A NOTE ON THE RECIPROCAL OF THE CONDITIONAL EXPECTATION OF A POSITIVE RANDOM VARIABLE¹

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1. Introduction and summary. Brunk [3] discusses conditional expectations given σ -lattices. This note is concerned with the observation that the reciprocal of the conditional expectation given a σ -lattice with respect to some measure μ of a positive random variable X is the conditional expectation of 1/X given the complementary σ -lattice with respect to another measure. (For the case of a σ -field this result is equivalent to the familiar result about the reciprocal of a Radon-Nikodym derivative.) Associated with this property of conditional expectations is a mapping of the set of all convex functions on $(0, \infty)$ into itself which leads to alternative ways of formulating certain extremum problems.

Let (Ω, α, μ) be a measure space with $\mu(\Omega) < \infty$. Let I_A denote the indicator function of a set A. We shall let $\mathfrak L$ denote a σ -lattice of subsets of $\Omega(\mathfrak L \subset \alpha)$. A σ -lattice is, by definition, closed under countable unions and intersections and contains both Ω and the null set \emptyset . If $\mathfrak L$ is such a σ -lattice then $\mathfrak L^{\sigma}$ will denote the σ -lattice of all subsets of Ω which are complements of members of $\mathfrak L$. We say that a random variable X is $\mathfrak L$ -measurable if $\{X > a\}$ $\mathfrak L$ for each real number a. Let L_2 denote the class of square integrable random variables and $L_2(\mathfrak L)$ the class of $\mathfrak L$ -measurable, square integrable random variables. Later we shall want to restrict our attention to strictly positive random variables, so for any set S of random variables we let S^+ denote the set of all those members of S which are strictly positive. Let $\mathfrak L$ denote the class of Borel subsets of the real line. The following is one of several available definitions for the conditional expectation, $E_{\mu}(X \mid \mathfrak L)$, of X given $\mathfrak L$ (see Brunk [3]).

DEFINITION. If $X \in L_2$ then $Y \in L_2(\mathfrak{L})$ is equal to $E_{\mu}(X \mid \mathfrak{L})$ if and only if Y has both of the following properties:

(1)
$$\int (X - Y)Z d\mu \leq 0 \quad \text{for each } Z \in L_2(\mathfrak{L})$$

(2)
$$\int_{B} (X - Y) d\mu = 0 \quad \text{for each } B \in Y^{-1}(\mathfrak{B})$$

(Brunk [3] shows the existence of such a Y and that it is unique in the sense that if W is any other member of $L_2(\mathfrak{L})$ having these properties then $W = Y[\mu]$.)

2. Some equivalent definitions for $E_{\mu}(X \mid \mathfrak{L})$. We may replace (2) in the definition of $E_{\mu}(X \mid \mathfrak{L})$ by:

(3)
$$\int (X - Y)\varphi(Y) d\mu = 0 \quad \text{for every Borel}$$
 function φ of a real variable such that $\varphi(Y) \varepsilon L_2$.

This can be shown by approximating φ by simple random variables.

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We also remark that:

(4)
$$\int_{A} (X - Y) d\mu \leq 0 \quad \text{for each } A \in \mathcal{L}$$

may replace (1) in the definition of $E_{\mu}(X \mid \mathfrak{L})$. This can be shown by approximating Z by simple random variables and using (2) and (4) together with the identity

$$\sum_{i=1}^{k} a_{i} I_{A_{i}} = \sum_{i=1}^{k-1} (a_{i} - a_{i+1}) I_{B_{i}} + a_{k} I_{B_{k}}$$

where $a_1 > a_2 > \cdots > a_k$, $A_i = [Z = a_i]$, $B_k = \Omega$ and $B_j = \sum_{i=1}^{j} A_i$ $(j = 1, 2, \dots, k-1)$.

3. The reciprocal of the conditional expectation of a positive random variable. Theorem 3.1. Suppose $X \in L_2^+$, $Y = E_{\mu}(X \mid \mathfrak{L})$ and $1/Y \in L_2$. If the measure γ is defined by $\gamma(A) = \int_A X d\mu$ for each $A \in \mathfrak{A}$ then

$$E_{\gamma}(1/X \mid \mathfrak{L}^{c}) = [E_{\mu}(X \mid \mathfrak{L})]^{-1}.$$

PROOF. Since $X \in L_2^+$ it follows from (2) that $Y \in L_2^+(\mathfrak{L})$. From this and the hypothesis we can conclude that $1/Y \in L_2^+(\mathfrak{L}^\circ)$. Suppose now that $B \in (1/Y)^{-1}(\mathfrak{B})$. Then B is also a member of $Y^{-1}(\mathfrak{B})$ and since X is the Radon-Nikodym derivative of γ with respect to μ we have:

$$\int_{B} (1/X - 1/Y) d\gamma = \int_{B} (1/X - 1/Y) X d\mu$$
$$= -\int (X - Y) (1/Y) I_{B} d\mu.$$

However, by (3) this integral must vanish so that:

$$\int_{B} (1/X - 1/Y) d\gamma = 0 \quad \text{for each } B \in (1/Y)^{-1}(\mathfrak{G}).$$

It remains only to verify that $\int (1/X - 1/Y)I_A d\gamma \leq 0$ for each $A \in \mathcal{L}^c$. We have:

$$\int (1/X - 1/Y) I_A d\gamma = \int (X - Y) (-I_A/Y) d\mu.$$

But, I_A is non-negative, 1/Y is strictly positive and both are measurable \mathcal{L}^{ϵ} so that $I_A/Y \in L_2(\mathcal{L}^{\epsilon})$ and $(-I_A/Y) \in L_2(\mathcal{L})$. However, by (1) this integral is non-positive and by (4) the proof is complete.

4. $E_{\mu}(X \mid \mathcal{L})$ as a solution to certain extremum problems. Suppose X is any member of L_2^+ , $Y = E_{\mu}(X \mid \mathcal{L})$ and $1/Y \in L_2$. Let Φ be any real valued function of a real variable whose domain includes $(0, \infty)$ and which is convex on this interval. Then for any $Z \in L_2^+$ we define the function $J_{\Phi}(Z; X, \mu)$ by:

(5)
$$J_{\Phi}(Z; X, \mu) = \int [\Phi(X) - \Phi(Z) - (X - Z)\varphi(Z)] d\mu$$

where $\varphi(Z)$ denotes any derivative, say for the sake of definiteness, the left derivative of Φ at Z. We still require that μ be totally finite. $J_{\Phi}(Z; X, \mu)$ always exists as a positive real number or $+\infty$. Brunk [2] shows the following:

(6)
$$\min_{Z \in L_2^+(\Omega)} J_{\Phi}(Z; X, \mu) = J_{\Phi}(Y; X, \mu).$$

However, by the same token, for any such convex Φ we have:

$$\min_{Z \in L_2^+(\mathfrak{L}^c)} J_{\Phi}(Z; 1/X, \gamma) = J_{\Phi}[E_{\gamma}(1/X \mid \mathfrak{L}^c); 1/X, \gamma].$$

But by Theorem 3.1, $E_{\gamma}(1/X \mid \mathfrak{L}^c) = [E_{\mu}(X \mid \mathfrak{L})]^{-1}$ so that $Y = E_{\mu}(X \mid \mathfrak{L})$ solves another minimum problem (other than (6)), namely:

(7)
$$\min_{Z \in L_2^+(\mathfrak{L})} J_{\Phi}(1/Z; 1/X, \gamma)$$

or equivalently:

$$\min_{Z \in L_2^+(\mathfrak{L})} \int X[\Phi(1/X) - \Phi(1/Z) - (1/X - 1/Z)\varphi(1/Z)] d\mu.$$

The question arises: is this actually a new minimum problem or is the functional $H_{\Phi}(Z; X, \mu) = J_{\Phi}(1/Z; 1/X, \gamma)$ of the form $J_{\Phi^*}(Z; X, \mu)$ for some other convex function Φ^* . The answer is the latter. Let $\Phi^*(\lambda) = \lambda \Phi(1/\lambda)$. It is easily verified that Φ^* is also convex on $(0, \infty)$ and that:

$$\Phi^*(X) - \Phi^*(Z) - (X - Z)\varphi^*(Z)$$

$$= X[\Phi(1/X) - \Phi(1/Z) - (1/X - 1/Z)\varphi(1/Z)].$$

Note that since Φ is convex if and only if $-\Phi$ is concave and since $J_{-\Phi}(Z; X, \mu) = -J_{\Phi}(Z; X, \mu)$, (6) implies that for any concave Θ :

(8)
$$\max_{Z \in L_2^+(\mathfrak{L})} J_{\Theta}(Z; X, \mu) = J_{\Theta}(Y, X, \mu).$$

Let θ denote the left derivative of Θ . Then $\int (X - Y)\theta(Y) d\mu = 0$ and if we let M denote the class of all those members Z of $L_2^+(\mathfrak{L})$ for which

$$\int (X-Z)\theta(Z) du = 0$$

and $\Theta(Z)$ is integrable then

(9)
$$\max_{Z \in \mathbf{M}} \left[- \int \Theta(Z) \ d\mu \right] = - \int \Theta(Y) \ d\mu.$$

It is assumed that $\Theta(X)$ and $\Theta(Y)$ are integrable.

To illustrate these comments suppose we are given two k-tuples (n_1, n_2, \dots, n_k) and (a_1, a_2, \dots, a_k) of positive real numbers and let $n_1 + n_2 + \dots + n_k = n$. In certain estimation problems (for example see [4]) we wish to find that k-tuple (p_1, p_2, \dots, p_k) , if it exists, of positive real numbers which satisfies

$$(10) p_1 \geq p_2 \geq \cdots \geq p_k$$

$$(11) a_1p_1 + a_2p_2 + \cdots + a_kp_k = 1$$

and $p_1^{n_1} \cdot p_2^{n_2} \cdot \cdots \cdot p_k^{n_k} \ge r_1^{n_1} \cdot r_2^{n_2} \cdot \cdots \cdot r_k^{n_k}$ for every other k-tuple (r_1, r_2, \dots, r_k) of positive real numbers which satisfies both (10) and (11).

Let $\Omega = \{1, 2, \dots, k\}$, α be the collection of all subsets of Ω and \mathcal{L} be the σ -lattice of left intervals in Ω :

$$\mathfrak{L} = \{\phi, \{1\}, \{1, 2\}, \cdots, \{1, 2, \cdots, k\}\}.$$

The k-tuples $p = (p_1, p_2, \dots, p_k)$ of reals can be thought of as functions on Ω and p satisfies (10) if and only if p is measurable \mathcal{L} .

Then using $\Theta(\lambda) = \log(\lambda)$ in (9) with X and μ defined on Ω and α respectively by $X(i) = n \cdot a_i/n_i$ and $\mu(i) = n_i/n$ it follows from the above remarks that the solution for Z is $E_{\mu}(X \mid \mathcal{L}^c)$, and hence for 1/p, $[E_{\mu}(X \mid \mathcal{L}^c)]^{-1}$. By Theorem 3.1, this coincides with $E_{\gamma}(1/X \mid \mathcal{L})$. In this approach to the problem we made use of the assumption that n_i 's were non-zero.

In some applications some of the n_i 's might be zero. Considerations mentioned above suggest rephrasing the problem in terms of $\Theta(\lambda) = \lambda \log (1/\lambda)$. With this approach using (8) instead of (9) our solution is given by $E_{\gamma}(Y \mid \mathfrak{L})$ where Y and γ are defined by $Y(i) = n_i/n \cdot a_i$ and $\gamma(i) = a_i$. Here we need to assume that the a_i 's are non-zero.

Representations of $E(X \mid \mathcal{L})$ and methods for calculating it are given in [1], in [5] and in [4] where an elegant graphical interpretation is given.

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REFERENCES

- [1] AYER, MIRIAM, BRUNK, H. D., EWING, G. M., REID, W. T. and SILVERMAN, EDWARD (1955). An empirical distribution function for sampling with incomplete information. Ann. Math. Statist. 26 641-647.
- [2] Brunk, H. D. (1961). Best fit to a random variable by a random variable measurable with respect to a σ-lattice. Pacific J. Math. 11 785-802.
- [3] BRUNK H. D. (1963). On an extension of the concept conditional expectation. Proc. Amer. Math. Soc. 14 298-304.
- [4] GRENANDER, ULF (1956). On the theory of mortality measurement, Part II. Skand. Aktuarietidskr. 39 126-153.
- [5] VAN EEDEN, CONSTANCE (1957). Maximum likelihood estimation of partially or completely ordered parameters. Nederl. Akad. Wetensch. Proc. Ser. A 60 128-136; Indag. Math. 19 201-211.