

59. Probability-theoretic Investigations on Inheritance.
X₁. Non-Paternity Concerning Mother-Child-Child
Combinations.

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(Comm. by T. FURUHATA, M.J.A., May 13, 1952.)

1. Non-paternity against a distinguished child.

Problems discussed in the preceding chapter¹⁾ have exclusively concerned two children belonging to the same family, that is, possessing a father also in common. There arise analogous problems concerning two children possessing a mother alone in common, which will be discussed in the present chapter. While the former problems have depended on *mother-children combinations*, the latter ones depend on *mother-child-child combinations*.

Now, consider a triple consisting of a mother A_{ij} , her first child A_{hk} and her second child A_{fg} , both children being assumed not to possess a common father. The probability of an event that such a triple appears and then a man chosen at random can assert his *non-paternity against second child at any rate* is, corresponding to a former expression (2.3) of IX, represented by

$$(1.1) \quad P_0(ij; hk, fg) \equiv \pi_0(ij; hk, fg) V(ij; fg);$$

the π_0 's denoting the probabilities of mother-child-child combination defined in (5.9) of IV and V 's the quantities introduced in (2.1) of VII. This is a basic quantity and can, in view of (5.7) of IV and (2.2) of VII, i.e.,

$$\pi_0(ij; hk, fg) = \pi(ij; hk) \pi(ij; fg) / \bar{A}_{ij}, \quad \pi(ij; fg) V(ij; fg) = P(ij; fg),$$

be expressed also in the form

$$(1.2) \quad P_0(ij; hk, fg) = P(ij; fg) \cdot \pi(ij; hk) / \bar{A}_{ij};$$

1) Y. Komatu, Probability-theoretic investigations on inheritance. I. Distribution of genes; II. Cross-breeding phenomena; III. Further discussions on cross-breeding; IV. Mother-child combinations; V. Brethren combinations; VI. Rate of danger in random blood transfusion; VII. Non-paternity problems; VIII. Further discussions on non-paternity problems; IX. Non-paternity concerning mother-children combinations. Proc. Japan Acad. **27** (1951), I. 371-377; II. 378-383, 384-387; III. 459-464, 466-471, 472-477, 478-483; IV. 587-592, 596-597, 598-603, 605-610, 611-614, 615-620; V. 689-693, 694-699; **28** (1952), VI. 54-58; VII. 102-104, 105-108, 109-111, 112-115, 116-120, 121-125; VIII. 162-164, 165-168, 169-171; IX. 207-212, 213-217, 218-223, 224-229. These papers will be referred to as I; II; III; IV; V; VI; VII; VIII; IX.

the coefficient of $P(ij;fg)$ in the right-hand side being nothing but the quantity noticed in (1.27) of IV, the fact which also implies immediately the relation (1.2) itself.

The coefficient in consideration being independent of the type of second child, the partial sum

$$(1.3) \quad I_0(ij;hk) = \sum_{f \leq g} P_0(ij;hk,fg),$$

which denotes the quantity corresponding to (2.5) of VII in the preceding case, can easily be determined. In fact, making use of the notation introduced in (2.3) of VII, we get an expression

$$(1.4) \quad I_0(ij;hk) = P(ij)\pi(ij;hk)/\bar{A}_{ij}.$$

If we sum up the last relations over all the possible pairs of suffices h and k , we then obtain

$$(1.5) \quad \sum_{h \leq k} I_0(ij;hk) = P(ij);$$

here use being made of an evident relation (1.20) of IV, i.e., $\sum_{h \leq k} \pi(ij;hk) = \bar{A}_{ij}$. If we further sum up the relations (1.5) over i and j , the whole probability (2.20) of VII will be reproduced (cf. also (2.19) of VIII and L in (4.33) below), namely,

$$(1.6) \quad I_0 \equiv \sum_{i \leq j, h \leq k} I_0(ij;hk) = 1 - 2S_2 + S_3 - 2S_2^2 + 2S_4 + 3S_2S_3 - 3S_5 = P.$$

Generalization of the above discussion to the mixed case is also immediate. We shall assume here that a putative father belongs to the same population as a true father. In general, given a quantity X dependent on $\{p_i\}$, let the quantity obtained from X by replacing all the $p_i (i=1, \dots, m)$ by the corresponding p_i'' be denoted by $[X]''$. Then, the quantity

$$(1.7) \quad V''(ij;fg) \equiv [V(ij;fg)]''$$

represents the probability of an event that, given a second child A_{fg} produced by a mother A_{ij} belonging to a population with distribution $\{p_i\}$ and by a father belonging to a population with distribution $\{p_i''\}$, a man belonging to the same population as a true father can assert his non-paternity. Thus we now get, instead of (1.1), a basic quantity

$$(1.8) \quad P^*(ij;hk,fg) = \pi^*(ij;hk,fg) V''(ij;fg),$$

the π^* 's being introduced in (5.6) of IV; here a father of first child is assumed to belong to a population with distribution $\{p_i\}$. Hence, we get, corresponding to (1.2),

$$(1.9) \quad P^*(ij;hk,fg) = P''(ij;fg)\pi'(ij;hk)/\bar{A}_{ij} = \pi'(ij;hk) [P(ij;fg)/\bar{A}_{ij}]''.$$

Thus, the relations corresponding to (1.3), (1.4); (1.5), (1.6) will become

$$(1.10) \quad I^*(ij;hk) \equiv \sum_{f \leq g} P^*(ij;hk,fg) = P''(ij)\pi'(ij;hk)/A_{ij};$$

$$(1.11) \quad \sum_{h \leq k} I^*(ij;hk) = P''(ij) \equiv \sum_{f \leq h} P''(ij;fg) = p_i p_j [P(ij)/p_i p_j]'',$$

$$(1.12) \quad I^* = P'',$$

respectively, P'' denoting the expression obtained from (4.12) of VII by replacing the p 's by the corresponding p'' 's.

2. Non-paternity against both children.

In discussion of the preceding section, it has been quite indifferent whether a man putative against second child is or is not a true father of the first child. We now consider a problem of determining the probability of proving non-paternity against both children of a mother-child-child combination. Given such a combination ($A_{ij}; A_{hk}, A_{fg}$), the probability of an event that a man chosen at random can assert his *non-paternity against both children* is represented by

$$(2.1) \quad Q_0(ij; hk, fg) \equiv \pi_0(ij; hk, fg) V(ij; hk, fg);$$

the π_0 's and the V 's being of the same meaning as in (5.9) of IV and (4.1) of IX, respectively. The partial sum

$$(2.2) \quad J_0(ij; hk) = \sum_{f \leq g} Q_0(ij; hk, fg),$$

corresponding to (1.3), can be represented, in view of (5.7) of IV, also in the form

$$(2.3) \quad J_0(ij; hk) = \frac{\pi(ij; hk)}{A_{ij}} \sum_{f \leq g} \pi(ij; fg) V(ij; hk, fg).$$

The values of (2.2) can, in separate cases, explicitly determined as follows:

$$(2.4) \quad \begin{aligned} J_0(ii; ii) &= Q_0(ii; ii, ii) + \sum_{k \neq i} Q_0(ii; ii, ik) \\ &= p_i^3(1 - 2S_2 + S_3 - 2(1 - S_2)p_i + 3p_i^2 - 3p_i^3), \end{aligned}$$

$$(2.5) \quad \begin{aligned} J_0(ii; ih) &= Q_0(ii; ih, ii) + Q_0(ii; ih, ih) + \sum_{k \neq i, h} Q_0(ii; ih, ik) \\ &= p_i^2 p_h (1 - 2S_2 + S_3 - 2(1 - S_2)p_h + 3p_h^2 - 3p_h^3) \quad (h \neq i); \end{aligned}$$

$$(2.6) \quad \begin{aligned} J_0(ij, ii) &= Q_0(ij; ii, ii) + Q_0(ij; ii, jj) + Q_0(ij; ii, ij) \\ &\quad + \sum_{k \neq i, j} (Q_0(ij; ii, ik) + Q_0(ij; ii, jk)) \\ &= p_i^2 p_j (1 - 2S_2 + S_3 - 2(1 - S_2)p_i + 3p_i^2 - p_i p_j \\ &\quad - 3p_i^3 + p_i p_j (p_i + \frac{1}{2} p_j)) \quad (i \neq j), \end{aligned}$$

$$(2.7) \quad \begin{aligned} J_0(ij; ij) &= Q_0(ij; ij, ii) + Q_0(ij; ij, jj) + Q_0(ij; ij, ij) \\ &\quad + \sum_{k \neq i, j} (Q_0(ij; ij, ik) + Q_0(ij; ij, jk)) \\ &= p_i p_j (p_i + p_j) (1 - 2S_2 + S_3 - 2(1 - S_2)(p_i + p_j) + 3(p_i^2 + p_j^2) \\ &\quad + 2p_i p_j - 3(p_i^3 + p_j^3) - 2p_i p_j (p_i + p_j)) \quad (i \neq j), \end{aligned}$$

$$(2.8) \quad \begin{aligned} J_0(ij; ih) &= Q_0(ij; ih, ii) + Q_0(ij; ih, jj) + Q_0(ij; ih, ij) + Q_0(ij; ih; ih) \\ &\quad + Q_0(ij; ih, jh) + \sum_{k \neq i, j, h} (Q_0(ij; ih, ik) + Q_0(ij; ih, jk)) \\ &= p_i p_j p_h (1 - 2S_2 + S_3 - 2p_i p_j + \frac{3}{2} p_i p_j (p_i + p_j) \\ &\quad - 2(1 - S_2 - p_i p_j) p_h + 3p_h^2 - 3p_h^3) \quad (i \neq j; h \neq i, j). \end{aligned}$$

These relations (2.4) to (2.8) correspond to (4.13) to (4.17) of IX, respectively. Corresponding to (4.18) to (4.22) of IX, we get the following results:

$$(2.9) \quad \sum_{i=1}^m J_0(ii; ii) = S_3 - 2S_4 - 2S_2S_3 + 3S_5 + S_3^2 + 2S_2S_4 - 3S_6,$$

$$(2.10) \quad \sum_{i=1}^m \sum_{h \neq i} J_0(ii; ih) = S_2 - S_3 - 4S_2^2 + 2S_4 + 6S_2S_3 \\ - 3S_5 + 2S_2^3 - S_3^2 - 5S_2S_4 + 3S_6;$$

$$(2.11) \quad \sum_{i,j}' (J_0(ij; ii) + J_0(ij; jj)) = S_2 - 3S_3 - 2S_2^2 + 5S_4 \\ + 4S_2S_3 - 5S_5 - \frac{1}{2}S_3^2 - S_2S_4 + \frac{3}{2}S_6,$$

$$(2.12) \quad \sum_{i,j}' J_0(ij; ij) = S_2 - 3S_3 - 4S_2^2 + 7S_4 + 10S_2S_3 - 11S_5 \\ + 2S_2^3 - 3S_3^2 - 9S_2S_4 + 10S_6,$$

$$(2.13) \quad \sum_{i,j}' \sum_{h \neq i,j} (J_0(ij; ih) + J_0(ij; jh)) = 1 - 7S_2 + 10S_3 + 8S_2^2 - 11S_4 \\ - 7S_2S_3 + 5S_5 - S_3^2 - 2S_2S_4 + 4S_6.$$

The sum of (2.9) and (2.10) yields the partial sum of probabilities of proving non-paternity against both children of a mother-child-child combination over homozygotic mothers:

$$(2.14) \quad S_2(1 - 4S_2 + 4S_3 + 2S_2^2 - 3S_4),$$

while the sum of (2.11) to (2.13) yields that over heterozygotic mothers:

$$(2.15) \quad 1 - 5S_2 + 4S_3 + 2S_2^2 + S_4 + 7S_2S_3 - 11S_5 + 2S_2^3 - \frac{9}{2}S_3^2 - 12S_2S_4 + \frac{31}{2}S_6.$$

The sum of the last two expressions (2.14) and (2.15) represents the *whole probability of proving non-paternity against both children of a mother-child-child combination*, stating

$$(2.16) \quad J_0 = 1 - 4S_2 + 4S_3 - 2S_2^2 + S_4 + 11S_2S_3 - 11S_5 \\ + 4S_2^3 - \frac{9}{2}S_3^2 - 15S_2S_4 + \frac{31}{2}S_6.$$

In case of *MN* blood type which may be regarded as a special case of general development, the whole probability becomes briefly

$$(2.17) \quad J_{0MN} = s^2t^2(2 - 3st).$$

The case where recessive genes are existent and phenotypes alone are available for judgement can also, in principle, quite similarly be treated. For instance, we obtain the following expressions for *whole probabilities in various blood types*:

$$(2.18) \quad J_{0ABO} = p^2(1-p)^4 + q^2(1-q)^4 + \frac{1}{2}pqr^2(1+7r^2),$$

$$(2.19) \quad J_{0Q} = u^2v^4,$$

$$(2.20) \quad J_{0Q\pm} = u^2v^4 + (2u + v_1)v_1v_1^4.$$

The results obtained hitherto in the present section can be generalized to mixed case. But, the detail will be left to the reader.

By comparing the probabilities J_0 's in (2.18), (2.19), (2.20) with the corresponding probabilities J 's in (4.26), (4.27), (4.28) of IX respectively, we know that *the inequalities*

$$(2.21) \quad J_{0ABO} \leq J_{ABO}, \quad J_{0Q} \leq J_Q, \quad J_{0Qq\pm} \leq J_{Qq\pm}$$

hold good, with inequality sign except trivial distributions. In fact, we can verify (2.21) as follows:

$$\begin{aligned} J_{ABO} - J_{0ABO} &= \frac{1}{2}p(1-p)^5 + \frac{1}{2}q(1-q)^5 + \frac{1}{4}pqr^2(2+r-7r^2) \\ &= \frac{1}{2}p(q+r)^5 + \frac{1}{2}q(p+r)^5 + \frac{1}{4}pqr^2(2+r-7r^2) \\ &\geq \frac{1}{2}p \cdot 5qr^4 + \frac{1}{2}q \cdot 5pr^4 + \frac{1}{4}pqr^2(2+r-7r^2) \\ &= \frac{1}{4}pqr^2(2+r+13r^2) \geq 0, \end{aligned}$$

$$J_Q - J_{0Q} = \frac{1}{2}uv^5 \geq 0, \quad J_{Qq\pm} - J_{0Qq\pm} = \frac{1}{2}(u(v^5 - v_1v_2^4) + v_1v_2^5) \geq 0.$$

Comparison of (2.17) with (8.1) of IX shows

$$(2.22) \quad J_{0MN} \leq J_{MN};$$

equality sign being exclusive unless $st=0$. In fact, remembering $0 \leq st \leq 1/4$, we have

$$J_{MN} - J_{0MN} = \frac{1}{4}st(1-2st)(2-3st) \geq 0.$$

An analogous inequality between (2.16) and (4.25) of IX, i.e.,

$$(2.23) \quad J_0 \leq J$$

does hold also good. We can conclude moreover that the corresponding inequalities between (2.14) and (4.23) of IX and between (2.15) and (4.24) of IX are also valid. For instance, the difference of (2.14) and (4.23) of IX becomes, in fact,

$$\begin{aligned} &S_2(1-3S_2 + \frac{5}{2}S_3 + S_2^2 - \frac{3}{2}S_4) - S_2(1-4S_2 + 4S_3 + 2S_2^2 - 3S_4) \\ &= S_2(S_2 - \frac{3}{2}S_3 - S_2^2 + \frac{3}{2}S_4) = S_2 \sum_{i=1}^m \sum_{h \neq i} p_i^2 p_h (1 - \frac{1}{2}p_i - p_h) \geq 0. \end{aligned}$$

The difference of (2.15) and (4.24) of IX can, though somewhat troublesome in calculation, be estimated in a similar way.

In conclusion, we remark that the whole probability of proving non-paternity against a distinguished child alone is given by

$$(2.24) \quad \begin{aligned} I_0 - J_0 &= 2S_2 - 3S_3 + S_4 - 8S_2S_3 + 8S_5 \\ &\quad - 4S_2^3 + \frac{9}{2}S_3^2 + 15S_2S_4 - \frac{31}{2}S_6, \end{aligned}$$

and that that against at least one child by

$$(2.25) \quad \begin{aligned} J_0 \equiv 2I_0 - J_0 &= 1 - 2S_3 - 2S_2^2 + 3S_4 - 5S_2S_3 + 5S_5 \\ &\quad - 4S_2^3 + \frac{9}{2}S_3^2 + 15S_2S_4 - \frac{31}{2}S_6; \end{aligned}$$

cf. (6.3) and (7.1) of IX.

—To be continued—