# A counterexample to monotonicity of relative mass in random walks* 

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#### Abstract

For a finite undirected graph $G=(V, E)$, let $p_{u, v}(t)$ denote the probability that a continuous-time random walk starting at vertex $u$ is in $v$ at time $t$. In this note we give an example of a Cayley graph $G$ and two vertices $u, v \in G$ for which the function $$
r_{u, v}(t)=\frac{p_{u, v}(t)}{p_{u, u}(t)} \quad t \geq 0
$$ is not monotonically non-decreasing. This answers a question asked by Peres in 2013.


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## 1 Introduction

Let $G=(V, E)$ be a finite undirected regular graph. Let $p_{u, v}(t)$ denote the probability that a continuous-time random walk starting at vertex $u$ is in $v$ at time $t$. In this note we are interested in the function

$$
r_{u, v}(t)=\frac{p_{u, v}(t)}{p_{u, u}(t)} \quad t \geq 0
$$

Clearly, in regular connected graphs for any $u \neq v$, we have $r_{u, v}(0)=0$ and $\lim _{t \rightarrow \infty} r_{u, v}(t)=1$. One might wonder if the function is monotonically non-decreasing. It is not difficult to see that there are regular graphs for which this is not the case. In fact, there are regular graphs such that $r_{u, v}(t)>1$ for some vertices $u, v$ and time $t$; in particular, $r_{u, v}(t)$ is not monotonically non-decreasing. We give an example of such a graph in Appendix A. We thank Jeff Cheeger [1] for pointing this out to us.

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For vertex-transitive graphs, however, it holds that $r_{u, v}(t) \leq 1$ for all vertices $u, v$ and all $t \geq 0$. Indeed, using Cauchy-Schwarz and the reversibility of the walk,

$$
\begin{aligned}
p_{u, v}(t) & =\sum_{w \in V} p_{u, w}(t / 2) p_{w, v}(t / 2) \\
& \leq\left(\sum_{w} p_{u, w}(t / 2)^{2}\right)^{1 / 2} \cdot\left(\sum_{w} p_{w, v}(t / 2)^{2}\right)^{1 / 2} \\
& =p_{u, u}(t)^{1 / 2} \cdot p_{v, v}(t)^{1 / 2}=p_{u, u}(t) .
\end{aligned}
$$

This motivates the following question, asked in 2013 by Peres [5]:
Is the function $r_{u, v}$ monotonically non-decreasing in $t$ for all vertex-transitive graphs and all vertices $u, v$ ?

More recently, a special case of that question was asked independently by Price [7]. Namely, Price asked whether for Brownian motion on flat tori (i.e., on $\mathbb{R}^{n}$ modulo a lattice), it holds that for any point $x$, the density at $x$ divided by the density at the starting point $x_{0}$ is monotonically non-decreasing in time. This would follow from a positive answer to Peres's question through a limit argument. Price gave a positive answer to his question for the case of a cycle ( $n=1$ ) and recently, a positive answer for arbitrary flat tori was found [8]. This can be seen as further evidence for a positive answer to Peres's question.

In this note we give a negative answer to Peres's question. In fact, we do so through a Cayley graph.

Theorem 1.1. There exists a Cayley graph $G=(V, E)$ and two vertices $u, v \in V$ such that the function $r_{u, v}$ is not monotonically non-decreasing.

One remaining open question is whether $r_{u, v}$ is monotonically non-decreasing for Abelian Cayley graphs. The positive result of [8] is a special case of that.

### 1.1 Some basic facts about continuous-time random walks

Given a weighted finite graph $G=(V, E)$ with weight function $w: E \rightarrow \mathbb{R}_{+}$a continuous-time random walk $X=\left(X_{t}\right)_{t \geq 0}$ on $G$ is defined by its heat kernel $H_{t}$, that at time $t>0$ is equal to

$$
H_{t}=e^{-t \cdot L}
$$

where $L$ is the Laplacian matrix of $G$ given by $L_{u, v}=-w(u, v)$ for $u \neq v$, and $L_{u, u}=$ $\sum_{v} w(u, v)$. As a result, for a random walk $X$ starting at a vertex $u$ the probability that $X$ is in $v$ at time $t$ is equal to $p_{u, v}(t):=H_{t}(u, v)$. When $G$ is a $d$-regular unweighted simple graph, we think of the edges as all having weight $1 / d$, in which case the Laplacian of $G$ is given by

$$
L_{u, v}= \begin{cases}-1 / d & \text { if }(u, v) \in E \\ 1 & \text { if } u=v \\ 0 & \text { otherwise }\end{cases}
$$

In this note we consider only vertex-transitive graphs, for which the sum $\sum_{v} w(u, v)$ is the same for all vertices $u$ of the graph. Note that we do not insist that this sum is equal to 1 , though this can be achieved by normalizing $L$, which corresponds to changing the speed of the random walk. For basic facts about continuous-time random walks see, e.g., [4].

If $G$ is a weighted Cayley graph with a generating set $S$ and a weight function $w: S \rightarrow \mathbb{R}_{+}$, then a continuous-time random walk $X=\left(X_{t}\right)_{t \geq 0}$ on $G$ is described by mutually independent Poisson processes of rate $w(g)$ for each group generator $g \in S$, where each process indicates the times when $X$ jumps along the corresponding edge.

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## 2 Non-monotonicity of time spent at the origin in the hypercube graph

For an integer $d \geq 1$ denote by $Q_{d}$ the $d$-dimensional hypercube graph. The vertices of $Q_{d}$ are $\{0,1\}^{d}$ and there is an edge between two vertices $u$ and $v$ if and only if they differ in exactly one coordinate. Let $X=\left(X_{t}\right)_{t>0}$ be a continuous-time random walk on $Q_{d}$ starting at the origin, denoted by $\mathbf{0}=(0, \ldots, 0) \in\{0,1\}^{d}$. Denote by $C_{d}(t)$ the expected time spent at the origin until time $t$, conditioned on the event that $X_{t}=\mathbf{0}$. That is,

$$
C_{d}(t)=\int_{0}^{t} \operatorname{Pr}\left[X_{s}=\mathbf{0} \mid X_{t}=\mathbf{0}\right] d s
$$

In this section we show that for $d$ sufficiently large $C_{d}(t)$ is not monotonically nondecreasing.
Lemma 2.1. Let $d \in \mathbb{N}$ be sufficiently large. Then, there are some $t_{1}<t_{2}$ such that $C_{d}\left(t_{1}\right)>C_{d}\left(t_{2}\right)$, and in particular, the function $C_{d}$ is not monotonically non-decreasing in $t$.
Remark 2.2. Numerically, one can see that the function $C_{d}$ is not monotone for $d \geq 5$. See Figure 1. Since $C_{d}$ has a closed form expression (as can be seen from the calculations below), one can probably show non-monotonicity directly for $C_{5}$ by analyzing the function, though doing so would likely be messy and not too illuminating.


Figure 1: $C_{d}(t)$ for $d=4,5,6,7$ (from top right to bottom right).
Before proving Lemma 2.1 we prove the following claim.
Claim 2.3. Let $d \geq 1$, and let $Q_{d}$ be the $d$-dimensional hypercube graph. Let $X=\left(X_{t}\right)_{t>0}$ be a continuous-time random walk on $Q_{d}$ starting at $\mathbf{0}$. Then,

$$
\operatorname{Pr}\left[X_{t}=\mathbf{0}\right]=\left(\frac{1+e^{-2 t / d}}{2}\right)^{d}
$$

Proof. Since $X$ moves in each coordinate with rate $1 / d$, it follows that for each $i \in[d]$ the number of steps in direction $i$ up to time $t$ is distributed like $\operatorname{Pois}(t / d)$. Therefore,

$$
\operatorname{Pr}\left[\left(X_{t}\right)_{i}=0\right]=\operatorname{Pr}[\operatorname{Pois}(t / d) \text { is even }]=\left(1+e^{-2 t / d}\right) / 2
$$

where we used that the probability that $\operatorname{Pois}(\lambda)$ is even is
$\operatorname{Pr}[\operatorname{Pois}(\lambda)$ is even $]=e^{-\lambda} \cdot \sum_{j \text { even }} \frac{\lambda^{j}}{j!}=e^{-\lambda} \cdot \frac{1}{2}\left(\sum_{j=0}^{\infty} \frac{\lambda^{j}}{j!}+\sum_{j=0}^{\infty} \frac{(-\lambda)^{j}}{j!}\right)=e^{-\lambda} \cdot \frac{1}{2}\left(e^{\lambda}+e^{-\lambda}\right)$.

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Since the coordinates of $X$ move independently the result follows.

We now prove Lemma 2.1.

Proof of Lemma 2.1. We show below that for all $d \geq 1$, it holds that

1. $C_{d}(\sqrt{d}) \geq e^{-1} \sqrt{d}$,
2. $C_{d}(d) \leq 6$.

This clearly proves the lemma for $d$ sufficiently large.
To prove Item 1, we show that if a walk starting from the origin is at the origin at time $\sqrt{d}$, then with constant probability it stayed at the origin throughout that time interval. Intuitively, this is because the probability of a coordinate flipping twice during that time is of order only $1 / d$ and so with constant probability none of the $d$ coordinates flips. In more detail, by Claim 2.3,

$$
\operatorname{Pr}\left[X_{\sqrt{d}}=\mathbf{0}\right]=\left(\frac{1+e^{-2 / \sqrt{d}}}{2}\right)^{d} \leq\left(1-\frac{1}{\sqrt{d}}+\frac{1}{d}\right)^{d}=\left(1-\frac{\sqrt{d}-1}{d}\right)^{d} \leq e^{-\sqrt{d}+1},
$$

where we used the inequality $e^{-x} \leq 1-x+x^{2} / 2$ valid for all $x \geq 0$. On the other hand, by definition of a continuous-time random walk the probability that $X$ stays in $\mathbf{0}$ during the entire time interval $[0, \sqrt{d}]$ is equal to $\operatorname{Pr}\left[X_{[0, \sqrt{d}]} \equiv \mathbf{0}\right]=e^{-\sqrt{d}}$. Therefore,

$$
\operatorname{Pr}\left[X_{[0, \sqrt{d}]} \equiv \mathbf{0} \mid X_{\sqrt{d}}=\mathbf{0}\right] \geq e^{-1},
$$

and hence the expected time spent at the origin conditioned on $X_{\sqrt{d}}=\mathbf{0}$ is as claimed in Item 1.

We next prove Item 2. Intuitively, here there is enough time for coordinates to flip twice, and only a very small part of the time will be spent at the origin. By definition of $C_{d}$ and Claim 2.3 we have

$$
\begin{aligned}
C_{d}(t) & =\int_{0}^{t} \frac{\operatorname{Pr}\left[X_{s}=\mathbf{0}\right] \cdot \operatorname{Pr}\left[X_{t-s}=\mathbf{0}\right]}{\operatorname{Pr}\left[X_{t}=\mathbf{0}\right]} d s \\
& =\int_{0}^{t}\left(h_{d}(t, s)\right)^{d} d s
\end{aligned}
$$

where

$$
h_{d}(t, s)=\frac{\left(1+e^{-2 s / d}\right)\left(1+e^{-2(t-s) / d}\right)}{2\left(1+e^{-2 t / d}\right)}=\frac{1+e^{-2 s / d}+e^{-2(t-s) / d}+e^{-2 t / d}}{2\left(1+e^{-2 t / d}\right)} .
$$

Since $h_{d}(t, s)$ is convex as a function of $s$, for all $0 \leq s \leq t / 2$ we have $h_{d}(t, s) \leq \ell(s)$ where $\ell$ is the unique linear function satisfying $\ell(0)=h_{d}(t, 0)$ and $\ell(t / 2)=h_{d}(t, t / 2)$. Therefore, taking $t=d$, we get

$$
h_{d}(d, s) \leq \ell(s / d)=1-\frac{c s}{d},
$$

where $c=\left(1-e^{-1}\right)^{2} /\left(1+e^{-2}\right)$. Noting that $h_{d}(t, s)=h_{d}(t, t-s)$, we get
$C_{d}(d)=\int_{0}^{d}\left(h_{d}(d, s)\right)^{d} d s=2 \int_{0}^{d / 2}\left(h_{d}(d, s)\right)^{d} d s<2 \int_{0}^{d / 2}\left(1-\frac{c s}{d}\right)^{d} d s<2 \int_{0}^{d / 2} e^{-c s} d s \leq \frac{2}{c}$.
This completes the proof of Lemma 2.1.

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## 3 Proof of Theorem 1.1

In this section we prove Theorem 1.1. We first give a proof for a weighted graph, and then remark on how to convert it into an unweighted graph. For $d \in \mathbb{N}$ sufficiently large we define the weighted graph $G$ to be the lamplighter graph on $Q_{d}$, whose edges corresponding to steps on $Q_{d}$ are of weight $1 / d$, and edges corresponding to toggling a lamp are of weight $\varepsilon$, for some $\varepsilon>0$ sufficiently small that depends on $d$ and $t_{1}, t_{2}$ from Lemma 2.1.

In more detail, the weighted lamplighter graph $G$ is described by placing a lamp at each vertex of $Q_{d}$ and a lamplighter walking on $Q_{d}$. A vertex of $G$ is described by the location $x \in\{0,1\}^{d}$ of the lamplighter, and a configuration $f:\{0,1\}^{d} \rightarrow\{0,1\}$ indicating which lamps are currently on. In each step the lamplighter either makes a step in the graph $Q_{d}$ or toggles the state of the lamp in the current vertex. More formally, we have an edge between $(x, f)$ and $(y, g)$ if and only if either

1. $(x, y) \in E_{d}$ and $f=g$ (this corresponds to a step in $Q_{d}$ ) or
2. $x=y$ and $f$ and $g$ differ on the input $x$ and are equal on all other inputs (this corresponds to toggling a lamp at $x$ ).

The edges of the first type are of weight $1 / d$, and those of the second type are of weight $\varepsilon$. Thus, in a random walk on $G$, the steps of the lamplighter are distributed as in a random walk on $Q_{d}$, and the number of times the lamps are toggled in a time interval of length $T$ is distributed like $\operatorname{Pois}(\varepsilon T)$ independently of the lamplighter's walk. Lamplighter graphs are well-studied objects (see, e.g., [6, 3, 2]), and are well known to be Cayley graphs.

Let $u$ be the vertex in $G$ corresponding to the lamplighter being at the origin with all lights off. Let $v$ be the vertex in $G$ corresponding to the lamplighter being at the origin with the light at the origin being on, and all other lights off. We show below that $r_{u, v}$ is not monotonically non-decreasing. More specifically, we show that $r_{u, v}\left(t_{1}\right)>r_{u, v}\left(t_{2}\right)$, where $t_{1}<t_{2}$ are from Lemma 2.1.

Let $X=\left(X_{t}\right)_{t \geq 0}$ be a continuous-time random walk on $G$ starting at $X_{0}=u$. Denote by $Y_{t}$ the number of times a toggle occurred during the time interval $[0, t]$. Denote by $Z=\left(Z_{t}\right)_{t \geq 0}$ the trajectory of the lamplighter, i.e., the projection of $X$ to the first coordinate. Note that by definition $Z$ is a continuous-time random walk on $Q_{d}$, and that $Z$ is independent of $Y_{t}$.
Claim 3.1. Let $u, v \in V$ be as above. Then, for all $t>0$ it holds that

$$
\begin{equation*}
0 \leq p_{u, u}(t)-e^{-\varepsilon t} \cdot \operatorname{Pr}\left[Z_{t}=\mathbf{0}\right] \leq \varepsilon^{2} t^{2}, \tag{3.1}
\end{equation*}
$$

and

$$
\begin{equation*}
0 \leq p_{u, v}(t)-\varepsilon e^{-\varepsilon t} \cdot C_{d}(t) \cdot \operatorname{Pr}\left[Z_{t}=\mathbf{0}\right] \leq \varepsilon^{2} t^{2} \tag{3.2}
\end{equation*}
$$

Using the claim,

$$
r_{u, v}(t)=\frac{p_{u, v}(t)}{p_{u, u}(t)}=\varepsilon \cdot C_{d}(t) \pm O\left(\varepsilon^{2}\right)
$$

where $O(\cdot)$ hides a constant that depends on $d$ and $t$. In particular, for $t_{1}<t_{2}$ from Lemma 2.1, and $\varepsilon>0$ sufficiently small we get that $r_{u, v}\left(t_{1}\right)>r_{u, v}\left(t_{2}\right)$, which proves Theorem 1.1.

Intuitively, (3.1) holds because the probability of toggling a lamp twice is very small, and hence $p_{u, u}(t)$ is approximately equal to the probability that no lamp has changed its state multiplied by the probability that a random walk on $Q_{d}$ will be at the origin at time $t$. The intuition for (3.2) is that in order to get from $u$ to $v$, in addition to getting back to the origin, the lamplighter must toggle the switch while being at the origin, and the probability of that is roughly $\varepsilon \cdot C_{d}(t)$.

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Proof of Claim 3.1. For $p_{u, u}$ we have

$$
p_{u, u}=\operatorname{Pr}\left[X_{t}=u \wedge Y_{t}=0\right]+\operatorname{Pr}\left[X_{t}=u \wedge Y_{t} \geq 2\right]
$$

Since $Y_{t}$ is distributed like $\operatorname{Pois}(\varepsilon t)$, the second term satisfies

$$
0 \leq \operatorname{Pr}\left[X_{t}=u \wedge Y_{t} \geq 2\right] \leq \operatorname{Pr}\left[Y_{t} \geq 2\right] \leq \varepsilon^{2} t^{2}
$$

and for the first term, by independence between $Y_{t}$ and $Z_{t}$ we have

$$
\operatorname{Pr}\left[X_{t}=u \wedge Y_{t}=0\right]=\operatorname{Pr}\left[Z_{t}=\mathbf{0} \wedge Y_{t}=0\right]=e^{-\varepsilon t} \cdot \operatorname{Pr}\left[Z_{t}=\mathbf{0}\right]
$$

proving (3.1).
For $p_{u, v}$ we similarly have

$$
p_{u, v}(t)=\operatorname{Pr}\left[X_{t}=v \wedge Y_{t}=1\right]+\operatorname{Pr}\left[X_{t}=v \wedge Y_{t} \geq 2\right]
$$

As above, the second term is at most $\varepsilon^{2} t^{2}$. For the first term, let $E_{t}$ be the event that $Y_{t}=1$, and the unique lamp that is on at time $t$ is the lamp at the origin. Denote by $T_{0}$ the time spent by $Z$ at the origin in the time interval $[0, t]$. Then, conditioning on $Z$, the event $E_{t}$ holds if and only if a unique switch happened during $T_{0}$ time, and zero switches in the remaining time. Therefore, by independence of a Poisson process in disjoint intervals
$\operatorname{Pr}\left[E_{t} \mid Z\right]=\operatorname{Pr}\left[\operatorname{Pois}\left(\varepsilon T_{0}\right)=1 \mid Z\right] \cdot \operatorname{Pr}\left[\operatorname{Pois}\left(\varepsilon\left(t-T_{0}\right)\right)=0 \mid Z\right]=\varepsilon T_{0} \cdot e^{-\varepsilon T_{0}} \cdot e^{-\varepsilon\left(t-T_{0}\right)}=\varepsilon e^{-\varepsilon t} \cdot T_{0}$.
This implies that

$$
\operatorname{Pr}\left[X_{t}=v \wedge Y_{t}=1\right]=\operatorname{Pr}\left[E_{t} \mid Z_{t}=\mathbf{0}\right] \cdot \operatorname{Pr}\left[Z_{t}=\mathbf{0}\right]=\varepsilon e^{-\varepsilon t} \cdot \mathbb{E}\left[T_{0} \mid Z_{t}=\mathbf{0}\right] \cdot \operatorname{Pr}\left[Z_{t}=\mathbf{0}\right]
$$

Therefore, since $C_{d}(t)=\mathbb{E}\left[T_{0} \mid Z_{t}=0\right]$ we get (3.2), and the claim follows.

Converting $G$ into an unweighted graph Below we show how to convert a weighted Cayley graph $G$ into an unweighted one, while preserving the property in Theorem 1.1. Let $\left(G, S_{G}\right)$ be a weighted Cayley graph with the generating set $S_{G}=\left\{g_{1}, \ldots, g_{k}\right\}$, and suppose that all the weights $w\left(g_{1}\right), \ldots, w\left(g_{k}\right)$ are integers. For $N \in \mathbb{N}$ sufficiently large define the graph $H$ by replacing each vertex $v \in G$ with a "cloud" of $N$ vertices $\left\{(v, i): i \in \mathbb{Z}_{N}\right\}$, adding edges between every pair of vertices in the cloud (i.e., replacing each vertex of $G$ with an $N$-clique), and replacing each edge ( $u, u g$ ) in $G$ of weight $w(g)$ with $w(g)$ perfect matchings $\left\{(u, i),(u g, i+j): i \in \mathbb{Z}_{N}\right\}_{j=1}^{w(g)}$. Formally, the graph $H$ is a Cayley graph, whose vertices are $G \times \mathbb{Z}_{N}=\left\{(v, i): v \in G, i \in \mathbb{Z}_{N}\right\}$, and the set of generators $S_{H}$ given by

$$
S_{H}=\left\{(0, i): i \in \mathbb{Z}_{N} \backslash\{0\}\right\} \bigcup \cup_{g \in S_{G}}\{(g, j): j \in\{1, \ldots, w(g)\}\}
$$

Note that the projection of a continuous-time random walk on $H$ to the first coordinate is a random walk on $G$, slowed down by $\operatorname{deg}(H)$. (The walk on the weighted graph $G$ makes $\sum_{g} w(g)$ steps per unit time on average, whereas by our convention, the walk on the unweighted graph $H$ makes one step per unit time on average, a $\sum_{g} w(g) / \operatorname{deg}(H)$ proportion of which is between the clouds.) Moreover, assuming $N$ is sufficiently large, after constant time the two coordinates become close to independent with the second coordinate being uniform. Therefore, if $u, v$ are vertices in $G$, and $x=(u, 0), y=(v, 0)$ are the corresponding vertices in $H$, then for any time $t>0$ and $t^{\prime}=\operatorname{deg}(H) \cdot t$ it holds that $p_{x, y}\left(t^{\prime}\right)=\frac{1}{N}\left(p_{u, v}(t) \pm o_{N}(1)\right)$ and hence $r_{x, y}\left(t^{\prime}\right)=r_{u, v}\left(t^{\prime}\right) \pm o_{N}(1)$.

For the graph $G$ given in the proof of Theorem 1.1 above, we may assume that $1 / \varepsilon$ is an integer, and so, by multiplying all weights by $d / \varepsilon$ we get a Cayley graph with integer weights. Hence, by applying the foregoing transformation we get a simple unweighted Cayley graph $H$ for which $r_{u, v}$ is not monotonically non-decreasing for some $u, v \in H$.

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## A Appendix: a counterexample in a regular non-transitive graph

Below we give a simple example of a regular non-transitive graph such that $r_{u, v}(t)>1$ for some vertices $u, v$ and some time $t$; in particular, $r_{u, v}(t)$ is not monotonically nondecreasing, since $r_{u, v}(t) \rightarrow 1$ as $t \rightarrow \infty$. We thank Jeff Cheeger [1] for pointing this out to us.
Proposition A.1. Let $L$ be the Laplacian of a regular graph on vertex set $V$. Denote its eigenvalues by $0=\lambda_{1} \leq \lambda_{2} \leq \cdots \leq \lambda_{|V|}$ and by $f_{i} \in \mathbb{R}^{V}$ the corresponding normalized eigenvectors. Suppose that $0<\lambda_{2}<\lambda_{3}$, and that $f_{2}$ is such that $f_{2}(v)>f_{2}(u)>0$ for some vertices $u, v$. Then, there is some $t>0$ such that $r_{u, v}(t)>1$.

Proof. Let $\pi_{u} \in \mathbb{R}^{V}$ be the vector with $\pi_{u}(u)=1$ and $\pi_{u}\left(u^{\prime}\right)=0$ for all $u^{\prime} \neq u$. Writing $\pi_{u}=\sum \alpha_{i} f_{i}$ for $\alpha_{i}=\left\langle\pi_{u}, f_{i}\right\rangle=f_{i}(u)$, for all $w \in V$ we have

$$
e^{-t L} \pi_{u}(w)=\sum_{i=1}^{|V|} e^{-t \lambda_{i}} \alpha_{i} \cdot f_{i}(w)=c+e^{-\lambda_{2} t} f_{2}(u) f_{2}(w)+O\left(e^{-\lambda_{3} t}\right)
$$

where $O()$ hides some constants that may depend on the graph, but not on $t$, and $c=\alpha_{1} \cdot f_{1}(w)$ is independent of $w$ since $f_{1}$ is a constant function. Using the facts that $f_{2}(v)>f_{2}(u)>0$ and $\lambda_{3}>\lambda_{2}$, it follows that for sufficiently large $t$,

$$
r_{u, v}(t)=\frac{e^{-t L} \pi_{u}(v)}{e^{-t L} \pi_{u}(u)}>1
$$

as desired.


Figure 2: A cube with two square pyramids attached.
Graphs satisfying the constraints in Proposition A. 1 are in abundance. As a concrete example, consider the 4 -regular graph on 10 vertices shown in Figure 2. Using Mathematica, we see that the second eigenvalue of the Laplacian of this graph is $\lambda_{2}=\frac{1}{8}(7-\sqrt{17}) \approx 0.36$, and it is a simple eigenvalue. The corresponding (non-normalized) eigenvector with vertices ordered from left to right is $(c, 1,1,1,1,-1,-1,-1,-1,-c)$, where $c=3-\frac{1}{2}(7-\sqrt{17}) \approx 1.56$. In particular, Proposition A. 1 is applicable to this graph.

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