SOME THEOREMS CONCERNING EXTREMA OF BROWNIAN MOTION WITH d-DIMENTIONAL TIME

Dedicated to Professor N. Ikeda on his 70th birthday

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Introduction

Let $X = \{X(x), x \in \mathbf{R}^d\}$ be a Lévy's Brownian motion with d-dimensional time ([2]) defined on a certain probability space (Ω, P) ; thus X is a centered Gaussian system with continuous sample functions satisfying X(0) = 0 and $E\{X(x)X(y)\} = (|x| + |y| - |x - y|)/2$. For a nonempty subset A of \mathbf{R}^d we put

$$\underline{X}(A) = \inf\{X(x) : x \in A\}, \quad \overline{X}(A) = \sup\{X(x) : x \in A\}.$$

We often use the notation X(A) to denote either $\underline{X}(A)$ or $\overline{X}(A)$. For example, X(A) - X(B) denotes any one of $\underline{X}(A) - \underline{X}(B)$, $\underline{X}(A) - \overline{X}(B)$, $\overline{X}(A) - \underline{X}(B)$ and $\overline{X}(A) - \overline{X}(B)$. A point x in \mathbf{R}^d is called a point of local minimum (resp. local maximum) of a sample function X if there exists a neighborhood U of X such that $X(X) = \underline{X}(U)$ (resp. $X(X) = \overline{X}(U)$). A point of either local minimum or local maximum is called an extreme-point.

The following are typical of those problems and theorems we discuss in this paper.

- (I) Under what condition on A does the probability distribution of X(A) admit a strictly positive C^{∞} -density?
- (II) Under what condition on A and B does the joint probability distribution of X(A) and X(B) admit a strictly positive C^{∞} -density?
- (III) Almost all sample functions X have the following property: There are no distinct extreme-points x and y with X(x) = X(y).

We give some sufficient conditions that will give positive answers to the problems (I) and (II) and then give a proof of (III). Formulating the problems somewhat generally we state our main results in the following theorems.

Theorem 1. Let A_k , $1 \le k \le n$, be nonempty bounded closed sets not containing the origin 0. Then for any constants c_k , $1 \le k \le n$, such that $c_1 + c_2 + \cdots + c_n \ne 0$, the probability distribution of

$$c_1X(A_1) + c_2X(A_2) + \cdots + c_nX(A_n)$$

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can be expressed as a convolution $\gamma * \mu$ where γ is a nondegenerate Gaussian distribution with mean 0 and μ is some probability distribution in \mathbf{R} . In particular, the distribution of each of $\underline{X}(A)$ and $\overline{X}(A)$ has a strictly positive C^{∞} -density provided that A is a nonempty bounded closed set not containing 0.

Theorem 2. Let $A_j, B_k, 1 \le j \le m, 1 \le k \le n$, be nonempty bounded closed sets such that $\bigcup_{j=1}^m A_j$ is separated from $\bigcup_{k=1}^n B_k$ by a certain (d-1)-dimensional hyperplane Π passing through the origin 0. Then for any constants $c_j, c'_k, 1 \le j \le m, 1 \le k \le n$, such that $\sum_{j=1}^m c_j \ne 0$ and $\sum_{k=1}^n c'_k \ne 0$, the joint distribution of

(1)
$$f_1(X) = \sum_{j=1}^m c_j X(A_j), \qquad f_2(X) = \sum_{k=1}^n c_k' X(B_k)$$

has a form $(\gamma_1 \otimes \gamma_2) * \nu$ where each γ_i is a nondegenerate Gaussian distribution with mean 0 and ν is some 2-dimensional probability distribution. In particular, the joint distribution of X(A) and X(B) has a strictly positive C^{∞} -density provided that A and B are nonempty bounded closed sets separated from each other by a certain (d-1)-dimensional hyperplane passing through 0.

Theorem 3. Let A_j , B_k , $1 \le j \le m$, $1 \le k \le n$, be nonempty bounded closed sets such that $\bigcup_{j=1}^n A_j$ is separated from $\bigcup_{k=1}^n B_k$ by a certain (d-1)-dimensional hyperplane. Then for any constants c_j , c'_k , $1 \le j \le m$, $1 \le k \le n$, such that $\sum_{j=1}^m c_j = \sum_{k=1}^n c'_k \ne 0$, the probability distribution of $f_1(X) - f_2(X)$, with f_1 and f_2 given by (1), has a form $\gamma * \mu$ where γ is a nondegenerate Gaussian distribution with mean 0 and μ is some distribution in R. In particular, the distribution of each of $\underline{X}(A) - \underline{X}(B)$, $\underline{X}(A) - \overline{X}(B)$ and $\overline{X}(A) - \overline{X}(B)$ has a strictly positive C^{∞} -density provided that A and B are nonempty bounded closed sets separated from each other by a certain (d-1)-dimensional hyperplane.

Theorem 4. Almost all sample functions X have the following property: There are no distinct extreme-points x and y of X such that X(x) = X(y).

An example of the applicability (or our motivation) of Theorem 4 will be given in the final section.

1. A lemma

Given a centered Gaussian system $\{X_{\lambda}, \lambda \in \Lambda\}$ defined on a certain probability space (Ω, P) , we denote by H the real Hilbert space spanned by $\{X_{\lambda}, \lambda \in \Lambda\}$ and by H_0 the closed linear span (abbreviation: c.l.s.) of $\{X_{\lambda} - X_{\mu}, \lambda, \mu \in \Lambda\}$. Clearly $H_0 \subset H \subset L^2(\Omega, P)$. We now introduce the following conditions.

Condition (A). There exists a nondegenerate Gaussian random variable Y_0 inde-

pendent of $\{X_{\lambda} - Y_0, \lambda \in \Lambda\}$.

Condition (B). There exists $\lambda \in \Lambda$ such that $X_{\lambda} \notin H_0$.

It is easy to see that the condition (B) implies that $X_{\lambda} \notin H_0$ for all $\lambda \in \Lambda$. Denote by \mathbf{R}^{Λ} the space of real valued functions on Λ ; it has a Borel structure defined in a natural way. Then we can regard $X_{\Lambda} = \{X_{\lambda}, \ \lambda \in \Lambda\}$ as a random variable taking values in \mathbf{R}^{Λ} . The following lemma is rather trivial; nevertheless, it plays a fundamental role in this paper.

Lemma 1. (i) Let f be a Borel function from \mathbf{R}^{Λ} to R such that

$$(1.1) f(w+t\mathbf{1}) = f(w) + ct$$

for any $w \in \mathbb{R}^{\Lambda}$ and $t \in \mathbb{R}$ where c is some nonzero constant and $\mathbf{1}$ denotes the function on Λ that identically equals 1. Then under the condition (A) we have $f(X_{\Lambda}) = cY_0 + Y$ with a suitable random variable Y independent of Y_0 ; in particular, the probability distribution of $f(X_{\Lambda})$ has a strictly positive C^{∞} -density.

- (ii) Suppose Λ is a locally compact space with a countable open base and assume that X_{λ} is continuous in λ with probability 1. We regard $X_{\Lambda} = \{X_{\lambda}, \lambda \in \Lambda\}$ as a random variable taking values in the space $C(\Lambda)$ of continuous functions on Λ , which is equipped with the compact uniform topology. Then, under the condition (A), the conclusion of (i) remains valid for any Borel function f from $C(\Lambda)$ to R satisfying (1.1) for $w \in C(\Lambda)$ and $t \in R$.
- (iii) The condition (B) implies the condition (A).

REMARK 1. Let Λ_k , $1 \leq k \leq n$, be subsets of Λ and let c_k , $1 \leq k \leq n$, be constants such that $c_1 + \cdots + c_n \neq 0$. Let $w(\Lambda_k)$ indicate either $\inf\{w(\lambda) : \lambda \in \Lambda_k\}$ or $\sup\{w(\lambda) : \lambda \in \Lambda_k\}$; the choice may depend on k but not on w. Then

$$f(w) = c_1 w(\Lambda_1) + \dots + c_n w(\Lambda_n)$$

is a typical example of f satisfying (1.1) with $c = c_1 + \cdots + c_n$ provided that f can be defined to be a Borel function.

REMARK 2. Let F be a class of functions defined on [0,1] and taking values in Λ (an example of such an F is the space of continuous paths in Λ connecting two given points of Λ). Then the function f defined by $f(w) = \inf\{g(w,u) : u \in F\}$ with $g(w,u) = \sup\{w(u(t)) : 0 \le t \le 1\}$ satisfies (1.1).

REMARK 3. If $\{X_{\lambda}, \lambda \in \Lambda\}$ satisfies (A) (resp. (B)) and if Λ_1 is a nonempty subset of Λ , then the sub-system $\{X_{\lambda}, \lambda \in \Lambda_1\}$ also satisfies (A) (resp. (B)).

Proof of Lemma 1. (i) Under the condition (A) $X_{\Lambda} - Y_0 \mathbf{1}$ and Y_0 are independent so $f(X_{\Lambda}) - cY_0 = f(X_{\Lambda} - Y_0 \mathbf{1})$ and Y_0 are independent. If we put $Y = f(X_{\Lambda}) - cY_0$,

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then we have the expression $f(X_{\Lambda}) = cY_0 + Y$ in which Y_0 and Y are independent and Y_0 is a nondegenerate Gaussian random variable. The assertion (ii) follows from (i). (iii) It is easy to see that $X_{\lambda} + H_0 = \{X_{\lambda} + Y : Y \in H_0\}$ does not depend on λ . The condition (B) means that $X_{\lambda} + H_0 \not\equiv 0$. Since $X_{\lambda} + H_0$ is a closed convex set, there exists a unique $Y_0 \in X_{\lambda} + H_0$ such that

$$\sqrt{E\{Y_0^2\}} = \min\left\{\sqrt{E\{|X_\lambda + Y|^2\}} : Y \in H_0\right\} > 0.$$

Then clearly $Y_0 \perp H_0$. Since $X_{\lambda} - Y_0 \in H_0$, $X_{\lambda} - Y_0 \perp Y_0$ for all λ . This implies that Y_0 is independent of $\{X_{\lambda} - Y_0, \lambda \in \Lambda\}$.

2. Proof of Theorem 1

As stated in Introduction let $X = \{X(x), x \in \mathbb{R}^d\}$ be a Brownian motion with d-dimensional time. For any fixed pair of real numbers t_1 and t_2 such that $0 < t_1 < t_2$ we put $\Lambda = \{x \in \mathbb{R}^d : t_1 \le |x| \le t_2\}$, $H = \text{c.l.s.}\{X(x), x \in \Lambda\}$ and $H_0 = \text{c.l.s.}\{X(x) - X(y), x, y \in \Lambda\}$. First we prepare the following lemma.

Lemma 2. The condition (B) is satisfied for $\{X(x), x \in \Lambda\}$, namely, there exists $x \in \Lambda$ such that $X(x) \notin H_0$.

Proof. (i) We consider the case where the dimension d is odd and $d \ge 3$. Denoting by $\hat{d}\theta$ the uniform distribution on $S^{d-1} = \{\theta \in \mathbb{R}^d : |\theta| = 1\}$, we put

$$R(t) = \int_{S^{d-1}} X(t\theta) \hat{d}\theta, \qquad t \ge 0,$$

$$H_1 = \text{c.l.s.} \{ R(t), \quad t_1 \le t \le t_2 \},$$

$$H_1^{\perp} = \text{the orthogonal complement of } H_1 \text{ in } H.$$

Then we have

(2.1)
$$X(x) - R(|x|) \in H_1^{\perp}$$
 for any $x \in \Lambda$.

In fact, it is easy to see that, for each fixed $t \ge 0$, $E\{(X(x) - R(|x|))R(t)\}$ depends only on |x| and hence it must vanish, which implies (2.1). We are going to prove that $X(t_1\theta) \not\in H_0$ for $\theta \in S^{d-1}$. The relation (2.1) implies that $X(t_1\theta) = R(t_1) + X'$ with $X' \in H_1^{\perp}$ and that $H_0 \subset H_{10} \oplus H_1^{\perp}$ where $H_{10} = \text{c.l.s.}\{R(t) - R(s), t, s \in [t_1, t_2]\}$. Therefore, for the proof of $X(t_1\theta) \not\in H_0$ it is enough to show that $R(t_1) \not\in H_{10}$. We now make use of the canonical representation of the Gaussian process $\{R(t), t \ge 0\}$ due to McKean [5], which means that

$$R(t) = \int_0^t f(t,r) dB(r), \quad t \ge 0,$$

where $\{B(r), r \ge 0\}$ is a one-dimensional standard Brownian motion and

(2.2)
$$f(t,r) = k(d) \int_{r/t}^{1} (1 - u^2)^{(d-3)/2} du, \qquad 0 \le r \le t,$$

k(d) being a suitable constant depending only on d. For any s and t with $t_1 \le s < t \le t_2$ we have

$$R(t) - R(s) = \int_0^{t_1} f_{ts}(r) dB(r) + \int_{t_1}^t g_{ts}(r) dB(r),$$

$$R(t_1) = \int_0^{t_1} f(r) dB(r),$$

where $f_{ts}(r) = f(t, r) - f(s, r)$, $f(r) = f(t_1, r)$ and $g_{ts}(r)$ is a suitable function. Therefore, if we put

$$\tilde{H}_0 = \text{c.l.s.} \left\{ \int_0^{t_1} f_{ts}(r) dB(r), t, s \in [t_1, t_2] \right\},$$

$$\tilde{H}_+ = \text{c.l.s.} \left\{ B(u) - B(r), r, u \in [t_1, t_2] \right\},$$

then $\tilde{H}_0 \perp \tilde{H}_+$, $H_{10} \subset \tilde{H}_0 \oplus \tilde{H}_+$ and $R(t_1) \perp \tilde{H}_+$. From these observations we see that for the proof of $R(t_1) \notin H_{10}$, it is enough to show

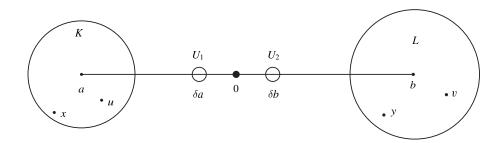
(2.3)
$$\int_0^{t_1} f(r)dB(r) \notin \tilde{H}_0.$$

Let L_0^2 be the subspace of $L^2[0,t_1]$ spanned by the functions $f_{ts}(\cdot),t,s\in[t_1,t_2]$. Then the Hilbert space \tilde{H}_0 is isomorphic to L_0^2 and (2.3) is equivalent to $f\not\in L_0^2$. Now the assumption that d is an odd integer ≥ 3 implies that $f_{ts}(r),t,s,\in[t_1,t_2]$, are polynomials of degree d-2 vanishing at r=0 (use (2.2)). Therefore all the functions in L_0^2 are also polynomials of degree at most d-2 vanishing at r=0. On the other hand it is easy to see that f is a polynomial of degree d-2 with f(0)>0. Therefore $f\not\in L_0^2$, which finally implies $X(t_1\theta)\not\in H_0$. This completes the proof in the case where d is odd and $d\geq 3$.

(ii) The proof in the case where d is even can be obtained by the method of descent in which a Brownian motion with d-dimensional time is viewed as the restriction of a Brownian motion with (d+1)-dimensional time to $\mathbb{R}^d \times \{0\} \subset \mathbb{R}^{d+1}$ and also by using Remark 3. The proof in the case d=1 is easy. The proof of Lemma 2 is finished.

We are now able to prove Theorem 1. From the assumption on A_k , $1 \le k \le n$, there exist t_1 and t_2 with $0 < t_1 < t_2$ such that $\Lambda = \{x \in \mathbb{R}^d : t_1 \le |x| \le t_2\}$ includes all A_k . Then, by Lemma 2 the condition (B) is satisfied for $X_{\Lambda} = \{X(x), x \in \Lambda\}$

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and by Remark 1 the condition (1.1) is satisfied for the function $f(w) = c_1w(A_1) + c_2w(A_1) + \cdots + c_nW(A_n)$, $w \in C(\Lambda)$, with $c = c_1 + \cdots + c_n$. Therefore by Lemma 1 the probability distribution of the random variable $f(X_{\Lambda}) = c_1X(A_1) + c_2X(A_2) + \cdots + c_nX(A_n)$ has a form $\gamma * \mu$. This completes the proof of Theorem 1.

3. Proof of Theorem 2

Under the assumption on A_j and B_k in Theorem 2 we can take disjoint closed balls K and L with the following properties:

- $(3.1) K \supset \bigcup_{j=1}^m A_j, \qquad L \supset \bigcup_{k=1}^n B_k.$
- (3.2) K is separated from L by the hyperplane Π .
- (3.3) The center a of K and the center b of L are on the straight line that passes through the origin 0 and is perpendicular to Π .

We consider open balls U_1 and U_2 with a common radius ε and with centers δa and δb , respectively, where $\delta > 0$ is chosen so that $\delta a \notin K$ and $\delta b \notin L$ (see the figure). We now make use of the Chentsov representation of X(x) ([1]), which asserts that

$$(3.4) X(x) = W(D_x),$$

where D_x is the open ball with center x/2 and radius |x|/2, and $\{W(d\xi)\}$ is a suitable white noise in \mathbb{R}^d associated with the measure $c_d|\xi|^{-d+1}d\xi$ (c_d is a suitable constant), namely, a Gaussian random measure in \mathbb{R}^d such that $E\{W(d\xi)\}=0$ and $E\{W(d\xi)^2\}=c_d|\xi|^{-d+1}d\xi$. By taking $\varepsilon>0$ small enough, we can assume

$$(3.5) U_1 \subset \left\{ \bigcap_{x \in K} D_x \right\} \bigcap \left\{ \bigcup_{y \in L} D_y \right\}^c, U_2 \subset \left\{ \bigcap_{y \in L} D_y \right\} \bigcap \left\{ \bigcup_{x \in K} D_x \right\}^c.$$

If we write $X(x) = W(D_x) = W(U_1) + \tilde{X}_x$ and $X(y) = W(D_y) = W(U_2) + \tilde{X}_y$, then (3.5) implies that the 2-dimensional random vector $(W(U_1), W(U_2))$ is independent of the Gaussian family $\{(\tilde{X}_x, \tilde{X}_y) : x \in K, y \in L\}$. Therefore we have

$$f_1(X) = cW(U_1) + \tilde{f}_1, \qquad f_2(X) = c'W(U_2) + \tilde{f}_2,$$

with $c = \sum_{j=1}^{m} c_j$, $c' = \sum_{k=1}^{n} c'_k$ and $(W(U_1), W(U_2))$ is independent of $(\tilde{f}_1, \tilde{f}_2)$. Since $W(U_1)$ and $W(U_2)$ are independent and each of them is a nondegenerate Gaussian random variable with mean 0, the joint distribution of $f_1(X)$ and $f_2(X)$ has a form $(\gamma_1 \otimes \gamma_2) * \nu$.

4. Proof of Theorem 3 and Theorem 4

By using the fact that $\{X(x)-X(x_0), x \in \mathbb{R}^d\}$ is identical in law to $\{X(x-x_0), x \in \mathbb{R}^d\}$ for each $x_0 \in \mathbb{R}^d$ and also by using the assumption $\sum_{j=1}^m c_j = \sum_{k=1}^n c_k'$, we see that the probability distribution of $f_1(X) - f_2(X)$ is invariant under any simultaneous shift of A_j and B_k . Therefore, in proving Theorem 3 we may assume that A_j . and B_k satisfy the same assumption as in Theorem 2. Then the joint distribution of $f_1(X)$ and $f_2(X)$ has a form $(\gamma_1 \otimes \gamma_2) * \nu$ by Theorem 2 and this implies the conclusion of Theorem 3.

Before going to the proof of Theorem 4 we introduce some notation. Denote by \mathcal{K} the set of all pairs (K_1,K_2) of disjoint closed balls K_1 and K_2 with rational centers and rational radii. We put $f(K_1,K_2;\sigma_1,\sigma_2)=X(K_1;\sigma_1)-X(K_2;\sigma_2)$ where each σ_i is either 0 or 1 and $X(K_i;\sigma_i)$ denotes either $\underline{X}(K_i)$ or $\overline{X}(K_i)$ according as $\sigma_i=0$ or 1. We also denote by $\mathcal{E}(K_1,K_2;\sigma_1,\sigma_2)$ the event $\{f(K_1,K_2;\sigma_1,\sigma_2)=0\}$ and then put $\mathcal{E}'=\cup\mathcal{E}(K_1,K_2;\sigma_1,\sigma_2)$ where the union is taken over all $(K_1,K_2)\in\mathcal{K}$ and all $(\sigma_1,\sigma_2)\in\{0,1\}^2$. Finally let \mathcal{E} be the event such that there exist distinct extremepoints x and y with X(x)=X(y). It is then easy to see that $\mathcal{E}\subset\mathcal{E}'$. On the other hand Theorem 3 implies $P\{\mathcal{E}(K_1,K_2;\sigma_1,\sigma_2)\}=0$ and hence $P\{\mathcal{E}'\}=0$. This implies $P\{\mathcal{E}\}=0$ as was to be proved.

5. Remarks on a diffusion process in a d-dimensional Brownian environment

This section is to supply an example for the applicability of Theorem 4. We change the notation for a Brownian motion with a d-dimensional time since we want to use X(t) for a diffusion process. Let \mathbf{W} be the space of continuous functions on \mathbf{R}^d vanishing at 0. In this section an element W of \mathbf{W} is called an environment. We consider the probability measure P on W such that $\{W(x), x \in \mathbf{R}^d, P\}$ is a Lévy's Brownian motion with a d-dimensional time. Let Ω be the space of continuous functions on $[0,\infty)$ taking values in \mathbf{R}^d . The value of $\omega(\in \Omega)$ at time t is denoted by $X(t) = X(t,\omega) = \omega(t)$. For each fixed environment W we consider the probability measure P_W on Ω such that $\{X(t), t \geq 0, P_W\}$ is a diffusion process in \mathbf{R}^d with generator

$$\frac{1}{2}(\Delta - \nabla W \cdot \nabla) = \frac{1}{2}e^{W} \sum_{k=1}^{d} \frac{\partial}{\partial x_{k}} \left(e^{-W} \frac{\partial}{\partial x_{k}} \right)$$

and starting from 0. Let \mathcal{P} be the probability measure on $W \times \Omega$ defined by $\mathcal{P}(dWd\omega) = P(dW)P_W(d\omega)$. Then $\{X(t), t \geq 0, \mathcal{P}\}$ can be regarded as a process defined on the probability space $(W \times \Omega, \mathcal{P})$, which we call a diffusion process in

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a d-dimensional Brownian environment. When d=1, this model is a diffusion analogue of well-known Sinai's random walk in a random environment(1982) and much is known about the long-term behavior of X(t) such as localization. When $d\geq 2$, a similar diffusion model appeared in [3]. Now our interest is the long-term behavior of $\{X(t), t\geq 0, \mathcal{P}\}$ in the case $d\geq 2$. Tanaka [6](see also [7]) proved that, for any dimension d, $\{X(t), t\geq 0, \mathcal{P}\}$ is recurrent for almost all Brownian sample environments W. Mathieu[4] proved that localization takes place for $\{X(t), t\geq 0, \mathcal{P}\}$, in the sense that

$$\lim_{N\to\infty} \overline{\lim_{\lambda\to\infty}} \, \mathcal{P}\left(\lambda^{-2} \max\{|X(t)| : 0 \le t \le e^{\lambda}\} > N\right) = 0.$$

However, in the case $d \geq 2$, it seems that the existence of the limiting distribution of $\{\lambda^{-2}X(e^{\lambda}), \mathcal{P}\}$ as $\lambda \to \infty$ is still an open problem. We give a remark on this problem. We notice the scaling relation

$$\{X(t), t \geq 0, P_{\lambda W_{\lambda}}\} \stackrel{d}{=} \{\lambda^{-2}X(\lambda^{4}t), t \geq 0, P_{W}\},\$$

where $\lambda > 0$ and $W \in \mathbf{W}$ are fixed, W_{λ} denotes an element of W defined by $W_{\lambda}(x) = \lambda^{-1}W(\lambda^2x)$, $x \in \mathbb{R}^d$, and $\stackrel{d}{=}$ means the equality in distribution. This scaling relation combined with $W_{\lambda} \stackrel{d}{=} W$ imply the following: If we can prove that $\{X(e^{r\lambda}), P_{\lambda W}\}$ has the limiting distribution as $\lambda \to \infty$ under the condition $r = r(\lambda) \to 1$, then so does $\{\lambda^{-2}X(e^{\lambda}), \mathcal{P}\}$. From now on we are interested in $\{X(t), P_{\lambda W}\}$. For $W \in \mathbf{W}$ we define the sub-level domain D as the connected component of the open set $\{x \in \mathbb{R}^d : W(x) < 1\}$ containing 0. Then it is easy to see that D is bounded, P-a.s. By making use of Theorem 4 we see that for W not belonging to some P-negligible subset of \mathbf{W} , there exists a point \tilde{b} of local (strict) minimum of W with depth > 1 inside D. Such a point \tilde{b} is characterized by (i) $W(\tilde{b}) < W(x)$ for $x \in U - \{\tilde{b}\}$ and (ii) $U \subset D$, where U denotes the connected component of the open set $\{x \in \mathbb{R}^d : W(x) - W(\tilde{b}) < 1\}$ containing \tilde{b} . It is obvious that the totality of such points \tilde{b} is a finite set, which is denoted by $\{b_k(W), 1 \le k \le l(W)\}$. Now suppose l(W) = 1 and put $b = b_1(W)$. Then from the argument of [4] we see that

(5.1)
$$X(e^{r\lambda}) \to b$$
 (in probability with respect to $P_{\lambda W}$)

as $\lambda \to \infty$ provided $r = r(\lambda)$ (non-random) tends to 1. If $l(W) \ge 2$, we do not know whether the limiting distribution of $X(e^{r\lambda})$ exists. Hoping for the best, we think it might be possible to define b, in one way or another, as *a single point* among $b_k(W)$, $1 \le k \le l(W)$, and to prove (5.1) even in the case $l(W) \ge 2$, for almost all w.

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