ON THE EXPECTED VALUES OF THE ELEMENTARY SYMMETRIC FUNCTIONS OF A NONCENTRAL WISHART MATRIX

BY D. J. DE WAAL

University of the Orange Free State

A conjecture is made for the expected value of the *j*th elementary symmetric function (ESF) of the roots of a noncentral Wishart matrix with covariance matrix $\sigma^2 I$. A conjecture is also made if j=p for any covariance matrix Σ . The expected value of the noncentral Wishart matrix is derived for any covariance matrix Σ and therefore also the expected value of the first ESF of the roots.

1. Introduction. If $X(p \times n)$ has independent normally distributed columns with covariance Σ and E(X) = M, then XX' = A is distributed as a noncentral Wishart distribution $W'(\Sigma, n, \Omega)$ with n degress of freedom and noncentrality parameter $\Omega = \Sigma^{-1}MM'$. The density is given by (Constantine (1963)).

(1.1)
$$(1/\Gamma_{p}(\frac{1}{2}n)|2\Sigma|^{\frac{1}{2}n}) \operatorname{etr}(-\frac{1}{2}\Omega)|A|^{\frac{1}{2}(n-p-1)} \\ \operatorname{etr}(-\frac{1}{2}\Sigma^{-1}A)_{0}F_{1}(\frac{1}{2}n;\frac{1}{4}\Omega\Sigma^{-1}A), \qquad A > 0$$

where etr $(B) = \exp(\operatorname{tr} B)$ and ${}_{0}F_{1}$ a hypergeometric function of matrix argument. Denote the *j*th elementary symmetric function (esf) of the roots of A by $\operatorname{tr}_{j} A$. We are interested in $E(\operatorname{tr}_{j} A)$ and will derive it in Section 2 if M=0, i.e. if A is distributed as a central Wishart distribution $W(\Sigma, n)$. In Section 3 we consider the noncentral Wishart matrix, but can only reach a conjecture for $E(\operatorname{tr}_{j} A)$, $j=3, \dots, p$ and $\Sigma=\sigma^{2}I_{n}$. We are able to prove it for j=1, 2.

2. Central Wishart. It has been proved (Constantine (1963)) that

(2.1)
$$E(C_{J}(A)) = 2^{j} (\frac{1}{2}n)_{J} C_{J}(2\Sigma)$$

where $(a)_J = \prod_{i=1}^p (a + \frac{1}{2}(1-i))_{j_i}$ and $(a)_j = a(a+1) \cdots (a+j-1)$. For every partition $J = (j_1, j_2, \dots, j_p)$ of $j, j_1 \ge j_2 \ge \dots \ge j_p$, the zonal polynomial $C_J(A)$ can be written as (James (1964))

(2.2)
$$C_{J}(A) = \chi_{(2J)}(1)2^{j}j! Z_{J}(A)/(2j)!$$

where $Z_j(A)$ has been tabulated by James (1964) in terms of the esf's of the roots of A for j = 1(1)6. The constant $\chi_{(2j)}^{(1)}$ is also tabulated. Since

$$(2.3) Z_{1j}(A) = j! \operatorname{tr}_{i} A$$

we have from (2.1)

(2.4)
$$E(\operatorname{tr}_{i} A) = n(n-1) \cdot \cdot \cdot (n-j+1) \operatorname{tr}_{i} \Sigma = (n)^{(j)} \operatorname{tr}_{i} \Sigma.$$

Received August 6, 1971.

3. Noncentral Wishart. It seems difficult to derive a similar result for the noncentral Wishart distribution. From (1.1)

(3.1)
$$E(C_{J}(A)) = \operatorname{etr}(-\frac{1}{2}\Omega) \sum_{k=0}^{\infty} \sum_{K} (1/(\frac{1}{2}n)_{K}) \int_{A>0} C_{K}(\frac{1}{4}\Omega \Sigma^{-1}A) \times C_{J}(A) dF(A; \Sigma, n)/k!$$

where $dF(A; \Sigma, n)$ is the $W(\Sigma, n)$ probability element and K is a partition on k. We shall find the solution to this integral only for the special case $\Sigma = \sigma^2 I_p$. In this case the integral is symmetric in Ω and we are allowed to transform $\Omega \to H'\Omega H$, $H \in O(p)$. Since James (1964) $\int_{O(p)} C_K(H'\Omega HA) dH = C_K(\Omega) C_K(A)/C_K(I_p)$, the integral can be written as

(3.2)
$$E(C_{J}(A)) = \operatorname{etr}(-\frac{1}{2}\Omega) \sum_{k} \sum_{K} (1/(\frac{1}{2}n)_{K}k!) (C_{K}(\frac{1}{2}\Omega)/C_{K}(I_{p})) \times \int_{A>0} C_{K}(\frac{1}{2}\sigma^{-2}A) C_{J}(A) dF(A; \sigma^{2}I_{p}, n).$$

Using the expression (Constantine (1966)) $C_K(A)C_J(A) = \Sigma_D g_{K,J}^D C_D(A)$, $g_{K,J}^D$ being a constant tabulated by Khatri and Pillai (1968) and D a partition of d = k + j, and the very useful integral (Constantine (1963))

$$\int_{A>0} |A|^{\frac{1}{2}(n-p-1)} \operatorname{etr}(-\frac{1}{2}PA)C_K(TA) dA
= \Gamma_n(\frac{1}{2}n)|2P^{-1}|^{\frac{1}{2}n}(\frac{1}{2}n)_K C_K(2TP^{-1}),$$

P and T symmetric matrices, it follows that

(3.3)
$$E(C_J(A)) = \operatorname{etr}(-\frac{1}{2}\Omega) \Sigma_k \Sigma_K \Sigma_D g_{K,J}^D(\frac{1}{2}n)_D (2\sigma^2)^j C_D(I_p) \times C_K(\frac{1}{2}\Omega) ((\frac{1}{2}n)_K C_K(I_p)k!)^{-1}.$$

Using (2.2), (2.3) and the result $Z_K(I_p)=2^k(\frac{1}{2}p)_K$ we have a result corresponding to (2.4) for the case $\Sigma=\sigma^2I_p$,

(3.4)
$$E(\operatorname{tr}_{j} A) = \operatorname{etr}(-\frac{1}{2}\Omega)a_{0}\Sigma_{k}a_{k}'\Sigma_{K}a_{K}''\Sigma_{D}a_{D}'''$$

where

$$\begin{split} a_0 &= \sigma^{2j}(2j)!/(j!j!\chi_{(2,1^j)}(1))\;,\\ a_{k'} &= 2^{2d}d!/(2^kk!\;(2d)!)\;,\quad a_{k''} = Z_{K}(\frac{1}{2}\Omega)/((\frac{1}{2}n)_{K}(\frac{1}{2}p)_{K})\;,\\ a_{D'''} &= g_{K,1}^{D}(j!2n)_{D}(\frac{1}{2}p)_{D}\chi_{(2D)}(1)\;. \end{split}$$

We will come back to this result after deriving $E(\operatorname{tr}_1 A)$ and obtaining a conjecture for $E(\operatorname{tr}_p A)$ using a simpler method. These two moments are considered for the general case $A \sim W'(\Sigma, n, \Omega)$. Since $A \sim XX'$ where $X \sim N(M, \Sigma \otimes I)$, $A \sim (U+M)(U+M)' = UU' + MU' + UM' + MM'$ where $U \sim N(0, \Sigma \otimes I)$. Hence

(3.5)
$$E(A) = n\Sigma + MM' = n\Sigma + \Sigma\Omega.$$

We have thus

(3.6)
$$E(\operatorname{tr}_{1} A) = n \operatorname{tr}_{1} \Sigma + \operatorname{tr}_{1} \Sigma \Omega.$$

It is also easy to see from (1.1) that

(3.7)
$$E(\operatorname{tr}_{p} A) = \operatorname{etr}(-\frac{1}{2}\Omega)(\Gamma_{p}(\frac{1}{2}n+1)/\Gamma_{p}(\frac{1}{2}n))|2\Sigma| \sum_{k=0}^{\infty} \Sigma_{K} \times (1/(\frac{1}{2}n)_{K}k!) \int_{A>0} C_{K}(\frac{1}{4}\Omega\Sigma^{-1}A) dF(A; \Sigma, n+2) = |\Sigma|(n)^{(p)} \operatorname{etr}(-\frac{1}{2}\Omega)_{1}F_{1}(\frac{1}{2}n+1; \frac{1}{2}n; \frac{1}{2}\Omega).$$

It seems that this expression can be written in a much simpler form, but the author has not succeeded in proving it. We shall however consider the two special cases p=2, 3 and make an obvious conjecture for the general case. (3.7) can be written as

(3.8)
$$E(\operatorname{tr}_{p} A) = |\Sigma|(n)^{(p)} \operatorname{etr}(-\frac{1}{2}\Omega) \Sigma_{k} \Sigma_{K} \prod_{i=1}^{p} (1 + 2k_{i}/(n+1-i)) \times C_{K}(\frac{1}{2}\Omega)/k!$$

Case
$$p = 2$$
. If $p = 2$,

$$\prod_{i=1}^{2} (1 + 2k_i/(n+1-i)) = 1 + 2k/n + 2k_2(1+2k_1)/n(n-1).$$

Therefore, using the fact $\Sigma_K C_K(S) = (\operatorname{tr} S)^k$, and the tables (James (1964)) expressing zonal polynomials in terms of esf's of the roots of a matrix,

(3.9)
$$E(tr_{p=2}A) = |\Sigma|((n)^{(2)} + (n-1)\operatorname{tr}_1\Omega + \operatorname{tr}_2\Omega).$$

Case p = 3. If p = 3,

$$\begin{split} \prod_{i=1}^{3} \left(1 + 2k_i/(n+1-i)\right) \\ &= 1 + 2k/n + (4k_1k_2 + 4k_1k_3 + 4k_2k_3 + 2k_2 + 4k_3)/n(n-1) \\ &+ 4(2k_1k_2k_3 + k_1k_3 + 2k_2k_3 + k_3)/n(n-1)(n-2) \; . \end{split}$$

Following the same procedure as for p = 2, it follows after substitution

(3.10)
$$E(\operatorname{tr}_{p=3} A) = |\Sigma|((n)^{(3)} + (n-1)^{(2)} \operatorname{tr}_1 \Omega + (n-2) \operatorname{tr}_2 \Omega + \operatorname{tr}_3 \Omega) .$$

Conjecture. An obvious conjecture for the general case follows immediately from (3.9) and (3.10), namely

(3.11)
$$E(\operatorname{tr}_{p} A) = |\Sigma|((n)^{(p)} + (n-1)^{(p-1)} \operatorname{tr}_{1} \Omega + (n-2)^{(p-2)} \operatorname{tr}_{2} \Omega + \cdots + \operatorname{tr}_{p} \Omega) .$$

From (3.6) and (3.11) we now have an intuitive idea that for the special case $\Sigma = \sigma^2 I_p$, $E(\operatorname{tr}_j A)$ can be written in terms of the first j esf's of the roots of Ω . By writing down the first few terms of the series (3.4) for j=2, we obtain a conjecture in general. It is only necessary to consider k=1(1)4 and for k=4 we need not have to go further than the partition K=(3,1) to show that

(3.12)
$$E(\operatorname{tr}_{2} A) = \sigma^{4}((n)^{(2)}\binom{p}{2} + (n-1)\binom{p-1}{1}\operatorname{tr}_{1}\Omega + \operatorname{tr}_{2}\Omega) .$$

This result leads us to the general result.

Conjecture.

(3.13)
$$E(\operatorname{tr}_{j} A) = \sigma^{2j}((n)^{(j)}\binom{p}{j} + (n-1)^{(j-1)}\binom{p-1}{j-1}\operatorname{tr}_{1}\Omega + (n-2)^{(j-2)}\binom{p-j}{j-2}\operatorname{tr}_{2}\Omega + \cdots + \operatorname{tr}_{j}\Omega).$$

The first term can be checked with (2.4). Further, if j = p we have (3.11) and if j = 1 we have (3.6) for $\Sigma = \sigma^2 I_p$. Equation (3.13) essentially says:

$$(3.14) E|A-\lambda I| = L(|\Omega-\lambda I|)$$

where L is a lower triangular linear operator acting on the vector of coefficients of the polynomial $|\Omega - \lambda I|$. The matrix L has the form

$$L_{jk} = \pm \binom{p-k}{j-k} (n-k)^{(j-k)}$$
.

With suitable normalization on L and defining

$$a_j = \operatorname{tr}_j A/(n)^{(j)}\binom{p}{j}$$
 and $w_j = \operatorname{tr}_j \Omega/(n)^{(j)}\binom{p}{j}$

the matrix L takes on the simple form

$$\tilde{L}_{jk} = \pm \binom{j}{k}$$
.

Acknowledgement. The author wishes to express his gratitude to the referees for their suggestions.

REFERENCES

Constantine, A. G. (1963). Some noncentral distribution problems in multivariate analysis. *Ann. Math. Statist.* 34 1270-1285.

Constantine, A. G. (1966). The distribution of Hotelling's generalised T_0^2 . Ann. Math. Statist. 37 215-225.

JAMES, A. T. (1964). Distribution of matrix variates and latent roots derived from normal samples. Ann. Math. Statist. 35 475-501.

KHATRI, C. G. and PILLAI, K. C. S. (1968). On the noncentral distribution of two test criteria in multivariate analysis of variance. *Ann. Math. Statist.* 39 215-226.