

Large deviations for the volume of k -nearest neighbor balls*

Christian Hirsch[†] Taegyu Kang[‡] Takashi Owada[§]

Abstract

This paper develops the large deviations theory for the point process associated with the Euclidean volume of k -nearest neighbor balls centered around the points of a homogeneous Poisson or a binomial point processes in the unit cube. Two different types of large deviation behaviors of such point processes are investigated. Our first result is the Donsker-Varadhan large deviation principle, under the assumption that the centering terms for the volume of k -nearest neighbor balls grow to infinity more slowly than those needed for Poisson convergence. Additionally, we also study large deviations based on the notion of \mathcal{M}_0 -topology, which takes place when the centering terms tend to infinity sufficiently fast, compared to those for Poisson convergence. As applications of our main theorems, we discuss large deviations for the number of Poisson or binomial points of degree at most k in a random geometric graph in the dense regime.

Keywords: large deviation principle; \mathcal{M}_0 -convergence; k -nearest neighbor ball; point process; stochastic geometry.

MSC2020 subject classifications: Primary 60F10, Secondary 60D05; 60G55.

Submitted to EJP on October 24, 2022, final version accepted on May 24, 2023.

1 Introduction

The main theme of this paper is to develop the large deviations theory for the point process associated with the Euclidean volume of k -nearest neighbor balls. We consider the unit cube $[0, 1]^d$ equipped with the toroidal metric

$$\text{dist}(x, y) = \min_{z \in \mathbb{Z}^d} \|x - y + z\|,$$

*CH would like to acknowledge the financial support of the CogniGron research center and the Ubbo Emmius Funds (Univ. of Groningen). TO's research was partially supported by the AFOSR grant FA9550-22-0238 and the NSF grant DMS-1811428.

[†]Department of Mathematics, Aarhus University, Ny Munkegade, 118, 8000, Aarhus C, Denmark.
E-mail: hirsch@math.au.dk

[‡]Department of Statistics, Purdue University, West Lafayette, 47907, USA. E-mail: kang426@purdue.edu

[§]Department of Statistics, Purdue University, West Lafayette, 47907, USA. E-mail: owada@purdue.edu

where $\|\cdot\|$ denotes the Euclidean metric in \mathbb{R}^d . Then, the unit cube $[0, 1]^d$ is considered as a flat torus with a periodic boundary. Let $B_r(x) = \{y \in [0, 1]^d : \text{dist}(x, y) \leq r\}$, $r > 0$, be the closed ball in $[0, 1]^d$ of radius r centered at $x \in [0, 1]^d$. Given a point measure ω in $[0, 1]^d$ and a point $x \in [0, 1]^d$, and a fixed integer $k \geq 1$, define

$$R_k(x, \omega) := \inf \{r > 0 : \omega(B_r(x) \setminus \{x\}) \geq k\} \tag{1.1}$$

to be the k -nearest neighbor distance of x ; it gives a radius r for which $B_r(x) \setminus \{x\}$ contains exactly k points of ω with one of those points lying on the boundary of $B_r(x)$.

Given a homogeneous Poisson point process \mathcal{P}_n on $[0, 1]^d$ with intensity n , we are interested in the stochastic behavior of the point process

$$L_{k,n} := \begin{cases} \sum_{X \in \mathcal{P}_n} \delta_{(X, n\kappa_d R_k(X, \mathcal{P}_n)^d - a_n)} & \text{if } |\mathcal{P}_n| > k, \\ \emptyset & \text{if } |\mathcal{P}_n| \leq k, \end{cases} \tag{1.2}$$

where $\delta_{(x,y)}$ is the Dirac measure at $(x, y) \in [0, 1]^d \times \mathbb{R}$, and κ_d is volume of the unit ball in \mathbb{R}^d , so that $\kappa_d R_k(X, \mathcal{P}_n)^d$ represents the volume of a k -nearest neighbor ball centered at $X \in \mathcal{P}_n$. Further, \emptyset represents the null measure, i.e., the measure assigning zeros to all Borel measurable sets. The process (1.2) dictates the magnitude of the normalized volume of k -nearest neighbor balls, as well as the location of their centers.

There have been a number of studies on the asymptotics of the process (1.2) or its similar variant, when the centering term a_n is given as

$$a_n = \log n + (k - 1) \log \log n + \text{constant}. \tag{1.3}$$

In this case, one may observe, asymptotically, at most *finitely many* k -nearest neighbor balls whose volume are approximately a_n/n up to the scale. As a consequence, the process (1.2) will have a Poissonian structure in the limit [16, 5, 2, 14, 1]. In particular, the rate of Poisson convergence has recently been derived in terms of the Kantorovich-Rubinstein distance [1] and the total variation distance [14].

In addition to these results on Poisson convergence, there have also been many attempts at deriving other limit theorems for the functional of a k -nearest neighbor distance in (1.1), among them, central limit theorems in [17, 19] and laws of large numbers in [22, 20, 23]. For example, Penrose [17] proved a (functional) central limit theorem for the number of inhomogeneous Poisson points X_i with density f , such that $f(X_i) R_{k_n}(X_i, \mathcal{P}_n)^d$ does not exceed certain thresholds (k_n is taken to be a function of n). Additionally, Penrose and Yukich [23] provided laws of large numbers for the sum of power weighted nearest neighbor distances. As for the large deviation results on the k -nearest neighbor distance, Schreiber and Yukich [26] obtained a *Donsker-Varadhan large deviation principle* (LDP) (see Section 1.2 in [3] for a precise definition) for the functional of the length of edges in the k -nearest neighbor graph in \mathbb{R}^d . This was obtained as an application of a more general LDP for spatial point processes satisfying a weak dependence condition characterized by a radius of stabilization. Moreover, Hirsch et al. [7] analyzed lower tail large deviations for general geometric functionals, including the power-weighted edge lengths in the k -nearest neighbor graph.

The primary objective of this paper is to provide comprehensive results on the asymptotics of the process (1.2), from the viewpoints of large deviations. We consider two distinct scenarios with respect to a divergence speed of the centering term (a_n). The first scenario examined in this paper is that

$$a_n \rightarrow \infty, \quad a_n - \log n - (k - 1) \log \log n \rightarrow -\infty, \quad \text{as } n \rightarrow \infty. \tag{1.4}$$

In this case, (a_n) grows to infinity more slowly than (1.3). Intrinsically, there appear *infinitely many* k -nearest neighbor balls as $n \rightarrow \infty$, whose volume are approximately

a_n/n up to the scale. Then, the process (1.2) “diverges” in the limit, in the sense that $L_{k,n}(A) \rightarrow \infty$ as $n \rightarrow \infty$, for all (nice) measurable sets A . Thus, in order to dictate its large deviation behavior, one has to scale the processes in (1.2) by some growing sequence $u_n \rightarrow \infty$. More concretely, we aim to establish the Donsker-Varadhan LDP for the properly scaled process $(L_{k,n}/u_n)_{n \geq 1}$.

In the second scenario of this paper, we consider the centering term (a_n) satisfying

$$a_n - \log n - (k - 1) \log \log n \rightarrow \infty, \quad a_n = o(n), \quad \text{as } n \rightarrow \infty. \quad (1.5)$$

Then, (a_n) tends to infinity more rapidly than in (1.3), so that the k -nearest neighbor balls centered around \mathcal{P}_n , whose volume are approximately a_n/n up to the scale, are even less likely to occur. In other words, the occurrence of such k -nearest neighbor balls is a “rare event”, in the sense of $\mathbb{P}(L_{k,n}(A) \geq 1) \rightarrow 0$ as $n \rightarrow \infty$, for all measurable sets A . In this setup, we present the other type of large deviation results, by detecting a sequence $v_n \rightarrow \infty$, such that

$$(v_n \mathbb{P}(L_{k,n} \in \cdot), n \geq 1) \quad (1.6)$$

converges to a (non-trivial) limit measure. The topology underlying the convergence of (1.6) is \mathcal{M}_0 -topology. The notion of \mathcal{M}_0 -topology was first developed by [9]. Since then, it has been used mainly for the study of regular variation of stochastic processes [10, 13, 4, 27, 15].

For the required LDP in Theorem 2.1 below, many of the techniques in our previous work [8] will be exploited. We first partition the unit cube $[0, 1]^d$ into smaller cubes of equal volume, and define a collection of i.i.d. point processes restricted to each of the small cubes. Next, using one of the main results in [1], Proposition 4.1 proves that the law of these point processes restricted to small cubes converges to the law of certain Poisson point processes, in terms of the Kantorovich-Rubinstein distance. Subsequently, Proposition 4.3 justifies that this approximation is still feasible even for the version of the empirical measures. The main machinery here is the notion of a maximal coupling, provided in [11, Lemma 4.32]. Other approximation arguments necessary for our proof will be completed in a series of results in Propositions 4.4–4.6. As a final note, we want to emphasize that the homogeneity assumption of \mathcal{P}_n is crucial throughout our proof. We anticipate that the LDP still holds even when \mathcal{P}_n in (1.2) is replaced by an inhomogeneous Poisson point process. It seems, however, that unlike the previous studies [16, 2, 14, 1], this extension should require much more involved machinery; this will be left as a topic of future research.

The rest of the paper is outlined as follows. In Section 2, under the assumption (1.4), we give a precise setup for the point process (1.2) and formalize the desired LDP. Section 3 assumes condition (1.5) and establishes the \mathcal{M}_0 -convergence for the sequence (1.6). In both settings, we also consider the case that the point processes are generated by a binomial point process. Unfortunately, in the context of LDPs, there are no unified results on the De-Poissonization scheme, such as [18, Section 2.5], which may allow us to extend the LDPs with a Poisson input to those with a binomial input. Alternatively, we have proved directly the desired exponential equivalence (in terms of the total variation distance) between processes with a Poisson input and those with a binomial input (see Corollary 2.2). Corollary 3.2 applies a similar kind of De-Poissonization machinery to the \mathcal{M}_0 -convergence in Theorem 3.1. Finally, as an application of Theorem 2.1 and Corollary 2.2, we deduce the LDP for the number of Poisson (or binomial) points of degree at most k in a random geometric graph in the dense regime. Additional applications in a similar vein from the perspective of \mathcal{M}_0 -convergence, can be found in Corollary 3.3.

2 Large deviation principle

Given a point measure ω in $[0, 1]^d$ and $x \in [0, 1]^d$, define the k -nearest neighbor distance $R_k(x, \omega)$ as in (1.1). Let $(a_n)_{n \geq 1}$ be a sequence tending to infinity, such that

$$a_n - \log n - (k - 1) \log \log n \rightarrow -\infty, \quad n \rightarrow \infty. \tag{2.1}$$

Given a homogeneous Poisson point process \mathcal{P}_n on $[0, 1]^d$ with intensity n , define the point process (1.2) on the space $M_p(E_0)$ of point measures on $E_0 := [0, 1]^d \times (s_0, \infty]$, where $s_0 \in \mathbb{R}$ is a fixed real number. Our aim is to explore the large deviation behavior of $(L_{k,n})_{n \geq 1}$. More specifically, we define the scaling constants

$$b_n = na_n^{k-1} e^{-a_n}, \quad n \geq 1, \tag{2.2}$$

and establish the LDP for $(L_{k,n}/b_n)_{n \geq 1}$ in the space $M_+(E_0)$ of finite non-negative Radon measures on E_0 . We will equip $M_+(E_0)$ with the weak topology; then $M_+(E_0)$ becomes a Polish space (See, e.g., [11, Lemma 4.5]). Under the assumption (2.1), it is elementary to show that $b_n \rightarrow \infty$ as $n \rightarrow \infty$.

To state our LDP more precisely, we introduce the measure

$$\tau_k(du) := \frac{e^{-u}}{(k-1)!} \mathbb{1}\{u \geq s_0\} du.$$

Here, we provide two equivalent representations of the rate function. The first is based on the relative entropy. More precisely, writing $\text{Leb} \otimes \tau_k$ for the product measure of the Lebesgue measure on $[0, 1]^d$ and τ_k , define the relative entropy of $\rho \in M_+(E_0)$ with respect to $\text{Leb} \otimes \tau_k$:

$$H_k(\rho | \text{Leb} \otimes \tau_k) := \int_{E_0} \log \left\{ \frac{d\rho(x, u)}{d(\text{Leb} \otimes \tau_k)} \right\} \rho(dx, du) - \rho(E_0) + (\text{Leb} \otimes \tau_k)(E_0), \quad \text{if } \rho \ll \text{Leb} \otimes \tau_k, \tag{2.3}$$

and $H_k(\rho | \text{Leb} \otimes \tau_k) = \infty$ otherwise. The second representation of the rate function is given in terms of the Legendre transform. More concretely,

$$\Lambda_k^*(\rho) := \sup_{f \in C_b(E_0)} \left\{ \int_{E_0} f(x, u) \rho(dx, du) - \int_{E_0} (e^{f(x, u)} - 1) dx \tau_k(du) \right\}, \quad \rho \in M_+(E_0),$$

where $C_b(E_0)$ is the space of continuous and bounded real-valued functions on E_0 .

The proofs of the results in this section are all deferred to Section 4.1.

Theorem 2.1. *The sequence $(L_{k,n}/b_n)_{n \geq 1}$ satisfies an LDP on $M_+(E_0)$ in the weak topology, with speed b_n and rate function $H_k(\cdot | \text{Leb} \otimes \tau_k) = \Lambda_k^*$.*

The corollary below extends Theorem 2.1 to the case that the process is generated by a binomial point process $\mathcal{B}_n := \{X_1, \dots, X_n\}$. Precisely, we define, in the space $M_p(E_0)$,

$$L_{k,n}^{\mathbb{B}} := \sum_{X \in \mathcal{B}_n} \delta_{(X, n\kappa_d R_k(X, \mathcal{B}_n)^d - a_n)}, \quad n > k,$$

and $L_{k,n}^{\mathbb{B}} \equiv \emptyset$ for $n \leq k$.

Corollary 2.2. *The sequence $(L_{k,n}^{\mathbb{B}}/b_n)_{n \geq 1}$ satisfies an LDP on $M_+(E_0)$ in the weak topology, with speed b_n and rate function $H_k(\cdot | \text{Leb} \otimes \tau_k) = \Lambda_k^*$.*

As an application of Theorem 2.1, we deduce the LDP for $(T_{k,n}/b_n)_{n \geq 1}$, where

$$T_{k,n} := \sum_{X \in \mathcal{P}_n} \mathbb{1}\left\{ \mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k \right\}, \quad n \geq 1. \tag{2.4}$$

The statistics (2.4) represents the number of Poisson points of degree at most k in a random geometric graph of vertex set \mathcal{P}_n and edges between X_i and X_j satisfying $\|X_i - X_j\| \leq r_n(s_0)$, where

$$r_n(s_0) := \left(\frac{a_n + s_0}{n\kappa_d} \right)^{1/d}. \tag{2.5}$$

The threshold radius in (2.5) ensures that the random geometric graph under consideration is of the *dense regime*, such that $nr_n(s_0)^d = (a_n + s_0)/\kappa_d \rightarrow \infty$ as $n \rightarrow \infty$. Replacing \mathcal{P}_n in (2.4) with its binomial counterpart \mathcal{B}_n , we also derive the LDP for $(T_{k,n}^B/b_n)_{n \geq 1}$, where

$$T_{k,n}^B := \sum_{X \in \mathcal{B}_n} \mathbb{1} \left\{ \mathcal{B}_n(B_{r_n(s_0)}(X)) \leq k \right\}, \quad n \geq 1.$$

Corollary 2.3. *The sequence $(T_{k,n}/b_n)_{n \geq 1}$ satisfies an LDP with speed b_n and rate function*

$$I_k(x) = \begin{cases} x \log(x/\alpha_k) - x + \alpha_k & \text{if } x \geq 0, \\ \infty & \text{if } x < 0, \end{cases}$$

where $\alpha_k = e^{-s_0}/(k-1)!$. Furthermore, $(T_{k,n}^B/b_n)_{n \geq 1}$ satisfies the same LDP as $(T_{k,n}/b_n)_{n \geq 1}$.

In the above, $I_k(x)$ coincides with a rate function in the LDP for $(n^{-1} \sum_{i=1}^n Y_i)_{n \geq 1}$ where the Y_i are i.i.d. Poisson with mean α_k . Such a coincidence happens because the Poisson approximation technique applied to the process $L_{k,n}$ indicates that the process $T_{k,n}$ can be approximated by the i.i.d. sum of Poisson random variables as well.

3 Large deviation under \mathcal{M}_0 -topology

In this section, we explore the large deviation behavior of the process (1.2), in the case that a_n tends to infinity more rapidly than in the last section. Namely, we assume that $(a_n)_{n \geq 1}$ satisfies

$$a_n - \log n - (k-1) \log \log n \rightarrow \infty, \quad n \rightarrow \infty.$$

We again introduce the sequence $b_n = na_n^{k-1}e^{-a_n}$ as in (2.2). However, unlike in the last section, $b_n \rightarrow 0$ as $n \rightarrow \infty$, because

$$\begin{aligned} b_n &= e^{-(a_n - \log n - (k-1) \log \log n)} \left(\frac{a_n}{\log n} \right)^{k-1} \\ &\leq C e^{-(a_n - \log n - (k-1) \log \log n)} (a_n - \log n - (k-1) \log \log n)^{k-1} \rightarrow 0, \end{aligned}$$

for some $C > 0$.

Our objective is to investigate large deviations for the sequence $(\mathbb{P} \circ L_{k,n}^{-1})_{n \geq 1}$ of probability distributions of $(L_{k,n})_{n \geq 1}$ on the space $M_p(E)$, where $E := [0, 1]^d \times (-\infty, \infty]$. Note that the vague topology on $M_p(E)$ is metrizable as a complete, separable metric space. The metric that induces the vague topology is called vague metric, and its explicit form is given in the proof of [24, Proposition 3.17]. A main challenge is that the space $M_p(E)$ is not locally compact, and therefore, the vague topology would no longer be applicable for the convergence of such probability distributions. To overcome this difficulty, we adopt the notion of \mathcal{M}_0 -topology. The main feature of \mathcal{M}_0 -topology is that the corresponding test functions are continuous and bounded real-valued functions on $M_p(E)$ that vanish in the neighborhood of the origin. For the space $M_p(E)$, one can take the null measure \emptyset as its origin. Let $B_{\emptyset,r}$ denote an open ball of radius $r > 0$ centered at \emptyset in the vague metric. Denote by $\mathcal{M}_0 = \mathcal{M}_0(M_p(E))$ the space of Borel measures on $M_p(E)$, the restriction of which to $M_p(E) \setminus B_{\emptyset,r}$ is finite for all $r > 0$.

Moreover, define $\mathcal{C}_0 = \mathcal{C}_0(M_p(E))$ to be the space of continuous and bounded real-valued functions on $M_p(E)$ that vanish in a neighborhood of \emptyset , i.e., a continuous and bounded real-valued function $f : M_p(E) \rightarrow \mathbb{R}$ is a member of \mathcal{C}_0 if and only if there exists a positive real number r such that $f(\xi) = 0$ for every $\xi \in B_{\emptyset,r}$. Given $\xi_n, \xi \in \mathcal{M}_0$, we say that ξ_n converges to ξ in the \mathcal{M}_0 -topology, denoted as $\xi_n \rightarrow \xi$ in \mathcal{M}_0 , if it holds that $\int_{M_p(E)} g(\eta) \xi_n(d\eta) \rightarrow \int_{M_p(E)} g(\eta) \xi(d\eta)$ for all $g \in \mathcal{C}_0$. For more information on \mathcal{M}_0 -topology we refer to [9].

Before stating the main theorem, we will impose an additional condition that $a_n = o(n)$ as $n \rightarrow \infty$. To see the necessity of this assumption, suppose, to the contrary, that $a_n/n \rightarrow \infty$ as $n \rightarrow \infty$. Then, it trivially holds that $L_{k,n}([0,1]^d \times (-M, \infty]) = 0$ a.s. for large enough n and any $M > 0$. By putting the assumption $a_n = o(n)$ as above, one can exclude such triviality.

The proofs of the results below are all given in Section 4.2.

Theorem 3.1. *In the above setting with $a_n = o(n)$, as $n \rightarrow \infty$,*

$$b_n^{-1} \mathbb{P}(L_{k,n} \in \cdot) \rightarrow \xi_k, \quad \text{in } \mathcal{M}_0, \tag{3.1}$$

where (b_n) is given as in (2.2) and

$$\xi_k(\cdot) := \frac{1}{(k-1)!} \int_E \mathbb{1}\{\delta_{(x,u)} \in \cdot\} e^{-u} dx du.$$

Additionally, we present an analogous result for the process $(L_{k,n}^B)_{n \geq 1}$ as well. For a precise statement, however, we need to put a more stringent condition on (a_n) for the purpose of proving (4.45) in Section 4.2.

Corollary 3.2. *In the above setting with $a_n = o(n^{1/3})$, as $n \rightarrow \infty$,*

$$b_n^{-1} \mathbb{P}(L_{k,n}^B \in \cdot) \rightarrow \xi_k, \quad \text{in } \mathcal{M}_0.$$

Finally, certain asymptotic results on $(T_{k,n})_{n \geq 1}$ in (2.4), as well as those on $(T_{k,n}^B)_{n \geq 1}$, are presented as a corollary of the above results. This corollary gives the exact rate (up to the scale) of a probability that the number of Poisson (or binomial) points of degree at most k becomes non-zero.

Corollary 3.3. *(i) If $a_n = o(n)$, then*

$$b_n^{-1} \mathbb{P}(T_{k,n} \geq 1) \rightarrow \alpha_k \quad n \rightarrow \infty,$$

where α_k is given in Corollary 2.3.

(ii) If $a_n = o(n^{1/3})$, then

$$b_n^{-1} \mathbb{P}(T_{k,n}^B \geq 1) \rightarrow \alpha_k, \quad n \rightarrow \infty.$$

4 Proofs

4.1 Proofs of Theorem 2.1, Corollary 2.2, and Corollary 2.3

First, let us generalize the radius (2.5) by

$$r_n(u) := \left(\frac{a_n + u}{n\kappa_d} \right)^{1/d}, \quad u \in \mathbb{R}.$$

For $x \in [0,1]^d$ and $\omega \in M_p([0,1]^d)$, define the following functions:

$$f(x, \omega) := n\theta_d R_k(x, \omega)^d - a_n, \quad \text{and} \quad g(x, \omega) := \mathbb{1}\{f(x, \omega) > s_0\}.$$

Then, the process (1.2) that is restricted here to the space $M_p(E_0)$, can be reformulated as

$$L_{k,n} = \begin{cases} \sum_{X \in \mathcal{P}_n} g(X, \mathcal{P}_n) \delta_{(X, f(X, \mathcal{P}_n))} & \text{if } |\mathcal{P}_n| > k, \\ \emptyset & \text{if } |\mathcal{P}_n| \leq k. \end{cases} \quad (4.1)$$

Next, the unit cube $[0, 1]^d$ is partitioned into smaller cubes Q_1, \dots, Q_{b_n} , so that $\text{Leb}(Q_\ell) = b_n^{-1}$ for all $\ell \in \{1, \dots, b_n\}$. Here, it is assumed, without loss of generality, that b_n takes only positive integers for all $n \geq 1$. To avoid unnecessary technicalities, we will put the same assumption on many of the sequences and functions throughout the proof. Fix a sequence $w_n \rightarrow \infty$ with $w_n = o(a_n)$ as $n \rightarrow \infty$. For each $\ell \in \{1, \dots, b_n\}$, define the “boundary part” of Q_ℓ by

$$M_\ell := \{x \in Q_\ell : \inf_{y \in \partial Q_\ell} \|x - y\| \leq r_n(w_n)\},$$

while the “internal region” of Q_ℓ is given as $K_\ell := Q_\ell \setminus M_\ell$.

We now consider the process

$$\eta_{k,n} := \sum_{\ell=1}^{b_n} \sum_{X \in \mathcal{P}_n|_{K_\ell}} g(X, \mathcal{P}_n|_{Q_\ell}) \delta_{(X, f(X, \mathcal{P}_n|_{Q_\ell}))} \in M_p(E_0), \quad (4.2)$$

where $\mathcal{P}_n|_{K_\ell}$ (resp. $\mathcal{P}_n|_{Q_\ell}$) represents the Poisson point process restricted to K_ℓ (resp. Q_ℓ). Setting up a “blocked” point process as in (4.2) is a standard approach in the literature (see, e.g., [28, 26]). Clearly, the process (4.2) is different from (1.2). For example, (4.2) removes all the points $X \in \mathcal{P}_n$ lying within the distance of $r_n(w_n)$ from the boundary of $(Q_\ell)_{\ell=1}^{b_n}$, whereas those points are possibly counted by (1.2). Even when the center $X \in \mathcal{P}_n$ is chosen from the inside of K_ℓ , i.e., $X \in \mathcal{P}_n|_{K_\ell}$, the processes (1.2) and (4.2) may exhibit different k -nearest neighbor balls, whenever $R_k(X, \mathcal{P}_n) \neq R_k(X, \mathcal{P}_n|_{Q_\ell})$. Despite such differences, it is justified later in Propositions 4.4 – 4.6 that the process (4.2) can be used to approximate the large deviation behavior of (1.2).

For later analyses, it is convenient to express (4.2) as a superposition of i.i.d. point processes on E_0 , which themselves are transformed by some homeomorphisms. For each $\ell \in \{1, \dots, b_n\}$, let $\gamma_{\ell,n} : [0, 1]^d \rightarrow Q_\ell$ be the homeomorphism defined by $\gamma_{\ell,n}(x) = b_n^{-1/d}x + z_{\ell,n}$, where $z_{\ell,n}$ is a lower-left corner of Q_ℓ . Further, define $\tilde{\gamma}_{\ell,n} : E_0 \rightarrow Q_\ell \times (s_0, \infty]$ by $\tilde{\gamma}_{\ell,n}(x, u) = (\gamma_{\ell,n}(x), u)$.

Using the homeomorphism $\tilde{\gamma}_{\ell,n}$, one can express $\eta_{k,n}$ as

$$\eta_{k,n} = \sum_{\ell=1}^{b_n} \eta_{k,n}^{(\ell)} \circ \tilde{\gamma}_{\ell,n}^{-1}, \quad (4.3)$$

where

$$\eta_{k,n}^{(\ell)} := \sum_{X \in \mathcal{P}_n|_{K_\ell}} g(X, \mathcal{P}_n|_{Q_\ell}) \delta_{(\gamma_{\ell,n}^{-1}(X), f(X, \mathcal{P}_n|_{Q_\ell}))} \in M_p(E_0).$$

Due to the spatial independence and homogeneity of \mathcal{P}_n , $(\eta_{k,n}^{(\ell)})_{\ell \geq 1}$ constitutes a sequence of i.i.d. point processes on E_0 .

Next, let $(\zeta_k^{(\ell)})_{\ell \geq 1}$ be a collection of i.i.d. Poisson point processes on E_0 with intensity $\text{Leb} \otimes \tau_k$. Then, the proposition below claims that for each $\ell \geq 1$, the law of $\eta_{k,n}^{(\ell)}$ converges to the law of $\zeta_k^{(\ell)}$ as $n \rightarrow \infty$, in terms of the *Kantorovich-Rubinstein distance*. Recall that the Kantorovich-Rubinstein distance between the distributions of two point processes ξ_i , $i = 1, 2$, is defined as

$$d_{\text{KR}}(\mathcal{L}(\xi_1), \mathcal{L}(\xi_2)) := \sup_h |\mathbb{E}[h(\xi_1)] - \mathbb{E}[h(\xi_2)]|, \quad (4.4)$$

where $\mathcal{L}(\xi_i)$ is a probability law of ξ_i , and h is taken over all measurable 1-Lipschitz functions with respect to the total variation distance on the space of point measures; see [1] for more information on the Kantorovich-Rubinstein distance. As a related notion, the *total variation distance* between two measures μ_1 and μ_2 on E_0 is defined as

$$d_{TV}(\mu_1, \mu_2) := \sup_{A \subset E_0} |\mu_1(A) - \mu_2(A)|.$$

As a final remark, we propose one key observation: for $x \in [0, 1]^d$, $\omega \in M_p([0, 1]^d)$ with $x \in \omega$, and $u \in \mathbb{R}$, the following conditions are equivalent.

$$f(x, \omega) > u \Leftrightarrow R_k(x, \omega) > r_n(u) \Leftrightarrow \omega(B_{r_n(u)}(x)) \leq k. \tag{4.5}$$

Throughout the proof, C^* denotes a generic, positive constant, which is independent of n but may vary from one line to another or even within the lines.

Proposition 4.1. *For every $\ell \geq 1$,*

$$d_{KR}(\mathcal{L}(\eta_{k,n}^{(\ell)}), \mathcal{L}(\zeta_k^{(\ell)})) \rightarrow 0, \text{ as } n \rightarrow \infty. \tag{4.6}$$

Proof. The proof is based on [1, Theorem 6.4]. Before applying this theorem, we need some preliminary works. First, define for $x \in [0, 1]^d$ and $\omega \in M_p([0, 1]^d)$,

$$\mathcal{S}(x, \omega) := B_{R_k(x, \omega)}(x);$$

then, f and g are *localized* to \mathcal{S} . Namely, for every $x \in \omega$ and all $S \supset \mathcal{S}(x, \omega)$, we have $g(x, \omega) = g(x, \omega \cap S)$, and also, $f(x, \omega) = f(x, \omega \cap S)$ if $g(x, \omega) = 1$. Moreover, $\mathcal{S}(x, \omega)$ is a *stopping set*; that is, for every compact $S \subset [0, 1]^d$,

$$\{\omega : B_{R_k(x, \omega)}(x) \subset S\} = \{\omega : B_{R_k(x, \omega \cap S)}(x) \subset S\}.$$

Finally, we set $S_x := B_{r_n(w_n)}(x)$ for $x \in K_\ell$.

According to [1, Theorem 6.4], (4.6) can be obtained as a direct consequence of the following conditions. First, one needs to show that

$$d_{TV}(\mathbb{E}[\eta_{k,n}^{(\ell)}(\cdot)], \text{Leb} \otimes \tau_k) \rightarrow 0, \quad n \rightarrow \infty, \tag{4.7}$$

where $\mathbb{E}[\eta_{k,n}^{(\ell)}(\cdot)]$ denotes the intensity measure of $\eta_{k,n}^{(\ell)}$. In addition to (4.7), we also have to show that as $n \rightarrow \infty$,

$$E_1 := 2n \int_{K_\ell} \mathbb{E}[g(x, \mathcal{P}_n|_{Q_\ell} + \delta_x) \mathbb{1}\{\mathcal{S}(x, \mathcal{P}_n|_{Q_\ell} + \delta_x) \not\subset S_x\}] dx \rightarrow 0, \tag{4.8}$$

$$E_2 := 2n^2 \int_{K_\ell} \int_{K_\ell} \mathbb{1}\{S_x \cap S_z \neq \emptyset\} \mathbb{E}[g(x, \mathcal{P}_n|_{Q_\ell} + \delta_x)] \mathbb{E}[g(z, \mathcal{P}_n|_{Q_\ell} + \delta_z)] dx dz \rightarrow 0, \tag{4.9}$$

and

$$E_3 := 2n^2 \int_{K_\ell} \int_{K_\ell} \mathbb{1}\{S_x \cap S_z \neq \emptyset\} \mathbb{E}[g(x, \mathcal{P}_n|_{Q_\ell} + \delta_x + \delta_z) g(z, \mathcal{P}_n|_{Q_\ell} + \delta_x + \delta_z)] dx dz \rightarrow 0. \tag{4.10}$$

Our goal in the sequel is to prove (4.7) – (4.10).

Proof of (4.7): We begin with calculating the measure $\mathbb{E}[\eta_{k,n}^{(\ell)}(\cdot)]$ more explicitly. For $A \subset [0, 1]^d$ and $u > s_0$, by (4.5) and the Mecke formula for Poisson point processes (see,

e.g., Chapter 4 in [12]), together with the fact that $\mathcal{P}_n(Q_\ell)$ is Poisson distributed with mean $n\text{Leb}(Q_\ell) = nb_n^{-1}$,

$$\begin{aligned} \mathbb{E}[\eta_{k,n}^{(\ell)}(A \times (u, \infty))] &= \mathbb{E}\left[\sum_{X \in \mathcal{P}_n|_{K_\ell}} \mathbb{1}\{\gamma_{\ell,n}^{-1}(X) \in A, f(X, \mathcal{P}_n|_{Q_\ell}) > u\}\right] \\ &= \mathbb{E}\left[\sum_{X \in \mathcal{P}_n|_{K_\ell}} \mathbb{1}\{\gamma_{\ell,n}^{-1}(X) \in A, \mathcal{P}_n|_{Q_\ell}(B_{r_n(u)}(X)) \leq k\}\right] \\ &= \mathbb{E}\left[\sum_{X \in \mathcal{P}_n|_{K_\ell}} \mathbb{1}\{\gamma_{\ell,n}^{-1}(X) \in A, \mathcal{P}_n(B_{r_n(u)}(X)) \leq k\}\right] \\ &= nb_n^{-1} \mathbb{P}\left(\gamma_{\ell,n}^{-1}(Y) \in A \setminus \gamma_{\ell,n}^{-1}(M_\ell), (\mathcal{P}_n + \delta_Y)(B_{r_n(u)}(Y)) \leq k\right), \end{aligned}$$

where Y is a uniform random variable on $[0, 1]^d$, independent of \mathcal{P}_n . At the third equality above, we have dropped the restriction of \mathcal{P}_n , i.e., $\mathcal{P}_n|_{Q_\ell} = \mathcal{P}_n$, because $B_{r_n(u)}(X) \subset Q_\ell$ for all $X \in \mathcal{P}_n|_{K_\ell}$. By the conditioning on Y , the last expression equals

$$\begin{aligned} &nb_n^{-1} \mathbb{E}\left[\mathbb{1}\{\gamma_{\ell,n}^{-1}(Y) \in A \setminus \gamma_{\ell,n}^{-1}(M_\ell)\} \mathbb{P}(\mathcal{P}_n(B_{r_n(u)}(Y)) \leq k - 1 | Y)\right] \\ &= nb_n^{-1} \text{Leb}(A \setminus \gamma_{\ell,n}^{-1}(M_\ell)) \sum_{i=0}^{k-1} e^{-(a_n+u)} \frac{(a_n+u)^i}{i!}. \end{aligned}$$

It thus turns out that $\mathbb{E}[\eta_{k,n}^{(\ell)}(\cdot)]$ has the density

$$q_k(x, u) := nb_n^{-1} \mathbb{1}\{x \notin \gamma_{\ell,n}^{-1}(M_\ell)\} \frac{e^{-(a_n+u)}(a_n+u)^{k-1}}{(k-1)!}, \quad x \in [0, 1]^d, u > s_0.$$

Therefore, we have

$$\begin{aligned} d_{\text{TV}}\left(\mathbb{E}[\eta_{k,n}^{(\ell)}(\cdot)], \text{Leb} \otimes \tau_k\right) &\leq \int_{E_0} \left|q_k(x, u) - \frac{e^{-u}}{(k-1)!}\right| dx du \\ &\leq \text{Leb}(\gamma_{\ell,n}^{-1}(M_\ell)) \frac{e^{-s_0}}{(k-1)!} + \int_{s_0}^{\infty} \left|\left(1 + \frac{u}{a_n}\right)^{k-1} - 1\right| e^{-u} du. \end{aligned}$$

The first term above tends to 0 as $n \rightarrow \infty$, because

$$\text{Leb}(\gamma_{\ell,n}^{-1}(M_\ell)) = b_n \text{Leb}(M_\ell) = 1 - (1 - b_n^{1/d} r_n(w_n))^d \rightarrow 0, \quad n \rightarrow \infty,$$

while the second term vanishes by the dominated convergence theorem.

Proof of (4.8): It follows from (4.5) that

$$\begin{aligned} \mathbb{1}\{S(x, \mathcal{P}_n|_{Q_\ell} + \delta_x) \not\subset S_x\} &= \mathbb{1}\{B_{R_k(x, \mathcal{P}_n|_{Q_\ell} + \delta_x)}(x) \supset B_{r_n(w_n)}(x)\} \\ &= \mathbb{1}\{(\mathcal{P}_n|_{Q_\ell} + \delta_x)(B_{r_n(w_n)}(x)) \leq k\} \\ &= \mathbb{1}\{\mathcal{P}_n|_{Q_\ell}(B_{r_n(w_n)}(x)) \leq k - 1\} \\ &= \mathbb{1}\{\mathcal{P}_n(B_{r_n(w_n)}(x)) \leq k - 1\}. \end{aligned}$$

At the fourth equality above, we have dropped the restriction of \mathcal{P}_n , due to the fact that $B_{r_n(w_n)}(x) \subset Q_\ell$ for all $x \in K_\ell$. Now, as $n \rightarrow \infty$,

$$\begin{aligned} E_1 &\leq 2n \int_{K_\ell} \mathbb{P}\left(\mathcal{P}_n(B_{r_n(w_n)}(x)) \leq k - 1\right) dx = 2n \text{Leb}(K_\ell) \sum_{i=0}^{k-1} e^{-(a_n+w_n)} \frac{(a_n+w_n)^i}{i!} \\ &\leq 2e^{-w_n} \sum_{i=0}^{k-1} \frac{(a_n+w_n)^i}{i! a_n^{k-1}} \leq C^* e^{-w_n} \rightarrow 0. \end{aligned}$$

Proof of (4.9): One can see that for $x, z \in K_\ell$,

$$\mathbb{1}\{S_x \cap S_z \neq \emptyset\} \leq \mathbb{1}\{\|x - z\| \leq 2r_n(w_n)\}, \tag{4.11}$$

and from (4.5),

$$\mathbb{E}[g(x, \mathcal{P}_n|_{Q_\ell} + \delta_x)] = \mathbb{P}\left(\mathcal{P}_n(B_{r_n(s_0)}(x)) \leq k - 1\right) = \sum_{i=0}^{k-1} e^{-(a_n+s_0)} \frac{(a_n + s_0)^i}{i!}.$$

Therefore, as $n \rightarrow \infty$,

$$\begin{aligned} E_2 &\leq 2n^2 \left\{ \sum_{i=0}^{k-1} e^{-(a_n+s_0)} \frac{(a_n + s_0)^i}{i!} \right\}^2 \int_{Q_\ell} \int_{Q_\ell} \mathbb{1}\{\|x - z\| \leq 2r_n(w_n)\} dx dz \\ &\leq C^* b_n^2 \int_{Q_\ell} \int_{Q_\ell} \mathbb{1}\{\|x - z\| \leq 2r_n(w_n)\} dx dz \\ &\leq C^* b_n^2 \text{Leb}(Q_\ell) r_n(w_n)^d = C^* b_n r_n(w_n)^d \rightarrow 0. \end{aligned}$$

Proof of (4.10): It follows from (4.5) and (4.11) that E_3 can be split into two terms:

$$\begin{aligned} E_3 &\leq 2n^2 \int_{K_\ell} \int_{K_\ell} \mathbb{1}\{\|x - z\| \leq 2r_n(w_n)\} \\ &\quad \times \mathbb{P}\left(\left(\mathcal{P}_n + \delta_z\right)(B_{r_n(s_0)}(x)) \leq k - 1, \left(\mathcal{P}_n + \delta_x\right)(B_{r_n(s_0)}(z)) \leq k - 1\right) dx dz \\ &= 2n^2 \int_{K_\ell} \int_{K_\ell} \mathbb{1}\{\|x - z\| \leq r_n(s_0)\} \\ &\quad \times \mathbb{P}\left(\left(\mathcal{P}_n + \delta_z\right)(B_{r_n(s_0)}(x)) \leq k - 1, \left(\mathcal{P}_n + \delta_x\right)(B_{r_n(s_0)}(z)) \leq k - 1\right) dx dz \\ &\quad + 2n^2 \int_{K_\ell} \int_{K_\ell} \mathbb{1}\{r_n(s_0) < \|x - z\| \leq 2r_n(w_n)\} \\ &\quad \times \mathbb{P}\left(\left(\mathcal{P}_n + \delta_z\right)(B_{r_n(s_0)}(x)) \leq k - 1, \left(\mathcal{P}_n + \delta_x\right)(B_{r_n(s_0)}(z)) \leq k - 1\right) dx dz \\ &=: E_{3,1} + E_{3,2}. \end{aligned} \tag{4.12}$$

For $E_{3,1}$, if $\|x - z\| \leq r_n(s_0)$ with $x, z \in K_\ell$, then $(\mathcal{P}_n + \delta_z)(B_{r_n(s_0)}(x)) = \mathcal{P}_n(B_{r_n(s_0)}(x)) + 1$. Because of the spatial independence of \mathcal{P}_n ,

$$\begin{aligned} E_{3,1} &\leq 2n^2 \int_{K_\ell} \int_{K_\ell} \mathbb{1}\{\|x - z\| \leq r_n(s_0)\} \mathbb{P}\left(\mathcal{P}_n(B_{r_n(s_0)}(x)) \leq k - 2\right) \\ &\quad \times \mathbb{P}\left(\mathcal{P}_n(B_{r_n(s_0)}(z) \setminus B_{r_n(s_0)}(x)) \leq k - 2\right) dx dz. \end{aligned} \tag{4.13}$$

Then, it is easy to see that

$$\mathbb{P}\left(\mathcal{P}_n(B_{r_n(s_0)}(x)) \leq k - 2\right) \leq C^* d_n^{k-2} e^{-a_n},$$

and also,

$$\mathbb{P}\left(\mathcal{P}_n(B_{r_n(s_0)}(z) \setminus B_{r_n(s_0)}(x)) \leq k - 2\right) \leq C^* e^{-\frac{n}{2} \text{Leb}(B_{r_n(s_0)}(z) \setminus B_{r_n(s_0)}(x))}.$$

Notice that

$$\text{Leb}(B_{r_n(s_0)}(z) \setminus B_{r_n(s_0)}(x)) \geq C^* r_n(s_0)^{d-1} \|x - z\|,$$

whenever $\|x - z\| \leq 2r_n(s_0)$ (see Equ. (7.5) in [21]); so, we have

$$\mathbb{P}\left(\mathcal{P}_n(B_{r_n(s_0)}(z) \setminus B_{r_n(s_0)}(x)) \leq k - 2\right) \leq C^* e^{-\frac{n}{2} C^* r_n(s_0)^{d-1} \|x - z\|}.$$

Referring these bounds back to (4.13),

$$\begin{aligned} E_{3,1} &\leq C^* n^2 \int_{Q_\ell} \int_{\mathbb{R}^d} a_n^{k-2} e^{-a_n} e^{-\frac{n}{2} C^* r_n(s_0)^{d-1} \|x-z\|} dx dz \\ &= C^* n^2 a_n^{k-2} e^{-a_n} b_n^{-1} \int_0^\infty e^{-\frac{n}{2} C^* r_n(s_0)^{d-1} \rho} \rho^{d-1} d\rho \\ &= \frac{C^* n^2 a_n^{k-2} e^{-a_n} b_n^{-1}}{(n r_n(s_0)^{d-1})^d} = C^* \left(1 + \frac{s_0}{a_n}\right) \frac{1}{(a_n + s_0)^d} \rightarrow 0, \quad n \rightarrow \infty. \end{aligned}$$

By the spatial independence of \mathcal{P}_n ,

$$\begin{aligned} E_{3,2} &\leq 2n^2 \int_{K_\ell} \int_{K_\ell} \mathbb{P}\left(\mathcal{P}_n(B_{r_n(s_0)}(x)) \leq k-1\right) \mathbb{P}\left(\mathcal{P}_n(B_{r_n(s_0)}(z) \setminus B_{r_n(s_0)}(x)) \leq k-1\right) \\ &\quad \times \mathbb{1}\{r_n(s_0) < \|x-z\| \leq 2r_n(w_n)\} dx dz. \end{aligned}$$

Then,

$$\mathbb{P}\left(\mathcal{P}_n(B_{r_n(s_0)}(x)) \leq k-1\right) \leq C^* a_n^{k-1} e^{-a_n}.$$

Moreover, if $\|x-z\| > r_n(s_0)$ with $x, z \in K_\ell$, then

$$n \text{Leb}(B_{r_n(s_0)}(z) \setminus B_{r_n(s_0)}(x)) \geq \frac{n}{2} \text{Leb}(B_{r_n(s_0)}(z)) = \frac{a_n + s_0}{2},$$

from which we have

$$\mathbb{P}\left(\mathcal{P}_n(B_{r_n(s_0)}(z) \setminus B_{r_n(s_0)}(x)) \leq k-1\right) \leq \sum_{i=0}^{k-1} e^{-\frac{a_n+s_0}{2}} \frac{(a_n + s_0)^i}{i!} \leq C^* a_n^{k-1} e^{-\frac{a_n}{2}}.$$

Appealing to these obtained bounds, as $n \rightarrow \infty$,

$$\begin{aligned} E_{3,2} &\leq C^* n^2 \int_{Q_\ell} \int_{\mathbb{R}^d} a_n^{k-1} e^{-a_n} \cdot a_n^{k-1} e^{-\frac{a_n}{2}} \mathbb{1}\{\|x-z\| \leq 2r_n(w_n)\} dx dz \\ &= C^* n^2 a_n^{2(k-1)} e^{-\frac{3a_n}{2}} b_n^{-1} r_n(w_n)^d \leq C^* a_n^k e^{-\frac{a_n}{2}} \rightarrow 0, \end{aligned}$$

as desired. □

Recall now that $(\zeta_k^{(\ell)})_{\ell \geq 1}$ are i.i.d. Poisson point processes on E_0 with intensity $\text{Leb} \otimes \tau_k$. The next proposition claims that the process

$$\zeta_{k,n} := \sum_{\ell=1}^{b_n} \zeta_k^{(\ell)} \circ \tilde{\gamma}_{\ell,n}^{-1} \tag{4.14}$$

satisfies the desired LDP in Theorem 2.1. Let $(\Omega', \mathcal{F}', \mathbb{P}')$ be the probability space on which (4.14) is defined.

Proposition 4.2. *The sequence $(\zeta_{k,n}/b_n)_{n \geq 1}$ satisfies an LDP in the space $M_+(E_0)$ equipped with weak topology, with speed b_n and rate function $H_k(\cdot | \text{Leb} \otimes \tau_k) = \Lambda_k^*$.*

Proof. By the transformation theorem, $\zeta_k^{(\ell)} \circ \tilde{\gamma}_{\ell,n}^{-1}$ becomes a Poisson point process on E_0 with intensity $(\text{Leb} \otimes \tau_k) \circ \tilde{\gamma}_{\ell,n}^{-1} = b_n (\text{Leb}|_{Q_\ell} \otimes \tau_k)$. As $(\zeta_k^{(\ell)})_{\ell \geq 1}$ are i.i.d., (4.14) turns out to be a Poisson point process on E_0 with intensity $b_n (\text{Leb} \otimes \tau_k)$; thus, there exists a sequence $(\xi_k^{(\ell)})_{\ell \geq 1}$ of i.i.d. Poisson point processes on E_0 with intensity $\text{Leb} \otimes \tau_k$, so that

$$\zeta_{k,n} \stackrel{d}{=} \sum_{\ell=1}^{b_n} \xi_k^{(\ell)}. \tag{4.15}$$

For convenience, we assume that $(\xi_k^{(\ell)})_{\ell \geq 1}$ are defined in the same probability space $(\Omega', \mathcal{F}', \mathbb{P}')$. By applying the Poisson variant of Sanov's theorem to (4.15), we conclude that $(\zeta_{k,n}/b_n)_{n \geq 1}$ satisfies an LDP with speed b_n and rate function $H_k(\cdot | \text{Leb} \otimes \tau_k)$; see [6, Proposition 3.6] and [3, Theorem 6.2.10] for details. To deduce the LDP with rate function Λ_k^* , we compute the logarithmic Laplace functional of $\zeta_k^{(1)}$: it follows from Theorem 5.1 in [25] that for every $f \in C_b(E_0)$,

$$\Lambda_k(f) := \log \mathbb{E}'[e^{\xi_k^{(1)}(f)}] = \int_{E_0} (e^{f(x,u)} - 1) \, dx \, \tau_k(du).$$

Recalling that the space $M_+(E_0)$ is Polish when it is equipped with weak topology, one can exploit Cramér's theorem for Polish spaces (see, e.g., Theorem 6.1.3 in [3]) to obtain the required LDP, for which Λ_k^* is given as the Legendre transform of Λ_k . \square

By the maximal coupling argument (see [11, Lemma 4.32]), for every $\ell \geq 1$ there exists a coupling $(\hat{\eta}_{k,n}^{(\ell)}, \hat{\zeta}_k^{(\ell)})$ defined on a probability space $(\hat{\Omega}_\ell, \hat{\mathcal{F}}_\ell, \hat{\mathbb{P}}_\ell)$, such that $\hat{\eta}_{k,n}^{(\ell)} \stackrel{d}{=} \eta_{k,n}^{(\ell)}$ and $\hat{\zeta}_k^{(\ell)} \stackrel{d}{=} \zeta_k^{(\ell)}$, and

$$\hat{\mathbb{P}}_\ell(\hat{\eta}_{k,n}^{(\ell)} \neq \hat{\zeta}_k^{(\ell)}) = d_{\text{TV}}(\mathcal{L}(\eta_{k,n}^{(\ell)}), \mathcal{L}(\zeta_k^{(\ell)})) \leq d_{\text{KR}}(\mathcal{L}(\eta_{k,n}^{(\ell)}), \mathcal{L}(\zeta_k^{(\ell)})) \rightarrow 0, \quad n \rightarrow \infty. \quad (4.16)$$

In particular, $(\hat{\eta}_{k,n}^{(\ell)}, \hat{\zeta}_k^{(\ell)})_{\ell \geq 1}$ constitutes a sequence of i.i.d. random vectors on the probability space $(\hat{\Omega}, \hat{\mathcal{F}}, \hat{\mathbb{P}})$, where $\hat{\Omega} = \prod_{\ell=1}^\infty \hat{\Omega}_\ell$, $\hat{\mathcal{F}} = \bigotimes_{\ell=1}^\infty \hat{\mathcal{F}}_\ell$, and $\hat{\mathbb{P}} = \bigotimes_{\ell=1}^\infty \hat{\mathbb{P}}_\ell$. Defining $\hat{\eta}_{k,n}$ and $\hat{\zeta}_{k,n}$ analogously to (4.3) and (4.14), Proposition 4.3 below demonstrates that $(\hat{\eta}_{k,n}/b_n)_{n \geq 1}$ and $(\hat{\zeta}_{k,n}/b_n)_{n \geq 1}$ are exponentially equivalent (in terms of the total variation distance) under the coupled probability measure $\hat{\mathbb{P}}$. Since the LDP for $(\zeta_{k,n}/b_n)_{n \geq 1}$ was already given by Proposition 4.2, this exponential equivalence allows us to conclude that $(\eta_{k,n}/b_n)_{n \geq 1}$ fulfills an LDP in Theorem 2.1.

Proposition 4.3. For every $\delta > 0$,

$$\frac{1}{b_n} \log \hat{\mathbb{P}}(d_{\text{TV}}(\hat{\eta}_{k,n}, \hat{\zeta}_{k,n}) \geq \delta b_n) \rightarrow -\infty, \quad n \rightarrow \infty.$$

Proof. The proof is highly related to [8, Lemma 5.5], but we still want to give a concise and self-contained argument. By Markov's inequality and the fact that $(\hat{\eta}_{k,n}^{(\ell)}, \hat{\zeta}_k^{(\ell)})_{\ell \geq 1}$ are i.i.d. processes, we have, for every $a > 0$,

$$\begin{aligned} \frac{1}{b_n} \log \hat{\mathbb{P}}(d_{\text{TV}}(\hat{\eta}_{k,n}, \hat{\zeta}_{k,n}) \geq \delta b_n) &\leq \frac{1}{b_n} \log \hat{\mathbb{P}}\left(\sum_{\ell=1}^{b_n} d_{\text{TV}}(\hat{\eta}_{k,n}^{(\ell)}, \hat{\zeta}_k^{(\ell)}) \geq \delta b_n\right) \\ &\leq -a\delta + \log \hat{\mathbb{E}}[e^{ad_{\text{TV}}(\hat{\eta}_{k,n}^{(1)}, \hat{\zeta}_k^{(1)})}], \end{aligned}$$

where $\hat{\mathbb{E}}$ denotes an expectation with respect to $(\hat{\Omega}, \hat{\mathcal{F}}, \hat{\mathbb{P}})$. Since $a > 0$ is arbitrary, the desired result immediately follows if we can prove that for every $a > 0$,

$$\lim_{n \rightarrow \infty} \hat{\mathbb{E}}[e^{ad_{\text{TV}}(\hat{\eta}_{k,n}^{(1)}, \hat{\zeta}_k^{(1)})}] = 1. \quad (4.17)$$

By virtue of (4.16), $d_{\text{TV}}(\hat{\eta}_{k,n}^{(1)}, \hat{\zeta}_k^{(1)})$ converges to 0 in probability with respect to $\hat{\mathbb{P}}$. By the Cauchy-Schwarz inequality, (4.5), and the fact that $\zeta_k^{(1)}(E_0)$ is Poisson with mean $e^{-s_0}/(k-1)!$,

$$\begin{aligned} \hat{\mathbb{E}}[e^{ad_{\text{TV}}(\hat{\eta}_{k,n}^{(1)}, \hat{\zeta}_k^{(1)})}] &\leq \left\{ \mathbb{E}[e^{2a\eta_{k,n}^{(1)}(E_0)}] \right\}^{1/2} \left\{ \mathbb{E}'[e^{2a\zeta_k^{(1)}(E_0)}] \right\}^{1/2} \\ &= \left\{ \mathbb{E}[e^{2a \sum_{X \in \mathcal{P}_n|_{K_1}} \mathbb{1}\{\mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k\}}] \right\}^{1/2} \exp\left\{ \frac{e^{-s_0}}{2(k-1)!} (e^{2a} - 1) \right\}. \end{aligned}$$

Now, the desired uniform integrability for (4.17) follows, provided that for every $a > 0$,

$$\limsup_{n \rightarrow \infty} \mathbb{E} \left[e^{a \sum_{X \in \mathcal{P}_n |_{K_1}} \mathbb{1} \{ \mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k \}} \right] < \infty. \tag{4.18}$$

As in the proof of [8, Lemma 5.5], by utilizing the diluted family of cubes

$$G = \{ 4r_n(s_0)z + [0, r_n(s_0)/\sqrt{d}]^d \subset Q_1 : z \in \mathbb{Z}^d \},$$

it turns out that (4.18) is obtained as a consequence of

$$\limsup_{n \rightarrow \infty} \left\{ \mathbb{E} \left[e^{a \sum_{X \in \mathcal{P}_n} \mathbb{1} \{ \mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k, X \in K_1 \cap J \}} \right] \right\}^{1/(b_n(4r_n(s_0))^d)} < \infty,$$

where $J = [0, r_n(s_0)/\sqrt{d}]^d$ and $1/(b_n(4r_n(s_0))^d)$ represents the number of cubes in G . Notice that

$$\sum_{X \in \mathcal{P}_n} \mathbb{1} \{ \mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k, X \in K_1 \cap J \} \in \{0, 1, \dots, k\}, \tag{4.19}$$

because if there exist more than k points inside J , these points never contribute to (4.19). Therefore, by Markov's inequality,

$$\begin{aligned} & \left\{ \mathbb{E} \left[e^{a \sum_{X \in \mathcal{P}_n} \mathbb{1} \{ \mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k, X \in K_1 \cap J \}} \right] \right\}^{1/(b_n(4r_n(s_0))^d)} \\ & \leq \left(1 + \sum_{\ell=1}^k e^{a\ell} \mathbb{E} \left[\sum_{X \in \mathcal{P}_n} \mathbb{1} \{ \mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k, X \in K_1 \cap J \} \right] \right)^{1/(b_n(4r_n(s_0))^d)} \\ & = \left(1 + \sum_{\ell=1}^k e^{a\ell} n \sum_{i=0}^{k-1} e^{-(a_n+s_0)} \frac{(a_n+s_0)^i}{i!} \mathbb{P}(X_1 \in K_1 \cap J) \right)^{1/(b_n(4r_n(s_0))^d)} \\ & \leq (1 + C^* e^{ak} b_n r_n(s_0)^d)^{1/(b_n(4r_n(s_0))^d)} \rightarrow e^{C^* e^{ak}/4^d} < \infty, \text{ as } n \rightarrow \infty. \quad \square \end{aligned} \tag{4.20}$$

As shown in the last two propositions, $(\eta_{k,n}/b_n)_{n \geq 1}$ satisfies the LDP in Theorem 2.1. Thus, our final task is to demonstrate that $(L_{k,n}/b_n)_{n \geq 1}$ exhibits the same LDP as $(\eta_{k,n}/b_n)_{n \geq 1}$. Although this can be done by establishing exponential equivalence between $(L_{k,n}/b_n)_{n \geq 1}$ and $(\eta_{k,n}/b_n)_{n \geq 1}$ in terms of the total variation distance, proving directly this exponential equivalence seems to be difficult. Alternatively, we set up an additional sequence

$$\eta'_{k,n} := \sum_{\ell=1}^{b_n} \sum_{X \in \mathcal{P}_n |_{Q_\ell}} g(X, \mathcal{P}_n |_{Q_\ell}) \mathbb{1} \{ R_k(X, \mathcal{P}_n |_{Q_\ell}) \leq r_n(w_n) \} \delta_{(X, f(X, \mathcal{P}_n |_{Q_\ell}))}, \tag{4.21}$$

and prove that $(\eta'_{k,n}/b_n)_{n \geq 1}$ shows the same LDP as $(\eta_{k,n}/b_n)_{n \geq 1}$ (see Proposition 4.4). Subsequently, we define

$$L'_{k,n} := \begin{cases} \sum_{X \in \mathcal{P}_n} g(X, \mathcal{P}_n) \mathbb{1} \{ R_k(X, \mathcal{P}_n) \leq \sqrt{d} b_n^{-1/d} \} \delta_{(X, f(X, \mathcal{P}_n))} & \text{if } |\mathcal{P}_n| > k, \\ \emptyset & \text{if } |\mathcal{P}_n| \leq k. \end{cases}$$

and prove that $(L'_{k,n}/b_n)_{n \geq 1}$ satisfies the same LDP as $(L_{k,n}/b_n)_{n \geq 1}$ (see Proposition 4.5). Finally, Proposition 4.6 gives exponential equivalence between $(L'_{k,n}/b_n)_{n \geq 1}$ and $(\eta'_{k,n}/b_n)_{n \geq 1}$. Combining Propositions 4.4 – 4.6 concludes the required exponential equivalence between $(L_{k,n}/b_n)_{n \geq 1}$ and $(\eta_{k,n}/b_n)_{n \geq 1}$.

Proposition 4.4. *The sequence $(\eta'_{k,n}/b_n)_{n \geq 1}$ satisfies the same LDP as $(\eta_{k,n}/b_n)_{n \geq 1}$.*

Proof. For our purpose, we demonstrate that for every $\delta > 0$,

$$\frac{1}{b_n} \log \mathbb{P}(d_{\text{TV}}(\eta_{k,n}, \eta'_{k,n}) \geq \delta b_n) \rightarrow -\infty, \quad n \rightarrow \infty. \tag{4.22}$$

First, we see from (4.5) that

$$\begin{aligned} d_{\text{TV}}(\eta_{k,n}, \eta'_{k,n}) &\leq \sum_{\ell=1}^{b_n} \sum_{X \in \mathcal{P}_n|_{K_\ell}} \mathbb{1}\{R_k(X, \mathcal{P}_n|_{Q_\ell}) > r_n(w_n)\} \\ &= \sum_{\ell=1}^{b_n} \sum_{X \in \mathcal{P}_n|_{K_\ell}} \mathbb{1}\{\mathcal{P}_n(B_{r_n(w_n)}(X)) \leq k\}, \end{aligned}$$

from which one can bound (4.22) by

$$-a\delta + \log \mathbb{E}\left[e^{a \sum_{X \in \mathcal{P}_n|_{K_1}} \mathbb{1}\{\mathcal{P}_n(B_{r_n(w_n)}(X)) \leq k\}}\right]$$

for every $a > 0$. Now, one has to show that for every $a > 0$,

$$\limsup_{n \rightarrow \infty} \mathbb{E}\left[e^{a \sum_{X \in \mathcal{P}_n|_{K_1}} \mathbb{1}\{\mathcal{P}_n(B_{r_n(w_n)}(X)) \leq k\}}\right] \leq 1. \tag{4.23}$$

As in the proof of Proposition 4.3, (4.23) can be implied by

$$\limsup_{n \rightarrow \infty} \left\{ \mathbb{E}\left[e^{a \sum_{X \in \mathcal{P}_n} \mathbb{1}\{\mathcal{P}_n(B_{r_n(w_n)}(X)) \leq k, X \in K_1 \cap J\}}\right] \right\}^{1/(b_n r_n(w_n)^d)} \leq 1,$$

for all $a > 0$, where $J = [0, r_n(w_n)/\sqrt{d}]^d$. Now, instead of (4.20), we have that

$$\begin{aligned} &\limsup_{n \rightarrow \infty} \left\{ \mathbb{E}\left[e^{a \sum_{X \in \mathcal{P}_n} \mathbb{1}\{\mathcal{P}_n(B_{r_n(w_n)}(X)) \leq k, X \in K_1 \cap J\}}\right] \right\}^{1/(b_n r_n(w_n)^d)} \\ &\leq \limsup_{n \rightarrow \infty} (1 + C^* e^{ak} b_n r_n(w_n)^d e^{-w_n})^{1/(b_n r_n(w_n)^d)} \\ &= \limsup_{n \rightarrow \infty} e^{C^* e^{ak} e^{-w_n}} = 1, \end{aligned}$$

as required. □

Proposition 4.5. *The sequence $(L'_{k,n}/b_n)_{n \geq 1}$ satisfies the same LDP as $(L_{k,n}/b_n)_{n \geq 1}$.*

Proof. Throughout the proof, we assume $|\mathcal{P}_n| > k$. Using the bound

$$d_{\text{TV}}(L_{k,n}, L'_{k,n}) \leq \sum_{X \in \mathcal{P}_n} \mathbb{1}\{R_k(X, \mathcal{P}_n) > \sqrt{db_n^{-1/d}}\},$$

it is sufficient to show that, for every $\delta > 0$,

$$\frac{1}{b_n} \log \mathbb{P}\left(\sum_{X \in \mathcal{P}_n} \mathbb{1}\{R_k(X, \mathcal{P}_n) > \sqrt{db_n^{-1/d}}\} \geq \delta b_n\right) \rightarrow -\infty, \quad n \rightarrow \infty.$$

Suppose there exists a point $X \in \mathcal{P}_n \cap Q_\ell$ for some $\ell \in \{1, \dots, b_n\}$ so that $R_k(X, \mathcal{P}_n) > \sqrt{db_n^{-1/d}}$. Then, $B_{R_k(X, \mathcal{P}_n)}(X) \cap Q_\ell$ contains at most $k + 1$ points of \mathcal{P}_n (including X itself). Thus,

$$\begin{aligned} \sum_{X \in \mathcal{P}_n} \mathbb{1}\{R_k(X, \mathcal{P}_n) > \sqrt{db_n^{-1/d}}\} &= \sum_{\ell=1}^{b_n} \sum_{X \in \mathcal{P}_n|_{Q_\ell}} \mathbb{1}\{R_k(X, \mathcal{P}_n) > \sqrt{db_n^{-1/d}}\} \\ &\leq (k + 1) \sum_{\ell=1}^{b_n} \mathbb{1}\{\mathcal{P}_n(Q_\ell) \leq k + 1\}. \end{aligned}$$

Now, we only have to demonstrate that

$$\frac{1}{b_n} \log \mathbb{P} \left(\sum_{\ell=1}^{b_n} \mathbb{1} \{ \mathcal{P}_n(Q_\ell) \leq k+1 \} \geq \delta b_n \right) \rightarrow -\infty, \quad n \rightarrow \infty. \quad (4.24)$$

By Markov's inequality, we have, for every $a > 0$,

$$\begin{aligned} \frac{1}{b_n} \log \mathbb{P} \left(\sum_{\ell=1}^{b_n} \mathbb{1} \{ \mathcal{P}_n(Q_\ell) \leq k+1 \} \geq \delta b_n \right) &\leq -a\delta + \log \mathbb{E} \left[e^{a \mathbb{1} \{ \mathcal{P}_n(Q_1) \leq k+1 \}} \right] \\ &\leq -a\delta + \log \left(1 + e^a \sum_{i=0}^{k+1} e^{-n \text{Leb}(Q_1)} \frac{(n \text{Leb}(Q_1))^i}{i!} \right) \\ &\leq -a\delta + \log \left(1 + (k+2) e^a e^{-nb_n^{-1}} (nb_n^{-1})^{k+1} \right) \\ &\rightarrow -a\delta, \quad \text{as } n \rightarrow \infty. \end{aligned}$$

As $a > 0$ is arbitrary, we have obtained (4.24). □

Proposition 4.6. *The sequence $(L'_{k,n}/b_n)_{n \geq 1}$ exhibits the same LDP as $(\eta'_{k,n}/b_n)_{n \geq 1}$.*

Proof. We prove exponential equivalence between the two sequences: for every $\delta > 0$,

$$\frac{1}{b_n} \log \mathbb{P} (d_{\text{TV}}(L'_{k,n}, \eta'_{k,n}) \geq \delta b_n) \rightarrow -\infty, \quad n \rightarrow \infty.$$

For convenience, let us slightly change the formulation of $\eta'_{k,n}$ given at (4.21). To begin, observe that if $X \in \mathcal{P}_n|_{K_\ell}$ with $R_k(X, \mathcal{P}_n|_{Q_\ell}) \leq r_n(w_n)$, then $R_k(X, \mathcal{P}_n|_{Q_\ell}) = R_k(X, \mathcal{P}_n)$. Hence, one can replace the restricted process $\mathcal{P}_n|_{Q_\ell}$ in (4.21) with \mathcal{P}_n ; that is,

$$\eta'_{k,n} = \sum_{\ell=1}^{b_n} \sum_{X \in \mathcal{P}_n|_{K_\ell}} g(X, \mathcal{P}_n) \mathbb{1} \{ R_k(X, \mathcal{P}_n) \leq r_n(w_n) \} \delta_{(X, f(X, \mathcal{P}_n))}.$$

Using this representation and assuming $|\mathcal{P}_n| > k$ allows us to express the total variation distance in such a way that

$$\begin{aligned} d_{\text{TV}}(L'_{k,n}, \eta'_{k,n}) &= \frac{1}{b_n} \sup_{A \subset E_0} \left\{ \sum_{X \in \mathcal{P}_n} \mathbb{1} \{ (X, f(X, \mathcal{P}_n)) \in A, R_k(X, \mathcal{P}_n) \leq \sqrt{db_n^{-1/d}} \} \right. \\ &\quad \left. - \sum_{\ell=1}^{b_n} \sum_{X \in \mathcal{P}_n|_{K_\ell}} \mathbb{1} \{ (X, f(X, \mathcal{P}_n)) \in A, R_k(X, \mathcal{P}_n) \leq r_n(w_n) \} \right\} \\ &=: \frac{1}{b_n} \sup_{A \subset E_0} (T_n^{(1)} - T_n^{(2)}). \end{aligned} \quad (4.25)$$

We now derive an upper bound of (4.25). We consider a k -nearest neighbor ball centered at $X \in \mathcal{P}_n$ with $(X, f(X, \mathcal{P}_n)) \in A$. Suppose this ball is counted by $T_n^{(1)}$, but not counted by $T_n^{(2)}$. Then, this ball must be of either Type 1 or Type 2 as defined below.

Type 1: The center X is located in Q_ℓ for some $\ell \in \{1, \dots, b_n\}$ such that

$$r_n(w_n) < R_k(X, \mathcal{P}_n) \leq \sqrt{db_n^{-1/d}}.$$

Type 2: The center X is in $Q_\ell^\partial(r_n(w_n))$ for some $\ell \in \{1, \dots, b_n\}$, where

$$Q_\ell^\partial(r) := \{x \in Q_\ell : \inf_{y \in \partial Q_\ell} \|x - y\| \leq r\}, \quad r > 0,$$

so that $r_n(s_0) < R_k(X, \mathcal{P}_n) \leq r_n(w_n)$.

Then, the number of k -nearest neighbor balls of Type 1 can be bounded by

$$U_n^{(1)} := \sum_{\ell=1}^{b_n} \sum_{X \in \mathcal{P}_n|_{Q_\ell}} \mathbb{1}\{R_k(X, \mathcal{P}_n) \in (r_n(w_n), \sqrt{db_n^{-1/d}}]\},$$

while the number of those of Type 2 is bounded by

$$U_n^{(2)} := \sum_{\ell=1}^{b_n} \sum_{X \in \mathcal{P}_n|_{Q_\ell^\partial(r_n(w_n))}} \mathbb{1}\{R_k(X, \mathcal{P}_n) \in (r_n(s_0), r_n(w_n)]\}.$$

Hence, the desired result will follow if one can show that for every $\delta > 0$ and $j = 1, 2$,

$$\frac{1}{b_n} \log \mathbb{P}(U_n^{(j)} \geq \delta b_n) \rightarrow -\infty, \quad n \rightarrow \infty. \tag{4.26}$$

We first deal with the case $j = 1$. Let S_ℓ be the collection of cubes $(Q_i)_{i=1}^{b_n}$ that intersect with

$$\text{Tube}(Q_\ell, \sqrt{db_n^{-1/d}}) := \{x \in [0, 1]^d : \inf_{y \in Q_\ell} \text{dist}(x, y) \leq \sqrt{db_n^{-1/d}}\}.$$

Then, the number D_d of such cubes in S_ℓ is finite, depending only on d . Now, we can offer the following bound:

$$U_n^{(1)} \leq \sum_{m=1}^{D_d} \sum_{\ell=1}^{\lfloor D_d^{-1}(b_n-m) \rfloor + 1} \sum_{X \in \mathcal{P}_n} \mathbb{1}\{X \in Q_{(\ell-1)D_d+m}\} \times \mathbb{1}\{R_k(X, \mathcal{P}_n) \in (r_n(w_n), \sqrt{db_n^{-1/d}}]\}.$$

By the homogeneity of \mathcal{P}_n on the torus, we consider only the case $m = D_d$ and obtain from Markov's inequality that, for every $a > 0$,

$$\begin{aligned} & \frac{1}{b_n} \log \mathbb{P}\left(\sum_{\ell=1}^{\lfloor D_d^{-1}b_n \rfloor} \sum_{X \in \mathcal{P}_n|_{Q_\ell D_d}} \mathbb{1}\{R_k(X, \mathcal{P}_n) \in (r_n(w_n), \sqrt{db_n^{-1/d}}]\} \geq \delta b_n\right) \\ & \leq -a\delta + \frac{1}{b_n} \log \mathbb{E}\left[e^{a \sum_{\ell=1}^{\lfloor D_d^{-1}b_n \rfloor} \sum_{X \in \mathcal{P}_n|_{Q_\ell D_d}} \mathbb{1}\{R_k(X, \mathcal{P}_n) \in (r_n(w_n), \sqrt{db_n^{-1/d}}]\}}\right]. \end{aligned} \tag{4.27}$$

Here, a key observation is that if $X \in \mathcal{P}_n|_{Q_\ell D_d}$ with $R_k(X, \mathcal{P}_n) \leq \sqrt{db_n^{-1/d}}$, then $R_k(X, \mathcal{P}_n) = R_k(X, \mathcal{P}_n|_{S_\ell D_d})$, such that $(\mathcal{P}_n|_{S_\ell D_d}, \ell = 1, \dots, \lfloor D_d^{-1}b_n \rfloor)$ are i.i.d. Poisson point processes. Hence, (4.27) can be further bounded by

$$\begin{aligned} & -a\delta + D_d^{-1} \log \mathbb{E}\left[e^{a \sum_{X \in \mathcal{P}_n|_{Q_1}} \mathbb{1}\{R_k(X, \mathcal{P}_n) \in (r_n(w_n), \sqrt{db_n^{-1/d}}]\}}\right] \\ & \leq -a\delta + D_d^{-1} \log \mathbb{E}\left[e^{a \sum_{X \in \mathcal{P}_n|_{Q_1}} \mathbb{1}\{\mathcal{P}_n(B_{r_n(w_n)}(X)) \leq k\}}\right]. \end{aligned}$$

For the inequality above, we have applied (4.5) and dropped the condition $R_k(X, \mathcal{P}_n) \leq \sqrt{db_n^{-1/d}}$. Now, it remains to show that for every $a > 0$,

$$\limsup_{n \rightarrow \infty} \mathbb{E}\left[e^{a \sum_{X \in \mathcal{P}_n|_{Q_1}} \mathbb{1}\{\mathcal{P}_n(B_{r_n(w_n)}(X)) \leq k\}}\right] \leq 1.$$

The proof is however a simple repetition of the argument for (4.23), so we skip it here.

Returning to (4.26), we next work with the case $j = 2$. To this aim, we exploit an argument similar to that in [8, Proposition 5.6]. For $1 \leq j \leq d$, define the collection of ordered j -tuples

$$\mathcal{I}_j = \{\ell = (\ell_1, \dots, \ell_j) : 1 \leq \ell_1 < \dots < \ell_j \leq d\}.$$

Given $\ell = (\ell_1, \dots, \ell_j) \in \mathcal{I}_j$, define also the collection of hyper-rectangles by

$$J_n(r) := \left\{ \left(b_n^{-1/d}z + [0, b_n^{-1/d}]^{\ell_1-1} \times [-r, r] \times [0, b_n^{-1/d}]^{\ell_2-\ell_1-1} \right. \right. \\ \times [-r, r] \times [0, b_n^{-1/d}]^{\ell_3-\ell_2-1} \times \dots \times [-r, r] \\ \left. \left. \times [0, b_n^{-1/d}]^{\ell_j-\ell_{j-1}-1} \times [-r, r] \times [0, b_n^{-1/d}]^{d-\ell_j} \right) \cap [0, 1]^d : z \in \mathbb{Z}_+^d \right\}, \quad r > 0.$$

By construction, all the rectangles in $J_n(r)$ are contained in $\bigcup_{\ell=1}^{b_n} Q_\ell^\partial(r)$, and the number of rectangles in $J_n(r)$ is b_n ; hence, we can enumerate these rectangles as

$$J_n(r) = (I_{p,n}^\ell(r), p = 1, \dots, b_n).$$

In this setting, one can bound $U_n^{(2)}$ by

$$\sum_{j=1}^d \sum_{\ell \in \mathcal{I}_j} \sum_{p=1}^{b_n} \sum_{X \in \mathcal{P}_n} \mathbb{1}\{X \in I_{p,n}^\ell(r_n(w_n))\} \\ \times \mathbb{1}\{R_k(X, \mathcal{P}_n) > r_n(s_0), B_{R_k(X, \mathcal{P}_n)}(X) \subset I_{p,n}^\ell(2r_n(w_n))\}.$$

Owing to this bound, we need to prove that for every $j \in \{1, \dots, d\}$, $\ell \in \mathcal{I}_j$, and $\delta > 0$,

$$\frac{1}{b_n} \log \mathbb{P} \left(\sum_{p=1}^{b_n} \sum_{X \in \mathcal{P}_n} \mathbb{1}\{X \in I_{p,n}^\ell(r_n(w_n))\} \right. \\ \left. \times \mathbb{1}\{R_k(X, \mathcal{P}_n) > r_n(s_0), B_{R_k(X, \mathcal{P}_n)}(X) \subset I_{p,n}^\ell(2r_n(w_n))\} \geq \delta b_n \right) \rightarrow -\infty, \quad n \rightarrow \infty. \tag{4.28}$$

In the above, $B_{R_k(X, \mathcal{P}_n)}(X) \subset I_{p,n}^\ell(2r_n(w_n))$ with $X \in I_{p,n}^\ell(r_n(w_n))$, implies that $R_k(X, \mathcal{P}_n) = R_k(X, \mathcal{P}_n|_{I_{p,n}^\ell(2r_n(w_n))})$. Additionally, $(I_{p,n}^\ell(2r_n(w_n)))_{p=1}^{b_n}$ are disjoint sets, so $(\mathcal{P}_n|_{I_{p,n}^\ell(2r_n(w_n))})_{p=1}^{b_n}$ becomes a sequence of i.i.d. Poisson point processes. Hence, by appealing to Markov's inequality as well as (4.5), one can bound (4.28) by

$$-a\delta + \log \mathbb{E} \left[e^{a \sum_{X \in \mathcal{P}_n} \mathbb{1}\{X \in I_{1,n}^\ell(r_n(w_n))\}} \times \mathbb{1}\{R_k(X, \mathcal{P}_n) > r_n(s_0), B_{R_k(X, \mathcal{P}_n)}(X) \subset I_{1,n}^\ell(2r_n(w_n))\} \right] \\ \leq -a\delta + \log \mathbb{E} \left[e^{a \sum_{X \in \mathcal{P}_n} \mathbb{1}\{X \in I_{1,n}^\ell(r_n(w_n))\}} \times \mathbb{1}\{\mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k\} \right].$$

It is now enough to demonstrate that, for every $a > 0$,

$$\limsup_{n \rightarrow \infty} \mathbb{E} \left[e^{a \sum_{X \in \mathcal{P}_n} \mathbb{1}\{X \in I_{1,n}^\ell(r_n(w_n))\}} \times \mathbb{1}\{\mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k\} \right] = 1.$$

Since the required uniform integrability has already been proven by (4.18), it suffices to show that as $n \rightarrow \infty$,

$$\mathbb{E} \left[\sum_{X \in \mathcal{P}_n} \mathbb{1}\{X \in I_{1,n}^\ell(r_n(w_n))\} \times \mathbb{1}\{\mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k\} \right] \rightarrow 0.$$

Proceeding as before, we obtain that

$$\begin{aligned} & \mathbb{E} \left[\sum_{X \in \mathcal{P}_n} \mathbb{1} \left\{ X \in I_{1,n}^\ell(r_n(w_n)) \right\} \times \mathbb{1} \left\{ \mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k \right\} \right] \\ &= n e^{-(a_n+s_0)} \sum_{i=0}^{k-1} \frac{(a_n+s_0)^i}{i!} \mathbb{P}(X_1 \in I_{1,n}^\ell(r_n(w_n))) \\ &= n e^{-(a_n+s_0)} \sum_{i=0}^{k-1} \frac{(a_n+s_0)^i}{i!} \times C^* r_n(w_n)^j b_n^{-(d-j)/d} \\ &\leq C^* (b_n r_n(w_n)^d)^{j/d} \rightarrow 0, \quad n \rightarrow \infty. \quad \square \end{aligned}$$

Proof of Corollary 2.2. Since the desired LDP has already been shown for the case of a Poisson input, it is sufficient to demonstrate that for every $\varepsilon_0 > 0$,

$$b_n^{-1} \log \mathbb{P}(d_{\text{TV}}(L_{k,n}, L_{k,n}^{\mathbb{B}}) \geq \varepsilon_0 b_n) \rightarrow -\infty, \quad \text{as } n \rightarrow \infty. \quad (4.29)$$

Our proof is inspired by Corollary 2.3 in [8]. First, define

$$G := \{3r_n(w_n)z + [0, r_n(s_0)/\sqrt{d}]^d \subset [0, 1]^d : z \in \mathbb{Z}^d\}, \quad (4.30)$$

and consider *finitely many* translates of G , denoted G_1, G_2, \dots, G_M for some M , such that $[0, 1]^d$ can be covered by the union of these translates. In particular, we set $G_1 = G$ and denote it specifically as $G = \{J_1, \dots, J_{b'_n}\}$, where $J_1 = [0, r_n(s_0)/\sqrt{d}]^d$ and $b'_n := (3r_n(w_n))^{-d}$ denotes the number of cubes in G . Since M is a finite constant, (4.29) follows if one can show that

$$\begin{aligned} & b_n^{-1} \log \mathbb{P} \left(\sup_{A \subset E_0} \left| \sum_{X \in \mathcal{P}_n} g(X, \mathcal{P}_n) \mathbb{1} \left\{ X \in \bigcup_{\ell=1}^{b'_n} J_\ell \right\} \delta_{(X, f(X, \mathcal{P}_n))}(A) \right. \right. \\ & \quad \left. \left. - \sum_{X \in \mathcal{B}_n} g(X, \mathcal{B}_n) \mathbb{1} \left\{ X \in \bigcup_{\ell=1}^{b'_n} J_\ell \right\} \delta_{(X, f(X, \mathcal{B}_n))}(A) \right| \geq \varepsilon_0 b_n \right) \rightarrow -\infty. \end{aligned}$$

We say that J_i is n -bad if one of the following events happens.

- (i) There exists $X \in \mathcal{P}_n \cap J_i$ such that $g(X, \mathcal{P}_n) = 1$ (equivalently, $\mathcal{P}_n(B_{r_n(s_0)}(X)) \leq k$; see (4.5)) and $X \notin \mathcal{B}_n$.
- (ii) There exists $X \in \mathcal{B}_n \cap J_i$ such that $g(X, \mathcal{B}_n) = 1$ and $X \notin \mathcal{P}_n$.
- (iii) There exist $X \in \mathcal{P}_n \cap \mathcal{B}_n \cap J_i$ and $u \geq s_0$ such that $\min \{ \mathcal{P}_n(B_{r_n(u)}(X)), \mathcal{B}_n(B_{r_n(u)}(X)) \} \leq k$ and $\max \{ \mathcal{P}_n(B_{r_n(u)}(X)), \mathcal{B}_n(B_{r_n(u)}(X)) \} > k$.

The key observation here is that

$$\begin{aligned} & \sup_{A \subset E_0} \left| \sum_{X \in \mathcal{P}_n} g(X, \mathcal{P}_n) \mathbb{1} \left\{ X \in \bigcup_{\ell=1}^{b'_n} J_\ell \right\} \delta_{(X, f(X, \mathcal{P}_n))}(A) \right. \\ & \quad \left. - \sum_{X \in \mathcal{B}_n} g(X, \mathcal{B}_n) \mathbb{1} \left\{ X \in \bigcup_{\ell=1}^{b'_n} J_\ell \right\} \delta_{(X, f(X, \mathcal{B}_n))}(A) \right| \leq (k+1) \sum_{i=1}^{b'_n} \mathbb{1} \{ J_i \text{ is } n\text{-bad} \}. \end{aligned}$$

Thus, it is enough to show that for every $\varepsilon_0 > 0$,

$$\frac{1}{b_n} \log \mathbb{P} \left(\sum_{i=1}^{b'_n} \mathbb{1} \{ J_i \text{ is } n\text{-bad} \} \geq \varepsilon_0 b_n \right) \rightarrow -\infty.$$

For $\eta \in (0, 1]$, let $\mathcal{P}_n^{(\eta)}$ be a homogeneous Poisson point process on $[0, 1]^d$ with intensity $n\eta$. We take $\mathcal{P}_n^{(\eta)}$ to be independent of \mathcal{P}_n . Then, $\mathcal{P}_n^{(\eta, a)} := \mathcal{P}_n \cup \mathcal{P}_n^{(\eta)}$ represents the *augmented* Poisson point process with intensity $n(1 + \eta)$. Moreover, let $\mathcal{P}_n^{(\eta, t)}$ denote a *thinned* version of \mathcal{P}_n obtained by removing each point of \mathcal{P}_n with probability η . If we denote by $\mathcal{D}_\eta(\mathcal{P}_n)$ a collection of deleted points of \mathcal{P}_n , one can write $\mathcal{P}_n^{(\eta, t)} = \mathcal{P}_n \setminus \mathcal{D}_\eta(\mathcal{P}_n)$. Notice that $\mathcal{P}_n^{(\eta, a)} \stackrel{d}{=} \mathcal{P}_{n(1+\eta)}$ and $\mathcal{P}_n^{(\eta, t)} \stackrel{d}{=} \mathcal{P}_{n(1-\eta)}$. Subsequently, for $\varepsilon > 0$ let

$$F_{n,\varepsilon} = \{\mathcal{P}_n^{(\varepsilon a_n^{-1}, t)} \subset \mathcal{B}_n \subset \mathcal{P}_n^{(\varepsilon a_n^{-1}, a)}\} \tag{4.31}$$

and claim that

$$b_n^{-1} \log \mathbb{P}(F_{n,\varepsilon}^c) \rightarrow -\infty, \quad n \rightarrow \infty. \tag{4.32}$$

For the proof we use Lemma 1.2 in [18] to get that

$$\mathbb{P}(F_{n,\varepsilon}^c) \leq e^{-n(1+\varepsilon a_n^{-1})H((1+\varepsilon a_n^{-1})^{-1})} + e^{-n(1-\varepsilon a_n^{-1})H((1-\varepsilon a_n^{-1})^{-1})}, \tag{4.33}$$

where $H(x) = x \log x + 1 - x$, $x > 0$. Applying the Taylor expansion to $H(\cdot)$, we have

$$\limsup_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{P}(F_{n,\varepsilon}^c) \leq - \lim_{n \rightarrow \infty} \frac{e^{a_n}}{2a_n^{k-1}} \left(\frac{\varepsilon a_n^{-1}}{1 - \varepsilon a_n^{-1}} \right)^2 = -\infty.$$

Suppose now that J_i is an n -bad cube and $F_{n,\varepsilon}$ holds, such that one of the events in case (i)–(iii) above occurs. Then, there exists $X \in \mathcal{P}_n^{(\varepsilon a_n^{-1}, a)} \cap J_i$, such that

$$\mathcal{P}_n^{(\varepsilon a_n^{-1}, t)}(B_{r_n(s_0)}(X)) \leq k, \quad \text{and} \quad (\mathcal{P}_n^{(\varepsilon a_n^{-1}, a)} \setminus \mathcal{P}_n^{(\varepsilon a_n^{-1}, t)})(B_{r_n(w_n)}(X)) \geq 1. \tag{4.34}$$

Since we work with the diluted cubes in (4.30), it follows from (4.34) and the spatial independence of Poisson processes that $\sum_{i=1}^{b_n} \mathbb{1}\{J_i \text{ is } n\text{-bad}\}$ is a binomial random variable. Below, we shall estimate its success probability $p_{n,\varepsilon}$ as follows:

$$\begin{aligned} p_{n,\varepsilon} &= \mathbb{P} \left(\bigcup_{X \in \mathcal{P}_n^{(\varepsilon a_n^{-1}, a)} \cap J_1} \left\{ \mathcal{P}_n^{(\varepsilon a_n^{-1}, t)}(B_{r_n(s_0)}(X)) \leq k, \right. \right. \\ &\quad \left. \left. (\mathcal{P}_n^{(\varepsilon a_n^{-1}, a)} \setminus \mathcal{P}_n^{(\varepsilon a_n^{-1}, t)})(B_{r_n(w_n)}(X)) \geq 1 \right\} \right) \\ &\leq \mathbb{E} \left[\sum_{X \in \mathcal{P}_n \cap J_1} \mathbb{1} \left\{ (\mathcal{P}_n \setminus \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n))(B_{r_n(s_0)}(X)) \leq k, \right. \right. \\ &\quad \left. \left. (\mathcal{P}_n^{(\varepsilon a_n^{-1})} \cup \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n))(B_{r_n(w_n)}(X)) \geq 1 \right\} \right] \\ &\quad + \mathbb{E} \left[\sum_{X \in \mathcal{P}_n^{(\varepsilon a_n^{-1})} \cap J_1} \mathbb{1} \left\{ (\mathcal{P}_n \setminus \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n))(B_{r_n(s_0)}(X)) \leq k \right\} \right] \\ &=: A_n + B_n. \end{aligned} \tag{4.35}$$

By the independence of $\mathcal{P}_n^{(\varepsilon a_n^{-1})}$ and \mathcal{P}_n ,

$$\begin{aligned} B_n &= n\varepsilon a_n^{-1} \mathbb{P} \left(Y \in J_1, (\mathcal{P}_n \setminus \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n))(B_{r_n(s_0)}(Y)) \leq k \right) \\ &= n\varepsilon a_n^{-1} \text{Leb}(J_1) \mathbb{P} \left(\mathcal{P}_{n(1-\varepsilon a_n^{-1})}(B_{r_n(s_0)}(Y)) \leq k \right) \\ &= n\varepsilon a_n^{-1} \text{Leb}(J_1) \sum_{i=0}^k \frac{(a_n(1 - \varepsilon a_n^{-1})(1 + s_0 a_n^{-1}))^i}{i!} e^{-a_n(1-\varepsilon a_n^{-1})(1+s_0 a_n^{-1})}. \end{aligned} \tag{4.36}$$

In the above and the following, Y is a uniform random variable on $[0, 1]^d$ that is independent of \mathcal{P}_n . Because of

$$a_n(1 - \varepsilon a_n^{-1})(1 + s_0 a_n^{-1}) \geq a_n - C^*, \tag{4.37}$$

we get that

$$B_n \leq C^* n \varepsilon a_n^{-1} \text{Leb}(J_1) a_n^k e^{-a_n} = C^* \varepsilon \text{Leb}(J_1) b_n.$$

Applying the Mecke formula for Poisson point processes,

$$A_n = n \mathbb{P} \left(Y \in J_1, \{(\mathcal{P}_n + \delta_Y) \setminus \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n + \delta_Y)\} (B_{r_n(s_0)}(Y)) \leq k, \right. \\ \left. \{ \mathcal{P}_n^{(\varepsilon a_n^{-1})} \cup \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n + \delta_Y) \} (B_{r_n(w_n)}(Y)) \geq 1 \right).$$

Denote $T(Y) = \{Y \text{ is deleted by thinning}\}$. Note that if $T(Y)$ holds, then $\mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n + \delta_Y) = \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n) + \delta_Y$ and if $T(Y)$ does not hold, we have $\mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n + \delta_Y) = \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n)$. Hence,

$$A_n \leq n \mathbb{P} \left(\left\{ Y \in J_1, (\mathcal{P}_n \setminus \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n)) (B_{r_n(s_0)}(Y)) \leq k \right\} \cap T(Y) \right) \\ + n \mathbb{P} \left(Y \in J_1, (\mathcal{P}_n \setminus \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n)) (B_{r_n(s_0)}(Y)) \leq k - 1, \right. \\ \left. (\mathcal{P}_n^{(\varepsilon a_n^{-1})} \cup \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n)) (B_{r_n(w_n)}(Y)) \geq 1 \right) =: C_n + D_n.$$

Repeating the same calculation as in (4.36) and using (4.37),

$$C_n \leq C^* n a_n^k e^{-a_n} \mathbb{P}(\{Y \in J_1\} \cap T(Y)) = C^* \varepsilon \text{Leb}(J_1) b_n.$$

Since $\mathcal{P}_n \setminus \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n)$ and $\mathcal{P}_n^{(\varepsilon a_n^{-1})} \cup \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n)$ are independent (see, e.g., Corollary 5.9 in [12]),

$$D_n = n \text{Leb}(J_1) \mathbb{P} \left((\mathcal{P}_n \setminus \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n)) (B_{r_n(s_0)}(Y)) \leq k - 1 \right) \\ \times \mathbb{P} \left((\mathcal{P}_n^{(\varepsilon a_n^{-1})} \cup \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n)) (B_{r_n(w_n)}(Y)) \geq 1 \right) \\ = n \text{Leb}(J_1) \mathbb{P} \left(\mathcal{P}_{n(1-\varepsilon a_n^{-1})} (B_{r_n(s_0)}(Y)) \leq k - 1 \right) \mathbb{P} \left(\mathcal{P}_{2n\varepsilon a_n^{-1}} (B_{r_n(w_n)}(Y)) \geq 1 \right).$$

By (4.37) and Markov's inequality,

$$D_n \leq C^* n \text{Leb}(J_1) a_n^{k-1} e^{-a_n} \mathbb{E} \left[\mathcal{P}_{2n\varepsilon a_n^{-1}} (B_{r_n(w_n)}(Y)) \right] \\ = C^* n \text{Leb}(J_1) a_n^{k-1} e^{-a_n} 2\varepsilon(1 + w_n a_n^{-1}) \leq C^* \varepsilon \text{Leb}(J_1) b_n.$$

Combining all these calculations concludes that $p_{n,\varepsilon} \leq C^* \varepsilon \text{Leb}(J_1) b_n$. If one takes sufficiently small $\varepsilon \in (0, \varepsilon_0)$, we have that $b'_n p_{n,\varepsilon} \leq C^* b'_n \varepsilon \text{Leb}(J_1) b_n = C^* \varepsilon b_n \leq \varepsilon_0 b_n$ for large n enough. Therefore, one can exploit the binomial concentration inequality (see, e.g., Lemma 1.1 in [18]) to obtain that

$$\limsup_{n \rightarrow \infty} b_n^{-1} \log \mathbb{P} \left(\text{Bin}(b'_n, C^* \varepsilon \text{Leb}(J_1) b_n) \geq \varepsilon_0 b_n \right) \leq -\frac{\varepsilon_0}{2} \lim_{n \rightarrow \infty} \log \left\{ \frac{\varepsilon_0}{C^* b'_n \varepsilon \text{Leb}(J_1)} \right\} \\ = -\frac{\varepsilon_0}{2} \log \left\{ \frac{(3\sqrt{d})^d \varepsilon_0}{C^* \varepsilon} \right\}.$$

The last term goes to $-\infty$ as $\varepsilon \rightarrow 0$. Combining this result with (4.32) concludes the proof. \square

Proof of Corollary 2.3. Define the map $S : M_+(E_0) \rightarrow [0, \infty)$ by $S(\rho) = \rho(E_0)$. Since S is continuous in the weak topology and $H_k(\cdot | \text{Leb} \otimes \tau_k) = \Lambda_k^*$ is a good rate function, the contraction principle (see, e.g., [3, Theorem 4.2.1]) is applied to the LDP in Theorem 2.1. In conclusion,

$$S\left(\frac{L_{k,n}}{b_n}\right) = \frac{1}{b_n} \sum_{X \in \mathcal{P}_n} g(X, \mathcal{P}_n) = \frac{T_{k,n}}{b_n}, \quad n \geq 1,$$

shows an LDP with speed b_n and rate function

$$\inf_{\nu \in M_+(E_0), \nu(E_0)=x} H_k(\nu | \text{Leb} \otimes \tau_k), \quad x \in \mathbb{R}, \tag{4.38}$$

where H_k is the relative entropy defined at (2.3). The rest of the argument must be devoted to verifying that (4.38) coincides with $I_k(x)$ for every $x \in \mathbb{R}$; this is however an immediate result as an analogue of Equ. (5.35) of [8].

The LDP for $(T_{k,n}^B/b_n)_{n \geq 1}$ is obtained by applying the contraction principle to Corollary 2.2. \square

4.2 Proofs of Theorem 3.1, Corollary 3.2, and Corollary 3.3

Proof of Theorem 3.1. Let $C_K^+(E)$ denote a collection of continuous and non-negative functions on E with compact support. Given two such functions $U_\ell \in C_K^+(E)$, $\ell = 1, 2$, together with $\varepsilon_1, \varepsilon_2 > 0$, we define $F_{U_1, U_2, \varepsilon_1, \varepsilon_2} : M_p(E) \rightarrow [0, 1]$ by

$$F_{U_1, U_2, \varepsilon_1, \varepsilon_2}(\eta) = \left(1 - e^{-(\eta(U_1) - \varepsilon_1)_+}\right) \left(1 - e^{-(\eta(U_2) - \varepsilon_2)_+}\right), \tag{4.39}$$

where $\eta(U_\ell) = \int_E U_\ell(x, u) \eta(dx, du)$, and $(a)_+ = a$ if $a \geq 0$ and 0 otherwise. Notice that $F_{U_1, U_2, \varepsilon_1, \varepsilon_2} \in \mathcal{C}_0$. In what follows, we fix U_1, U_2 and $\varepsilon_1, \varepsilon_2$, and simply write $F = F_{U_1, U_2, \varepsilon_1, \varepsilon_2}$. Define $\xi_{k,n}(\cdot) := b_n^{-1} \mathbb{P}(L_{k,n} \in \cdot)$. Then, according to Theorem A.2 in [10], (3.1) follows if one can show that

$$\xi_{k,n}(F) \rightarrow \xi_k(F), \quad \text{as } n \rightarrow \infty.$$

First, note that

$$\xi_{k,n}(F) = \int_{M_p(E)} F(\eta) \xi_{k,n}(d\eta) = b_n^{-1} \mathbb{E}[F(L_{k,n})].$$

Let $\zeta_{k,n}$ denote a Poisson point process on E with mean measure

$$\frac{b_n}{(k-1)!} e^{-u} dx du, \quad x \in [0, 1]^d, \quad u \in \mathbb{R}. \tag{4.40}$$

In this setting, our proof breaks down into two parts:

$$b_n^{-1} \left| \mathbb{E}[F(L_{k,n})] - \mathbb{E}[F(\zeta_{k,n})] \right| \rightarrow 0, \quad n \rightarrow \infty, \tag{4.41}$$

$$b_n^{-1} \mathbb{E}[F(\zeta_{k,n})] \rightarrow \xi_k(F), \quad n \rightarrow \infty. \tag{4.42}$$

Proof of (4.41): Since U_ℓ has compact support on E , there exists $s_0 \in \mathbb{R}$, so that

$$\text{supp}(U_1) \cup \text{supp}(U_2) \subset [0, 1]^d \times (s_0, \infty],$$

where $\text{supp}(U_\ell)$ represents the support of U_ℓ . Hence, we may assume, without loss of generality, that $L_{k,n}$ and $\zeta_{k,n}$ are both random elements of the *restricted* state space $M_p([0, 1]^d \times (s_0, \infty])$. Equivalently, one can reformulate $L_{k,n}$ by

$$L_{k,n} = \begin{cases} \sum_{X \in \mathcal{P}_n} g(X, \mathcal{P}_n) \delta_{(X, f(X, \mathcal{P}_n))} & \text{if } |\mathcal{P}_n| > k, \\ \emptyset & \text{if } |\mathcal{P}_n| \leq k, \end{cases} \tag{4.43}$$

in the same way as (4.1). Similarly, $\zeta_{k,n}$ can be defined as the Poisson point process whose mean measure is given by the restricted version of (4.40); that is,

$$(\text{Leb} \otimes \tau_{k,n})(dx, du) := \frac{b_n}{(k-1)!} e^{-u} \mathbb{1}\{u \geq s_0\} dx du, \quad x \in [0, 1]^d, u \in \mathbb{R}.$$

Next, it is not hard to prove that F in (4.39) is a 1-Lipschitz function with respect to the total variation distance on the space of point measures. Namely, for $\eta_1, \eta_2 \in M_p(E)$,

$$|F(\eta_1) - F(\eta_2)| \leq 2d_{\text{TV}}(\eta_1, \eta_2).$$

Thus, by (4.4),

$$|\mathbb{E}[F(L_{k,n})] - \mathbb{E}[F(\zeta_{k,n})]| \leq d_{\text{KR}}(\mathcal{L}(L_{k,n}), \mathcal{L}(\zeta_{k,n})).$$

Proposition 4.7. *We have, as $n \rightarrow \infty$,*

$$b_n^{-1} d_{\text{KR}}(\mathcal{L}(L_{k,n}), \mathcal{L}(\zeta_{k,n})) \rightarrow 0.$$

Proof of Proposition 4.7. The proof is analogous to that of Proposition 4.1. Precisely, we first need to show that

$$b_n^{-1} d_{\text{TV}}(\mathbb{E}[L_{k,n}(\cdot)], \text{Leb} \otimes \tau_{k,n}) \rightarrow 0, \quad n \rightarrow \infty,$$

and verify also that $b_n^{-1} E_i \rightarrow 0$, $n \rightarrow \infty$, for $i = 1, 2, 3$, where E_i 's are defined analogously to (4.8), (4.9), and (4.10). More concretely, they are respectively defined as

$$E_1 := 2n \int_{[0,1]^d} \mathbb{E}[g(x, \mathcal{P}_n + \delta_x) \mathbb{1}\{\mathcal{S}(x, \mathcal{P}_n + \delta_x) \not\subseteq S_x\}] dx,$$

$$E_2 := 2n^2 \int_{[0,1]^d} \int_{[0,1]^d} \mathbb{1}\{S_x \cap S_z \neq \emptyset\} \mathbb{E}[g(x, \mathcal{P}_n + \delta_x)] \mathbb{E}[g(z, \mathcal{P}_n + \delta_z)] dx dz,$$

and

$$E_3 := 2n^2 \int_{[0,1]^d} \int_{[0,1]^d} \mathbb{1}\{S_x \cap S_z \neq \emptyset\} \mathbb{E}[g(x, \mathcal{P}_n + \delta_x + \delta_z) g(z, \mathcal{P}_n + \delta_x + \delta_z)] dx dz,$$

for which $\mathcal{S}(x, \omega) = B_{R_k(x, \omega)}(x)$ for $x \in [0, 1]^d$ and $\omega \in M_p([0, 1]^d)$, and $S_x = B_{r_n(w_n)}(x)$ for some sequence $w_n \rightarrow \infty$ with $w_n = o(a_n)$, $n \rightarrow \infty$.

First, for $B \subset [0, 1]^d$ and $u > s_0$, by the Mecke formula for Poisson point processes and (4.5),

$$\mathbb{E}[L_{k,n}(B \times (u, \infty))] = n\mathbb{P}(Y \in B, (\mathcal{P}_n + \delta_Y)(B_{r_n(u)}(Y)) \leq k),$$

where Y is a uniform random variable on $[0, 1]^d$, independent of \mathcal{P}_n . By the conditioning on Y ,

$$\mathbb{E}[L_{k,n}(B \times (u, \infty))] = n \text{Leb}(B) \sum_{i=0}^{k-1} e^{-(a_n+u)} \frac{(a_n+u)^i}{i!}.$$

This means that $\mathbb{E}[L_{k,n}(\cdot)]$ has the density

$$n \frac{e^{-(a_n+u)} (a_n+u)^{k-1}}{(k-1)!}, \quad x \in [0, 1]^d, u > s_0,$$

and hence, it follows from the dominated convergence theorem that

$$\begin{aligned} & b_n^{-1} d_{\text{TV}}(\mathbb{E}[L_{k,n}(\cdot)], \text{Leb} \otimes \tau_{k,n}) \\ & \leq b_n^{-1} \int_{[0,1]^d \times (s_0, \infty)} \left| n e^{-(a_n+u)} \frac{(a_n+u)^{k-1}}{(k-1)!} - b_n \frac{e^{-u}}{(k-1)!} \right| dx du \\ & = \frac{1}{(k-1)!} \int_{s_0}^{\infty} \left| \left(1 + \frac{u}{a_n}\right)^{k-1} - 1 \right| e^{-u} du \rightarrow 0, \quad n \rightarrow \infty. \end{aligned}$$

Subsequently,

$$\begin{aligned} b_n^{-1} E_1 & \leq 2nb_n^{-1} \int_{[0,1]^d} \mathbb{P}(\mathcal{P}_n(B_{r_n(w_n)}(x)) \leq k-1) dx \\ & = 2nb_n^{-1} \sum_{i=0}^{k-1} e^{-(a_n+w_n)} \frac{(a_n+w_n)^i}{i!} \leq C^* e^{-w_n} \rightarrow 0, \quad n \rightarrow \infty, \end{aligned}$$

while we also have

$$b_n^{-1} E_2 \leq 2n^2 b_n^{-1} \left\{ \sum_{i=0}^{k-1} e^{-(a_n+s_0)} \frac{(a_n+s_0)^i}{i!} \right\}^2 \leq C^* b_n \rightarrow 0, \quad n \rightarrow \infty.$$

Similarly to (4.12), E_3 can be split into two additional terms:

$$\begin{aligned} E_3 & \leq 2n^2 \int_{[0,1]^d} \int_{[0,1]^d} \mathbb{1}\{\|x-z\| \leq r_n(s_0)\} \\ & \quad \times \mathbb{P}\left((\mathcal{P}_n + \delta_z)(B_{r_n(s_0)}(x)) \leq k-1, (\mathcal{P}_n + \delta_x)(B_{r_n(s_0)}(z)) \leq k-1\right) dx dz \\ & + 2n^2 \int_{[0,1]^d} \int_{[0,1]^d} \mathbb{1}\{r_n(s_0) < \|x-z\| \leq 2r_n(w_n)\} \\ & \quad \times \mathbb{P}\left((\mathcal{P}_n + \delta_z)(B_{r_n(s_0)}(x)) \leq k-1, (\mathcal{P}_n + \delta_x)(B_{r_n(s_0)}(z)) \leq k-1\right) dx dz \\ & =: E_{3,1} + E_{3,2}. \end{aligned}$$

Although we shall skip detailed discussions, one can still demonstrate that $b_n^{-1} E_{3,1} \rightarrow 0$ and $b_n^{-1} E_{3,2} \rightarrow 0$, by the arguments nearly identical to those for Proposition 4.1. \square

Now, the proof of Proposition 4.7 has been completed, which in turn concludes (4.41). Our next goal is to prove (4.42).

Proof of (4.42): Note that $\zeta_{k,n}$ can be written as

$$\zeta_{k,n} = \sum_{i=1}^{N_n} \delta_{(T_i, Z_i)},$$

where (T_i, Z_i) are i.i.d. random variables on E with density given by $e^{-(u-s_0)} \mathbb{1}\{u \geq s_0\} dx du$, and N_n is Poisson distributed with mean $b_n e^{-s_0} / (k-1)!$. Furthermore, (T_i, Z_i) and N_n are taken to be independent. Substituting this representation,

$$\begin{aligned} b_n^{-1} \mathbb{E}[F(\zeta_{k,n})] & = b_n^{-1} \mathbb{E}\left[\prod_{\ell=1}^2 \left(1 - e^{-\left(\sum_{i=1}^{N_n} U_\ell(T_i, Z_i) - \varepsilon_\ell\right)_+}\right)\right] \\ & = b_n^{-1} \mathbb{E}\left[\prod_{\ell=1}^2 \left(1 - e^{-\left(U_\ell(T_1, Z_1) - \varepsilon_\ell\right)_+}\right) \mathbb{1}\{N_n = 1\}\right] \\ & \quad + b_n^{-1} \mathbb{E}\left[\prod_{\ell=1}^2 \left(1 - e^{-\left(\sum_{i=1}^{N_n} U_\ell(T_i, Z_i) - \varepsilon_\ell\right)_+}\right) \mathbb{1}\{N_n \geq 2\}\right] \\ & =: A_n + B_n. \end{aligned}$$

Of the last two terms, one can immediately show that

$$B_n \leq b_n^{-1} \mathbb{P}(N_n \geq 2) \leq \left(\frac{e^{-s_0}}{(k-1)!} \right)^2 b_n \rightarrow 0, \text{ as } n \rightarrow \infty.$$

By the independence of (T_1, Z_1) and N_n , we have as $n \rightarrow \infty$,

$$\begin{aligned} A_n &= b_n^{-1} \mathbb{E} \left[\prod_{\ell=1}^2 \left(1 - e^{-\left(U_\ell(T_1, Z_1) - \varepsilon_\ell \right)_+} \right) \right] \mathbb{P}(N_n = 1) \\ &= \frac{e^{-s_0}}{(k-1)!} e^{-e^{-s_0} b_n / (k-1)!} \int_E \prod_{\ell=1}^2 \left(1 - e^{-\left(U_\ell(x, u) - \varepsilon_\ell \right)_+} \right) e^{-(u-s_0)} \mathbb{1}\{u \geq s_0\} dx du \\ &\rightarrow \frac{1}{(k-1)!} \int_E \prod_{\ell=1}^2 \left(1 - e^{-\left(U_\ell(x, u) - \varepsilon_\ell \right)_+} \right) e^{-u} dx du = \xi_k(F). \end{aligned}$$

We thus conclude that $A_n + B_n \rightarrow \xi_k(F)$, $n \rightarrow \infty$, as required. □

Proof of Corollary 3.2. Because of (4.41) and (4.42), it is sufficient to show that

$$b_n^{-1} \mathbb{E} \left[\left| F(L_{k,n}) - F(L_{k,n}^B) \right| \right] \rightarrow 0, \quad n \rightarrow \infty,$$

where F is defined at (4.39). Under the map F , one can represent $L_{k,n}$ as in (4.43). Clearly, $L_{k,n}^B$ has the same representation as an element of $M_p([0, 1]^d \times (s_0, \infty])$. Since F is bounded,

$$b_n^{-1} \mathbb{E} \left[\left| F(L_{k,n}) - F(L_{k,n}^B) \right| \right] \leq 2b_n^{-1} \mathbb{P}(L_{k,n} \neq L_{k,n}^B).$$

We now claim that $b_n^{-1} \mathbb{P}(L_{k,n} \neq L_{k,n}^B) \rightarrow 0$ as $n \rightarrow \infty$. The proof is analogous to that of (4.29) by borrowing the idea of n -bad cubes. Specifically, we say that $[0, 1]^d$ is n -bad if one of the following events occurs.

- (i) There exists $X \in \mathcal{P}_n$ such that $g(X, \mathcal{P}_n) = 1$ and $X \notin \mathcal{B}_n$.
- (ii) There exists $X \in \mathcal{B}_n$ such that $g(X, \mathcal{B}_n) = 1$ and $X \notin \mathcal{P}_n$.
- (iii) There exist $X \in \mathcal{P}_n \cap \mathcal{B}_n$ and $u \geq s_0$ such that $\min \{ \mathcal{P}_n(B_{r_n(u)}(X)), \mathcal{B}_n(B_{r_n(u)}(X)) \} \leq k$ and $\max \{ \mathcal{P}_n(B_{r_n(u)}(X)), \mathcal{B}_n(B_{r_n(u)}(X)) \} > k$.

The key observation is that $[0, 1]^d$ becomes n -bad whenever $L_{k,n} \neq L_{k,n}^B$. Using this fact, we now need to show that

$$b_n^{-1} \mathbb{P}([0, 1]^d \text{ is } n\text{-bad}) \rightarrow 0, \quad n \rightarrow \infty. \tag{4.44}$$

The first step for the proof of (4.44) is to demonstrate that

$$b_n^{-1} \mathbb{P}(F_{n,\varepsilon}^c) \rightarrow 0, \quad n \rightarrow \infty, \tag{4.45}$$

where $F_{n,\varepsilon}$ is given in (4.31). By virtue of the bound in (4.33), it follows that

$$\begin{aligned} b_n^{-1} \mathbb{P}(F_{n,\varepsilon}^c) &\leq b_n^{-1} \left(e^{-n(1+\varepsilon a_n^{-1})H((1+\varepsilon a_n^{-1})^{-1})} + e^{-n(1-\varepsilon a_n^{-1})H((1-\varepsilon a_n^{-1})^{-1})} \right) \\ &= \frac{e^{a_n}}{n a_n^{k-1}} \left(e^{n \log(1+\varepsilon a_n^{-1}) - n \varepsilon a_n^{-1}} + e^{n \log(1-\varepsilon a_n^{-1}) + n \varepsilon a_n^{-1}} \right). \end{aligned}$$

Applying the Taylor expansion of the logarithm function at 1, the first term on the last expression is bounded by $(n a_n^{k-1})^{-1} e^{a_n - C n a_n^{-2}}$ for some constant $C > 0$. Under the assumption $a_n = o(n^{1/3})$, this quantity converges to 0 as n goes to infinity. By the same argument, the other term also goes to 0 under the same assumption on a_n . Therefore,

one can get (4.45) as desired. Observe also that if $[0, 1]^d$ is n -bad under $F_{n,\varepsilon}$, then there exists $X \in \mathcal{P}_n^{(\varepsilon a_n^{-1}, a)}$ such that (4.34) holds. Hence, by (4.45) and Markov's inequality,

$$\begin{aligned} b_n^{-1} \mathbb{P}([0, 1]^d \text{ is } n\text{-bad}) &\leq b_n^{-1} \mathbb{P}(\{[0, 1]^d \text{ is } n\text{-bad}\} \cap F_{n,\varepsilon}) + o(1) \\ &\leq b_n^{-1} \mathbb{P}\left(\bigcup_{X \in \mathcal{P}_n^{(\varepsilon a_n^{-1}, a)}} \left\{ \mathcal{P}_n^{(\varepsilon a_n^{-1}, t)}(B_{r_n(s_0)}(X)) \leq k, \right. \right. \\ &\quad \left. \left. (\mathcal{P}_n^{(\varepsilon a_n^{-1}, a)} \setminus \mathcal{P}_n^{(\varepsilon a_n^{-1}, t)})(B_{r_n(w_n)}(X)) \geq 1 \right\}\right) + o(1) \\ &\leq b_n^{-1} \left\{ \mathbb{E}\left[\sum_{X \in \mathcal{P}_n} \mathbb{1}\left\{ (\mathcal{P}_n \setminus \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n))(B_{r_n(s_0)}(X)) \leq k, \right. \right. \right. \\ &\quad \left. \left. (\mathcal{P}_n^{(\varepsilon a_n^{-1})} \cup \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n))(B_{r_n(w_n)}(X)) \geq 1 \right\} \right] \right. \\ &\quad \left. + \mathbb{E}\left[\sum_{X \in \mathcal{P}_n^{(\varepsilon a_n^{-1})}} \mathbb{1}\left\{ (\mathcal{P}_n \setminus \mathcal{D}_{\varepsilon a_n^{-1}}(\mathcal{P}_n))(B_{r_n(s_0)}(X)) \leq k \right\} \right] \right\} + o(1) \\ &=: b_n^{-1} (A'_n + B'_n) + o(1). \end{aligned}$$

where $\varepsilon \in (0, 1)$ is an arbitrary constant. Repeating the calculations very similar to those bounding A_n, B_n in (4.35), one can see that $A'_n + B'_n \leq C^* \varepsilon b_n$. Thus,

$$\limsup_{n \rightarrow \infty} b_n^{-1} \mathbb{P}([0, 1]^d \text{ is } n\text{-bad}) \leq C^* \varepsilon,$$

and letting $\varepsilon \rightarrow 0$ completes the proof of Corollary 3.2. □

Proof of Corollary 3.3. We prove only the first statement. By a straightforward modification of Theorem 3.1 by restricting the state space from E to $E_0 = [0, 1]^d \times (s_0, \infty]$, we have, as $n \rightarrow \infty$,

$$\xi_{k,n}(\cdot) := b_n^{-1} \mathbb{P}(L_{k,n} \in \cdot) \rightarrow \xi_k \text{ in } \mathcal{M}_0.$$

Due to the change of the state space, $L_{k,n}$ is now formulated as in (4.43), while the limit ξ_k is taken to be

$$\xi_k(\cdot) = \frac{1}{(k-1)!} \int_{E_0} \mathbb{1}\{\delta_{(x,u)} \in \cdot\} e^{-u} dx du.$$

Now, we define a map $V : M_p(E_0) \rightarrow \mathbb{N} := \{0, 1, 2, \dots\}$ by $V(\rho) = \rho(E_0)$. Here, \mathbb{N} is equipped with the discrete topology. Since V is continuous in the weak topology, it follows from [9, Theorem 2.5] that

$$\xi_{k,n} \circ V^{-1} \rightarrow \xi_k \circ V^{-1}, \text{ in } \mathcal{M}_0, \quad n \rightarrow \infty. \tag{4.46}$$

Note that $\mathbb{1}_{[1,\infty)}(x)$ is continuous and bounded on \mathbb{N} (in terms of the discrete topology), vanishing in the neighborhood of 0 (i.e., the origin of \mathbb{N}). Thus, the \mathcal{M}_0 -convergence in (4.46) implies that

$$b_n^{-1} \mathbb{P}(T_{k,n} \geq 1) = \int_{\mathbb{N}} \mathbb{1}_{[1,\infty)}(x) \xi_{k,n} \circ V^{-1}(dx) \rightarrow \int_{\mathbb{N}} \mathbb{1}_{[1,\infty)}(x) \xi_k \circ V^{-1}(dx) = \alpha_k,$$

as desired. □

References

- [1] O. Bobrowski, M. Schulte, and D. Yogeshwaran, *Poisson process approximation under stabilization and Palm coupling*, arXiv:2104.13261, 2021. MR4526259

- [2] N. Chenavier, N. Henze, and M. Otto, *Limit laws for large k th-nearest neighbor balls*, Journal of Applied Probability **59** (2022), 880–894. MR4480085
- [3] A. Dembo and O. Zeitouni, *Large deviations techniques and applications*, Springer, New York, 1998. MR2571413
- [4] V. Fasen and P. Roy, *Stable random fields, point processes and large deviations*, Stochastic Processes and their Applications **126** (2016), 832–856. MR3452814
- [5] L. Györfi, N. Henze, and H. Walk, *The limit distribution of the maximum probability nearest-neighbour ball*, Journal of Applied Probability **56** (2019), 574–589. MR3986953
- [6] C. Hirsch, B. Jahnel, and R. Patterson, *Space-time large deviations in capacity-constrained relay networks*, ALEA. Latin American Journal of Probability and Mathematical Statistics **15** (2018), 587–615. MR3808899
- [7] C. Hirsch, B. Jahnel, and A. Tóbiás, *Lower large deviations for geometric functionals*, Electronic Communications in Probability **25** (2020), 1–12. MR4112772
- [8] C. Hirsch and T. Owada, *Large deviation principle for geometric and topological functionals and associated point processes*, arXiv:2201.07276, 2022.
- [9] H. Hult and F. Lindskog, *Regular variation for measures on metric spaces*, Publications de l’Institut Mathématique **80** (2006), 121–140. MR2281910
- [10] H. Hult and G. Samorodnitsky, *Large deviations for point processes based on stationary sequences with heavy tails*, Journal of Applied Probability **47** (2010), 1–40. MR2654756
- [11] O. Kallenberg, *Random measures, theory and applications*, Springer, Cham, 2017. MR3642325
- [12] G. Last and M. D. Penrose, *Lectures on the poisson process*, Cambridge University Press, 2017. MR3791470
- [13] F. Lindskog, S. I. Resnick, and J. Roy, *Regularly varying measures on metric spaces: Hidden regular variation and hidden jumps*, Probability Surveys **11** (2014), 270–314. MR3271332
- [14] M. Otto, *Poisson approximation of Poisson-driven point processes and extreme values in stochastic geometry*, arXiv:2005.10116, 2020.
- [15] T. Owada, *Limit theory for U -statistics under geometric and topological constraints with rare events*, To appear in *Journal of Applied Probability*, 2022. MR4546125
- [16] M. D. Penrose, *The longest edge of the random minimal spanning tree*, The Annals of Applied Probability **7** (1997), 340–361. MR1442317
- [17] M. D. Penrose, *Central limit theorems for k -nearest neighbour distances*, Stochastic Processes and their Applications **85** (2000), 295–320. MR1731028
- [18] M. D. Penrose, *Random geometric graphs, oxford studies in probability 5*, Oxford University Press, Oxford, 2003. MR1986198
- [19] M. D. Penrose, *Gaussian limits for random geometric measures*, Electronic Journal of Probability **12** (2007), 989–1035. MR2336596
- [20] M. D. Penrose, *Laws of large numbers in stochastic geometry with statistical applications*, Bernoulli **13** (2007), 1124–1150. MR2364229
- [21] M. D. Penrose and L. Goldstein, *Normal approximation for coverage models over binomial point processes*, The Annals of Applied Probability **20** (2010), 696–721. MR2650046
- [22] M. D. Penrose and J. E. Yukich, *Weak laws of large numbers in geometric probability*, The Annals of Applied Probability **13** (2003), 277–303. MR1952000
- [23] M. D. Penrose and J. E. Yukich, *Laws of large numbers and nearest neighbor distances*, Advances in Directional and Linear Statistics (M.T. Wells and A. SenGupta, eds.), Physica-Verlag, 2011, pp. 189–199. MR2767541
- [24] S. Resnick, *Extreme values, regular variation and point processes*, Springer, New York, 1987. MR900810
- [25] S. Resnick, *Heavy-tail phenomena: Probabilistic and statistical modeling*, Springer, New York, 2007. MR2271424
- [26] T. Schreiber and J. E. Yukich, *Large deviations for functionals of spatial point processes with applications to random packing and spatial graphs*, Stochastic Processes and their Applications **115** (2005), 1332–1356. MR2152378

- [27] J. Segers, Y. Zhao, and T. Meinguet, *Polar decomposition of regularly varying time series in star-shaped metric spaces*, *Extremes* **20** (2017), 539–566. MR3679981
- [28] T. Seppäläinen and J. E. Yukich, *Large deviation principles for Euclidean functionals and other nearly additive processes*, *Probability Theory and Related Fields* **120** (2001), 309–345. MR1843178

Acknowledgments. The authors are very grateful for valuable comments received from an anonymous referee and an anonymous Associate Editor. These comments helped the authors to improve the quality of the paper. The third author is thankful for fruitful discussions with Yogeshwaran D. and Z. Wei, which has helped him to deduce the desired \mathcal{M}_0 -convergence in Theorem 3.1 by means of [1, Theorem 6.4].

Electronic Journal of Probability

Electronic Communications in Probability

Advantages of publishing in EJP-ECP

- Very high standards
- Free for authors, free for readers
- Quick publication (no backlog)
- Secure publication (LOCKSS¹)
- Easy interface (EJMS²)

Economical model of EJP-ECP

- Non profit, sponsored by IMS³, BS⁴, ProjectEuclid⁵
- Purely electronic

Help keep the journal free and vigorous

- Donate to the IMS open access fund⁶ (click here to donate!)
- Submit your best articles to EJP-ECP
- Choose EJP-ECP over for-profit journals

¹LOCKSS: Lots of Copies Keep Stuff Safe <http://www.lockss.org/>

²EJMS: Electronic Journal Management System: <https://vtex.lt/services/ejms-peer-review/>

³IMS: Institute of Mathematical Statistics <http://www.imstat.org/>

⁴BS: Bernoulli Society <http://www.bernoulli-society.org/>

⁵Project Euclid: <https://projecteuclid.org/>

⁶IMS Open Access Fund: <https://imstat.org/shop/donation/>