CONTINUUM PERCOLATION WITH STEPS IN AN ANNULUS

BY PAUL BALISTER,¹ BÉLA BOLLOBÁS² AND MARK WALTERS¹

University of Memphis

Let *A* be the annulus in \mathbb{R}^2 centered at the origin with inner and outer radii $r(1 - \varepsilon)$ and *r*, respectively. Place points $\{x_i\}$ in \mathbb{R}^2 according to a Poisson process with intensity 1 and let \mathcal{G}_A be the random graph with vertex set $\{x_i\}$ and edges $x_i x_j$ whenever $x_i - x_j \in A$. We show that if the area of *A* is large, then \mathcal{G}_A almost surely has an infinite component. Moreover, if we fix ε , increase *r* and let $n_c = n_c(\varepsilon)$ be the area of *A* when this infinite component appears, then $n_c \to 1$ as $\varepsilon \to 0$. This is in contrast to the case of a "square" annulus where we show that n_c is bounded away from 1.

1. Introduction. Consider the following percolation process. Fix r > 0 and ε with $0 \le \varepsilon \le 1$. Place points $\{x_i\}$ in \mathbb{R}^2 according to a Poisson process with intensity 1. Join points x_i and x_j provided the Euclidean distance $||x_i - x_j||_2$ lie between $r(1 - \varepsilon)$ and r. We wish to know whether the resulting graph has an infinite component.

More, generally, let A be a bounded region in \mathbb{R}^2 and define an infinite random graph \mathcal{G}_A with vertex set $V(\mathcal{G}_A) \subseteq \mathbb{R}^2$ given by a Poisson process with intensity 1 and edges $xy \in E(\mathcal{G}_A)$ when $x \in y + A$ (where $y + A = \{y + a : a \in A\}$). Define, for $x \in \mathbb{R}^2$, $A_{\varepsilon}(x, r) = \{y : (1 - \varepsilon)r \leq ||y - x||_2 \leq r\}$ to be the annulus centered at x with inner radius $r(1 - \varepsilon)$ and outer radius r. Then the percolation process above is given by \mathcal{G}_A , where $A = A_{\varepsilon}(0, r)$.

For convex A, the model \mathcal{G}_A has been widely studied; see, for example, Grimmett (1999), Häggström and Meester (1996), Jonasson (2001) and especially Meester and Roy (1996). However, very little is known about the model for non-convex bodies such as the annulus A_{ε} .

The models \mathcal{G}_A occur very naturally as ad hoc wireless networks. Suppose that transmitters are distributed at random in the plane and broadcast to neighbors if they are inside a certain region relative to the transmitter (i.e., the set *A*). It is then very natural to ask whether a message can propogate through the network (i.e., does an infinite connected component exist?).

The model $\mathcal{G}_{A_{\varepsilon}(r)}$ is monotonic in *r* since, by scaling, it is equivalent to one in which *r* (and hence *A*) is fixed and the intensity of the Poisson process

Received June 2003; revised September 2003.

¹Supported by NSF Grant EIA-01-30352.

²Supported by NSF Grants DMS-99-70404 and EIA-01-30352 and DARPA Grant F33615-01-C1900.

AMS 2000 subject classifications. Primary 60K35; secondary 82B43.

Key words and phrases. Continuum percolation, continuous percolation, annulus.

varies. Hence, by standard results, for any ε there is a critical value $n_c(\varepsilon)$ of the area $|A| = \pi r^2 \varepsilon (2 - \varepsilon)$ (or, equivalently, the radius r) above which an infinite component exists with probability 1 and below which there is almost surely no infinite component.

Our main result is the following:

THEOREM 1.

$$1 + \frac{\varepsilon}{\pi\sqrt{3}} \le n_c(\varepsilon) \le 1 + o(\varepsilon^{1/15}).$$

In particular, the critical area tends to 1 as $\varepsilon \to 0$. The area |A| is just the expected number of neighbors of a point, so for small ε , an infinite component appears when the average degree of \mathcal{G}_A is slightly more than 1. This is not quite as surprising as it may seem: The giant component in a random graph appears when the average degree of the graph is 1 [see, e.g., Bollobás (2001)].

By comparison, if instead of the annulus $A = A_{\varepsilon}(0, r)$ we use the "square" annulus defined by the l_{∞} -norm, then, as we show, $n_{c}(\varepsilon)$ is bounded away from 1 independently of ε .

We abuse notation by writing |S| for the standard Lebesgue measure in either \mathbb{R}^2 or \mathbb{R} , or for the number of elements of *S*, depending on context.

The layout of this article is as follows: In the next section we prove the lower bound for Theorem 1 and the lower bound for the square annulus. In Section 3 we prove the upper bound in Theorem 1 modulo a key proposition (Proposition 6). In the final section we prove Proposition 6.

After this work was done, we found out that, independently and simultaneously, Franceschetti, Booth, Cook, Meester and Bruck (2003) proved the same result.

2. Lower bound. We first prove the lower bound on $n_c(\varepsilon)$ from Theorem 1. We start with a simple geometric lemma.

LEMMA 2. Suppose $y - x \in A = A_{\varepsilon}(0, r)$. Then $|A_{\varepsilon}(x, r) \cap A_{\varepsilon}(y, r)| \ge \frac{\varepsilon}{\pi \sqrt{3}} |A|$.

PROOF. The proof is easy but tedious, so we shall not spell it out. For small ε , the worst case occurs when $||x - y||_2 = r$. \Box

PROOF OF THE LOWER BOUND OF THEOREM 1. Suppose $|A| \le 1 + \frac{\varepsilon}{\pi\sqrt{3}}$ and fix a vertex x_1 of the Poisson process. Let E_n be the expected number of induced paths of length *n* starting at x_1 . We show that $E_n \to 0$ as $n \to \infty$. Suppose $n \ge 2$ and we have an induced path x_1, \ldots, x_n of length *n*. If this is the initial segment of an induced path x_1, \ldots, x_{n+1} of length n + 1, then $x_{n+1} \in A_{\varepsilon}(x_n, r) \setminus \bigcup_{i=1}^{n-1} A_{\varepsilon}(x_i, r)$, so the expected number of points x_{n+1} ,

1870

conditioned on the points x_1, \ldots, x_n and on the points in $\bigcup_{i=1}^{n-1} A_{\varepsilon}(x_i, r)$ is at most $|A_{\varepsilon}(x_n, r) \setminus \bigcup_{i=1}^{n-1} A_{\varepsilon}(x_i, r)| \le |A_{\varepsilon}(x_n, r) \setminus A_{\varepsilon}(x_{n-1}, r)|$. However, $x_n - x_{n-1} \in A$, so by Lemma 2, this is at most $\alpha = |A|(1 - \frac{\varepsilon}{\pi\sqrt{3}}) < 1$. Hence the expected number of points x_{n+1} conditioned on the points x_1, \ldots, x_n is at most α and $E_{n+1} \le \alpha E_n$. It follows that $E_n \to 0$ as $n \to \infty$, but

 $\mathbb{P}(x_1 \text{ is in an infinite component}) \leq \mathbb{P}(\exists \text{ induced path } x_1, \dots, x_n) \leq E_n$

and hence, with probability 1, x_1 is not in an infinite component. Since x_1 was chosen arbitrarily and there are almost surely countably many points in the Poisson process, \mathcal{G}_A almost surely has no infinite component. \Box

We show in the next section that $n_c \to 1$ as $\varepsilon \to 0$. However, first we show that this is not true for the square annulus, defined by the l_{∞} -norm.

LEMMA 3. Let A be a region in \mathbb{R}^2 with A = -A. Define a percolation process by joining x_i and x_j when $x_i - x_j \in A$. Then the Poisson process with density 1 almost surely does not have an infinite component if $|A|^3 - \int_{A \times A} |(x + A) \cap (y + A)| dx dy < 1$.

PROOF. As before, we count the expected number of induced paths of length *n* starting at x_1 . Suppose x_1, \ldots, x_n are fixed. We count the expected number of choices of x_{n+1}, x_{n+2} and x_{n+3} that give an induced path x_1, \ldots, x_{n+3} , conditional on x_1, \ldots, x_n . Write $x_{n+1} = x_n + x$ and $x_{n+2} = x_{n+1} - y$, so that $x, y \in A = -A$. Now $x_{n+3} \in A' = (x_{n+2} + A) \setminus (x_n + A)$. The area of this set is $|A| - |(x + A) \cap (y + A)|$. From this we see that

$$E_{n+3} \leq \left(\int_{A \times A} |A'| \, dx \, dy\right) E_n = \left(|A|^3 - \int_{A \times A} |(x+A) \cap (y+A)| \, dx \, dy\right) E_n.$$

The result now follows as before. \Box

LEMMA 4. For any interval $I = [a, b], \int_{I \times I} |(x + I) \cap (y + I)| dx dy = \frac{2}{3} |I|^3$.

PROOF. Both sides are unchanged if we replace *I* with I - a = [0, c], where c = b - a. So $\int_{I \times I} |(x + I) \cap (y + I)| dx dy = \int_0^c \int_0^c (c - |x - y|) dx dy = 2 \int_0^c \int_0^y (c - y + x) dx dy = \frac{2}{3}c^3$. \Box

THEOREM 5. Let $A = \{x \in \mathbb{R}^2 : r(1-\varepsilon) \le ||x||_{\infty} \le r\}$. Then \mathcal{G}_A almost surely has no infinite component when $|A| \le 1.014$.

PROOF. Define the intervals $I = [r(1 - \varepsilon), r]$ and J = [-r, r], and the rectangle $S = I \times J$. Then $A \supseteq S \cup -S$. Now

$$\int_{A \times A} |(x+A) \cap (y+A)| \, dx \, dy \ge 6 \int_{S \times S} |(x+S) \cap (y+S)| \, dx \, dy,$$

where we have used symmetry to bound the integral over $x, y \in \pm S$ of $|(x + (S \cup -S)) \cap (y + (S \cup -S))|$ by six equal terms given by the six choices of sign in $x \in \pm S$, $y \in \pm S$, $(x \pm S)$ and $(y \pm S)$ that give the same contribution as $\int_{S \times S} |(x + S) \cap (y + S)| dx dy$. However, *S* is a product $I \times J$, so we can separate variables to obtain

$$\begin{split} &\int_{A \times A} |(x+A) \cap (y+A)| \, dx \, dy \\ &\geq 6 \Big(\int_{I \times I} |(x+I) \cap (y+I)| \, dx \, dy \Big) \Big(\int_{J \times J} |(x+J) \cap (y+J)| \, dx \, dy \Big) \\ &= 6 \big(\frac{2}{3} |I|^3 \big) \big(\frac{2}{3} |J|^3 \big) \geq \frac{1}{24} |A|^3, \end{split}$$

where we have used Lemma 4 and the bound $|A| \le 4|I||J|$. Hence if $|A|^3 \times (1 - \frac{1}{24}) < 1$, then there is no percolation. This occurs if $|A| \le 1.014$. \Box

3. Upper bound. We prove the upper bound of Theorem 1, namely that $n_c = 1 + o(\varepsilon^{1/15})$ as $\varepsilon \to 0$. The strategy is to compare the percolation process with an oriented bond percolation on \mathbb{Z}^2 . Fix a large *R* and integers n < N. Partition \mathbb{R}^2 into $6R \times 6R$ squares, and let the site $x \in \mathbb{Z}^2$ correspond to the square $S_x = 6Rx + [-3R, 3R)^2$ in \mathbb{R}^2 . Our bonds xy in \mathbb{Z}^2 correspond to certain good events in the corresponding $6R \times 12R$ rectangle $S_x \cup S_y$ of \mathbb{R}^2 .

Roughly speaking, these good events comprise the ability to get from some set of *n* points near the middle of S_x to at least *n* points in the middle of S_y by paths that lie entirely within $S_x \cup S_y$. However, we insist on more. We "test" annuli around at most *N* points when we construct the joining paths. Also, there are up to 3*N* points in $S_x \cup S_y$ that have been tested when constructing earlier bonds, so we wish to avoid these points and the annuli about them.

Fix sets of points $P = \{x_1, \ldots, x_n\}$ and $Q = \{y_1, \ldots, y_K\}$, $K \le 3N$, in the rectangle $S_x \cup S_y$ that corresponds to bond xy of \mathbb{Z}^2 . We assume that no two points of $P \cup Q$ lie within $r\sqrt{\varepsilon}$ of each other. Assume further that x_1, \ldots, x_n lie in the middle $4R \times 4R$ square $6Rx + [-2R, 2R)^2$ of S_x . Suppose y is one of the two points x + (1, 0) or x + (0, 1). We construct subsets P' and Q' of the Poisson process with the following properties:

(a) $P' \subseteq 6Ry + [-2R, 2R)^2$ and $|P'| \le n$.

(b) $Q' \subseteq S_x \cup S_y$, every element of Q' is at distance at least *r* from the boundary of $S_x \cup S_y$ and $|Q'| \le N$.

(c) All points of $P \cup Q \cup P' \cup Q'$ are distinct and are at least $r\sqrt{\varepsilon}$ from each other.

(d) $(P' \cup Q') \cap \bigcup_{y \in Q} A_{\varepsilon}(y, r) = \emptyset$.

(e) Every $x' \in P'$ is joined to some point $x \in P$ by a sequence of points y'_1, \ldots, y'_k of Q'.

We declare the oriented bond \overrightarrow{xy} to be *open with respect to P and Q* if, in this construction, |P'| = n.

The idea is that *P* is a set of points in S_x that we can get to from the origin, and we wish to find a set of points *P'* in S_y that we can get to from *P*. The set *Q'* consists of all the points we need to look at when constructing paths from *P* to *P'* and *Q* is the set of points we have looked at previously. When constructing paths from *P* to *P'* it is important that we avoid annuli about points in *Q*, since this would introduce uncontrollable dependencies on our previous bonds.

Note that the openness of \vec{xy} depends on the choice of *P* and *Q* as well as the restriction of the Poisson process to the rectangle $S_x \cup S_y$. The sets *P* and *Q* depend on the construction of previous bonds, which will introduce dependencies on the bonds. The key element of the proof is the following:

PROPOSITION 6. Fix $\varepsilon > 0$ small and write $|A| = 1 + \eta$, where $\eta = \eta(r, \varepsilon)$. If r is large enough that $\varepsilon \le c\eta^{14} |\log \eta|^{-10}$, then there exist N, n and R such that the probability of a bond being open with respect to P and Q is at least 0.9.

The proof of this proposition is deferred until the next section.

PROOF OF THE UPPER BOUND OF THEOREM 1. Using Proposition 6 we complete the proof. We define the oriented percolation on \mathbb{Z}^2 . Order the bonds in the first quadrant of \mathbb{Z}^2 by their l_1 distance from the origin and, for each distance k, order the bonds at distance k from the origin as

$$(0,k)(0,k+1)$$
 $(0,k)(1,k)$ $(1,k-1)(1,k)$... $(k,0)(k+1,0)$.

Suppose the bonds in this order are $\{e_1, e_2, ...\}$. We declare some bonds open in such a way that if there is an infinite directed open path in \mathbb{Z}^2 from (0, 0), then (with positive probability) there is an infinite path in \mathcal{G}_A . To this end, we inductively define, for each edge e_i , a subset Q_i of the Poisson process with $Q_{i-1} \subseteq Q_i$ and, for each vertex x joined by an open path to the origin in \mathbb{Z}^2 , a subset P_x of the Poisson process inside S_x .

Initially set $Q_0 = \emptyset$ and set $P_{(0,0)}$ to be any subset of *n* points of the Poisson process that lie in $[-2R, 2R)^2$ and such that no two points of $P_{(0,0)}$ lie within $r\sqrt{\varepsilon}$ of each other. It is easy to check that such a set exists with high probability.

Now suppose we have defined the openness of the bonds e_j for j < i and the set Q_i and all relevant P_x . We now consider the bond $e_i = xy$.

(a) If there is no directed path consisting of open bonds from (0, 0) to x, set xy to be open and $Q_{i+1} = Q_i$.

(b) If there is a directed open path from (0, 0) to x, declare xy to be open if it is open with respect to $P = P_x$ and $Q = Q_i \cap (S_x \cup S_y)$. Set $Q_{i+1} = Q_i \cup Q'$ and, if xy is open and P_y is not yet defined, set $P_y = P'$.

Condition (a) is a technical condition which clearly does not affect whether or not (0, 0) is in an infinite cluster.

By (b), at each step $|Q_i \cap S_z| \le kN$, where k is the number of edges e_j , j < i, meeting z. Thus, given the ordering of bonds as above, if e_i is a vertical bond, $|Q_i \cap S_x| \le 2N$ and $|Q_i \cap S_y| \le N$, whereas if e_i is a horizontal bond, $|Q_i \cap S_x| \le 3N$ and $|Q_i \cap S_y| = 0$. Hence the set Q in (b) always satisfies $|Q| \le 3N$.

There are two edges $e_i = xy$ with a given value of y, so there are two chances for P_y to be defined in (b). Clearly P_y is defined iff y is joined to (0, 0) by an open path. Also, by construction, all points in $(\bigcup_i Q_i) \cup (\bigcup_x P_x)$ are at least $r\sqrt{\varepsilon}$ from each other. If x is joined to (0, 0) by an open path $x = x_k, \ldots, x_0 = (0, 0)$, then each point in $P_{x_{i+1}}$ is joined by a path in \mathcal{G}_A to a point in P_{x_i} . Hence there is a path from any point of P_x to one of the n points of $P_{(0,0)}$.

Finally, by Proposition 6, if $\varepsilon \leq c\eta^{14} |\log \eta|^{-10}$ [which is satisfied if $\eta = \Omega(\varepsilon^{1/15})$], then each bond is open with probability bounded below by 0.9 even conditioned on the state of all previous edges and regions of \mathbb{R}^2 on which they depend (the annuli around the points of Q_i). Thus the percolation process on \mathbb{Z}^2 stochastically dominates an independent oriented bond percolation with bond probability 0.9. However, (0, 0) is then in an infinite cluster with positive probability [see, e.g., Balister, Bollobás and Stacey (1994)]. The result now follows. \Box

4. Proof of Proposition 6. The proof of the bound is complicated by the fact that the annuli intersect, so we first consider the simpler case when we ignore these intersections and model the percolation by a branching process [see, e.g., Athreya and Ney (1972)]. We generally refer to the points of a branching process as *nodes* to avoid confusion with the points of our Poisson process.

LEMMA 7. Consider a branching process where at each step each node branches into several new nodes according to independent identical Poisson distributions with mean $1 + \eta$. Let N_t be the number of nodes at time t > 0. Then $\mathbb{P}(N_t \ge (1 + \eta)^t) \ge \eta \mathbb{P}(N_t = 0)^2 \ge \eta e^{-2(1+\eta)}$.

PROOF. It is easy to show by induction on *t* that

$$\mathbb{E}N_t = (1+\eta)^t$$

and

Var
$$N_t = (1 + \eta)^t ((1 + \eta)^t - 1)/\eta < (\mathbb{E}N_t)^2/\eta$$
.

Let $X_t = N_t / \mathbb{E}N_t$, so that $\mathbb{E}X_t = 1$ and $\operatorname{Var} X_t < 1/\eta$. By Cauchy–Schwarz,

$$\mathbb{E}(I_{X_t\geq 1})\mathbb{E}((X_t-1)^2)\geq (\mathbb{E}((X_t-1)I_{X_t\geq 1}))^2.$$

$$\mathbb{E}((X_t - 1)^2) = \operatorname{Var} X_t < 1/\eta,$$

$$\mathbb{E}((X_t - 1)I_{X_t \ge 1}) = \mathbb{E}((1 - X_t)I_{X_t \le 1})$$

$$\ge \mathbb{P}(X_t = 0) = \mathbb{P}(N_t = 0).$$

Hence $\mathbb{P}(N_t \ge (1+\eta)^t) = \mathbb{E}(I_{X_t \ge 1}) \ge \eta \mathbb{P}(N_t = 0)^2$. Finally $\mathbb{P}(N_t = 0) \ge \mathbb{P}(N_1 = 0) = e^{-(1+\eta)}$. \Box

LEMMA 8. Assume $\eta > 0$ is sufficiently small and consider a branching process where at each step each node branches into several new nodes independently according to identical Poisson distributions with mean $1 + \eta$. Suppose also that we remove nodes so there are at most K nodes at each step and assume $T \le e^{\eta K} \eta/3$. Then the probability that there is at least one node at time T is at least $\eta/3$.

PROOF. Let N_t be the number of nodes at time *t* and consider the random variable $X_t = \exp(-\lambda N_t)$, where λ is the positive solution to the equation $(1 - e^{-\lambda})(1 + \eta) = \lambda$ [unique since $(1 - e^{-\lambda})/\lambda$ monotonically decreases from 1 to 0]. Now

$$N_{t+1} = \min(N'_{t+1}, K)$$
 where $N'_{t+1} = \sum_{i=1}^{N_t} Y_i$

and Y_i are independent identical Poisson random variables of mean $1 + \eta$. Also

$$\mathbb{E}\exp(-\lambda Y_i) = e^{-(1+\eta)} \sum_{n=0}^{\infty} \frac{(1+\eta)^n}{n!} e^{-\lambda n} = \exp((e^{-\lambda} - 1)(1+\eta)) = e^{-\lambda}$$

Hence

$$\mathbb{E}\left(\exp(-\lambda N_{t+1}')|N_t\right) = \prod_{i=1}^{N_t} \mathbb{E}\exp(-\lambda Y_i) = \exp(-\lambda N_t) = X_t,$$

but if $N_{t+1} \neq N'_{t+1}$, then $K = N_{t+1} < N'_{t+1}$. Hence

$$0 \leq \mathbb{E} \left(\exp(-\lambda N_{t+1}) - \exp(-\lambda N_{t+1}') | N_t \right) \leq e^{-\lambda K}.$$

So $\mathbb{E}(X_{t+1} | N_t) \le X_t + e^{-\lambda K}$ and thus $\mathbb{E}X_{t+1} \le \mathbb{E}X_t + e^{-\lambda K}$. However, $\mathbb{E}X_t \ge \mathbb{P}(N_t = 0)$ and $\mathbb{E}X_0 = \exp(-\lambda N_0) = e^{-\lambda}$. Hence

$$\mathbb{P}(N_T = 0) \le \mathbb{E}X_T \le e^{-\lambda} + Te^{-\lambda K}.$$

It remains to bound λ . Let $f(x) = (1 - e^{-x})(1 + \eta) - x$, so that λ is a solution of $f(\lambda) = 0$. Then $f(\eta) = (\eta - \eta^2/2 + O(\eta^3))(1 + \eta) - \eta$ is positive for small η ,

but $f(1 + \eta) < 0$. Hence $\lambda > \eta$ and thus $\mathbb{P}(N_t = 0) \le e^{-\eta} + Te^{-\eta K} \le e^{-\eta} + \eta/3$, which is at most $1 - \eta/3$ for sufficiently small η . \Box

We now consider a simplified version of our percolation process in which each step is independent of all previous steps. Consider a branching process where at each step each node branches into several new nodes according to independent identical Poisson distributions with mean $1 + \eta$. Assign to each node v (other than the root node) a random variable A_v uniformly distributed in $A = A_{\varepsilon}(0, r)$ independently of the branching process and all other A_u 's. Fix a position z_0 in \mathbb{R}^2 for the root node and define the position z_v of a node v to be $z_0 + \sum_u A_u$, where the sum runs over all ancestors u of v back to the root node. Let \mathcal{T}_A be the random graph with vertices z_v and edges $z_v z_{v'}$ for all child nodes v' of v. Also, define \mathcal{T}_A^t to be the set of nodes that occur at time t. The process \mathcal{T}_A clearly approximates the percolation process \mathcal{G}_A , but it differs in that the distribution of points (child nodes) in $A_{\varepsilon}(z_v, r)$ is independent of the process up to that point, whereas in \mathcal{G}_A the points in $A_{\varepsilon}(z_v, r)$ depend on points in previous annuli where they intersect.

LEMMA 9. Pick $z_0 \in [-2R, 2R)^2$ and consider the branching process \mathcal{T}_A defined above, except that at each step (if necessary) we remove nodes (randomly, independent of their locations) so there are at most K new nodes at each step. Run this process up to time $T = (R/r)^2$. Define the event \mathcal{E} to be the event that the process has not died out at time T (i.e., $\mathcal{T}_A^T \neq \emptyset$) and that, picking a node at random from \mathcal{T}_A^T , this node lies in the square $(0, 6R) + [-R, R)^2$ and all its ancestors lie in $[-3R + r, 3R - r) \times [-3R + r, 9R - r)$. Then there exists an absolute constant $c_0 > 0$, independent of η , ε , r and R, such that for η sufficiently small and R/r sufficiently large, $\mathbb{P}(\mathcal{E}|\mathcal{T}_A^T \neq \emptyset) \ge c_0$.

PROOF. Conditioning on $\mathcal{T}_A^T \neq \emptyset$, pick a node v at random from \mathcal{T}_A^T . Note that the choice of v is independent of the locations of the nodes. Consider the locations of the nodes on the path from the root node to v. These form a random walk with steps taken uniformly from $A = A_{\varepsilon}(0, r)$. The probability we are interested in is bounded below by the probability of this random walk meeting $(0, 6R) + [-R, R)^2$ before hitting the boundary of the rectangle $[-3R+r, 3R-r) \times [-3R+r, 9R-r)$ and before time T. Each step has mean 0 and variance in either coordinate direction of cr^2 , where $c = c(\varepsilon)$ can be bounded above and below by positive constants independently of ε . By scaling the dimensions by 1/R and time by $c'(r/R)^2$, this random walk for large R/r can be approximated by a two-dimensional Brownian motion, run for constant time t_0 , starting in $[-2, 2]^2$. The probability of the Brownian motion lying in $[-0.9, 0.9] \times [5.1, 6.9]$ at time t_0 without hitting the boundary of $[-2.9, 2.9] \times [-2.9, 8.9]$ before time t_0 is bounded below by a constant c'' > 0. Choosing $c_0 < c''$ we see that provided R/r is larger than some absolute constant, the probability that v lies in the square

 $(0, 6R) + [-R, R)^2$ and all its ancestors lie in the rectangle $[-3R + r, 3R - r) \times [-3R + r, 9R - r)$ is at least c_0 . Moreover, the bound on R/r and the value of c_0 can be chosen independently of r, R, η and ε . \Box

We now need to deal with the dependencies caused by the intersections of annuli. We start with a couple of geometric lemmas.

LEMMA 10. There exists a constant $c_1 > 0$ such that for any r, ε and x, y, with $||x - y||_2 \ge r\sqrt{\varepsilon}$, we have $|A_{\varepsilon}(x, r) \cap A_{\varepsilon}(y, r)| \le c_1 |A|\sqrt{\varepsilon}$.

PROOF. Tedious, but straightforward verification. Note that there are two cases when the bound is tight: one when $||x - y||_2 \approx r\sqrt{\varepsilon}$ and the other when $||x - y||_2 \approx r(2 - \varepsilon)$.

Note that this lemma fails for the square annulus when $||x - y||_{\infty} \approx r(2 - \varepsilon)$.

LEMMA 11. Let x_1, \ldots, x_k be points, no two of which are within $r\sqrt{\varepsilon}$ of each other. Let $A_i = A_{\varepsilon}(x_i, r)$ and let $B_i = A_1(x_i, r\sqrt{\varepsilon})$ be the ball around x_i of radius $r\sqrt{\varepsilon}$. Then

$$\left|A_i \cap \left(\bigcup_{j \neq i} (A_j \cup B_j)\right)\right| \le c_2 k |A| \sqrt{\varepsilon}.$$

PROOF. The region $A_i \cap B_j$ has area $O((r\varepsilon)(r\sqrt{\varepsilon})) = O(|A|\sqrt{\varepsilon})$ since A_i is of "width" $r\varepsilon$ and B_j is of diameter $2r\sqrt{\varepsilon}$. The region $A_i \cap A_j$ has area $O(|A|\sqrt{\varepsilon})$ by Lemma 10. The result follows. \Box

PROOF OF PROPOSITION 6. The strategy of the proof is to find some points x_i'' in $6Ry + [-R, R)^2$ by comparing the process \mathcal{G}_A with \mathcal{T}_A . There will, however, be rather fewer than *n* such points, so we then run the percolation for R/r further steps in \mathcal{G}_A to obtain sufficiently many points in $6Ry + [-2R, 2R)^2$ (note that we cannot travel more than distance *R* in R/r steps).

Pick each $x_i \in P$ in turn and run a truncated branching process \mathcal{T}_A as in Lemma 9, starting at x_i . We call a node *bad* if it lies in any annulus $A_{\varepsilon}(z, r)$ or ball $A_1(z, r\sqrt{\varepsilon})$, where z is in $Q, P \setminus \{x_i\}$ or any one of the nodes of the branching process other than the parent of v, or any one of the points in the branching processes already constructed for $x_j \in P$, j < i. There are at most $3N + n + nK(R/r)^2$ such values of z. Set

(1)
$$K = \frac{1}{\eta} \log \left(\frac{3R^2}{\eta r^2} \right).$$

It is clear that we can run the same branching process in the percolation model \mathcal{G}_A , coupling \mathcal{G}_A with \mathcal{T}_A so that they agree up until we hit a bad node of \mathcal{T}_A . By

Lemma 8, the branching process has not died out by time $T = (R/r)^2 = e^{\eta K} \eta/3$ with probability at least $\eta/3$. Moreover, the probability that a given node in this process is bad conditioned on all its predecessors being good is at most $c_2(3N + n + nK(R/r)^2)|A|\sqrt{\varepsilon}/|A|$ by Lemma 11. Also this is independent of any event involving the existence of its decendants in the branching process. Thus, if we condition on $\mathcal{T}_A^T \neq \emptyset$ and we pick a path of length T (independently of locations of the nodes), the probability that this path contains a bad node is at most $(R/r)^2 c_2(3N + n + nK(R/r)^2)\sqrt{\varepsilon}$. We require

(2)
$$(R/r)^2 c_2 (3N + n + nK(R/r)^2) \sqrt{\varepsilon} \le c_0/2.$$

Thus by Lemma 9, conditioning on $\mathcal{T}_A^T \neq \emptyset$, we obtain a point x_i'' in $6Ry + [-R, R)^2$ joined to x_i in the percolation process within $S_x \cup S_y$ with probability at least $c_0 - c_0/2 = c_0/2$. Thus, by Lemma 8, the probability of finding such a point x_i'' is at least $(c_0/2)(\eta/3) = c_0\eta/6$, independently of all previously found points. The number of points x_i'' found is hence stochastically bounded below by a binomial variable X with mean $c_0n\eta/6$. Now take the points $V_0 = \{x_1'', \ldots, x_X''\}$ and construct inductively sets V_i by taking all points joined to some point in V_{i-1} which are not in $A_{\varepsilon}(z, r) \cup A_1(z, r\sqrt{\varepsilon})$ for any point z already considered. Repeat for R/r steps or until $|V_i| \ge n$ if this occurs earlier. All points of V_i must then lie in $6Ry + [-2R, 2R)^2$ and the cardinalities $|V_i|$ are stochastically dominated by a Poisson branching process with mean $1 + \eta/2$ provided

(3)
$$c_2(3N+n+nK(R/r)^2+n(R/r))\sqrt{\varepsilon} \le \eta/2$$

Define Q' to be all the new points encountered above which are in $6Ry + [-3R, 3R)^2$ except those of the last V_i and let P' consist of n points from this last V_i , if they exist. Let

(4)
$$N = nK(R/r)^2 + n(R/r).$$

Then N > |Q'|. Finally, the probability of the bond being open is bounded below by the probability that a branching process with parameter $1 + \eta/2$ starting from *X* points will have at least *n* points by time R/r. If

$$(5) n \le (1+\eta/2)^{R/r}$$

then by Lemma 7, the probability that such a branching process starting with one node has at least *n* nodes by time R/r is at least $(\eta/2)e^{-2(1+\eta/2)}$, which is at least $\eta/20$ for small η . Now the probability of not having at least *n* nodes by time R/r starting from X points is bounded above by $(1 - \eta/20)^X$, which has expected value

$$\sum_{i} {n \choose i} (c_0 \eta / 6)^i (1 - c_0 \eta / 6)^{n-i} (1 - \eta / 20)^i$$

= $(1 - c_0 \eta^2 / 120)^n$
 $\leq \exp(-c_0 n \eta^2 / 120).$

1878

The result follows provided

(6)
$$c_0 n \eta^2 / 120 \ge -\log(0.1).$$

It remains to choose the parameters ε , n, N and R so that equations (1)–(6) are satisfied. For (6) we can define $n = c_3 \eta^{-2}$ for some $c_3 > 0$. Then (5) is satisfied if we take $R/r = c_4 \eta^{-1} |\log \eta|$. From (1) and (4) we get $N = O(\eta^{-5} |\log \eta|^3)$, so (2) and (3) are now satisfied if $\sqrt{\varepsilon} \le c_5 \eta^7 |\log \eta|^{-5}$ or, equivalently, $\varepsilon \le c\eta^{14} |\log \eta|^{-10}$. \Box

Note that the sets P', Q' depend only on the points of the Poisson process inside the region $(S_x \cup S_y) \cap (\bigcup_{z \in P \cup Q'} A_{\varepsilon}(z, r)) \setminus (\bigcup_{z \in Q} A_{\varepsilon}(z, r))$, so each bond is independent of the regions tested when constructing previous bonds.

REFERENCES

ATHREYA, K. B. and NEY, P. (1972). Branching Processes. Springer, Berlin.

BALISTER, P., BOLLOBÁS, B. and STACEY, A. (1994). Improved upper bounds for the critical probability of oriented percolation in two dimensions. *Random Structures Algorithms* 5 573–589.

BOLLOBÁS, B. (2001). Random Graphs, 2nd ed. Cambridge Univ. Press.

FRANCESCHETTI, M., BOOTH, L., COOK, M., MEESTER, R. and BRUCK, J. (2003). Unpublished manuscript.

GRIMMETT, G. R. (1999). Percolation, 2nd ed. Springer, New York.

HÄGGSTRÖM, O. and MEESTER, R. (1996). Nearest neighbor and hard sphere models in continuum percolation. *Random Structures Algorithms* **9** 295–315.

JONASSON, J. (2001). Optimization of shape in continuum percolation. *Ann. Probab.* **29** 624–635. MEESTER, R. and ROY, R. (1996). *Continuum Percolation*. Cambridge Univ. Press.

DEPARTMENT OF MATHEMATICS UNIVERSITY OF MEMPHIS DUNN HALL 3725 NORISWOOD MEMPHIS, TENNESSEE 38152 USA E-MAIL: balistep@msci.memphis.edu bollobas@msci.memphis.edu mjw1009@cam.ac.uk