}{(q)_{n-k} (q)_k} & , n \geq k \geq 0 \\ 0 & , otherwise \end{cases} . \text{ The above mentioned quantities were defined for } |q| < 1 .$$

Note that for $q = 1$ $[n]_1 = n$, $[n_1]! = n!$, $(a|1)_n = (1 - a)^n$ and $\begin{bmatrix} n \\ i \end{bmatrix}_1 = \binom{n}{i}$.

Let us also introduce two functionals defined on functions $g : \mathbb{R} \rightarrow \mathbb{C}$,

$$\|g\|_L^2 = \int_{\mathbb{R}} |g(x)|^2 f_N(x) dx, \quad \|g\|_{CL}^2 = \int_{\mathbb{R}} |g(x)|^2 f_{CN}(x|y, \rho, q) dx$$

and sets:

$$\begin{aligned} L(q) &= \{g : \mathbb{R} \rightarrow \mathbb{C} : \|g\|_L < \infty\}, \\ CL(y, \rho, q) &= \{g : \mathbb{R} \rightarrow \mathbb{C} : \|g\|_{CL} < \infty\}. \end{aligned}$$

Spaces $(L(q), \|\cdot\|_L)$ and $(CL(y, \rho, q), \|\cdot\|_{CL})$ are Hilbert spaces with the usual definition of scalar product.

Let us also define the following two sets of polynomials:

-the q -Hermite polynomials defined by

$$H_{n+1}(x|q) = xH_n(x|q) - [n]_q H_{n-1}(x|q), \quad (2.1)$$

for $n \geq 1$ with $H_{-1}(x|q) = 0$, $H_0(x|q) = 1$, and

-the so called Al-Salam-Chihara polynomials defined by the relationship for $n \geq 0$:

$$P_{n+1}(x|y, \rho, q) = (x - \rho y q^n) P_n(x|y, \rho, q) - (1 - \rho^2 q^{n-1}) [n]_q P_{n-1}(x|y, \rho, q), \quad (2.2)$$

with $P_{-1}(x|y, \rho, q) = 0$, $P_0(x|y, \rho, q) = 1$.

Polynomials (2.1) satisfy the following very useful identity originally formulated for so called continuous q -Hermite polynomials h_n (can be found in e.g. [7] Thm. 13.1.5) and here below presented for polynomials H_n using the relationship

$$h_n(x|q) = (1-q)^{n/2} H_n\left(\frac{2x}{\sqrt{1-q}}|q\right), \quad n \geq 1, \quad (2.3)$$

$$H_n(x|q) H_m(x|q) = \sum_{j=0}^{\min(n,m)} \begin{bmatrix} m \\ j \end{bmatrix}_q \begin{bmatrix} n \\ j \end{bmatrix}_q [j]_q! H_{n+m-2k}(x|q). \quad (2.4)$$

It is known (see e.g. [1]) that q -Hermite polynomials constitute an orthogonal base of $L(q)$ while from [3] one can deduce that $\{P_n(x|y, \rho, q)\}_{n \geq -1}$ constitute an orthogonal base of $CL(y, \rho, q)$. Thus in particular $0 = \int_{S(q)} P_1(x|y, \rho, q) f_{CN}(x|y, \rho, q) dx = \mathbb{E}(X|Y=y) - \rho y$. Consequently, if Y has also q -Normal distribution, then $\mathbb{E}XY = \rho$.

It is known (see e.g. [7] formula 13.1.10) that

$$\sup_{x \in S(q)} |H_n(x|q)| \leq W_n(q) (1-q)^{-n/2}, \quad (2.5)$$

where

$$W_n(q) = \sum_{i=0}^n \begin{bmatrix} n \\ i \end{bmatrix}_q. \quad (2.6)$$

We will also use Chebyshev polynomials of the second kind $U_n(x)$, that is $U_n(\cos \theta) = \frac{\sin(n+1)\theta}{\sin \theta}$ and ordinary (probabilistic) Hermite polynomials $H_n(x)$ i.e. polynomials orthogonal with respect to $\frac{1}{\sqrt{2\pi}} \exp(-x^2/2)$. They satisfy 3-term recurrences:

$$2xU_n(x) = U_{n+1}(x) + U_{n-1}(x), \quad (2.7)$$

$$xH_n(x) = H_{n+1}(x) + nH_{n-1}(x) \quad (2.8)$$

with $U_{-1}(x) = H_{-1}(x) = 0$, $U_0(x) = H_1(x) = 1$.

Some immediate observations concerning q -Normal and (y, ρ, q) -Conditional Normal distributions are collected in the following Proposition:

Proposition 1. 1. $f_{CN}(x|y, 0, q) = f_N(x|q)$.

2. $\forall n \geq 0 : H_n(x|0) = U_n(x/2)$, $H_n(x|1) = H_n(x)$.

3. $\forall n \geq 0 : P_n(x|y, 0, q) = H_n(x|q)$, $P_n(x|y, \rho, 1) = (1 - \rho^2)^{n/2} H_n\left(\frac{x - \rho y}{\sqrt{1 - \rho^2}}\right)$, $P_n(x|y, \rho, 0) = U_n(x/2) - \rho y U_{n-1}(x/2) + \rho^2 U_{n-2}(x/2)$.

4. $f_N(x|0) = \frac{1}{2\pi} \sqrt{4 - x^2} I_{<-2,2>}(x)$, $f_N(x|q) \xrightarrow{q \rightarrow 1^-} \frac{1}{\sqrt{2\pi}} \exp(-x^2/2)$ pointwise.

5. $f_{CN}(x|y, \rho, 0) = \frac{(1 - \rho^2) \sqrt{4 - x^2}}{2\pi((1 - \rho^2)^2 - \rho(1 + \rho^2)xy + \rho^2(x^2 + y^2))} I_{<-2,2>}(x)$, $f_{CN}(x|y, \rho, q) \xrightarrow{q \rightarrow 1^-}$

$\frac{1}{\sqrt{2\pi(1 - \rho^2)}} \exp\left(-\frac{(x - \rho y)^2}{2(1 - \rho^2)}\right)$ pointwise.

Proof. 1. Is obvious. 2. Follows observation that (2.1) simplifies to (2.7) and (2.8) for $q = 0$ and $q = 1$ respectively. 3. First two assertions follow either direct observation in case of $P_n(x|y, \rho, 0)$ or comparison of (2.2) and (2.8) considered for $x \rightarrow (x - \rho y)/\sqrt{1 - \rho^2}$ and then multiplication of both sides by $(1 - \rho^2)^{(n+1)/2}$. Third assertion follows following observations: $P_{-1}(x|y, \rho, 0) = 0$, $P_0(x|y, \rho, 0) = 1$, $P_1(x|y, \rho, 0) = x - \rho y$, $P_2(x|y, \rho, 0) = x(x - \rho y) - (1 - \rho^2)$, $P_{n+1}(x|y, \rho, 0) = xP_n(x|y, \rho, 0) - P_{n-1}(x; y, \rho, 0)$ for $n \geq 1$ which is an equation (2.7) with x replaced by $x/2$.

4. 5. First assertions are obvious. Rigorous prove of pointwise convergence of respective densities can be found in work of [9]. To support intuition we will sketch the proof of convergence in distribution of respective distributions. To do this we apply 2. and 3. and see that $\forall n \geq 1 H_n(x|q) \rightarrow H_n(x)$, and $P_n(x|y, \rho, q) \rightarrow (1 - \rho^2)^{n/2} H_n\left(\frac{x - \rho y}{\sqrt{1 - \rho^2}}\right)$ as $q \rightarrow 1^-$. Now keeping in mind that families $\{H_n(x|q)\}_{n \geq 0}$ and $\{P_n(x|y, \rho, q)\}_{n \geq 0}$ are orthogonal with respect to distributions defined by respectively f_N and f_{CN} we deduce that distributions defined by f_N and f_{CN} tend to normal $N(0, 1)$ and $N(\rho y, (1 - \rho^2))$ distributions weakly as $q \rightarrow 1^-$ since both $N(0, 1)$ and $N(\rho y, (1 - \rho^2))$ are defined by their moments, which are defined by polynomials H_n , and P_n . \square

2.2 Multidimensional q -Normal and related distributions

Before we present definition of the multidimensional q -Normal and related distributions, let us generalize the two discussed above one-dimensional distributions by introducing (m, σ^2, q) -Normal distribution as the distribution with the density $f_N((x - m)/\sigma|q)/\sigma$ for $m \in \mathbb{R}$, $\sigma > 0$, $q \in (-1, 1]$. That is if $X \sim (m, \sigma^2, q)$ -Normal then $(X - m)/\sigma \sim q$ -Normal.

Similarly let us extend definition of (y, ρ, q) -Conditional Normal by introducing for $m \in \mathbb{R}$, $\sigma > 0$, $q \in (-1, 1]$, $|\rho| < 1$, $(m, \sigma^2, y, \rho, q)$ -Conditional Normal distribution as the distribution whose density is equal to $f_{CN}((x - m)/\sigma|y, \rho, q)/\sigma$.

Let $\mathbf{m}, \boldsymbol{\sigma} \in \mathbb{R}^d$ and $\boldsymbol{\rho} \in (-1, 1)^{d-1}$, $q \in (-1, 1]$. Now we are ready to introduce a multidimensional q -Normal distribution $N_d(\mathbf{m}, \boldsymbol{\sigma}^2, \boldsymbol{\rho}|q)$.

Definition 1. Multidimensional q -Normal distribution $N_d(\mathbf{m}, \boldsymbol{\sigma}^2, \boldsymbol{\rho}|q)$, is the continuous distribution in \mathbb{R}^d that has density equal to

$$g(\mathbf{x}|\mathbf{m}, \boldsymbol{\sigma}^2, \boldsymbol{\rho}, q) = f_N\left(\frac{(x_1 - m_1)}{\sigma_1}|q\right) \prod_{i=1}^{d-1} f_{CN}\left(\frac{x_{i+1} - m_{i+1}}{\sigma_{i+1}} \middle| \frac{x_i - m_i}{\sigma_i}, \rho_i, q\right) / \prod_{i=1}^d \sigma_i$$

where $\mathbf{x} = (x_1, \dots, x_d)^d$, $\mathbf{m} = (m_1, \dots, m_d)$, $\boldsymbol{\sigma}^2 = (\sigma_1^2, \dots, \sigma_d^2)$, $\boldsymbol{\rho} = (\rho_1, \dots, \rho_{d-1})$.

As an immediate consequence of the definition we see that $\text{supp}(N_d(\mathbf{m}, \boldsymbol{\sigma}^2|q)) = \mathbf{m} + \mathbf{S}(q)$. One can also easily see that \mathbf{m} is a shift parameter and $\boldsymbol{\sigma}$ is a scale parameter. Hence in particular $\mathbb{E}\mathbf{X} = \mathbf{m}$. In the sequel we will be mostly concerned with distributions $N_d(\mathbf{0}, \mathbf{1}, \boldsymbol{\rho}|q)$.

Remark 1. Following assertion 1. of Proposition 1 we see that distribution $N_d(\mathbf{0}, \mathbf{1}, \mathbf{0}|q)$ is the product distribution of d i.i.d. q -Normal distributions. Another words "lack of correlation means independence" in the case of multidimensional q -Normal distributions. More generally if the sequence $\boldsymbol{\rho} = (\rho_1, \dots, \rho_{d-1})$ contain, say, r zeros at, say, positions t_1, \dots, t_r then the distribution of $N_d(\mathbf{0}, \mathbf{1}, \boldsymbol{\rho})$ is a product distribution of $r+1$ independent multidimensional q -Normal distributions: $N_{t_1}(\mathbf{0}, \mathbf{1}, (\rho_1, \dots, \rho_{t_1-1}))$, \dots , $N_{d-t_r}(\mathbf{0}, \mathbf{1}, (\rho_{t_r+1}, \dots, \rho_{t_d}))$.

Thus in the sequel all considered vectors ρ will be assumed to contain only nonzero elements. Let us introduce the following functions (generating functions of the families of polynomials):

$$\varphi(x, t|q) = \sum_{i=0}^{\infty} \frac{t^i}{[i]_q!} H_i(x|q), \quad (2.9)$$

$$\tau(x, t|y, \rho, q) = \sum_{i=0}^{\infty} \frac{t^i}{[i]_q!} P_i(x|y, \rho, q). \quad (2.10)$$

The basic properties of the discussed distributions will be collected in the following Lemma that contains facts from mostly [7] and the paper [3].

Lemma 1. *i) For $n, m \geq 0$:*

$$\int_{S(q)} H_n(x|q) H_m(x|q) f_N(x|q) dx = \begin{cases} 0 & \text{when } n \neq m \\ [n]_q! & \text{when } n = m \end{cases}.$$

ii) For $n \geq 0$:

$$\int_{S(q)} H_n(x|q) f_{CN}(x|y, \rho, q) dx = \rho^n H_n(y|q).$$

iii) For $n, m \geq 0$:

$$\int_{S(q)} P_n(x|y, \rho, q) P_m(x|y, \rho, q) f_{CN}(x|y, \rho, q) dx = \begin{cases} 0 & \text{when } n \neq m \\ (\rho^2)_n [n]_q! & \text{when } n = m \end{cases}.$$

iv)

$$\int_{S(q)} f_{CN}(x|y, \rho_1, q) f_{CN}(y|z, \rho_2, q) dy = f_{CN}(x|z, \rho_1 \rho_2, q).$$

v) For $|t|, |q| < 1$:

$$\sum_{i=0}^{\infty} \frac{W_i(q) t^i}{(q)_i} = \frac{1}{(t)_{\infty}^2}, \quad \sum_{i=0}^{\infty} \frac{W_i^2(q) t^i}{(q)_i} = \frac{(t^2)_{\infty}}{(t)_{\infty}^4},$$

convergence is absolute, where $W_i(q)$ is defined by (2.6).

vi) For $(1-q)x^2 \leq 2$ and $\forall (1-q)t^2 < 1$:

$$\varphi(x, t|q) = \prod_{k=0}^{\infty} (1 - (1-q)xtq^k + (1-q)t^2q^{2k})^{-1},$$

convergence (2.9) is absolute in t & x and uniform in x . Moreover $\varphi(x, t|q)$ is positive and $\int_{S(q)} \varphi(x, t|q) f_N(x|q) dx = 1$. $\varphi(t, x|1) = \exp(xt - t^2/2)$.

vii) For $(1-q)\max(x^2, y^2) \leq 2$, $|\rho| < 1$ and $\forall (1-q)t^2 < 1$:

$$\tau(x, t|y, \rho, q) = \prod_{k=0}^{\infty} \frac{(1 - (1-q)\rho y t q^k + (1-q)\rho^2 t^2 q^{2k})}{(1 - (1-q)xtq^k + (1-q)t^2 q^{2k})},$$

convergence (2.10) is absolute in t & x and uniform in x . Moreover $\tau(x, t|\theta, \rho, q)$ is positive and $\int_{S(q)} \tau(x, t|y, \rho, q) f_{CN}(x|y, \rho, q) dx = 1$.

$$\tau(x, t|y, \rho, 1) = \exp\left(t(x - \rho y) - t^2(1 - \rho^2)/2\right).$$

viii) For $(1 - q)\max(x^2, y^2) \leq 2$, $|\rho| < 1$:

$$f_{CN}(x|y, \rho, q) = f_N(x|q) \sum_{n=0}^{\infty} \frac{\rho^n}{[n]_q!} H_n(x|q) H_n(y|q) \quad (2.11)$$

and convergence is absolute t , y & x and uniform in x and y .

Proof. i) It is formula 13.1.11 of [7] with obvious modification for polynomials H_n instead of h_n (compare (2.3)) and normalized weight function (i.e. f_N) ii) Exercise 15.7 of [7] also in [1], iii) Formula 15.1.5 of [7] with obvious modification for polynomials P_n instead of $p_n(x|y, \rho, q) = (1 - q)^{n/2} P\left(\frac{2x}{\sqrt{1-q}} \middle| \frac{2y}{\sqrt{1-q}}, \rho, q\right)$ and normalized weight function (i.e. f_{CN}), iv) see (2.6) of [3]. v) Exercise 12.2(b) and 12.2(c) of [7]. vi)-viii) The exact formulae are known and are given in e.g. [7] (Thm. 13.1.1, 13.1.6) and [10] (3.6, 3.10). Absolute convergence of φ and τ follow (2.5) and v). Positivity of φ and τ follow formulae $1 - (1 - q)xtq^k + (1 - q)t^2q^{2k} = (1 - q)(tq^k - x/2)^2 + 1 - (1 - q)x^2/4$ and $1 - (1 - q)\rho y t q^k + (1 - q)\rho^2 t^2 q^{2k} = (1 - q)\rho^2(q^k t - y/(2\rho))^2 + 1 - (1 - q)y^2/4$. Values of integrals follow (2.9) and (2.10) and the fact that $\{H_n\}$ and $\{P_n\}$ are orthogonal bases in spaces $L(q)$ and $CL(y, \rho, q)$. \square

Corollary 1. Every marginal distribution of multidimensional q -Normal distribution $N_d(\mathbf{m}, \sigma^2, \rho|q)$ is multidimensional q -Normal. In particular every one-dimensional distribution is q -Normal. More precisely i -th coordinate of $N_d(\mathbf{m}, \sigma^2, \rho|q)$ - vector has (m_i, σ_i^2, q) - Normal distribution.

Proof. By considering transformation $(X_1, \dots, X_d) \longrightarrow \left(\frac{X_1 - m_1}{\sigma_1}, \dots, \frac{X_d - m_d}{\sigma_d}\right)$ we reduce considerations to the case $N_d(\mathbf{0}, \mathbf{1}, \rho|q)$. First let us consider $d - 1$ dimensional marginal distributions. The assertion of Corollary 1 is obviously true since we have assertion iv) of the Lemma 1. We can repeat this reasoning and deduce that all $d - 2, d - 3, \dots, 2$ dimensional distributions are multidimensional q -Normal. The fact that 1-dimensional marginal distributions are q -normal follows the fact that $f_{CN}(y|x, \rho, q)$ is a one-dimensional density and integrates to 1. \square

Corollary 2. If $\mathbf{X} = (X_1, \dots, X_d) \sim N_d(\mathbf{m}, \mathbf{1}, \rho|q)$, then

i) $\forall n \in \mathbb{N}, 1 \leq j_1 < j_2 < \dots < j_m < i \leq d$:

$X_i | X_{j_1}, \dots, X_{j_m} \sim f_{CN}\left(x_i | x_{j_m}, \prod_{k=j_m}^{i-1} \rho_k, q\right)$. Thus in particular

$$\mathbb{E}\left(H_n(X_i - m_i) | X_{j_1}, \dots, X_{j_m}\right) = \left(\prod_{k=j_m}^{i-1} \rho_k\right)^n H_n(X_{j_m} - m_{j_m})$$

and $\text{var}(X_i | X_{j_1}, \dots, X_{j_m}) = 1 - \left(\prod_{k=j_m}^{i-1} \rho_k\right)^2$.

ii) $\forall n \in \mathbb{N}, 1 \leq j_1 < \dots < j_k < i < j_m < \dots < j_h \leq d$:

$$X_i | X_{j_1}, \dots, X_{j_k}, X_{j_m}, \dots, X_{j_h} \sim f_N(x_i | q) \prod_{l=0}^{\infty} \frac{h_l(x_{j_k}, x_{j_m}, \rho_k^* \rho_m^*, q)}{h_l(x_i, x_{j_k}, \rho_k^*, q) h_l(x_i, x_{j_m}, \rho_m^*, q)},$$

where $h_l(x, y, \rho, q) = ((1 - \rho^2 q^{2l})^2 - (1 - q)\rho q^l (1 + \rho^2 q^{2l})xy + (1 - q)\rho^2 q^{2l}(x^2 + y^2))$, $\rho_k^* = \prod_{i=j_k}^{i-1} \rho_i$, $\rho_m^* = \prod_{i=i}^{j_m-1} \rho_i$. Thus in particular this density depends only on X_{j_k} and X_{j_m} .

Proof. i) As before, by suitable change of variables we can work with distribution $N_d(\mathbf{0}, \mathbf{1}, \rho|q)$. Then following assertion iii) of the Lemma 1 and the fact that m - dimensional marginal, with respect to which we have to integrate is also multidimensional q -Normal and that the last factor in the product representing density of this distribution is $f_{CN}(x_i|x_{j_m}, \prod_{k=j_m}^{i-1} \rho_k, q)$ we get i).

ii) First of all notice that joint distribution of $(X_{j_1}, \dots, X_{j_k}, X_i, X_{j_m}, \dots, X_{j_h})$ depends only on x_{j_k}, x_i, x_{j_m} since sequence $X_i, i = 1 \dots, n$ is Markov. It is also obvious that the density of this distribution exist and can be found as a ratio of joint distribution of (X_{j_k}, X_i, X_{j_m}) divided by the joint density of (X_{j_k}, X_{j_m}) . Keeping in mind that X_{j_k}, X_i, X_{j_m} have the same marginal f_N and because of assertion iv of Lemma 1 we get the postulated form. \square

Having Lemma 1 we can present Proposition concerning mutual relationship between spaces $L(q)$ and $CL(y, \rho, q)$ defined at the beginning of previous section.

Proposition 2. $\forall q \in (-1, 1), y \in S(q), |\rho| < 1 : L(q) = CL(y, \rho, q)$. Besides $\exists C_1(y, \rho, q), C_2(y, \rho, q) : \|g\|_L \leq C_1 \|g\|_{CL}$ and $\|g\|_{CL} \leq C_2 \|g\|_L$ for every $g \in L(q)$.

Proof. Firstly observe that : $(1 - \rho^2 q^{2k})^2 - (1 - q)\rho q^k(1 + \rho^2 q^{2k})xy + (1 - q)\rho^2(x^2 + y^2)q^{2k} = (1 - q)\rho^2 q^{2k} \left(x - y \frac{\rho q^k + \rho^{-1} q^{-k}}{2} \right)^2 + (1 - (1 - q)y^2/4)(1 - \rho^2 q^{2k})^2$ which is elementary to prove. We will use modification of the formula (2.11) that is obtained from it by dividing both sides by $f_N(x|q)$. That is formula:

$$\begin{aligned} & \prod_{k=0}^{\infty} \frac{(1 - \rho^2 q^k)}{(1 - \rho^2 q^{2k})^2 - (1 - q)\rho q^k(1 + \rho^2 q^{2k})xy + (1 - q)\rho^2(x^2 + y^2)q^{2k}} \\ &= \sum_{n=0}^{\infty} \frac{\rho^n}{[n]_q!} H_n(x|q) H_n(y|q). \end{aligned}$$

Now we use (2.5) and assertion v) of Lemma 1 and get $\forall x, y \in S(q)$:

$$f_{CN}(x|y, \rho, q) \leq f_N(x|q) \frac{(\rho^2)_{\infty}}{(\rho)_{\infty}^4}.$$

Hence $C_2 = \frac{(\rho^2)_{\infty}}{(\rho)_{\infty}^4}$ and $\|g\|_{CL}^2 \leq \|g\|_L^2$, for every $g \in L(q)$. Thus $g \in CL(y, \rho, q)$.

Conversely to take a function $g \in CL(y, \rho, q)$. We have

$$\infty > \int_{S(q)} |g(x)|^2 f_{CN}(x|y, \rho, q) dx.$$

Now we keeping in mind that $(1 - \rho^2 q^{2k})^2 - (1 - q)\rho q^k(1 + \rho^2 q^{2k})xy + (1 - q)\rho^2(x^2 + y^2)q^{2k}$ is a quadratic function in x , we deduce that it reaches its maximum for $x \in S(q)$ on the end points of

$S(q)$. Hence we have

$$\begin{aligned} & (1 - \rho^2 q^{2k})^2 - (1 - q)\rho q^k(1 + \rho^2 q^{2k})xy + (1 - q)\rho^2(x^2 + y^2)q^{2k} \\ & \leq (1 + \rho^2 q^{2k} + \sqrt{1 - q}|\rho y||q|^k)^2. \end{aligned}$$

Since for $\forall y \in S(q), |\rho|, |q| < 1$:

$$\prod_{k=0}^{\infty} (1 + \rho^2 q^{2k} + \sqrt{1 - q}|\rho y||q|^k)^2 < \infty$$

and we see that

$$\begin{aligned} & \infty > \int_{S(q)} |g(x)|^2 f_{CN}(x|y, \rho, q) dx = \int_{S(q)} |g(x)|^2 f_N(x|q) \\ & \times \prod_{k=0}^{\infty} \frac{1 - \rho^2 q^k}{(1 - \rho^2 q^{2k})^2 - (1 - q)\rho q^k(1 + \rho^2 q^{2k})xy + (1 - q)\rho^2(x^2 + y^2)q^{2k}} dx \\ & \geq \frac{(\rho^2)_{\infty}}{\prod_{k=0}^{\infty} (1 + \rho^2 q^{2k} + \sqrt{1 - q}|\rho y||q|^k)^2} \int_{S(q)} |g(x)|^2 f_N(x|q) dx. \end{aligned}$$

So $g \in L(q)$. □

Remark 2. Notice that the assertion of Proposition 2 is not true for $q = 1$ since then the respective densities are $N(0, 1)$ and $N(\rho y, 1 - \rho^2)$.

Remark 3. Using assertion of Proposition 2 we can rephrase Corollary 2 in terms of contraction $\mathcal{R}(\rho, q)$, (defined by (2.12), below). For $g \in L(q)$ we have

$$\mathbb{E}(g(X_i) | X_{j_1}, \dots, X_{j_m}) = \mathcal{R} \left(\prod_{k=j_m}^{i-1} \rho_k, q \right) (g(X_{j_m})),$$

where $\mathcal{R}(\rho, q)$ is a contraction on the space $L(q)$ defined by the formula (using polynomials H_n for $|\rho|, |q| < 1$):

$$L(q) \ni f = \sum_{i=0}^{\infty} a_i H_i(x|q) \longrightarrow \mathcal{R}(\rho, q)(f) = \sum_{i=0}^{\infty} a_i \rho^i H_i(x|q). \quad (2.12)$$

By the way it is known that \mathcal{R} is not only contraction but also ultra contraction i.e. mapping L_2 on L_{∞} (Bożejko).

We have also the following almost obvious observation that follows, in fact, from assertion iii) of the Lemma 1.

Proposition 3. Suppose that $\mathbf{X}=(X_1, \dots, X_d) \sim \mathbf{N}_d(0, \mathbf{1}, \boldsymbol{\rho}|q)$ and $g \in L(q)$. Assume that for some $n \in \mathbb{N}$, and $1 \leq j_1 < j_2 \dots < j_m < i \leq d$.

i) If $\mathbb{E}(g(X_i) | X_{j_1}, \dots, X_{j_m}) = \text{polynomial of degree at most } n \text{ of } X_{j_m}$, then function g must be also a polynomial of degree at most n .

ii) If additionally $\mathbb{E}g(X_i) = 0$

$$\mathbb{E}((\mathbb{E}(g(X_i)|X_{j_1}, \dots, X_{j_m}))^2) \leq r^2 \mathbb{E}g^2(X_i), \quad (\text{Generalized Gebelein's inequality})$$

where $r = \prod_{k=j_m}^{i-1} \rho_k$.

Proof. i) The fact that $\mathbb{E}(g(X_i)|X_{j_1}, \dots, X_{j_m})$ is a function of X_{j_m} only, is obvious. Since $g \in L(q)$ we can expand it in the series $g(x) = \sum_{i \geq 0} c_i H_i(x|q)$. By Corollary 2 we know that $\mathbb{E}(g(X_i)|X_{j_1}, \dots, X_{j_m}) = \sum_{i \geq 0} c_i r^i H(X_{j_m}|q)$ for $r = \prod_{k=j_m}^{i-1} \rho_k$. Now since $c_i r^i = 0$ for $i > n$ and $r \neq 0$ we deduce that $c_i = 0$ for $i > n$.

ii) Suppose $g(x) = \sum_{i=1}^{\infty} g_i H_i(x)$. We have $\mathbb{E}(g(X_i)|X_{j_1}, \dots, X_{j_m}) = \sum_{i=1}^{\infty} g_i r^i H_i(Y_{j_m})$. Hence $\mathbb{E}((\mathbb{E}(g(X_i)|X_{j_1}, \dots, X_{j_m}))^2) = \sum_{i=1}^{\infty} g_i^2 r^{2i} [i]_q! \leq r^2 \sum_{i=1}^{\infty} g_i^2 [i]_q! = r^2 \mathbb{E}g^2(X_i)$. \square

Remark 4. As it follows from the above mentioned definition, the multidimensional q -Normal distribution is not a true generalization of n -dimensional Normal law $\mathbf{N}_n(\mathbf{m}, \sigma)$. It a generalization of distribution $\mathbf{N}_n(\mathbf{m}, \sigma)$ with very specific matrix σ namely with entries equal to $\sigma_{ii} = \sigma_i^2; \sigma_{ij} = \sigma_i \sigma_j \prod_{k=i}^{j-1} \rho_k$ for $i < j$ and $\sigma_{ij} = \sigma_{ji}$ for $i > j$ where $\sigma_i; i = 1, \dots, n$ are some positive numbers and $|\rho_i| < 1, i = 1, \dots, n-1$.

Proof. Follows the fact that two dimensional q -Normal distribution of say $(X_i/\sigma_i, X_j/\sigma_j)$ has density $f_N(x_i|q) f_{CN}(x_j|x_i, \prod_{k=i}^{j-1} \rho_k, q)$ if $i < j$. \square

Remark 5. Suppose that $(X_1, \dots, X_n) \sim N_n(\mathbf{m}, \sigma, \rho|q)$ then X_1, \dots, X_n form a finite Markov chain with $X_i \sim (m_i, \sigma_i^2, q)$ -Normal and transition density $X_i|X_{i-1} = y \sim (m_i, \sigma_i^2, y, \rho_{i-1}, q)$ -Conditional Normal distribution

Following assertions vi) and vii) of Lemma 1 we deduce that for $\forall t^2 < 1/(1-q), |\rho| < 1$ functions $\varphi(x, t|q) f_N(x|q)$ and $\tau(x, t|y, \rho, q) f_{CN}(x|y, \rho, q)$ are densities. Hence we obtain new densities with additional parameter t . This observation leads to the following definitions:

Definition 2. Let $|q| \in (-1, 1], t^2 < 1/(1-q), x \in S(q)$. A distribution with the density $\varphi(x, t|q) f_N(x|q)$ will be called *modified* (t, q) -Normal (briefly (t, q) -MN distribution).

We have immediate observation that follows from assertion vi) of Lemma 1.

Proposition 4. i) $\int_{S(q)} \varphi(y, t|q) f_N(y|q) f_{CN}(x|y, \rho, q) dy = \varphi(x, t\rho|q) f_N(x|q)$

ii) Let $X \sim (t, q)$ -MN. Then for $n \in \mathbb{N} : \mathbb{E}(H_n(X|q)) = t^n$.

Proof. i) Using assertions vi) and viii) of the Lemma 1 we get:

$$\begin{aligned} & \int_{S(q)} \varphi(y, t|q) f_N(y|q) f_{CN}(x|y, \rho, q) dy \\ &= f_N(x|q) \int_{S(q)} f_N(y|q) \sum_{i=1}^{\infty} \frac{t^i}{[i]_q!} H_i(y|q) \sum_{j=0}^{\infty} \frac{\rho^j}{[j]_q!} H_j(y|q) H_j(x|q) dy. \end{aligned}$$

Now utilizing assertion ii) of the same Lemma we get:

$$\begin{aligned} & \int_{S(q)} \varphi(y, t|q) f_N(y|q) f_{CN}(x|y, \rho, q) dy \\ &= f_N(x|q) \sum_{i=0}^{\infty} \frac{(t\rho)^i}{[i]_q!} H_i(x|q) = \varphi(x, t\rho|q) f_N(x|q). \end{aligned}$$

To get ii) we utilize assertion vi) of the Lemma 1. \square

In particular we have:

Corollary 3. *If $X \sim (t, q) - MN$, then $\mathbb{E}X = t$, $\text{var}(X) = 1$, $\mathbb{E}\left((X - t)^3\right) = -t(1 - q)$, $\mathbb{E}(X - t)^4 = 2 + q - t^2(5 + 6q + q^2)$.*

Proof. We have $x^3 = H_3(x|q) + (2 + q)H_1(x)$ and $H_4(x|q) + (3 + 2q + q^2)H_2(x|q) + 2 + q = x^4$ so $\mathbb{E}(X - t)^3 = \mathbb{E}(X^3) - 3\mathbb{E}(X^2)t + 3\mathbb{E}(X)t^2 - t^3 = t^3 + (2 + q)t - 3t(1 + t^2) + 3t^3 - t^3 = -t(1 - q)$ and $\mathbb{E}(X - t)^4 = \mathbb{E}(X^4) - 4t\mathbb{E}(X^3) + 6t^2\mathbb{E}(X^2) - 4t^3\mathbb{E}(X) + t^4 = t^4 - (3 + 2q + q^2)t^2 + 2 + q - 4t(t^3 + (2 + q)t) + 6t^2(t^2 + 1) - 4t^4 + t^4$ which reduces to $2 + q - t^2(5 + 6q + q^2)$. \square

In particular kurtosis of $(t, q) - MN$ distributions is equal to $-(1 - q) - t^2(5 + 6q + q^2)$. Hence it is negative and less, for $t \neq 0$, than that of q -Normal which is also negative (equal to $-(1 - q)$).

Assertion i) of the Proposition 4 leads to the generalization of the multidimensional q -Normal distribution that allows different one-dimensional and other marginals.

Definition 3. A distribution in \mathbb{R}^d having density equal to

$$\varphi(x_1, t|q) f_N(x_1|q) \prod_{i=1}^{d-1} f_{CN}(x_{i+1}|x_i, \rho_i, q),$$

where $x_i \in S(q)$, $\rho_i \in (-1, 1) \setminus \{0\}$, $i = 1, \dots, d - 1$, $|q| \in (-1, 1]$, $t^2 < 1/(1 - q)$ will be called *modified multidimensional q -Normal distribution* (briefly $MMN_d(\rho|q, t)$).

Reasoning in the similar way as in the proof of Corollary 1 and utilizing observation following from Proposition 4, we have immediately the following observation.

Proposition 5. *Let $(X_1, \dots, X_d) \sim MMN_d(\rho|q, t)$. Then every marginal of it is also modified multidimensional q -Normal. In particular $\forall i = 1, \dots, d : X_i \sim \left(t \prod_{k=1}^{i-1} \rho_k, q\right) - MN$*

Remark 6. Suppose that $(X_1, \dots, X_d) \sim MMN_d(\rho|q, t)$ and define $\rho_0 = 1$, then the sequence X_1, \dots, X_d form a non-stationary Markov chain such that $X_i \sim \varphi\left(x|t \prod_{k=1}^{i-1} \rho_k, q\right) f_N(x|q)$ with transitional probability $X_{i+1}|X_i = y \sim f_{CN}(x|y, \rho_i, q)$.

We can define another one-dimensional distribution depending on 4 parameters. We have:

Definition 4. Let $|q| \in (-1, 1]$, $t^2 < 1/(1 - q)$, $x, y \in S(q)$, $|\rho| < 1$. A distribution with the density $\tau(x, t|y, \rho, q) f_{CN}(x|y, \rho, q)$ will be called *modified (y, ρ, t, q) -Conditional Normal* (briefly (y, ρ, t, q) -MCN).

We have immediate observation that follows from assertion vii) of Lemma 1.

Proposition 6. Let $X \sim (y, \rho, t, q)$ -MCN. Then for $n \in \mathbb{N}$: $\mathbb{E}(P_n(X|y, \rho, q)) = (\rho^2)_n t^n$.

Hence in particular one can state the following Corollary.

Corollary 4. $\mathbb{E}X = \rho y + (1 - \rho^2)t$, $\text{var}(X) = (1 - \rho^2)(1 - (1 - q)t\rho + (1 - q)t^2\rho^2)$.

Proof. Follows expressions for first two Al-Salam-Chihara polynomials. Namely we have: $P_1(x|y, \rho, q) = x - \rho y$, $P_2(x|y, \rho, q) = x^2 - 1 + \rho^2 + q\rho^2 y^2 - x\rho y(1 + q)$. \square

We can define two formulae for densities of multidimensional distributions in \mathbb{R}^d . Namely one of them would have density of the form

$$\varphi(x_1, t|q) f_N(x_1|q) \tau(x_2, t|x_1, \rho_1, q) \prod_{i=1}^{d-1} f_{CN}(x_{i+1}|x_i, \rho_i, q)$$

and the other of the form

$$f_N(x_1|q) \tau(x_2, t|x_1, \rho_1, q) \prod_{i=1}^{d-1} f_{CN}(x_{i+1}|x_i, \rho_i, q).$$

However to find marginals of such families of distributions is a challenge and an open question. In particular are they also of modified conditional normal type?

3 Main Results

In this section we are going to study properties of 3 dimensional case of multidimensional normal distribution. To simplify notation we will consider vector (Y, X, Z) having distribution $\mathbf{N}_3((0, 0, 0), (1, 1, 1), (\rho_1, \rho_2)|q)$ that is having density $f_{CN}(y|x, \rho_1, q) f_{CN}(x|z, \rho_2, q) f_N(z|q)$. We start with the following obvious result:

Remark 7. Conditional distribution $X|Y, Z$ has density

$$\phi(x|y, z, \rho_1, \rho_2, q) = f_N(x|q) \frac{(\rho_1^2, \rho_2^2)_\infty}{(\rho_1^2 \rho_2^2)_\infty} \prod_{i=0}^{\infty} \frac{w_k(y, z, \rho_1 \rho_2, q)}{w_k(x, y, \rho_1, q) w_k(x, z, \rho_2, q)}, \quad (3.1)$$

where we denoted $w_k(s, t, \rho, q) = (1 - \rho^2 q^{2k})^2 - (1 - q)\rho q^k(1 + \rho^2 q^{2k})st + (1 - q)\rho^2(s^2 + t^2)q^{2k}$.

Proof. It is in fact rewritten version of the proof of assertion ii) of Corollary 2. \square

Remark 8. Notice that $\phi(x|y, z, \rho_1, \rho_2, q)$ is the re-scaled Askey-Wilson density. Namely $\phi(x|y, z, \rho_1, \rho_2, q) = \psi(\frac{\sqrt{1-q}}{2}x|a, b, c, d)$ where

$$a = \frac{\sqrt{1-q}}{2} \rho_1(y - i\sqrt{\frac{4}{1-q} - y^2}), b = \frac{\sqrt{1-q}}{2} \rho_1(y + i\sqrt{\frac{4}{1-q} - y^2}),$$

$$c = \frac{\sqrt{1-q}}{2} \rho_2(z - i\sqrt{\frac{4}{1-q} - z^2}), d = \frac{\sqrt{1-q}}{2} \rho_2(z + i\sqrt{\frac{4}{1-q} - z^2}).$$

and $\psi(t|a, b, c, d)$ is a normalized (that is multiplied by a constant so that its integral is 1) weight function of Askey-Wilson polynomials. Compare e.g. [5] and [7]. Hence our results would concern properties of Askey - Wilson density and Askey-Wilson polynomials.

Let us denote

$$g_n(y, z, \rho_1, \rho_2, q) = \int_{S(q)} H_n(x|q) \phi(x|y, z, \rho_1, \rho_2, q).$$

Remark 9. Following probabilistic interpretation of function g_n (as conditional expectation) and (2.5) we deduce that for $|q| < 1, y, z \in S(q)$

$$|g_n(y, z, \rho_1, \rho_2, q)| \leq W_n(q) (1-q)^{-n/2}. \quad (3.2)$$

Our main result is the following

Theorem 1. i) $\phi(x|y, z, \rho_1, \rho_2|q) = f_N(x|q) \times \sum_{n=0}^{\infty} \frac{H_n(x|q)}{[n]_q!} g_n(y, z, \rho_1, \rho_2, q)$, where g_n is a polynomial of order n in (y, z) .

ii) More over polynomial g_n has the following structure $\forall n \geq 1 : g_n(y, z, \rho_1, \rho_2, q) = \sum_{i=0}^n \binom{n}{i}_q \rho_1^i \rho_2^{n-i} \Theta_{i, n-i}(y, z, \rho_1 \rho_2|q)$, where $\Theta_{k, l}(y, z, \rho_1 \rho_2|q)$ is a polynomial in y of order k and in z of order l . Moreover $\Theta_{0, n}(y, z, 0|q) = H_n(z|q)$ and $\Theta_{n, 0}(y, z, 0, q) = H_n(y|q)$.

Remark 10. Assertion i) of the Theorem 1 is in fact a generalization of Poisson-Mehler formula (that is assertion viii) of Lemma 1) for Askey-Wilson density.

Remark 11. Notice also that for $q = 1$, ϕ is a density function of normal distribution $N\left(\frac{y\rho_1(1-\rho_2^2)+z\rho_2(1-\rho_1^2)}{1-\rho_1^2\rho_2^2}, \frac{(1-\rho_1^2)(1-\rho_2^2)}{1-\rho_1^2\rho_2^2}\right)$ and it is obvious that expectation of any polynomial is a polynomial in $\frac{y\rho_1(1-\rho_2^2)+z\rho_2(1-\rho_1^2)}{1-\rho_1^2\rho_2^2}$. Hence it turns out that this is true for all q -Normal distributions for $q \in (-1, 1]$.

As a Corollary we have the following result.

Corollary 5. Let $\mathbf{X} = (X_1, \dots, X_d) \sim \mathbf{N}_d(0, \mathbf{1}, \boldsymbol{\rho}|q)$. Let us select indices $1 \leq j_1 < \dots < j_k < i < j_m < \dots < j_h \leq d$. Then

$$\forall n \in \mathbb{N} : \mathbb{E}(H_n(X_i|q) | X_{j_1}, \dots, X_{j_k}, X_{j_m}, \dots, X_{j_h}) = \sum_{r=0}^{\lfloor n/2 \rfloor} \sum_{l=0}^{n-2r} A_{r, -\lfloor n/2 \rfloor + r + l}^{(n)} H_l(X_{j_k}|q) H_{n-2r-l}(X_{j_m}|q), \quad (3.3)$$

for $\lfloor \frac{n+2}{2} \rfloor \lfloor \frac{n+3}{2} \rfloor$ constants (depending only on $n, q, \boldsymbol{\rho}$ and numbers i, j_k, j_m) $A_{r,s}^{(n)}$, $r = 0, \dots, \lfloor n/2 \rfloor, s = -\lfloor n/2 \rfloor + r, \dots, -\lfloor n/2 \rfloor + r + n - 2r$.

Corollary below gives detailed form of coefficients $A_{r,s}^{(n)}$ for $n = 1, \dots, 4$.

Corollary 6. Let $\mathbf{X} = (X_1, \dots, X_d) \sim \mathbf{N}_d(0, \mathbf{1}, \rho|q)$. Let $1 \leq i-1 < i < i+1 \leq d$. Then:

$$\mathbb{E}(H_n(X_i) | \mathcal{F}_{\neq i}) = \sum_{r=0}^{\lfloor n/2 \rfloor} \sum_{l=0}^{n-2r} A_{r, -\lfloor n/2 \rfloor + r + l}^{(n)} H_l(X_{i-1}|q) H_{n-2r-l}(X_{i+1}|q),$$

where $A_{0, -\lfloor n/2 \rfloor + l}^{(n)} = \begin{bmatrix} n \\ l \end{bmatrix}_q \frac{\rho_{i-1}^{n-l} (\rho_i^2)_{n-l} \rho_i^l (\rho_{i-1}^2)_l}{(\rho_{i-1}^2 \rho_i^2)_n}$, $l = 0, \dots, n$, $n = 1, \dots, 4$. If $n \leq 3$ then $A_{1, -\lfloor n/2 \rfloor + l}^{(n)} = -[n-1]_q \rho_{i-1} \rho_i A_{0, -\lfloor n/2 \rfloor + l}^{(n)}$, $l = 1, \dots, n-1$. If $n = 4$ then $A_{1,j}^{(4)} = -[3]_q \rho_{i-1} \rho_i A_{0,j}^{(4)}$, $j = -1, 1$ and $A_{1,0}^{(4)} = -[2]_q^2 \rho_{i-1} \rho_i A_{0,0}^{(4)}$, $A_{2,0}^{(4)} = q(1+q) \rho_{i-1}^2 \rho_i^2 A_{0,0}^{(4)}$.

In particular :

$$\text{var}(X_i | X_1, \dots, X_{i-1}, X_{i+1}, \dots, X_d) = \frac{(1 - \rho_{i-1}^2)(1 - \rho_i^2)}{(1 - q\rho_{i-1}^2\rho_i^2)} \quad (3.4)$$

$$\times \left(1 - \frac{(1-q)(X_{i-1} - \rho_{i-1}\rho_i X_{i+1})(X_i - X_{i-1}\rho_{i-1}\rho_i)}{(1 - \rho_{i-1}^2\rho_i^2)^2}\right). \quad (3.5)$$

Remark 12. Notice that in general conditional variance

$\text{var}(X_i | X_1, \dots, X_{i-1}, X_{i+1}, \dots, X_d)$ is not nonrandom indicating that q -Normal distribution does not behave as Normal in this case, however if we set $q = 1$ in (3.4) then we get $\text{var}(X_i | X_1, \dots, X_{i-1}, X_{i+1}, \dots, X_d) = \frac{(1 - \rho_{i-1}^2)(1 - \rho_i^2)}{(1 - \rho_{i-1}^2\rho_i^2)}$ as it should be in the normal case.

Notice that examining the form of coefficients $A_{m,k}^{(n)}$ for $n = 1, \dots, 4$ we can formulate the following Hypothesis concerning general form of them:

Conjecture 1. For $n \geq 1$, we have: $A_{0, -\lfloor n/2 \rfloor + l}^{(n)} = \begin{bmatrix} n \\ l \end{bmatrix}_q \frac{\rho_{i-1}^{n-l} (\rho_i^2)_{n-l} \rho_i^l (\rho_{i-1}^2)_l}{(\rho_{i-1}^2 \rho_i^2)_n}$, $l = 0, \dots, n$. Moreover for $r = 0, \dots, \lfloor n/2 \rfloor$ and $l = 0, \dots, n - 2r$ $A_{r, -\lfloor n/2 \rfloor + r + l}^{(n)} / A_{0, -\lfloor n/2 \rfloor + l}^{(n)} = \rho_{i-1}^r \rho_i^r Q_{r,l}(q)$, where $Q_{r,l}(q)$ is a polynomial in q with coefficients depending only on r and l .

4 Proofs

Proof of the Theorem 1 is based on the properties of the following function $G_{l,k}(y, z, t|q) = \sum_{m \geq 0} \frac{t^m}{[m]_q!} H_{m+l}(y|q) H_{m+k}(z|q)$. We will need some of its properties. Namely we will prove the following Proposition which is in fact a generalization and reformulation (in terms of polynomials H_n) of an old result of Carlitz. Original result of Carlitz concerned polynomials $w_n(x|q) = \sum_{i=0}^n \begin{bmatrix} n \\ i \end{bmatrix}_q x^i$ and expressions of the form $\sum_{i=0}^{\infty} \frac{w_i(x|q) w_{i+k}(x|q) t^i}{(q)_n}$, compare [7], Exercise 12.3(d) or [6].

Proposition 7. i) $\forall k, l \geq 0 : G_{k,l}(y, z, t|q) = G_{l,k}(z, y, t|q)$

ii) for $1 \leq j \leq k$:

$$G_{k,l}(y, z, t|q) = \sum_{i=0}^{j-1} (-1)^i \begin{bmatrix} k \\ i \end{bmatrix}_q q^{\binom{i}{2}} t^i H_{k-i}(y|q) G_{0,i+l}(y, z, t|q) \quad (4.1)$$

$$+ (-1)^j q^{\binom{j}{2}} \sum_{i=j}^k \begin{bmatrix} k \\ i \end{bmatrix}_q \begin{bmatrix} i-1 \\ j-1 \end{bmatrix}_q t^i G_{k-i,i+l}(y, z, t|q).$$

iii)

$$G_{k,0}(y, z, t|q) = \sum_{i=0}^k (-1)^i \begin{bmatrix} k \\ i \end{bmatrix}_q q^{\binom{i}{2}} t^i H_{k-i}(y|q) G_{0,i}(y, z, t|q). \quad (4.2)$$

iv)

$$G_{k,0}(y, z, t|q) = \frac{1}{1 - q^{k(k-1)} t^{2k}} \sum_{i=0}^{k-1} (-1)^i q^{\binom{i}{2}} \begin{bmatrix} k \\ i \end{bmatrix}_q t^i \tau_{k,i}(y, z, t|q), \quad (4.3)$$

where $\tau_{k,i}(y, z, t|q) = (H_{k-i}(y|q) G_{0,i}(y, z, t|q) + (-1)^k q^{\binom{k}{2}} t^k H_{k-i}(z|q) G_{i,0}(y, z, t|q))$.

Proof. i) is obvious.

iii) Take $j = k$ and $l = 0$ in ii).

ii) To prove (4.1) we will use formula

$H_{n+m}(x|q) = H_n(x|q) H_m(x|q) - \sum_{j=1}^{\min(n,m)} \begin{bmatrix} n \\ j \end{bmatrix}_q \begin{bmatrix} m \\ j \end{bmatrix}_q \begin{bmatrix} j \end{bmatrix}_q! H_{n+m-2j}(x|q)$. We have

$$\begin{aligned} G_{k,l}(y, z, t) &= \sum_{m \geq 0} \frac{t^m}{[m]_q!} H_{m+k}(y|q) H_{m+l}(z|q) \\ &= \sum_{m \geq 0} \frac{t^m}{[m]_q!} (H_k(y|q) H_m(y|q) \\ &\quad - \sum_{j=1}^{\min(k,m)} \begin{bmatrix} k \\ j \end{bmatrix}_q \begin{bmatrix} m \\ j \end{bmatrix}_q \begin{bmatrix} j \end{bmatrix}_q! H_{k+m-2j}(y|q)) H_{m+l}(z|q) \\ &= H_k(y|q) G_{0,l}(y, z, t) - \sum_{j=1}^k t^j \begin{bmatrix} k \\ j \end{bmatrix}_q \sum_{m=j}^{\infty} \frac{t^{m-j}}{[m-j]_q!} H_{k+(m-j)-j}(y|q) H_{l+j+(m-j)}(z|q) \\ &= H_k(y|q) G_{0,l}(y, z, t) - \sum_{j=1}^k t^j \begin{bmatrix} k \\ j \end{bmatrix}_q G_{k-j,l+j}(y, z, t). \end{aligned}$$

Hence let us assume that (4.1) is true for $j = 1, 2, \dots, m$. We have after applying just obtained

formula (for $G_{k,l}$) applied for $k \rightarrow k - m$ and $l \rightarrow m + l$:

$$\begin{aligned}
G_{k,l}(y,z,t) &= \sum_{i=0}^{m-1} (-1)^i \begin{bmatrix} k \\ i \end{bmatrix}_q q^{\binom{i}{2}} t^i H_{k-i}(y|q) G_{0,i+l}(y,z,t) \\
&\quad + (-1)^m q^{\binom{m}{2}} \sum_{i=m}^k \begin{bmatrix} k \\ i \end{bmatrix}_q \begin{bmatrix} i-1 \\ m-1 \end{bmatrix}_q t^i G_{k-i,i+l}(y,z,t) \\
&= \sum_{i=0}^{m-1} (-1)^i \begin{bmatrix} k \\ i \end{bmatrix}_q q^{\binom{i}{2}} t^i H_{k-i}(y|q) G_{0,i+l}(y,z,t) \\
&\quad + (-1)^m q^{\binom{m}{2}} \begin{bmatrix} k \\ m \end{bmatrix}_q t^m (H_{k-m}(y|q) G_{0,m+l}(y,z,t) \\
&\quad - \sum_{i=1}^{k-m} \begin{bmatrix} k-m \\ i \end{bmatrix}_q t^i G_{k-m-i,l+m+i}(y,z,t)) \\
&\quad + (-1)^m q^{\binom{m}{2}} \sum_{i=m+1}^k \begin{bmatrix} k \\ i \end{bmatrix}_q \begin{bmatrix} i-1 \\ m-1 \end{bmatrix}_q t^i G_{k-i,i+l}(y,z,t)
\end{aligned}$$

Now since $\begin{bmatrix} k \\ m \end{bmatrix}_q \begin{bmatrix} k-m \\ i-m \end{bmatrix}_q - \begin{bmatrix} k \\ i \end{bmatrix}_q \begin{bmatrix} i-1 \\ m-1 \end{bmatrix}_q = \begin{bmatrix} k \\ i \end{bmatrix}_q (\begin{bmatrix} i \\ m \end{bmatrix}_q - \begin{bmatrix} i-1 \\ m-1 \end{bmatrix}_q) = q^m \begin{bmatrix} k \\ i \end{bmatrix}_q \begin{bmatrix} i-1 \\ m \end{bmatrix}_q$ and $\binom{m}{2} + m = \binom{m+1}{2}$ we have

$$\begin{aligned}
G_{k,l}(y,z,t) &= \sum_{i=0}^m (-1)^i \begin{bmatrix} k \\ i \end{bmatrix}_q q^{\binom{i}{2}} t^i H_{k-i}(y|q) G_{0,i+l}(y,z,t) \\
&\quad - (-1)^m q^{\binom{m}{2}} \sum_{i=m+1}^k (\begin{bmatrix} k \\ m \end{bmatrix}_q \begin{bmatrix} k-m \\ i-m \end{bmatrix}_q - \begin{bmatrix} k \\ i \end{bmatrix}_q \begin{bmatrix} i-1 \\ m-1 \end{bmatrix}_q) t^i G_{k-i,i+l}(y,z,t) \\
&= \sum_{i=0}^m (-1)^i \begin{bmatrix} k \\ i \end{bmatrix}_q q^{\binom{i}{2}} t^i H_{k-i}(y|q) G_{0,i+l}(y,z,t) \\
&\quad + (-1)^{m+1} q^{\binom{m+1}{2}} \sum_{i=m+1}^k \begin{bmatrix} k \\ i \end{bmatrix}_q \begin{bmatrix} i-1 \\ m \end{bmatrix}_q t^i G_{k-i,i+l}(y,z,t)
\end{aligned}$$

iv) For $k = 0$ this is obviously true. Now let us iterate (4.2) once, applied however, for $G_{0,k}$. We will get then

$$\begin{aligned}
G_{k,0}(y,z,t|q) &= \sum_{i=0}^{k-1} (-1)^i \begin{bmatrix} k \\ i \end{bmatrix}_q q^{\binom{i}{2}} t^i H_{k-i}(y|q) G_{0,i}(y,z,t|q) + (-1)^k q^{\binom{k}{2}} t^k G_{0,k}(y,z,t|q) = \\
&= \sum_{i=0}^{k-1} (-1)^i \begin{bmatrix} k \\ i \end{bmatrix}_q q^{\binom{i}{2}} t^i H_{k-i}(y|q) G_{0,i}(y,z,t|q) + \\
&(-1)^k q^{\binom{k}{2}} t^k (\sum_{i=0}^{k-1} (-1)^i \begin{bmatrix} k \\ i \end{bmatrix}_q q^{\binom{i}{2}} t^i H_{k-i}(z|q) G_{i,0}(y,z,t|q)) + q^{k(k-1)} t^{2k} G_{k,0}(y,z,t|q).
\end{aligned}$$

Thus we see that since for all $i \leq k-1$ $G_{i,0}$ and $G_{0,i}$ and are of the claimed form then from (4.3) it follows that $G_{k,0}$ has the claimed form. \square

Now we are ready to present the proof of Theorem 1.

Proof of the Theorem 1. First notice that following (3.2) and assertion v of Lemma 1 we see that series $\sum_{n \geq 0} \frac{H_n(x|q) g_n(y,z,\rho_1,\rho_2,q)}{[n]_q!}$ converges absolutely. To prove i) we will use formula viii) of Lemma

1, that is Poisson-Mehler expansion formula. Following (3.1) we see that

$$\begin{aligned} \phi(x|y, z, \rho_1, \rho_2, q) &= f_N(x|q) \sum_{i=0}^{\infty} \frac{\rho_1^i}{[i]_q!} H_i(x|q) H_i(y|q) \times \sum_{i=0}^{\infty} \frac{\rho_2^i}{[i]_q!} H_i(x|q) H_i(z|q) \\ &\quad / \sum_{i=0}^{\infty} \frac{\rho_1^i \rho_2^i}{[i]_q!} H_i(y|q) H_i(z|q). \end{aligned}$$

First, let us concentrate on the quantity:

$$R(x, y, z, \rho_1, \rho_2|q) = \sum_{n \geq 0} \frac{\rho_1^n}{[n]_q!} H_n(x|q) H_n(y|q) \times \sum_{m \geq 0} \frac{\rho_2^m}{[m]_q!} H_m(x|q) H_m(z|q).$$

We will apply identity (2.4), distinguish two cases $n+m$ is even and $n+m$ is odd, denote $n+m-2j = 2k$ or $n+m-2j = 2k+1$ depending on the case and sum over the set of $\{(n, m) : n+m-2k \leq 2 \min(n, m), m, n \geq 0\} \cup \{(n, m) : n+m-2k-1 \leq 2 \min(n, m), m, n \geq 0\}$. We have

$$\begin{aligned} R(x, y, z, \rho_1, \rho_2|q) &= \sum_{n, m \geq 0} \frac{\rho_1^n \rho_2^m}{[n]_q! [m]_q!} H_n(y|q) H_m(z|q) \sum_{j=0}^{\min(n, m)} \begin{bmatrix} n \\ j \end{bmatrix}_q \begin{bmatrix} m \\ j \end{bmatrix}_q [j]_q! H_{n+m-2j}(x|q) \\ &= \sum_{k=0}^{\infty} \frac{H_{2k}(x|q)}{[2k]_q!} \sum_{j=0}^{\infty} \sum_{m=j}^{2k+j} \frac{[2k]_q! \rho_1^m \rho_2^{2k+2j-m}}{[m-j]_q! [2k-(m-j)]_q! [j]_q!} H_m(y|q) H_{2k+2j-m}(z|q) \\ &\quad + \sum_{k=0}^{\infty} \frac{H_{2k+1}(x|q)}{[2k+1]_q!} \sum_{j=0}^{\infty} \sum_{m=j}^{2k+1+j} \frac{[2k+1]_q! \rho_1^{i+j} \rho_2^{2k+1-i+j}}{[j]_q! [m-j]_q! [2k+1-(m-j)]_q!} H_m(y|q) H_{2k+1+2j-m}(z|q) \\ &= \sum_{k=0}^{\infty} \frac{H_{2k}(x|q)}{[2k]_q!} \sum_{j=0}^{\infty} \sum_{i=0}^{2k} \frac{[2k]_q! \rho_1^{i+j} \rho_2^{2k-i+j}}{[j]_q! [i]_q! [2k-i]_q!} H_{i+j}(y|q) H_{2k-i+j}(z|q) \\ &\quad + \sum_{k=0}^{\infty} \frac{H_{2k+1}(x|q)}{[2k+1]_q!} \sum_{j=0}^{\infty} \sum_{i=0}^{2k+1} \frac{[2k+1]_q! \rho_1^{i+j} \rho_2^{2k+1-i+j}}{[j]_q! [i]_q! [2k+1-i]_q!} H_{i+j}(y|q) H_{2k+1+j-i}(z|q) \end{aligned}$$

We get then

$$\begin{aligned} R(x, y, z, \rho_1, \rho_2|q) &= \sum_{k=0}^{\infty} \frac{H_{2k}(x|q)}{[2k]_q!} \sum_{i=0}^{2k} \begin{bmatrix} 2k \\ i \end{bmatrix}_q \rho_1^i \rho_2^{2k-i} \sum_{j=0}^{\infty} \frac{(\rho_1 \rho_2)^j}{[j]_q!} H_{i+j}(y|q) H_{2k-i+j}(z|q) \\ &\quad + \sum_{k=0}^{\infty} \frac{H_{2k+1}(x|q)}{[2k+1]_q!} \sum_{i=0}^{2k+1} \begin{bmatrix} 2k+1 \\ i \end{bmatrix}_q \rho_1^i \rho_2^{2k+1-i} \sum_{j=0}^{\infty} \frac{(\rho_1 \rho_2)^j}{[j]_q!} H_{i+j}(y|q) H_{2k+1-i+j}(z|q) \\ &= \sum_{n=0}^{\infty} \frac{H_n(x|q)}{[n]_q!} \sum_{i=0}^n \begin{bmatrix} n \\ i \end{bmatrix}_q \rho_1^i \rho_2^{n-i} \sum_{j=0}^{\infty} \frac{(\rho_1 \rho_2)^j}{[j]_q!} H_{i+j}(y|q) H_{n-i+j}(z|q) \end{aligned}$$

Using introduced in Proposition 7 function $G_{l,k}(y, z, t|q) = \sum_{m \geq 0} \frac{t^m}{[m]_q!} H_{m+l}(y|q) H_{m+k}(z|q)$ we can express both

$$\sum_{i=0}^{\infty} \frac{\rho_1^i \rho_2^i}{[i]_q!} H_i(y|q) H_i(z|q) = G_{0,0}(y, z, \rho_1 \rho_2|q) \text{ and}$$

$R(x, y, z, \rho_1, \rho_2 | q) = \sum_{n=0}^{\infty} \frac{H_n(x|q)}{[n]_q!} \sum_{i=0}^n [i]_q \rho_1^i \rho_2^{n-i} G_{i, n-i}(y, z, \rho_1 \rho_2 | q)$. Our Theorem will be proved if we will be able to show that $\forall l, k \geq 0 : G_{l, k}(y, z, t | q) = G_{0, 0}(y, z, t | q) \Theta_{l, k}(y, z, t | q)$ where $\Theta_{l, k}$ is a polynomial of order l in y k in z . This fact follows by induction from formula (4.3) of assertion iv) of the Proposition 7 since it expresses $G_{k, 0}$ in terms of k functions $G_{i, 0}$ and $G_{0, i}$ for $i = 0, \dots, k - 1$ and the fact that all $G_{l, k}$ can be expressed by $G_{i, 0}$ and $G_{0, i}$; $i \leq k + l$. \square

Proof of Corollary 5. By Theorem 1 we know that regression

$\mathbb{E}(H_n(X_i | q) | X_{j_1}, \dots, X_{j_k}, X_{j_m}, \dots, X_{j_n})$ is a polynomial in X_{j_k} and X_{j_m} of order at most n . To analyze the structure of this polynomial let us present it in the form $\sum_{s=0}^n a_{s, n} H_s(X_{j_m} | q)$ where coefficients $a_{s, n}$ are some polynomials of X_{j_k} . Now let us take conditional expectation with respect to X_{j_1}, \dots, X_{j_k} of both sides. On one hand we get

$$\mathbb{E}(H_n(X_i | q) | X_{j_1}, \dots, X_{j_k}) = \left(\prod_{m=j_k}^{i-1} \sigma_m \right)^n H_n(X_{j_k})$$

on the other we get

$$\sum_{s=0}^n a_{s, n} \left(\prod_{m=j_k}^{j_m-1} \sigma_m \right)^s H_s(X_{j_k} | q).$$

Since $a_{s, n}$ are polynomials in X_{j_k} of order at most n , we can present them in the form

$$a_{s, n} = \sum_{t=0}^n \beta_{t, s} H_t(x_{j_k} | q).$$

Thus we have equality:

$$\left(\prod_{m=j_k}^{i-1} \sigma_m \right)^n H_n(x_{j_k}) = \sum_{s=0}^n \left(\prod_{m=j_k}^{j_m-1} \sigma_m \right)^s \sum_{t=0}^n \beta_{t, s} H_t(x_{j_k}) H_s(x_{j_k}).$$

Now we use the identity (2.4) and get

$$\begin{aligned} \left(\prod_{m=j_k}^{i-1} \sigma_m \right)^n H_n(x_{j_k} | q) &= \sum_{s=0}^n \sum_{t=0}^n \beta_{t, s} \left(\prod_{m=j_k}^{j_m-1} \sigma_m \right)^s \\ &\times \sum_{m=0}^{\min(t, s)} \begin{bmatrix} t \\ m \end{bmatrix}_q \begin{bmatrix} s \\ m \end{bmatrix}_q [m]_q! H_{t+s-2m}(x_{j_k} | q). \end{aligned}$$

Hence we deduce that $\beta_{t, s} = 0$ for $t + s > n$, $t + s = n - 1, n - 3, \dots$. To count the number of coefficients $A_{j, k}^{(n)}$ observe that we have $n + 1$ coefficients $A_{0, k}^{(n)}$ since k ranges from $-\lfloor \frac{n}{2} \rfloor$ to $-\lfloor \frac{n}{2} \rfloor + n$, $n - 1$ coefficients $A_{1, k}^{(n)}$ where k ranges from $-\lfloor n/2 \rfloor + 1$ to $-\lfloor n/2 \rfloor + n - 1$ and so on. \square

Proof of Corollary 6. The proof is based on the idea of writing down system of $\lfloor \frac{n+2}{2} \rfloor \lfloor \frac{n+3}{2} \rfloor$ ($n = 1, \dots, 4$) linear equations satisfied by coefficients $A_{m, k}^{(n)}$. These equations are obtained according

to the similar pattern. Namely we multiply both sides of identity (3.3) by $H_m(X_{i-1})$ and $H_k(X_i)$ and calculate conditional expectation of both sides with respect to $\mathcal{F}_{<i}$ or with respect to $\mathcal{F}_{>i}$ remembering that $\mathbb{E}(H_m(X_{i+1})|\mathcal{F}_{<i}) = \rho_{i-1}^m \rho_i^m H_m(X_{i-1})$ and $\mathbb{E}(H_m(X_i)|\mathcal{F}_{<i}) = \rho_{i-1}^m H_m(X_{i-1})$ and similar formulae for $\mathcal{F}_{>i}$. We expand both sides with respect to H_s , $s = n + m + k - 2t$, $t = 0, \dots, \lfloor (n + m + k)/2 \rfloor$. On the way we utilize (2.4) and compare coefficients standing by H_s on both sides. Thus each obtained equation involving coefficients $A_{i,j}^{(n)}$ can be indexed by s, m, j and r if we calculate conditional expectation with respect to $\mathcal{F}_{<i}$ of l if we conditional expectation is calculated with respect to $\mathcal{F}_{>i}$. Of course if $s = 0$ then r and l lead to the same result. Formulae for $A_{i,j}^{(n)}$, for $n = 1$ are obtained by taking $s = 1$, $m = 0$, $j = 0$ and applying r and l . For $n = 2$ first we consider $m = 0$, $j = 0$ and $s = 2$ and applying r and l and then $m = 0$, $j = 0$ and $s = 0$. In this way we get 3 equations. The fourth one is obtained by taking $m = 0$, $j = 1$, $s = 1$ and r . Denote $\mathbf{X} = (A_{0,-1}^{(2)}, A_{0,0}^{(2)}, A_{0,1}^{(2)}, A_{1,0}^{(2)})^T$, then \mathbf{X} satisfies system of linear equation with matrix

$$\begin{pmatrix} 1 & \rho_{i-1}\rho_i & \rho_{i-1}^2\rho_i^2 & 0 \\ \rho_{i-1}^2\rho_i^2 & \rho_{i-1}\rho_i & 1 & 0 \\ 0 & \rho_{i-1}\rho_i & 0 & 1 \\ [2]_q \rho_{i-1}\rho_i & 1 + [2]_q \rho_{i-1}^2\rho_i^2 & [2]_q \rho_{i-1}\rho_i & \rho_{i-1}\rho_i \end{pmatrix}$$

and with right side vector equal to : $\begin{pmatrix} \rho_{i-1}^2 \\ \rho_i^2 \\ 0 \\ [2]_q \rho_{i-1}\rho_i \end{pmatrix}$. Besides formulae for coefficients $A_{m,k}^{(n)}$,

for $n = 1, 2$ can be obtained from formulae scattered in the literature like e.g. [1], [12] or [19]. To get equations satisfied by coefficients $A_{i,j}^{(n)}$, for $n = 3, 4$ First $n + 1$ equations are obtained by taking $m = 0$, $k = 0$ and $s = 3, 1$ if $n = 3$ and $s = 4, 2, 0$ if $n = 4$ and then applying operations r and l . Then, in order to get remaining 2 (in case of $n = 3$) or 4 (in case of $n = 4$) equations one has to be more careful since it often turns out that many equations obtained for some m and k are linearly dependent on the previously obtained equations. In the case of $n = 3$ to get remaining two linearly independent equations we took $m = 2$, $k = 0$, $s = 3$ and applied operations r and l . In this way we obtained system of 6 linear equations with matrix

$$\begin{pmatrix} 1 & \rho_{i-1}\rho_i & \rho_{i-1}^2\rho_i^2 & \rho_{i-1}^3\rho_i^3 & 0 & 0 \\ \rho_{i-1}^3\rho_i^3 & \rho_{i-1}^2\rho_i^2 & \rho_{i-1}\rho_i & 1 & 0 & 0 \\ 0 & (1+q)\rho_{i-1}\rho_i & (1+q)\rho_{i-1}^2\rho_i^2 & 0 & 1 & \rho_{i-1}\rho_i \\ 0 & (1+q)\rho_{i-1}^2\rho_i^2 & (1+q)\rho_{i-1}\rho_i & 0 & \rho_{i-1}\rho_i & 1 \\ [3]_q \rho_{i-1}\rho_i & 1 + [2]_q^2 \rho_{i-1}^2\rho_i^2 & [2]_q \rho_{i-1}\rho_i + [3]_q \rho_{i-1}^3\rho_i^3 & [3]_q \rho_{i-1}^2\rho_i^2 & \rho_{i-1}\rho_i & \rho_{i-1}^2\rho_i^2 \\ [3]_q \rho_{i-1}^2\rho_i^2 & [2]_q \rho_{i-1}\rho_i + [3]_q \rho_{i-1}^3\rho_i^3 & 1 + [2]_q^2 \rho_{i-1}^2\rho_i^2 & [3]_q \rho_{i-1}\rho_i & \rho_{i-1}^2\rho_i^2 & \rho_{i-1}\rho_i \end{pmatrix}$$

right hand side vector $\begin{pmatrix} \rho_{i-1}^3 \\ \rho_i^3 \\ 0 \\ 0 \\ [3]_q \rho_{i-1}^2\rho_i^2 \\ [3]_q \rho_{i-1}\rho_i \end{pmatrix}$ if the vector of unknowns is the following

$(A_{0,-1}^{(3)}, \dots, A_{0,2}^{(2)}, A_{1,0}^{(3)}, A_{1,1}^{(3)})^T$. For $n = 4$ remaining 4 equations we obtained by taking: $(m = 1, k = 4, s = 3, r)$, $(m = 4, k = 1, s = 1, r)$, $(m = 4, k = 2, s = 4, r)$ and $(m = 2, k = 4, s = 4, r)$. Recall that in this case we have 9 equations. Matrix of this system has 81 entries. That is why we will skip writing down the whole system of equations. To get the scent of how complicated these equations are we

will present one equation. For $n = 4$ one of the equations (referring to the case $m = 4, k = 1, s = 1, r$) is $[2]_q^2 (1 + q^2) [3]_q ([4]_q + [5]_q) \rho_{i-1} \rho_i A_{0,-2}^{(4)} + [2]_q^2 (1 + q^2) [3]_q \times (1 + \rho_{i-1}^2 \rho_i^2 + 3q \rho_{i-1}^2 \rho_i^2 + 3q^2 \rho_{i-1}^2 \rho_i^2 + q^3 \rho_{i-1}^2 \rho_i^2 + q^4 \rho_{i-1}^2 \rho_i^2) A_{0,-1}^{(4)} + [2]_q^2 (1 + q^2) [3]_q \rho_{i-1} \rho_i ([2]_q + \rho_{i-1}^2 \rho_i^2 + 3q \rho_{i-1}^2 \rho_i^2 + 3q^2 \rho_{i-1}^2 \rho_i^2 + q^3 \rho_{i-1}^2 \rho_i^2 + q^4 \rho_{i-1}^2 \rho_i^2) A_{0,0}^{(4)} + [2]_q^2 (1 + q^2) [3]_q \rho_{i-1}^2 \rho_i^2 ([3]_q + ([4]_q + [5]_q) \rho_{i-1}^2 \rho_i^2) A_{0,1}^{(4)} + [2]_q^2 (1 + q^2) [3]_q \rho_{i-1}^3 \rho_i^2 (1 + q + q^2 + q^3 + \rho_{i-1}^2 \rho_i^2 + q \rho_{i-1}^2 \rho_i^2 + q^2 \rho_{i-1}^2 \rho_i^2 + q^3 \rho_{i-1}^2 \rho_i^2 + q^4 \rho_{i-1}^2 \rho_i^2) A_{0,2}^{(4)} + [2]_q^2 (1 + q^2) [3]_q \rho_{i-1} \rho_i A_{1,-1}^{(4)} + [2]_q^2 (1 + q^2) [3]_q \rho_{i-1}^2 \rho_i^2 A_{1,0}^{(4)} + [2]_q^2 (1 + q^2) [3]_q \rho_{i-1}^3 \rho_i^3 A_{1,1}^{(4)} = [2]_q^2 (1 + q^2) [3]_q \rho_{i-1}^3 ([4]_q + [5]_q \rho_{i-1}^2) \rho_i. \quad \square$

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