

A CENTRAL LIMIT THEOREM FOR GENERAL WEIGHTED SUMS OF LNQD RANDOM VARIABLES AND ITS APPLICATION

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ABSTRACT. In this paper we derive the central limit theorem for $\sum_{i=1}^n a_{ni} \xi_i$, where $\{a_{ni}, 1 \leq i \leq n\}$ is a triangular array of nonnegative numbers such that $\sup_n \sum_{i=1}^n a_{ni}^2 < \infty$, $\max_{1 \leq i \leq n} a_{ni} \rightarrow 0$ as $n \rightarrow \infty$ and ξ_i 's are a linearly negative quadrant dependent sequence. We also apply this result to consider a central limit theorem for a partial sum of a generalized linear process of the form $X_n = \sum_{j=-\infty}^{\infty} a_{k+j} \xi_j$.

1. Introduction and results. Lehmann [8] introduced a simple and natural definition of positive (negative) dependence: A sequence $\{\xi_i, 1 \leq i \leq n\}$ of random variables is said to be pairwise positive (negative) quadrant dependent (pairwise PQD (NQD)) if, for any real α_i, α_j and $i \neq j$ $P(\xi_i > \alpha_i, \xi_j > \alpha_j) \geq (\leq) P(\xi_i > \alpha_i)P(\xi_j > \alpha_j)$. Much stronger dependent concepts than PQD and NQD were considered by Esary, Proschan and Walkup [4] and Joag-Dev and Proschan [6], respectively. A sequence $\{\xi_i, 1 \leq i \leq n\}$ of random variables is said to be associated if, for any real coordinatewise increasing functions f, g on \mathbf{R}^n , $\text{Cov}(f(\xi_1, \dots, \xi_n), g(\xi_1, \dots, \xi_n)) \geq 0$ and $\{\xi_i, 1 \leq i \leq n\}$ is said to be negatively associated if, for any disjoint subsets, $A, B \subset \{1, 2, \dots, n\}$ and any real coordinatewise increasing functions f on \mathbf{R}^A and g on \mathbf{R}^B , $\text{Cov}(f(\xi_i, i \in A), g(\xi_i \in B)) \leq 0$.

Instead of association (negative association) Newman's [10] central limit theorem requires only that positive linear combinations of the random variables are PQD (NQD). The definition of positive (negative) dependence introduced by Newman [10] is the following: A sequence $\{\xi_i, 1 \leq i \leq n\}$ of random variables is said to be linearly positive

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(negative) quadrant dependent (LPQD (LNQD)) if, for every pair of disjoint subsets $A, B \subset \{1, 2, \dots, n\}$ and positive r_j 's

$$(1.1) \quad \sum_{i \in A} r_i \xi_i \quad \text{and} \quad \sum_{j \in B} r_j \xi_j \quad \text{are PQD(NQD)}.$$

Let us remark that LPQD (LNQD) is between pairwise PQD (NQD) and association (negative association) and it is well known, see, for example, [10, p. 131] that association (negative association) implies LPQD (LNQD) and LPQD (LNQD) implies PQD (NQD).

Newman [10] established the central limit theorem for a strictly stationary LPQD (LNQD) process and Birkel [2] also obtained a functional central limit theorem for LPQD processes which can be used to obtain the functional central limit theorem for LNQD processes. Kim and Baek [7] extended this result to a stationary linear process of the form $X_k = \sum_{j=0}^{\infty} a_j \xi_{k-j}$, where $\{a_j\}$ is a sequence of real numbers with $\sum_{j=0}^{\infty} |a_j| < \infty$ and $\{\xi_k\}$ is a strict stationary LPQD process with $E \xi_i = 0$, $0 < E \xi_i^2 < \infty$; this result can be extended to the LNQD case by a similar method.

In this paper we derive a central limit theorem for a linearly negative quadrant dependent sequence in a double array, replacing the strict stationarity assumption with uniform integrability, see Theorem 1.1 below. We apply this result to obtain a central limit theorem for a partial sum of a linear process of the form $X_n = \sum_{j=-\infty}^{\infty} a_{k+j} \xi_j$ generated by linearly negative quadrant dependent sequence $\{\xi_j\}$, see Theorem 1.2 below.

Theorem 1.1. *Let $\{a_{ni}, 1 \leq i \leq n\}$ be a triangular array of nonnegative numbers such that*

$$(1.2) \quad \sup_n \sum_{i=1}^n a_{ni}^2 < \infty$$

and

$$(1.3) \quad \max_{1 \leq i \leq n} a_{ni} \longrightarrow 0 \quad \text{as} \quad n \rightarrow \infty$$

Let $\{\xi_i\}$ be a centered sequence of linearly negative quadrant dependent random variables such that

$$(1.4) \quad \{\xi_i^2\} \text{ is a uniformly integrable family,}$$

$$(1.5) \quad \text{Var} \left(\sum_{i=1}^n a_{ni} \xi_i \right) = 1$$

and

$$(1.6) \quad \sum_{j: |i-j| \geq u} \text{Cov}(\xi_i, \xi_j)^- \longrightarrow 0 \quad \text{as } u \rightarrow \infty \quad \text{uniformly in } i \geq 1,$$

see [3]. Then

$$\sum_{i=1}^n a_{ni} \xi_i \xrightarrow{\mathcal{D}} N(0, 1) \quad \text{as } n \rightarrow \infty.$$

Remark. Theorem 1.1 extends Newman's [10] central limit theorem for strictly stationary LNQD sequences from equal weights to general weights, while at the same time weakening the assumption of stationarity.

Corollary 1.1. *Let $\{\xi_i\}$ be a centered sequence of linearly negative quadrant dependent random variables such that $\{\xi_i^2\}$ is a uniformly integrable family, and let $\{a_{ni}, 1 \leq i \leq n\}$ be a triangular array of nonnegative numbers such that*

$$(1.7) \quad \sup_n \sum_{i=1}^n \frac{a_{ni}^2}{\sigma_n^2} < \infty,$$

$$(1.8) \quad \max_{1 \leq i \leq n} \frac{a_{ni}}{\sigma_n} \longrightarrow 0 \quad \text{as } n \rightarrow \infty,$$

where $\sigma_n^2 = \text{Var}(\sum_{i=1}^n a_{ni} \xi_i)$. If (1.6) holds, then as $n \rightarrow \infty$,

$$(1.9) \quad \frac{1}{\sigma_n} \sum_{i=1}^n a_{ni} \xi_i \xrightarrow{\mathcal{D}} N(0, 1).$$

Theorem 1.2. *Let $\{a_j, j \in \mathbb{Z}\}$ be a sequence of nonnegative numbers such that $\sum_j a_j < \infty$, and let $\{\xi_j, j \in \mathbb{Z}\}$ be a centered sequence of linearly negative quadrant dependent random variables which is uniformly integrable in L_2 and satisfying (1.6). Let*

$$X_k = \sum_{j=-\infty}^{\infty} a_{k+j} \xi_j \quad \text{and} \quad S_n = \sum_{i=1}^n X_i.$$

Assume

$$(1.10) \quad \inf_{n \geq 1} n^{-1} \sigma_n^2 > 0$$

where $\sigma_n^2 = \text{Var}(S_n)$. Then

$$(1.11) \quad \frac{S_n}{\sigma_n} \xrightarrow{\mathcal{D}} N(0, 1) \quad \text{as } n \rightarrow \infty.$$

This result extends Theorem 18.6.5 in [5] from the i.i.d. case to the linearly negative quadrant dependence case by adding condition (1.6) and improves the central limit theorem of Kim and Baek [7] for linear processes generated by LNQD sequences.

2. Proofs. We start with the following lemma.

Lemma 2.1 [9]. *Let $\{Z_i, 1 \leq i \leq n\}$ be a sequence of linearly negative quadrant dependent random variables with finite second moments. Then*

$$\begin{aligned} \left| E \exp \left(it \sum_{j=1}^n Z_j \right) - \prod_{j=1}^n E \exp(itZ_j) \right| \\ \leq Ct^2 \left| \text{Var} \left(\sum_{j=1}^n Z_j \right) - \sum_{j=1}^n \text{Var}(Z_j) \right| \end{aligned}$$

for all $t \in \mathbf{R}$, where $C > 0$ is an arbitrary constant, not depending on n .

Proof of Theorem 1.1. Without loss of generality, we assume that $a_{ni} = 0$ for all $i > n$ and $\sup_{n \geq 1} E \xi_n^2 = M < \infty$. For every $1 \leq a < b \leq n$ and $1 \leq u \leq b - a$, we have, after some manipulations,

$$(2.1) \quad \begin{aligned} 0 &\leq \sum_{i=a}^{b-u} a_{ni} \sum_{j=i+u}^b a_{nj} \text{Cov}(\xi_i, \xi_j)^- \\ &\leq \sup_k \left(\sum_{j: |k-j| \geq u} \text{Cov}(\xi_k, \xi_j)^- \right) \left(\sum_{i=a}^b a_{ni}^2 \right). \end{aligned}$$

By the definition of LNQD, we also have, for every $1 \leq a \leq b \leq n$,

$$\text{Var} \left(\sum_{i=a}^b a_{ni} \xi_i \right) \leq M \sum_{i=a}^b a_{ni}^2.$$

We shall construct now a triangular array of random variables $\{Z_{ni}, 1 \leq i \leq n\}$ for which we shall make use of Lemma 2.1. Fix a small positive ε and find a positive integer $u = u_\varepsilon$ such that, for every $n \geq u + 1$,

$$(2.2) \quad \begin{aligned} 0 &\leq \left(\sum_{i=1}^{n-u} a_{ni} \sum_{j=i+u}^n a_{nj} \text{Cov}(\xi_i, \xi_j)^- \right) \\ &\leq \varepsilon. \end{aligned}$$

This is possible because of (2.1) and (1.6). Denote by $[x]$ the integer part of x , and define

$$\begin{aligned} K &= \left[\frac{1}{\varepsilon} \right]; \\ Y_{nj} &= \sum_{i=u_j+1}^{u(j+1)} a_{ni} \xi_i, \quad j = 0, 1, \dots, \end{aligned}$$

$$A_j = \left\{ i : 2Kj \leq i < 2Kj + K, \text{Cov}(Y_{ni}, Y_{n,i+1})^- \leq \frac{2}{K} \sum_{i=2Kj}^{2Kj+K} \text{Var}(Y_{ni}) \right\}.$$

Since $2 \text{Cov}(Y_{ni}, Y_{n,i+1})^- \leq \text{Var}(Y_{ni}) + \text{Var}(Y_{n,i+1})$, we get that for every j the set A_j is not empty. Now we define the integers m_1, m_2, \dots, m_n , recursively. Let $m_0 = 0$ and

$$m_{j+1} = \min\{m : m > m_j, m \in A_j\}$$

and define

$$Z_{nj} = \sum_{i=m_j+1}^{m_{j+1}} Y_{ni}, \quad j = 0, 1, \dots,$$

$$\mathcal{D}_j = \{u(m_j + 1) + 1, \dots, u(m_{j+1} + 1)\}.$$

We observe that

$$Z_{nj} = \sum_{k \in \mathcal{D}_j} a_{nk} \xi_k, \quad j = 0, 1, \dots$$

By the definition of LNQD the random variables $\{Z_{nj}\}$ are LNQD. From the fact that $m_j \geq 2K(j-1)$ and $m_{j+1} \leq K(2j+1)$ every set \mathcal{D}_j contains no more than $3Ku$ elements and $m_{j+1}/m_j \rightarrow 1$ as $j \rightarrow \infty$. Hence, for every fixed positive ε by (1.2)–(1.5) the array $\{Z_{ni} : i = 0, 1, \dots, n; n \geq 1\}$ satisfies the Lindeberg condition, see Petrov [11, Theorem 22, p. 100], that is, $\{Z_{nj}\}$ satisfies

$$(2.3) \quad \sigma_n^{-1} \sum_{j=1}^n E Z_{nj}^2 I(|Z_{nj}| > \varepsilon \sigma_n) \longrightarrow 0 \quad \text{as } n \rightarrow \infty$$

where $\sigma_n^2 = \text{Var}(\sum_{j=1}^n Z_{nj})$.

We can observe that, by Lemma 2.1 and the construction,

(2.4)

$$\begin{aligned} & \left| E \exp \left(it \sum_{j=1}^n Z_{nj} \right) - \prod_{j=1}^n E \exp(it Z_{nj}) \right| \\ & \leq Ct^2 \left| \left\{ \text{Var} \left(\sum_{j=1}^n Z_{nj} \right) - \sum_{j=1}^n \text{Var}(Z_{nj}) \right\} \right| \\ & \leq Ct^2 \left\{ 2 \left(\sum_{i=1}^n \text{Cov}(Z_{ni}, Z_{n,i+1})^- \right) + 2 \left(\sum_{i=1}^{n-2} \sum_{j=i+2}^n \text{Cov}(Z_{ni}, Z_{nj})^- \right) \right\} \end{aligned}$$

$$\begin{aligned}
&\leq Ct^2 \left\{ 2 \sum_{j=1}^n \text{Cov}(Y_{n,m_j}, Y_{n,m_{j+1}})^- + 2 \sum_{i=1}^{n-u} a_{ni} \sum_{j=i+u}^n a_{nj} \text{Cov}(\xi_i, \xi_j)^- \right\} \\
&\quad + \left\{ 2 \sum_{i=1}^{n-u} a_{ni} \sum_{j=i+u}^n a_{nj} \text{Cov}(\xi_i, \xi_j)^- \right\} \\
&= Ct^2 \left\{ 4 \sum_{i=1}^{n-u} a_{ni} \sum_{j=i+u}^n a_{nj} \text{Cov}(\xi_i, \xi_j)^- + 2 \sum_{j=1}^n \text{Cov}(Y_{n,m_j}, Y_{n,m_{j+1}})^- \right\} \\
&\leq Ct^2 \left\{ 4\varepsilon + \frac{8}{K} \sum_{i=1}^n \text{Var}(Y_{ni}) \right\} \\
&= Ct^2 \left\{ 4\varepsilon + \frac{8}{K} \sum_{j=1}^n \text{Var} \left(\sum_{i=u_j+1}^{u(j+1)} a_{ni} \xi_i \right) \right\} \\
&\leq Ct^2 \left\{ 4\varepsilon + \frac{8M}{K} \sum_{j=1}^n \sum_{i=u_j+1}^{u(j+1)} a_{ni} \xi_i \right\} \\
&\leq C_1 t^2 \varepsilon \left\{ 1 + \sup_n \sum_{i=1}^n a_{ni}^2 \right\} \\
&\leq C_2 t^2 \varepsilon \quad \text{for every positive } \varepsilon.
\end{aligned}$$

Therefore the problem is now reduced to the study of the central limit theorem of a decoupled sequence $\{\tilde{Z}_{nj}\}$ of independent random variables such that, for each n and j , the variable \tilde{Z}_{nj} is distributed as Z_{nj} .

By (2.3) $\{\tilde{Z}_{nj}\}$ also satisfies the Lindeberg condition, that is, $\{\tilde{Z}_{nj}\}$ satisfies $\tilde{\sigma}_n^{-1} \sum_{j=1}^n E \tilde{Z}_{nj}^2 I(|\tilde{Z}_{nj}| > \varepsilon \tilde{\sigma}_n) \rightarrow 0$ as $n \rightarrow \infty$ where $\tilde{\sigma}_n^2 = \text{Var}(\sum_{j=1}^n \tilde{Z}_{nj})$, and hence by [1, Theorem 7.2]

$$(2.5) \quad \tilde{\sigma}_n^{-1} \sum_{j=1}^n \tilde{Z}_{nj} \xrightarrow{\mathcal{D}} N(0, 1) \quad \text{as } n \rightarrow \infty$$

where $\tilde{\sigma}_n^2 = \text{Var}(\sum_{j=1}^n \tilde{Z}_{nj})$. It follows from (2.3), (2.4) and (2.5) that

$$(2.6) \quad \sigma_n^{-1} \sum_{j=1}^n Z_{nj} \xrightarrow{\mathcal{D}} N(0, 1) \quad \text{as } n \rightarrow \infty$$

where $\sigma_n^2 = \text{Var}(\sum_{j=1}^n Z_{nj})$, and now the proof is complete by (2.5), (2.6) and [1, Theorem 4.2].

Proof of Corollary 1.1. Let $A_{ni} = a_{ni}/\sigma_n$. Then we have

$$\begin{aligned} \max_{1 \leq i \leq n} A_{ni} &\longrightarrow 0 \quad \text{as } n \rightarrow \infty, \\ \sup_n \sum_{i=1}^n A_{ni}^2 &< \infty, \\ \text{Var} \left(\sum_{i=1}^n A_{ni} \xi_i \right) &= 1. \end{aligned}$$

Hence, by Theorem 1.1 the desired result (1.11) follows.

Proof of Theorem 1.2. First note that $\sum_j a_j^2 < \infty$ and, without loss of generality, we can assume $\sup E \xi_k^2 = 1$. Let

$$S_n = \sum_{k=1}^n X_k = \sum_{j=-\infty}^{\infty} \left(\sum_{k=1}^n a_{k+j} \right) \xi_j.$$

In order to apply Theorem 1.1, fix W_n such that $\sum_{|j| > W_n} a_j^2 < n^{-3}$, and take $k_n = W_n + n$. Then

$$\frac{S_n}{\sigma_n} = \sum_{|j| \leq k_n} \left(\sum_{k=1}^n a_{k+j} \right) \frac{\xi_j}{\sigma_n} + \sum_{|j| > k_n} \left(\sum_{k=1}^n a_{k+j} \right) \frac{\xi_j}{\sigma_n} = T_n + U_n.$$

By the Cauchy-Schwarz inequality and the assumptions we have the following estimate

$$\begin{aligned} \text{Var}(U_n) &\leq \sum_{|j| > k_n} \text{Var} \left(\sum_{k=1}^n a_{k+j} \frac{\xi_j}{\sigma_n} \right) \\ &\leq \sum_{|j| > k_n} \left(\sum_{k=1}^n a_{k+j} / \sigma_n \right)^2 E \xi_j^2 \leq n \sigma_n^{-2} \sum_{|j| > k_n} \left(\sum_{k=1}^n a_{k+j}^2 \right) \\ &\leq n^2 \sigma_n^{-2} \sum_{|j| > k_n - n} a_j^2 \leq n^2 \sigma_n^{-2} \sum_{|j| > W_n} a_j^2 \\ &\leq n^{-1} \sigma_n^{-2} \longrightarrow 0 \quad \text{as } n \rightarrow \infty \end{aligned}$$

which yields

$$(2.7) \quad U_n \longrightarrow 0 \quad \text{in probability as } n \rightarrow \infty.$$

By [1, Theorem 4.1], it remains to prove that $T_n \xrightarrow{\mathcal{D}} N(0, 1)$. Put

$$(2.8) \quad a_{nk} = \frac{\sum_{j=1}^n a_{k+j}}{\sigma_n}.$$

From the assumption $\sum_j a_j < \infty$ ($a_j > 0$), (1.10) and (2.8) we obtain

$$\begin{aligned} \frac{\sup_{-\infty < k < \infty} \sum_{j=1}^n a_{k+j}}{\sigma_n} &\longrightarrow 0 \quad \text{as } n \rightarrow \infty, \\ \max_{1 \leq k \leq n} a_{nk} &\longrightarrow 0 \quad \text{as } n \rightarrow \infty, \\ \sup_n \sum_{k=1}^n a_{nk}^2 &< \infty. \end{aligned}$$

Hence, by Theorem 1.1,

$$(2.9) \quad T_n \xrightarrow{\mathcal{D}} N(0, 1)$$

and from (2.7) and (2.9) the desired result (1.10) follows.

REFERENCES

1. P. Billingsley, *Convergence of probability measures*, Wiley, New York, 1968.
2. T. Birkel, A functional central limit theorem for positively dependent random variables, *J. Multivariate Anal.* **44** (1993), 314–320.
3. J.T. Cox and G. Grimmett, *Central limit theorems for associated random variables and the percolation model*, *Ann. Probab.* **12** (1984), 514–528.
4. J. Esary, F. Proschan and D. Walkup, *Association of random variables with applications*, *Ann. Math. Statist.* **38** (1967), 1466–1474.
5. I.A. Ibragimov and Yu.V. Linnik, *Independent and stationary sequences of random variables*, Volters, Groningen, 1971.
6. K. Joag-Dev and F. Proschan, Negative association of random variables with applications, *Ann. Statist.* **11** (1983), 1037–1041.
7. T.S. Kim and J.L. Baek, *A central limit theorem for stationary linear processes generated by linearly positive quadrant dependent process*, *Statist. Prob. Letters* **5** (2001), 299–305.

8. E.L. Lehmann, *Some concepts of dependence*, Ann. Math. Statist. **37** (1966), 1137–1153.
9. C.M. Newman, *Normal fluctuations and the FKG inequalities*, Comm. Math. Phys. **91** (1980), 75–80.
10. ———, *Asymptotic independence and limit theorems for positively and negatively dependent random variables*, in *Stochastics and probability* (Y.L. Tong, ed.), vol. 5, Inst. Math. Statist., Hayward, CA, 1984, pp. 127–140.
11. V.V. Petrov, *Sums of independent random variables*, Berlin, 1975.

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