

Research Article

Optimal Kalman Filtering for a Class of State Delay Systems with Randomly Multiple Sensor Delays

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The optimal Kalman filtering problem is investigated for a class of discrete state delay stochastic systems with randomly multiple sensor delays. The phenomenon of measurement delay occurs in a random way and the delay rate for each sensor is described by a Bernoulli distributed random variable with known conditional probability. Based on the innovative analysis approach and recursive projection formula, a new linear optimal filter is designed such that, for the state delay and randomly multiple sensor delays with different delay rates, the filtering error is minimized in the sense of mean square and the filter gain is designed by solving the recursive matrix equation. Finally, a simulation example is given to illustrate the feasibility and effectiveness of the proposed filtering scheme.

1. Introduction

Over the past decades, the estimation/filtering problems have received considerable attention due to their wide application in many practical systems [1–5]. Accordingly, the optimal state estimation/filtering problems have been one of the mainstream research topics in the signal processing field and a large number of results have been reported to design the optimal estimators; see, for example, [6–9]. Generally speaking, the aim of the optimal state estimation is to estimate the internal state of a dynamic system based on the available measurement data and the estimation principle is in the sense of the minimum mean square error (MMSE). Note that the dynamic behaviours of the engineering systems are described by the internal state variable of the systems. Hence, it is important to design the filter so as to further understand and control a practical system and then achieve the desired performance requirements [10–12]. It is worth mentioning that, based on the MMSE principle and the projection theory, the famous Kalman filter has been constructed in [13] for linear discrete stochastic systems and the recursive estimation method has been developed. Subsequently, many important results have been published based on this pioneering method. Moreover, some effective yet easy-to-implement filtering

algorithms have been developed in [14, 15] for complex dynamical systems with the prevalently network-induced phenomena.

It is well known that the time delay is inevitable in many industrial process systems [16–22]. Also, it is necessary to deal with the time delay to improve the control performance for the practical systems [23–26]. In the past years, a great deal of effort has been devoted to address the problems of the optimal state estimation for time delay systems. To mention a few, by applying the state augmentation approach, the problem of the optimal state estimation has been investigated in [17] for linear discrete stochastic systems with measurement delay. In [19–22], the optimal filters have been designed for linear state delay systems. In particular, by using the state augmentation approach, the optimal filter has been designed in [19]. Without using the state augmentation approach, a new optimal filter has been constructed in [20] for linear discrete state delay stochastic systems by applying the projection theory and recursive projection formula. It should be noted that the dimension of the proposed filter in [20] is the same as the original system state and then the computational burden can be reduced. Based on the method in [20], the problem of optimal filtering has been studied in [21] for linear discrete state delay systems under uncertain observations. In [22], an

effective robust Kalman filter has been designed for a class of uncertain state delay systems with random observation delays and missing measurements.

Due to the sudden changes in the environment and unreliability of the communication network, the sensor measurement of the system may experience the unexpected measurement delays in reality [27–29]. However, it is worthwhile mentioning that most of the published results have tackled the deterministic delays only. In fact, it is necessary to deal with the random sensor delays where the communication transmission is commonly unreliable and the filtering performance would be degraded. Recently, the problem of linear minimum variance estimation has been investigated in [30] by applying the state augmentation approach for systems with bounded random measurement delays and packet dropouts. In [31], a new optimal filtering scheme has been developed for networked control system subject to random delay and packet dropouts. By applying the quasi Markov-chain approach, the optimal Kalman filtering problem has been studied in [32] for networked control systems with random measurement delays, packet dropouts, and missing measurement. Recently, the recursive filters have been designed in [33, 34] for discrete-time systems with different delay rates. However, to the best of authors' knowledge, the problem of the linear optimal filtering has not been thoroughly investigated for discrete state delay systems with randomly multiple sensors delay which constitutes our research motivation.

Motivated by the above discussions, in this paper, we aim to investigate the linear optimal filtering problem for a class of discrete state delay systems measured by multiple sensors with different delay rates. The time delay exists in the system state and the measurement output may experience the random one-step sensor delay probably due to the unreliable communication transmissions. The considered phenomena of multiple measurement delays are characterized by a set of Bernoulli distributed random variables with known conditional probabilities. Based on the MMSE estimation principle, the linear optimal filter is designed which can deal with the effects from the state delay and randomly multiple sensor delays in a unified framework. The main contribution of this paper is to make first attempt to design the optimal filter for state delay systems with randomly multiple sensor delays. Accordingly, a new filtering algorithm is developed and filter gain is obtained recursively by the solutions to the matrix equations. Without resorting to the state augmentation approach, the dimension of the developed filter is the same as the system state and then the proposed filtering algorithm can reduce the computational burden. Finally, a simulation example is shown to verify the feasibility and usefulness of the proposed filtering approach.

Notations. The notations used throughout the paper are standard. \mathbb{R}^n denotes the n -dimensional Euclidean space. For a matrix A , A^T represents its transpose. $\mathbb{E}\{x\}$ represents the expectation of a random variable x . I and 0 represent the identity matrix and the zero matrix with appropriate dimensions, respectively. $\text{diag}\{X_1, X_2, \dots, X_N\}$ stands for a diagonal matrix with elements X_1, X_2, \dots, X_N in the diagonal. The Hadamard product is defined as $[T \circ S]_{p \times p} = [t_{ij} \times$

$s_{ij}]_{p \times p}$. Matrices are assumed to be compatible for algebraic operations if their dimensions are not explicitly stated.

2. Problem Formulation and Preliminaries

We consider the following class of linear discrete stochastic systems with state delay:

$$x_{k+1} = A_k x_{k-d} + B_k \omega_k, \quad (1)$$

$$z_k^i = C_k^i x_k + v_k^i, \quad (2)$$

$$y_k^i = \lambda_k^i z_k^i + (1 - \lambda_k^i) z_{k-1}^i, \quad (3)$$

where $i = 1, 2, \dots, N$, $x_k \in \mathbb{R}^n$ is the state vector, d is the state delay, $z_k^i \in \mathbb{R}^m$ is the i th actual output, and $y_k^i \in \mathbb{R}^m$ is the measured output of the i th sensor. $\omega_k \in \mathbb{R}^r$ and $v_k^i \in \mathbb{R}^m$ are uncorrelated zero-mean Gaussian white noises with covariances $Q_k \geq 0$ and $R_k^i > 0$, respectively. A_k , B_k , and C_k^i are known matrices with appropriate dimensions. The random variables λ_k^i obey the Bernoulli distribution and have the following statistical properties:

$$P\{\lambda_k^i = 1\} = \mathbb{E}\{\lambda_k^i\} = \alpha_i, \quad (4)$$

$$P\{\lambda_k^i = 0\} = 1 - \mathbb{E}\{\lambda_k^i\} = 1 - \alpha_i,$$

where $i = 1, 2, \dots, N$, $\alpha_i \in [0, 1]$ are known positive scalars, and λ_k^i are uncorrelated with other noise signals.

Remark 1. In model (3), if $\lambda_k^i = 1$, $y_k^i = z_k^i$, it represents that the i th sensor receives successfully the data at time instant k . If $\lambda_k^i = 0$, $y_k^i = z_{k-1}^i$, it stands for the fact that there exists one-step delay.

Setting

$$\begin{aligned} z_k &= \begin{bmatrix} z_k^1 \\ z_k^2 \\ \vdots \\ z_k^N \end{bmatrix}, \quad C_k = \begin{bmatrix} C_k^1 \\ C_k^2 \\ \vdots \\ C_k^N \end{bmatrix}, \quad v_k = \begin{bmatrix} v_k^1 \\ v_k^2 \\ \vdots \\ v_k^N \end{bmatrix}, \\ y_k &= \begin{bmatrix} y_k^1 \\ y_k^2 \\ \vdots \\ y_k^N \end{bmatrix}, \quad \Lambda_k = \begin{bmatrix} \lambda_k^1 & & & \\ & \lambda_k^2 & & \\ & & \ddots & \\ & & & \lambda_k^N \end{bmatrix}, \end{aligned} \quad (5)$$

then, the systems (1)–(3) can be rewritten as follows:

$$x_{k+1} = A_k x_{k-d} + B_k \omega_k, \quad (6)$$

$$z_k = C_k x_k + v_k, \quad (7)$$

$$y_k = \Lambda_k z_k + (I - \Lambda_k) z_{k-1}. \quad (8)$$

Substituting (7) into (8), we have

$$x_{k+1} = A_k x_{k-d} + B_k \omega_k, \quad (9)$$

$$y_k = \Lambda_k C_k x_k + (I - \Lambda_k) C_{k-1} x_{k-1} + \Lambda_k v_k + (I - \Lambda_k) v_{k-1}. \quad (10)$$

It is not difficult to verify that the following statistical properties are true:

$$\begin{aligned} \Lambda &= \mathbb{E} \{\Lambda_k\} = \text{diag} \{\alpha_1, \alpha_2, \dots, \alpha_N\}, \\ R_k &= \mathbb{E} \{v_k v_k^T\} = \text{diag} \{R_k^1, R_k^2, \dots, R_k^N\}, \\ R_{k-1} &= \mathbb{E} \{v_{k-1} v_{k-1}^T\} = \text{diag} \{R_{k-1}^1, R_{k-1}^2, \dots, R_{k-1}^N\}. \end{aligned} \quad (11)$$

The purpose of this paper is to design the linear optimal filter of state x_{k+1} in the sense of minimum variance for the discrete-time delay stochastic systems (1)–(3) based on the observation sequence $\{y_1, y_2, \dots, y_k\}$.

3. Main Results

In order to facilitate the subsequent developments, we introduce the following definitions.

Definition 2. Let $\tilde{x}_{i|k} = x_i - \hat{x}_{i|k}$. Then, define $\Phi_{k(i,j)} = \mathbb{E}\{\tilde{x}_{i|k} \tilde{x}_{j|k}^T\}$, where $i \neq j$. Particularly, $P_{i|k} = \Phi_{k(i,i)} = \mathbb{E}\{\tilde{x}_{i|k} \tilde{x}_{i|k}^T\}$, when $i = j$. In addition, $\Phi_{k(i,j)} = \Phi_{k(j,i)}^T$.

Definition 3. Define $\Theta_{(k,j)} = \mathbb{E}\{x_k x_j^T\}$, where $k \neq j$. Particularly, $\Theta_{(k,k)} = \mathbb{E}\{x_k x_k^T\}$, when $k = j$. Also, $\Theta_{(k,j)} = \Theta_{(j,k)}^T$. Then, $\Theta_{(k,k)}$ can be calculated as follows:

$$\Theta_{(k+1,k+1)} = A_k \Theta_{(k-d,k-d)} A_k^T + B_k Q_k B_k^T. \quad (12)$$

Definition 4. Define $\Sigma_k^t = \Theta_{(k-t,k-d)}$, where $t = 0, 1, \dots, d$. Then, Σ_k^t can be calculated as follows:

$$\Sigma_k^t = \Theta_{(k-t,k-d)} = A_{k-t-1} \Theta_{(k-t-1-d,k-d)} = A_{k-t-1} \left(\Sigma_{k-t-1}^{d-t-1} \right)^T, \quad (13)$$

where $\Sigma_k^d = \Theta_{(k-d,k-d)}$ can be computed by (12).

Definition 5. Define a series of the matrices as follows:

$$\begin{aligned} \Gamma_k^i &= \mathbb{E} \{ \varepsilon_k x_{k-i}^T \}, & Q_{\varepsilon_k} &= \mathbb{E} \{ \varepsilon_k \varepsilon_k^T \}, \\ K_k^i &= \mathbb{E} \{ x_{k-i} \varepsilon_k^T \} Q_{\varepsilon_k}^{-1} = \left(\Gamma_k^i \right)^T Q_{\varepsilon_k}^{-1}, \\ \Psi_k^i &= \Phi_{k(k,k-i+1)}, & \Pi_{(k,j)}^\mu &= \mathbb{E} \{ x_{k+1-j} x_{k-j-\mu}^T \}, \\ \Omega_{(k,j)}^\mu &= \Phi_{k(k+1-j,k-j-\mu)}, & \Xi_k^i &= \Phi_{k(k-d,k-i+1)}, \end{aligned} \quad (14)$$

where $\varepsilon_k = y_k - \hat{y}_{k|k-1}$ is the innovation sequence.

Next, we are ready to introduce the following lemmas which will be used for the further developments.

Lemma 6. Ψ_k^i and Ξ_k^i satisfy the following recursive equation:

$$\Psi_k^i = A_{k-1} \Xi_{k-1}^{i-1} - K_k \Gamma_{k-1}^{i-1}, \quad i = 2, \dots, d+1, \quad (15)$$

$$\Xi_k^i = \begin{cases} \left(\Psi_{k-d+1}^{d+2-i} \right)^T - \sum_{\rho=0}^{i-2} K_{k-\rho}^{d-\rho} \Gamma_{k-\rho}^{i-\rho-1}, & i = 2, \dots, d, \\ P_{k-d|k}, & i = d+1, \end{cases} \quad (16)$$

where $K_k^0 = K_k$, $\Psi_k^1 = P_{k|k}$, and $\Xi_k^1 = (\Psi_k^{d+1})^T$.

Proof. From (10), we can define the innovation sequence by the projection theory as follows:

$$\begin{aligned} \varepsilon_k &= y_k - \Lambda C_k \hat{x}_{k|k-1} - (I - \Lambda) C_{k-1} \hat{x}_{k-1|k-1} \\ &= (\Lambda_k - \Lambda) C_k x_k + \Lambda C_k \tilde{x}_{k|k-1} \\ &\quad - (\Lambda_k - \Lambda) C_{k-1} x_{k-1} + (I - \Lambda) C_{k-1} \tilde{x}_{k-1|k-1} \\ &\quad + \Lambda_k v_k + (I - \Lambda_k) v_{k-1}. \end{aligned} \quad (17)$$

By the definitions of Ψ_k^i and Ξ_k^i , the following equations can be obtained:

$$\Psi_k^1 = P_{k|k}, \quad \Xi_k^1 = \left(\Psi_k^{d+1} \right)^T, \quad \Xi_k^{d+1} = P_{k-d|k}. \quad (18)$$

Subsequently, by employing the projection theory, the recursive equations can be obtained as follows:

$$\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k \varepsilon_k, \quad (19)$$

$$K_k = \mathbb{E} \{ x_k \varepsilon_k^T \} \left[\mathbb{E} \{ \varepsilon_k \varepsilon_k^T \} \right]^{-1},$$

$$\hat{x}_{k-i+1|k} = \hat{x}_{k-i+1|k-1} + K_k^{i-1} \varepsilon_k, \quad (20)$$

$$K_k^{i-1} = \mathbb{E} \{ x_{k-i+1} \varepsilon_k^T \} \left[\mathbb{E} \{ \varepsilon_k \varepsilon_k^T \} \right]^{-1}.$$

Then, the following equation can be obtained by (19) and (20):

$$\begin{aligned} \Phi_{k(k,k-i+1)} &= \Phi_{k-1(k,k-i+1)} - K_k \mathbb{E} \{ \varepsilon_k \tilde{x}_{k-i+1|k-1}^T \} \\ &\quad - \mathbb{E} \{ \tilde{x}_{k|k-1} \varepsilon_k^T \} \left(K_k^{i-1} \right)^T + K_k \mathbb{E} \{ \varepsilon_k \varepsilon_k^T \} \left(K_k^{i-1} \right)^T. \end{aligned} \quad (21)$$

Noting that the following fact is true.

$$\begin{aligned} \mathbb{E} \{ \Lambda_k - \Lambda \} &= 0, & \hat{x}_{k-i+1|k-1} &\perp \tilde{x}_{k|k-1}, \\ \hat{x}_{k-i+1|k-1} &\perp \tilde{x}_{k-1|k-1}. \end{aligned} \quad (22)$$

We can establish the following equation:

$$\begin{aligned} \mathbb{E} \{ \varepsilon_k \tilde{x}_{k-i+1|k-1}^T \} &= \mathbb{E} \{ \varepsilon_k (\hat{x}_{k-i+1|k-1} + \tilde{x}_{k-i+1|k-1})^T \} \\ &= \mathbb{E} \{ \varepsilon_k x_{k-i+1}^T \} = \Gamma_k^{i-1}. \end{aligned} \quad (23)$$

Substituting K_k of (19) into (21) leads to

$$\Phi_{k(k,k-i+1)} = \Phi_{k-1(k,k-i+1)} - K_k \Gamma_{k-1}^{i-1}. \quad (24)$$

Similarly, we have

$$\Phi_{k-\rho(k-d,k-i+1)} = \Phi_{k-\rho-1(k-d,k-i+1)} - K_{k-\rho}^{d-\rho} \Gamma_{k-\rho}^{i-\rho-1}. \quad (25)$$

From (9), we have

$$x_k = A_{k-1}x_{k-d-1} + B_{k-1}\omega_{k-1}. \quad (26)$$

Taking projection on both sides of (26), one has

$$\hat{x}_{k|k-1} = A_{k-1}\hat{x}_{k-d-1|k-1}. \quad (27)$$

Subtracting both sides of (27) by (26) yields

$$\tilde{x}_{k|k-1} = A_{k-1}\tilde{x}_{k-d-1|k-1} + B_{k-1}\omega_{k-1}. \quad (28)$$

Since $\omega_{k-1} \perp \tilde{x}_{k-i+1|k-1}$, the following equation can be derived:

$$\Phi_{k-1(k,k-i+1)} = A_{k-1}\Phi_{k-1(k-d-1,k-i+1)}. \quad (29)$$

By the definitions of Ψ_k^i and Ξ_k^i , substituting (29) into (24) yields

$$\Psi_k^i = A_{k-1}\Xi_{k-1}^{i-1} - K_k\Gamma_k^{i-1}. \quad (30)$$

According to (25), we can obtain the following equation:

$$\Xi_k^i = \left(\Psi_{k-d+1}^{d+2-i}\right)^T - \sum_{\rho=0}^{i-2} K_{k-\rho}^{d-\rho} \Gamma_k^{i-\rho-1}, \quad i = 2, \dots, d. \quad (31)$$

Therefore, (15) can be derived by (18) and (30). Moreover, it follows from (18) and (31) that (16) holds. Then, the proof of this lemma is complete. \square

Lemma 7. For $j = 0, 1$ and $\mu = 0, 1, \dots, d-1$, $\Pi_{(k,j)}^\mu$ and $\Omega_{(k,j)}^\mu$ satisfy the following recursive equations:

$$\Pi_{(k,j)}^\mu = A_{k-j} \left(\Sigma_{k-j}^\mu \right)^T, \quad (32)$$

$$\Omega_{(k,j)}^\mu = \begin{cases} A_k \Xi_k^{\mu+1}, & j = 0, \\ \Psi_k^{\mu+2}, & j = 1, \end{cases} \quad (33)$$

where Σ_{k-j}^μ is computed by (13) and $\Psi_k^{\mu+2}$ and $\Xi_k^{\mu+1}$ are calculated by (15) and (16), respectively.

Proof. From (9), we have

$$x_{k+1-j} = A_{k-j}x_{k-j-d} + B_{k-j}\omega_{k-j}. \quad (34)$$

By the fact that $x_{k-j-\mu} \perp \omega_{k-j}$ as well as the definitions of $\Pi_{(k,j)}^\mu$ and Σ_k^t , we have

$$\begin{aligned} \mathbb{E} \{x_{k+1-j}x_{k-j-\mu}^T\} &= A_{k-j} \mathbb{E} \{x_{k-j-d}x_{k-j-\mu}^T\}, \\ \Pi_{(k,j)}^\mu &= A_{k-j} \Theta_{(k-j-d,k-j-\mu)} = A_{k-j} \left(\Sigma_{k-j}^\mu \right)^T. \end{aligned} \quad (35)$$

Thus, (32) is true.

At the same time, by the definition of $\Omega_{(k,j)}^\mu$, we have

$$\Omega_{(k,0)}^\mu = \Phi_{k(k+1,k-\mu)}, \quad j = 0, \quad (36)$$

$$\Omega_{(k,1)}^\mu = \Phi_{k(k,k-1-\mu)}, \quad j = 1. \quad (37)$$

For (36), by the same idea of (29) and the definition of Ξ_k^i , it can be obtained that

$$\Omega_{(k,0)}^\mu = \Phi_{k(k+1,k-\mu)} = A_k \Phi_{k(k-d,k-\mu)} = A_k \Xi_k^{\mu+1}. \quad (38)$$

For (37), by the definition of Ψ_k^i , it can be obtained that

$$\Omega_{(k,1)}^\mu = \Phi_{k(k,k-1-\mu)} = \Psi_k^{\mu+2}. \quad (39)$$

Therefore, it can be shown that (33) holds according to (38) and (39). The proof of this lemma is complete. \square

Lemma 8. For $j = 0, 1$, one has

$$\Phi_{k-1(k-j,k-i)} = \begin{cases} A_{k-1}\Xi_{k-1}^i, & j = 0, \\ \Psi_{k-1}^i, & j = 1. \end{cases} \quad (40)$$

Proof. When $j = 0$, by the same line of (29) and the definition of Ξ_k^i , we obtain

$$\Phi_{k-1(k,k-i)} = A_{k-1}\Phi_{k-1(k-d-1,k-i)} = A_{k-1}\Xi_{k-1}^i. \quad (41)$$

When $j = 1$, by the definition of Ψ_k^i , we have

$$\Phi_{k-1(k-1,k-i)} = \Psi_{k-1}^i. \quad (42)$$

Therefore, it follows from (41) and (42) that (40) is true. The proof of this lemma is complete. \square

Lemma 9 (see [35]). Let $T = [t_{ij}]_{p \times p}$ be a real matrix and $S = \text{diag}\{s_1, s_2, \dots, s_p\}$ a diagonal random matrix. Then

$$\mathbb{E} \{STS^T\} = \begin{bmatrix} \mathbb{E} \{s_1^2\} & \mathbb{E} \{s_1 s_2\} & \cdots & \mathbb{E} \{s_1 s_p\} \\ \mathbb{E} \{s_2 s_1\} & \mathbb{E} \{s_2^2\} & \cdots & \mathbb{E} \{s_2 s_p\} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbb{E} \{s_p s_1\} & \mathbb{E} \{s_p s_2\} & \cdots & \mathbb{E} \{s_p^2\} \end{bmatrix} \circ T, \quad (43)$$

where \circ is the Hadamard product.

Now, we are ready to design the linear recursive optimal filter for systems (9)-(10) by employing the observation sequence $\{y_1, y_2, \dots, y_k\}$. Based on the above lemmas and motivated by [22], we have the following theorem.

Theorem 10. The recursive optimal filter for systems (9)-(10) is given as follows:

$$\hat{x}_{k+1|k+1} = \hat{x}_{k+1|k} + K_{k+1}\varepsilon_{k+1}, \quad (44)$$

$$\varepsilon_{k+1} = y_{k+1} - \Lambda C_{k+1}\hat{x}_{k+1|k} - (I - \Lambda)C_k\hat{x}_{k|k}, \quad (45)$$

$$K_{k+1} = [P_{k+1|k}C_{k+1}^T\Lambda + A_k\Xi_k^1C_k^T(I - \Lambda)]Q_{\varepsilon_{k+1}}^{-1}, \quad (46)$$

$$\begin{aligned} Q_{\varepsilon_{k+1}} = & \Lambda C_{k+1}P_{k+1|k}C_{k+1}^T\Lambda + (I - \Lambda)C_kP_{k|k}C_k^T(I - \Lambda) \\ & + \Lambda C_{k+1}\Omega_{(k,0)}^0C_k^T(I - \Lambda) \\ & + (\Lambda C_{k+1}\Omega_{(k,0)}^0C_k^T(I - \Lambda))^T \\ & + \Lambda R_{k+1}\Lambda + (I - \Lambda)R_k(I - \Lambda) \\ & + \mathcal{H}_1 + \mathcal{H}_2 - \mathcal{H}_3 - \mathcal{H}_3^T, \end{aligned} \quad (47)$$

$$\begin{aligned} P_{k+1|k+1} = & P_{k+1|k} \\ & - K_{k+1}[\Lambda C_{k+1}P_{k+1|k} + (I - \Lambda)C_k(A_k\Xi_k^1)^T], \end{aligned} \quad (48)$$

$$\hat{x}_{k+1|k} = A_k\hat{x}_{k-d|k}, \quad (49)$$

$$P_{k+1|k} = A_kP_{k-d|k}A_k^T + B_kQ_kB_k^T, \quad (50)$$

$$\hat{x}_{k-d|k} = \hat{x}_{k-d|k-d} + \sum_{\rho=0}^{d-1} K_{k-\rho}^{d-\rho}\varepsilon_{k-\rho}, \quad (51)$$

$$P_{k-d|k} = P_{k-d|k-d} - \sum_{\rho=0}^{d-1} K_{k-\rho}^{d-\rho}\Gamma_{k-\rho}^{d-\rho}, \quad (52)$$

where

$$\begin{aligned} \Gamma_k^i = & \begin{cases} \Lambda C_k\Omega_{(k-1,0)}^0 + (I - \Lambda)C_{k-1}P_{k-1|k-1}, & i = 1, \\ \Lambda C_k\Omega_{(k-1,0)}^{i-1} + (I - \Lambda)C_{k-1}\Omega_{(k-1,1)}^{i-2}, & i = 2, \dots, d, \end{cases} \\ \mathcal{H}_1 = & \Lambda(I - \Lambda) \circ C_{k+1}\Theta_{(k+1,k+1)}C_{k+1}^T, \\ \mathcal{H}_2 = & \Lambda(I - \Lambda) \circ C_k\Theta_{(k,k)}C_k^T, \\ \mathcal{H}_3 = & \Lambda(I - \Lambda) \circ C_{k+1}\Pi_{(k,0)}^0C_k^T \end{aligned} \quad (53)$$

and $K_{k-\rho}^{d-\rho}$ is computed by Definition 5, $\Gamma_{k-\rho}^{d-\rho}$ is computed by (53), $\Theta_{(k,k)}$ and Ξ_k^1 are calculated by (12) and (16), respectively, $\Pi_{(k,0)}^0$ and $\Omega_{(k,0)}^0$ are computed by Lemma 7, and \mathcal{H}_1 , \mathcal{H}_2 , and \mathcal{H}_3 are calculated by Lemma 9.

Proof. By the projection formula, we have

$$\hat{x}_{k+1|k+1} = \hat{x}_{k+1|k} + K_{k+1}\varepsilon_{k+1}, \quad (54)$$

$$K_{k+1} = \mathbb{E}\{x_{k+1}\varepsilon_{k+1}^T\}Q_{\varepsilon_{k+1}}^{-1}. \quad (55)$$

According to (10), we have the innovation equation as follows:

$$\begin{aligned} \varepsilon_{k+1} = & y_{k+1} - \Lambda C_{k+1}\hat{x}_{k+1|k} - (I - \Lambda)C_k\hat{x}_{k|k} \\ = & (\Lambda_{k+1} - \Lambda)C_{k+1}x_{k+1} + \Lambda C_{k+1}\tilde{x}_{k+1|k} \\ & - (\Lambda_{k+1} - \Lambda)C_kx_k + (I - \Lambda)C_k\tilde{x}_{k|k} \\ & + \Lambda_{k+1}v_{k+1} + (I - \Lambda_{k+1})v_k. \end{aligned} \quad (56)$$

Note that $\mathbb{E}\{\Lambda_{k+1} - \Lambda\} = 0$, v_{k+1} , and v_k are uncorrelated with other terms. Then, we have

$$\begin{aligned} Q_{\varepsilon_{k+1}} = & \mathbb{E}\{\varepsilon_{k+1} \cdot \varepsilon_{k+1}^T\} \\ = & \Lambda C_{k+1}P_{k+1|k}(\Lambda C_{k+1})^T \\ & + (I - \Lambda)C_kP_{k|k}[(I - \Lambda)C_k]^T \\ & + \Lambda C_{k+1}\Phi_{k(k+1,k)}[(I - \Lambda)C_k]^T \\ & + (\Lambda C_{k+1}\Phi_{k(k+1,k)}[(I - \Lambda)C_k]^T)^T \\ & + \Lambda R_{k+1}\Lambda^T + (I - \Lambda)R_k(I - \Lambda)^T \\ & + \mathcal{H}_1 + \mathcal{H}_2 - \mathcal{H}_3 - \mathcal{H}_3^T, \end{aligned} \quad (57)$$

where

$$\begin{aligned} \mathcal{H}_1 = & \mathbb{E}\{(\Lambda_{k+1} - \Lambda)C_{k+1}x_{k+1}x_{k+1}^TC_{k+1}^T(\Lambda_{k+1} - \Lambda)^T\}, \\ \mathcal{H}_2 = & \mathbb{E}\{(\Lambda_{k+1} - \Lambda)C_kx_kx_k^TC_k^T(\Lambda_{k+1} - \Lambda)^T\}, \\ \mathcal{H}_3 = & \mathbb{E}\{(\Lambda_{k+1} - \Lambda)C_{k+1}x_{k+1}x_k^TC_k^T(\Lambda_{k+1} - \Lambda)^T\}. \end{aligned} \quad (58)$$

By applying Lemma 9, the following \mathcal{H}_1 can be obtained:

$$\begin{aligned} \mathcal{H}_1 = & \mathbb{E}\{(\Lambda_{k+1} - \Lambda)C_{k+1}x_{k+1}x_{k+1}^TC_{k+1}^T(\Lambda_{k+1} - \Lambda)^T\} \\ = & \begin{bmatrix} \mathbb{E}\{(\lambda_{k+1}^1 - \alpha_1)^2\} & \cdots & \mathbb{E}\{(\lambda_{k+1}^1 - \alpha_1)(\lambda_{k+1}^N - \alpha_N)\} \\ \vdots & \ddots & \vdots \\ \mathbb{E}\{(\lambda_{k+1}^N - \alpha_N)(\lambda_{k+1}^1 - \alpha_1)\} & \cdots & \mathbb{E}\{(\lambda_{k+1}^N - \alpha_N)^2\} \end{bmatrix} \circ C_{k+1}\Theta_{(k+1,k+1)}C_{k+1}^T \\ = & \text{diag}\{\alpha_1(1 - \alpha_1), \alpha_2(1 - \alpha_2), \dots, \alpha_N(1 - \alpha_N)\} \circ C_{k+1}\Theta_{(k+1,k+1)}C_{k+1}^T \\ = & (\text{diag}\{\alpha_1, \alpha_2, \dots, \alpha_N\} \text{diag}\{1 - \alpha_1, 1 - \alpha_2, \dots, 1 - \alpha_N\}) \circ C_{k+1}\Theta_{(k+1,k+1)}C_{k+1}^T \\ = & \Lambda(I - \Lambda) \circ C_{k+1}\Theta_{(k+1,k+1)}C_{k+1}^T. \end{aligned} \quad (59)$$

By the same derivation of \mathcal{H}_1 , we have

$$\begin{aligned}\mathcal{H}_2 &= \Lambda (I - \Lambda) \circ C_k \Theta_{(k,k)} C_k^T, \\ \mathcal{H}_3 &= \Lambda (I - \Lambda) \circ C_{k+1} \Pi_{(k,0)}^0 C_k^T.\end{aligned}\quad (60)$$

Then, by the definitions of $\Pi_{(k,j)}^\mu$ and $\Omega_{(k,j)}^\mu$, \mathcal{H}_1 , \mathcal{H}_2 , \mathcal{H}_3 , and (47) can be obtained.

By the orthogonality of the projection and $\mathbb{E}\{\Lambda_{k+1} - \Lambda\} = 0$, $\hat{x}_{k+1|k} \perp \tilde{x}_{k+1|k}$, $\hat{x}_{k+1|k} \perp \tilde{x}_{k|k}$, we have

$$\begin{aligned}\mathbb{E}\{x_{k+1} \varepsilon_{k+1}^T\} &= \mathbb{E}\{x_{k+1} (\tilde{x}_{k+1|k} C_{k+1}^T \Lambda + \tilde{x}_{k|k} C_k^T (I - \Lambda))\} \\ &= \mathbb{E}\{\tilde{x}_{k+1|k} (\tilde{x}_{k+1|k} C_{k+1}^T \Lambda + \tilde{x}_{k|k} C_k^T (I - \Lambda))\} \\ &= \Phi_{k(k+1,k+1)} C_{k+1}^T \Lambda + \Phi_{k(k+1,k)} C_k^T (I - \Lambda).\end{aligned}\quad (61)$$

Along the same method of the derivation of (29) as well as the definition of Ξ_k^i , we obtain

$$\Phi_{k(k+1,k)} = A_k \Phi_{k(k-d,k)} = A_k \Xi_k^1. \quad (62)$$

Substituting (57), (61), and (62) into (55) yields (46).

On the other hand, according to (54), we have

$$\tilde{x}_{k+1|k+1} = \tilde{x}_{k+1|k} - K_{k+1} \varepsilon_{k+1}. \quad (63)$$

And then

$$\begin{aligned}P_{k+1|k+1} &= P_{k+1|k} + K_{k+1} Q_{\varepsilon_{k+1}} K_{k+1}^T - \mathbb{E}\{\tilde{x}_{k+1|k} \varepsilon_{k+1}^T\} K_{k+1}^T \\ &\quad - K_{k+1} \mathbb{E}\{\varepsilon_{k+1} \tilde{x}_{k+1|k}^T\}.\end{aligned}\quad (64)$$

By the same line of (21), (48) can be obtained.

Thirdly, the following equations can be obtained by the same line of the derivation of (27) and (28):

$$\hat{x}_{k+1|k} = A_k \hat{x}_{k-d|k}, \quad (65)$$

$$\tilde{x}_{k+1|k} = A_k \tilde{x}_{k-d|k} + B_k \omega_k. \quad (66)$$

By the fact that $\omega_k \perp \tilde{x}_{k-d|k}$, it can be deduced that

$$P_{k+1|k} = A_k P_{k-d|k} A_k^T + B_k Q_k B_k^T. \quad (67)$$

Therefore, (49) and (50) can be obtained by (65) and (67).

Subsequently, the following derivations are given to obtain $\hat{x}_{k-d|k}$ and $P_{k-d|k}$. By using the projection theory, we have

$$\hat{x}_{k-d|k} = \hat{x}_{k-d|k-1} + K_k^d \varepsilon_k. \quad (68)$$

Then

$$P_{k-d|k} = P_{k-d|k-1} - K_k^d \Gamma_k^d. \quad (69)$$

Hence, it follows from (68) and (69) that we have (51) and (52).

Finally, note that

$$\Gamma_k^i = \mathbb{E}\{\varepsilon_k x_{k-i}^T\} = \Lambda C_k \Phi_{k-1(k,k-i)} + (I - \Lambda) C_{k-1} \Phi_{k-1(k-1,k-i)}. \quad (70)$$

When $i = 1$, we have

$$\Gamma_k^1 = \Lambda C_k \Phi_{k-1(k,k-1)} + (I - \Lambda) C_{k-1} P_{k-1|k-1}. \quad (71)$$

When $i \neq 1$, by (40) and (33), the following recursive equations can be obtained:

$$\Phi_{k-1(k,k-i)} = \Omega_{(k-1,0)}^{i-1}, \quad \Phi_{k-1(k-1,k-i)} = \Omega_{(k-1,1)}^{i-2}. \quad (72)$$

From (70), (71), and (72), we have (53). The proof of this theorem is now complete. \square

Remark 11. In Theorem 10, the linear recursive optimal filter is designed for the addressed discrete state delay stochastic systems with random multiple sensor delays. A unified framework is established to address complexities from the state delay and the random multiple sensor delays. The proposed filtering algorithm is of a recursive form suitable for online applications. On the other hand, it is worth mentioning that the proposed linear optimal filter can be reduced to the traditional Kalman filter when $\Lambda = I$.

Remark 12. Note that there has not been much work concerning the design of the optimal filter for systems with state delay and randomly multiple sensor delays. It is well known that, due to the uncertain influence of the practical environment, it is the case and more reasonable to deal with the problem of sensor measurements with different delay rates. Hence, we have made the first attempt to tackle the optimal filtering problem for discrete stochastic systems subject to the state time delay and randomly multiple sensor delays with different delay rates. Compared with the existing results, the developed filtering algorithm can better deal with the engineering practice in a more effective way especially for the case of the different delay rates.

According to Theorem 10, a new recursive algorithm can be established to obtain the linear optimal filter for the addressed discrete state time delay stochastic systems. The following algorithm shows how to design the linear optimal recursive filter in Theorem 10.

Algorithm 13. The steps of the design of the linear optimal recursive filter are shown as follows:

Step 1. Give the following initial values $\Theta_{(-d,-d)}, \dots, \Theta_{(0,0)}$; x_{-d}, \dots, x_0 ; $\hat{x}_{-d|d}, \dots, \hat{x}_{0|0}$; $P_{-d|d}, \dots, P_{0|0}$; $\Gamma_0^1, \dots, \Gamma_0^d$; $\Psi_0^1, \dots, \Psi_0^d$ and Ξ_0^1, \dots, Ξ_0^d .

Step 2. In the time period of $[k-d, d]$, by the value of the previous time, $\Phi_{k-1(k-j,k-d)}$ can be computed.

Step 3. Substituting (40) into (53), we can obtain Γ_k^i . Also, K_k^i can be computed by (53).

Step 4. Calculate Ψ_k^i by substituting (53) into (15). Calculate $\hat{x}_{k-d|k}$ and $P_{k-d|k}$ by (51) and (52). Calculate $\hat{x}_{k+1|k}$ and $P_{k+1|k}$ by substituting (51) and (52) into (49) and (50).

Step 5. Compute $\Theta_{(k,k)}$ by (12). Σ_k^t is calculated by substituting (12) into (13). Then, (32) is computed by (12) and (13).

Step 6. Substituting (15) and (53) into (16), we have Ξ_k^i . Then, (33) is computed by (15) and (16).

Step 7. Calculate $\mathcal{H}_1, \mathcal{H}_2$ and \mathcal{H}_3 . Substituting $\mathcal{H}_1, \mathcal{H}_2, \mathcal{H}_3$, (50), and (33) into (47), we obtain $Q_{\varepsilon_{k+1}}$.

Step 8. Substituting (16) and (47) into (46) and (48), we have K_{k+1} and $P_{k+1|k+1}$, respectively.

Step 9. Substituting (51), (49), (46) into (45) and (44), we obtain the optimal estimation $\hat{x}_{k+1|k+1}$.

Remark 14. According to the above algorithm, the filter gain K_{k+1} can be computed recursively. It is worth pointing out that, when deriving the filter gain, additional efforts should be made to derive the terms $P_{k+1|k+1}$ and $Q_{\varepsilon_{k+1}}$ due to the simultaneous consideration of the randomly multiple sensor delays and the state delay. After having obtained these terms, the filter gain K_{k+1} can be constructed and then the estimation $\hat{x}_{k+1|k+1}$ can be computed. In the following, an illustrative example will be provided to show the feasibility of the proposed filtering scheme.

4. An Illustrative Example

In this section, a numerical example is proposed to show the feasibility and effectiveness of the developed main results.

Consider the following linear discrete-time delay stochastic systems:

$$\begin{aligned} x_{k+1} &= A_k x_{k-2} + B_k \omega_k, \\ z_k^i &= C_k^i x_k + v_k^i, \\ y_k^i &= \lambda_k^i z_k^i + (1 - \lambda_k^i) z_{k-1}^i, \end{aligned} \quad (73)$$

where

$$\begin{aligned} A_k &= \begin{bmatrix} 0.35 & -0.15 & 0.15 \\ 0.35 & 0 & 0.1 \\ 0 & 0.2 & 0.1 \end{bmatrix}, \quad B_k = \begin{bmatrix} 1.5 \\ 2 \\ 2.5 \end{bmatrix}, \\ C_k^1 &= \begin{bmatrix} 0.9 \\ 0 \\ 0 \end{bmatrix}^T, \quad C_k^2 = \begin{bmatrix} 0 \\ 0.85 \\ 0 \end{bmatrix}^T, \quad C_k^3 = \begin{bmatrix} 0 \\ 0 \\ 0.8 \end{bmatrix}^T \end{aligned} \quad (74)$$

and $x_k = [x_{k,1} \ x_{k,2} \ x_{k,3}]^T$, ω_k , and v_k^i , $i = 1, 2, 3$, are uncorrelated zero-mean Gaussian white noises with covariances 0.2 and 0.1, respectively.

Let

$$\Theta_{(-2,-2)} = \Theta_{(-1,-1)} = \Theta_{(0,0)} = \text{diag}\{1, 1, 1\},$$

$$P_{-2|-2} = P_{-1|-1} = P_{0|0} = \text{diag}\{5, 5, 5\},$$

$$x_{-2} = [-0.3 \ -0.1 \ 0.1]^T, \quad x_{-1} = [-0.2 \ 0 \ 0.2]^T,$$

$$x_0 = [-0.1 \ 0.1 \ 0.3]^T, \quad \hat{x}_{-2|-2} = [-0.4 \ -0.2 \ 0]^T,$$

$$\hat{x}_{-1|-1} = [-0.3 \ -0.1 \ 0.1]^T,$$

$$\hat{x}_{0|0} = [-0.2 \ 0 \ 0.2]^T,$$

$$\Gamma_0^1 = \Gamma_0^2 = \text{diag}\{1, 1, 1\}, \quad \Psi_0^1 = \Psi_0^2 = \text{diag}\{1, 1, 1\},$$

$$\Xi_0^1 = \Xi_0^2 = \text{diag}\{1, 1, 1\}, \quad \mathbb{E}\{\lambda_k^1\} = \alpha_1 = 0.95,$$

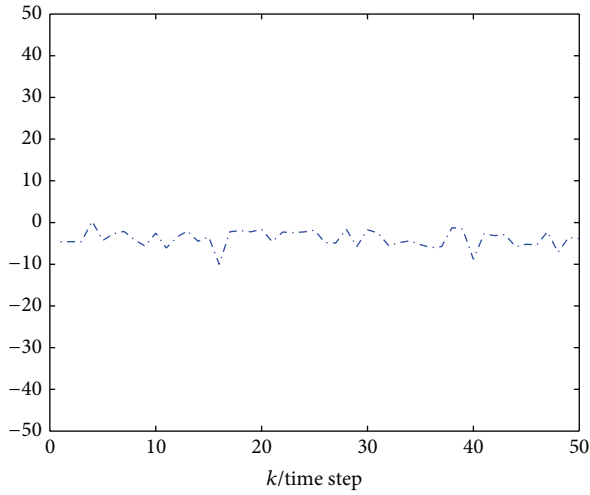
$$\mathbb{E}\{\lambda_k^2\} = \alpha_2 = 0.9, \quad \mathbb{E}\{\lambda_k^3\} = \alpha_3 = 0.95 \quad (75)$$

and let MSE_i denote the mean-square error for the estimation of $x_{k,i}$, that is, $(1/M) \sum_{j=1}^M (x_{k,i}^{(j)} - \hat{x}_{k|k,i}^{(j)})^2$, where $i = 1, 2, 3$ and M is the number of simulation test.

According to Theorem 10, the linear optimal recursive filter can be constructed by applying the innovative analysis approach and MMSE estimation principle. The values of the filter gains are given as in Table 1. The simulations are shown in Figures 1–6. Among them, Figures 1, 2, and 3 plot the $\log(\text{MSE}_i)$ ($i = 1, 2, 3$) of the proposed filtering algorithm. The actual system states and the newly designed estimation are plotted in Figures 4, 5, and 6. From the simulations, we can see that the proposed filter can estimate the system state well irrespective of the state delay and the occurrence of the randomly multiple sensor delays with different delay rates. The reason is that, when deriving the recursive optimal filter, we have made additional efforts to compensate the effects from the randomly multiple sensor delays and state delay.

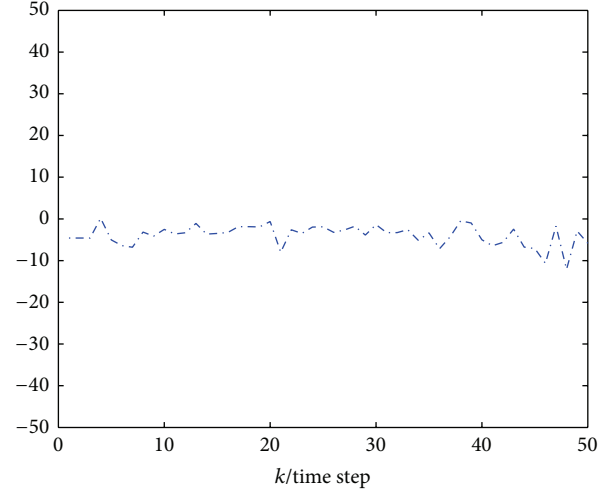
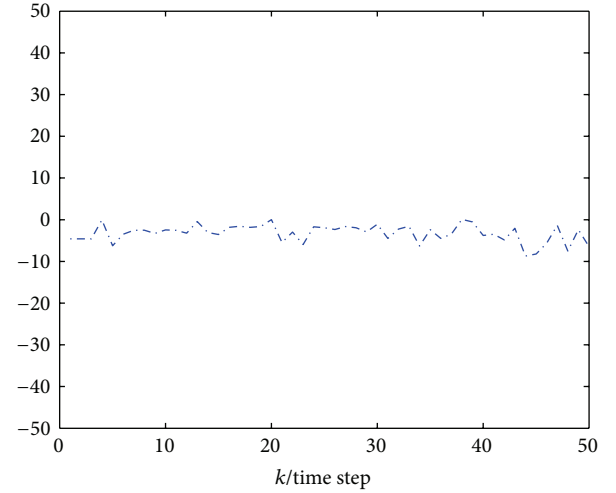
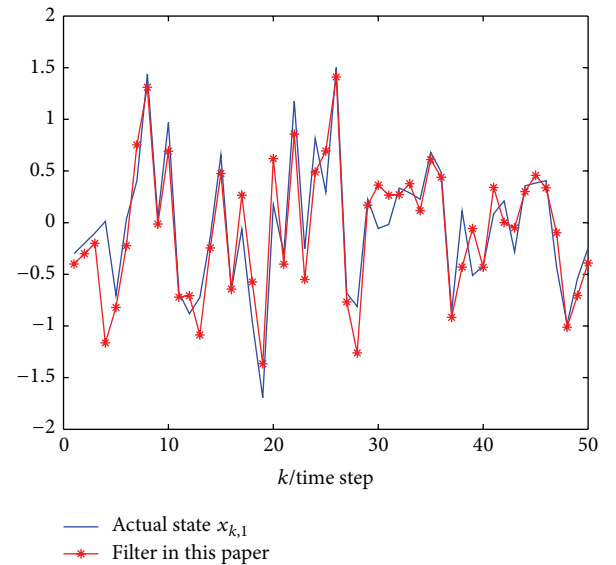
TABLE 1: Filter gains K_{k+1} .

k	K_{k+1}
$k = 0$	$\begin{bmatrix} 3.7383 & 0.3829 & -1.2625 \\ 3.3140 & 0.4513 & -0.9463 \\ 2.8896 & 0.5198 & -0.6302 \end{bmatrix}$
	$\begin{bmatrix} 1.2523 & 0.2830 & -0.1191 \\ 0.9912 & 0.4258 & 0.1461 \\ 0.7301 & 0.5686 & 0.4114 \end{bmatrix}$
	$\begin{bmatrix} 0.7535 & 0.2712 & 0.0502 \\ 0.4415 & 0.4235 & 0.3401 \\ 0.1296 & 0.5758 & 0.6300 \end{bmatrix}$
$k = 1$	$\begin{bmatrix} -0.4412 & 0.4183 & 0.6398 \\ 0.6148 & 0.3916 & 0.2305 \\ 1.6707 & 0.3649 & -0.1789 \end{bmatrix}$
	$\begin{bmatrix} 0.4101 & 0.3031 & 0.1925 \\ 0.4747 & 0.4138 & 0.2985 \\ 0.5392 & 0.5245 & 0.4045 \end{bmatrix}$
	$\begin{bmatrix} 0.4101 & 0.3031 & 0.1925 \\ 0.4747 & 0.4138 & 0.2985 \\ 0.5392 & 0.5245 & 0.4045 \end{bmatrix}$
\vdots	\vdots

FIGURE 1: $\log(\text{MSE1})$.

5. Conclusion

The problem of linear optimal estimation has been investigated for discrete state delay stochastic systems measured by multiple sensors with different delay rates. Based on the innovative analysis approach and MMSE estimation principle, a new linear optimal filter has been constructed. Compared with the state augmentation approach, the computational burden of the proposed method has been decreased due to the fact that the dimension of the filter is equal to the state vector. Future research topics include the extension of the proposed main results to the filter design for the data-driven systems [36, 37] and the networked control systems [38–42]. Also, it would be interesting to develop the smoother

FIGURE 2: $\log(\text{MSE2})$.FIGURE 3: $\log(\text{MSE3})$.FIGURE 4: The trajectories of $x_{k,1}$ and $\hat{x}_{k|k,1}$.

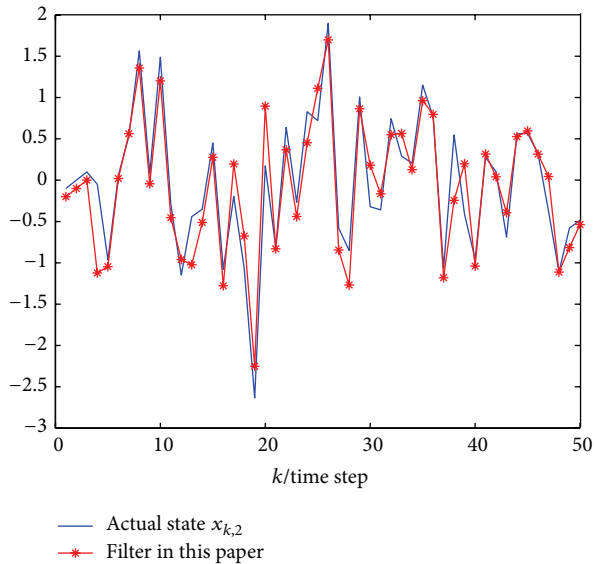


FIGURE 5: The trajectories of $x_{k,2}$ and $\hat{x}_{k|k,2}$.

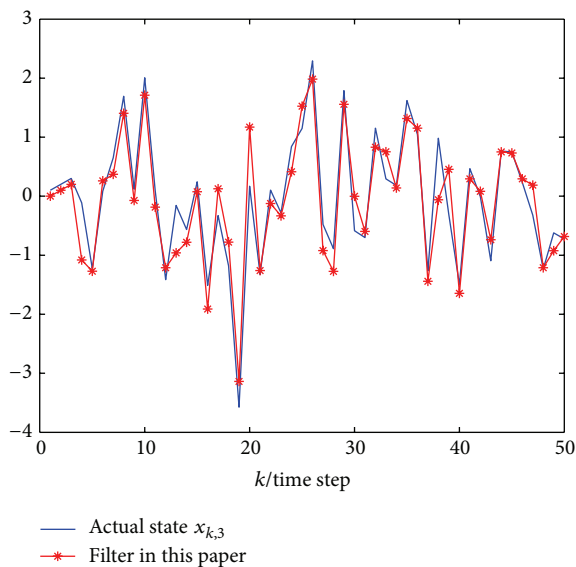


FIGURE 6: The trajectories of $x_{k,3}$ and $\hat{x}_{k|k,3}$.

to the addressed networked systems with different delays and discuss the steady-state filter.

Conflict of Interests

The authors declare that there is no conflict of interests regarding the publication of this paper.

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