A NOTE ON THE MOMENTS OF THE NUMBER OF AXIS-CROSSINGS BY A STOCHASTIC PROCESS

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1. Introduction. A general formula, for moments of arbitrary order of the number of upcrossings of a level u by a stationary normal process in unit time, was obtained by Cramér and Leadbetter [1], using a combination of techniques due to Kac [3], and Ylvisaker [6]. Ylvisaker [7] has weakened the conditions of this result slightly by a proof which depends on interesting applications of martingale convergence theory and which may be applied also to nonstationary normal situations. In this note we give a somewhat different direct procedure, under the weakened conditions, for the calculation of these moments. This procedure gives an alternative to that of Ylvisaker [7] for normal processes, without the use of martingale theory, and may be also applied to nonnormal situations in the same way as the discussion in [4] for the first moment.

We shall here give the "counting procedure" used to obtain the number of upcrossings, sketching the derivation, and indicating the extension to nonnormal cases. A detailed proof along these lines (for the stationary normal case) will be given elsewhere (Cramér and Leadbetter [2]).

2. A general result. We shall consider a process x(t) possessing, a.s., continuous sample functions and, for a given integer k, absolutely continuous 2k-dimensional distributions with corresponding densities of the form $f_{t_1...t_{2k}}(x_1 \cdot \cdot \cdot x_{2k})$. There will be no loss of generality in considering the number N of upcrossings of the zero level by x(t) in $0 \le t \le 1$, which is a well-defined random variable (cf. [4]).

For $t = (t_1 \cdots t_k)$ lying in the k-dimensional unit cube, let m_r denote the unique integer such that $m_r/2^n \le t_r < (m_r+1)/2^n$. Write $E_n(t)$ for the k-dimensional cube whose sides are the intervals $[m_r/2^n, (m_r+1)/2^n)$. For $\epsilon > 0$, let $A_{n\epsilon}$ denote the set of all points t in the unit cube such that for all $s = (s_1 \cdots s_k) \in E_n(t)$, we have $|s_i - s_j| > \epsilon$ whenever $i \ne j$, and write $\lambda_{n\epsilon}(t)$ for the characteristic function of the set $A_{n\epsilon}$. Finally let the random variable $\chi_{i,n} = 1$ if $x(i/2^n) < 0 < x[(i+1)/2^n]$, $\chi_{i,n} = 0$ otherwise. The following lemma

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gives the basic properties of the counting procedure used to obtain the factorial moments of N.

LEMMA. Let $M_{n\epsilon} = \sum_{x_{i_1,n},\dots,x_{i_k,n}} \lambda_{n\epsilon}(i_1/2^n \cdots i_k/2^n)$, where the summation is extended over all ordered sets of distinct integers $i_1 \cdots i_k$, $0 \le i_r \le 2^n - 1$. Then, with probability one,

- (i) $M_{n\epsilon}$ is nondecreasing as n increases or as ϵ decreases,
- (ii) $\lim_{n\to\infty}\lim_{\epsilon\to 0} M_{n\epsilon} = N(N-1)\cdot\cdot\cdot(N-k+1)$.

The proof of this lemma is accomplished by arguments extending those in [1, Part B]. From the monotonicity properties stated in (i) of the lemma it follows that the order of the ϵ and n-limits in (ii) may be interchanged. Hence, writing $M_k = \mathcal{E}N(N-1) \cdot \cdot \cdot (N-k+1)$, two applications of monotone convergence show that $M_k = \lim_{\epsilon \to 0} \lim_{n \to \infty} \mathcal{E} M_{n\epsilon}$. From the definition of $M_{n\epsilon}$ and a simple transformation of variables we thus have

(1)
$$M_k = \lim_{\epsilon \to 0} \lim_{n \to \infty} \sum_{n \to \infty} (i_1/2^n \cdot \cdot \cdot i_k/2^n)$$

 $\cdot P\{x_{i_r} < 0 < x_{i_r} + 2^{-n}y_{i_r}, r = 1, 2, \cdot \cdot \cdot, k\}$

in which $x_r = x(r/2^n)$ and $y_r = 2^n(x_{r+1} - x_r)$. In fact, the only nonzero terms in the sum on the right correspond to integers $i_1 \cdots i_k$ satisfying $|i_r - i_s| > 1$ for $r \neq s$. For such integer sets, the random variables $x_{i_1} \cdots x_{i_k}$, $y_{i_1} \cdots y_{i_k}$ possess a joint density. Write $\psi_{n_{t_s}}(x_1 \cdots x_k, y_1 \cdots y_k)$ to be equal to this joint density for all $t = (t_1 \cdots t_k)$ of A_{n_s} lying in the cube $E_n(i_1/2^n \cdots i_k/2^n)$, and $\psi_{n_{t_s}} = 0$ outside such cubes. From (1) we then obtain the following result by straightforward calculation.

THEOREM. For the process x(t) considered, the kth factorial moment of N is

(2)
$$M_{k} = \lim_{\epsilon \to 0} \lim_{n \to \infty} \int_{0}^{1} \cdots \int dt \int_{0}^{\infty} \cdots \int dy \int_{-y_{1}}^{0} \cdots \int_{-y_{k}}^{0} \cdot \psi_{nt\epsilon}(2^{-n}x_{1} \cdots 2^{-n}x_{k}, y_{1} \cdots y_{k}) dx.$$

3. Normal processes and generalizations. It can easily be seen that the above assumptions are satisfied for a (separable) stationary normal process x(t) whose covariance function has a finite second derivative at the origin. Further if $p_t(x, y)$ denotes the joint density for $x(t_1) \cdots x(t_k)$ and the quadratic mean derivatives $x'(t_1) \cdots x'(t_k)$, it can be shown (cf. [5]) by convergence of covariances that $\psi_{nt}(2^{-n}x_1 \cdots 2^{-n}x_k, y_1 \cdots y_k) \rightarrow p_t(0, y)$ as $n \rightarrow \infty$, for all t in the

region $D(\epsilon) = \lim_{n \to \infty} A_{n\epsilon}$. It can be shown by dominated convergence that the *n*-limit in (2) can be taken inside all the integral signs, and it is then an easy application of monotone convergence as $\epsilon \to 0$ to obtain the result:

$$(3) M_k = \int_0^1 \cdots \int dt \int_0^{\infty} \cdots \int y_1 \cdots v_k p_t(0, y) dy \leq \infty.$$

Certain nonnormal processes may be treated in a similar way from (2) to obtain (3), (see, for example, the derivation of &N in [4] for the *envelope* of a stationary normal process.) In general we may obtain a result corresponding to that given in [4] for the mean. To that end write, for $t = (t_1 \cdot \cdot \cdot t_k)$,

$$g_{t,\tau}(x, y) = \tau^k f_{t_1 \cdots t_k, t_1 + \tau \cdots t_k + \tau}(x_1 \cdots x_k, x_1 + \tau y_1 \cdots x_k + \tau y_k).$$

That is $g_{t,\tau}$ is the joint density for the $x(t_i)$ and the incrementary ratios $(x_{t_i+\tau}-x_{t_i})/\tau$. Then we have the following result.

THEOREM. Consider points $t = (t_1 \cdot \cdot \cdot t_k)$ such $t_i \neq t_j$ for $i \neq j$ and suppose that,

- (i) $g_{t\tau}(x, y)$ is continuous in (t, x) for each y, τ ,
- (ii) For each $\epsilon > 0$, $g_{t\tau}(\mathbf{x}, \mathbf{y}) \rightarrow p_t(\mathbf{x}, \mathbf{y})$ as $\tau \rightarrow 0$ uniformly in (t, \mathbf{x}) for $t \in D(\epsilon)$ and each \mathbf{y} .
 - (iii) For each $\epsilon > 0$, there is a function $h_{\epsilon}(y)$ such that for $t \in D(\epsilon)$,

$$g_{t,\tau}(\mathbf{x},\mathbf{y}) \leq h_{\epsilon}(\mathbf{y})$$
 and $\int_{0}^{\infty} \cdots \int y_{1} \cdots y_{k} h_{\epsilon}(\mathbf{y}) d\mathbf{y} < \infty$.

Then (3) holds for the process x(t).

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