# Absolute continuity of similar translations

Dedicated to Professor Shinzo Watanabe on his sixtieth birthday

Ву

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### § 1. Introduction

Let  $X = \{X_k\}_k$  be an IID, let  $Y = \{Y_k\}_k$  be an independent random sequence defined on a probability space  $(\Omega, \mathcal{F}, P)$ , and assume that X and Y are independent. Denote by  $\mu_x$ ,  $\mu_y$  and  $\mu_{x+y}$  the probability measures on  $\mathbf{R}^{N}$  (the space of all real sequences), induced by  $\mathbf{X}$ ,  $\mathbf{Y}$  and  $\mathbf{X} + \mathbf{Y} = \{X_{k} + Y_{k}\}_{k}$ , respectively. Since X and X + Y are independent random sequences,  $\mu_X$  and  $\mu_{X+Y}$  are product measures:

$$\mu_{\mathbf{X}} = \prod_{\mathbf{k}} \mu_{X_{\mathbf{k}}}$$
 and  $\mu_{\mathbf{X}+\mathbf{Y}} = \prod_{\mathbf{k}} \mu_{X_{\mathbf{k}}+Y_{\mathbf{k}}}$ ,

where  $\mu_{X_k}$  and  $\mu_{X_k+Y_k}$  are marginal distributions. When  $\mu_{X_1}$  is absolutely continuous with respect to the Lebesgue measure dx, define  $f(x) = \frac{d\mu_{x_1}}{dx}(x)$ . If f is an absolutely continuous function, f' denotes the derivative of f in the distribution sense, and if f' is an absolutely continuous function, f'' denotes the derivative of f' in the same sense. In these cases, define

$$I_1(\mathbf{X}) = \int_{-\infty}^{+\infty} \frac{f'(x)^2}{f(x)} dx$$
 and  $I_2(\mathbf{X}) = \int_{-\infty}^{+\infty} \frac{f''(x)^2}{f(x)} dx$  if  $f > 0$  a.e..

Sato and Watari [8, Theorem 1] proved the relation  $I_1(\mathbf{X}) \leq \frac{3}{2} \sqrt{I_2(\mathbf{X})}$ , so that  $I_2(\mathbf{X}) < \infty$  implies  $I_1(\mathbf{X}) < \infty$ .

Several authors have investigated the conditions for satisfying  $\mu_{x+y}$ ~  $\mu_{\rm X}$  (mutually absolutely continuous) in terms of the distribution of Y, but necessary and sufficient conditions are not yet known in general (see Sato [7]). In the present paper we concentrate on the case in which Y is a similar random sequence, that is,  $Y = a\Theta = \{a_k \Theta_k\}_k$ , where  $\Theta = \{\Theta_k\}_k$  are independent copies of a random variable  $\Theta$ , and  $\mathbf{a} = \{a_k\}_k$  is a real sequence. In the following with the exception of Section 2, we fix the above notation and assume  $P(\Theta \neq 0) > 0$ . The following results are known.

**Theorem A** (Shepp [9]). Assume  $\Theta \equiv 1$  a. s.. Then we have:

- (1)  $\mu_{X+a\theta} \sim \mu_X \text{ implies } \sum_k a_k^2 < \infty.$
- (2) Assume  $I_1(\mathbf{X}) < \infty$ . Then  $\sum_k a_k^2 < \infty$  implies  $\mu_{\mathbf{X}+\mathbf{a}\mathbf{\theta}} \sim \mu_{\mathbf{X}}$ .
- (3) If  $\sum_{k} a_k^2 < \infty$  implies  $\mu_{X+a\theta} \sim \mu_X$ , then  $I_1(X) < \infty$  holds.

**Theorem B** (Okazaki and Sato [6], Sato and Watari [8], and Okazaki [5]). Let  $\Theta = \{\Theta_k\}_k$  be the Rademacher sequence, that is,  $P(\Theta = 1) = P(\Theta = -1) = \frac{1}{2}$ . Then we have:

- (1)  $\mu_{X+a\theta} \sim \mu_X \text{ implies } \sum_k a_k^4 < \infty.$
- (2) Assume  $I_2(\mathbf{X}) < \infty$ . Then  $\sum_k a_k^4 < \infty$  implies  $\mu_{\mathbf{X}+\mathbf{a}\theta} \sim \mu_{\mathbf{X}}$ .
- (3) If  $\sum_{k} a_{k}^{4} < \infty$  implies  $\mu_{X+a\theta} \sim \mu_{X}$ , then  $I_{2}(X) < \infty$  holds.

**Theorem C** (Kakutani [3]). Let  $X = \{X_k\}_k$  be a standard Gaussian sequence and  $\Theta$  be a standard Gaussian random variable. Then  $\mu_{X+a\Theta} \sim \mu_X$  holds if and only if  $\sum_k a_k^4 < \infty$ .

In this paper we first prove, without assumption of the similarity of Y, a variation of Theorem 3 of Sato and Watari [8], and then prove the following theorems for similar  $Y = a\theta$ . We begin with necessary conditions for the relation  $\mu_{X+a\theta} \sim \mu_X$ .

Theorem 1. (1)  $\mu_{X+a\theta} \sim \mu_X \text{ implies } \sum_k a_k^4 < \infty.$ 

- (2) If  $\lim \inf_{x\to\infty} | E[\Theta: |\Theta| \le x] | > 0$ , then  $\mu_{X+a\Theta} \sim \mu_X$  implies  $\sum_k a_k^2 < \infty$ .
- (3) If  $\lim \inf_{x\to\infty} x^p P(\mid \Theta \mid >x) > 0$  for some p>0, then  $\mu_{X+a\theta} \sim \mu_X$  implies  $\sum_k \mid a_k \mid^{2p} < \infty$ .

The following corollary is an immediate consequence of Theorem 1 (2).

Corollary 1. If  $\Theta$  is integrable and  $\mathbb{E}[\Theta] \neq 0$ , then  $\mu_{X+u\Theta} \sim \mu_X$  implies  $\sum_k a_k^2 < \infty$ .

Then sufficient conditions are:

- Theorem 2. (1) Assume  $I_1(\mathbf{X}) < \infty$  and  $E[\Theta^2] < \infty$ . Then  $\sum_k a_k^2 < \infty$  implies  $\mu_{\mathbf{X}+\mathbf{a}\mathbf{\theta}} \sim \mu_{\mathbf{X}}$ .
- (2) Assume  $I_2(\mathbf{X}) < \infty$ ,  $E[ | \Theta |^p] < \infty$  for some  $p \ge 2$ , and  $E[\Theta] = 0$ . Then  $\sum_k |a_k|^{p \land 4} < \infty$  implies  $\mu_{\mathbf{X} + n \Theta} \sim \mu_{\mathbf{X}}$ , where  $p \land 4 = \min(p, 4)$ .

Combining Theorems 1 and 2, we obtain necessary and sufficient conditions for several cases, extending (1) and (2) of both Theorems A

and B, and Theorem C.

**Theorem 3.** (1) Assume  $I_1(\mathbf{X}) < \infty$ ,  $E[\Theta^2] < \infty$ , and  $E[\Theta] \neq 0$ . Then  $\mu_{\mathbf{X}+\mathbf{a}\mathbf{e}} \sim \mu_{\mathbf{X}}$  holds if and only if  $\sum_k a_k^2 < \infty$ .

(2) Assume  $I_2(\mathbf{X}) < \infty$ ,  $E[\Theta^4] < \infty$ , and  $E[\Theta] = 0$ . Then  $\mu_{\mathbf{X}+\mathbf{a}\Theta} \sim \mu_{\mathbf{X}}$  holds if and only if  $\sum_k a_k^4 < \infty$ .

We now refine certain sufficient conditions. In the following Proposition 1, we weaken the assumption  $E[\mid \Theta \mid^{p}] < \infty$  to  $\sup_{x>0} x^{p} P(\mid \Theta \mid >x) < \infty$ .

**Proposition 1.** Assume  $I_2(\mathbf{X}) < \infty$  and  $\sup_{x \ge 0} x^p P(\mid \Theta \mid > x) < \infty$  for some p > 0. If one of the following (1)  $\sim$  (4) holds, then  $\sum_k \mid a_k \mid^{p \land 4} < \infty$  implies  $\mu_{\mathbf{X}+\mathbf{a}\mathbf{0}} \sim \mu_{\mathbf{X}}$ :

- (1) 0
- (2)  $2 and <math>E[\Theta] = 0$ .
- (3) p=4,  $E[\Theta]=0$ , and there exists  $\varepsilon > 0$  such that

$$\sup_{|z|<\varepsilon} (\varepsilon - |z|)^2 \int_{-\infty}^{+\infty} \frac{f''(x+z)^2}{f(x)} dx < \infty, \quad \text{where} \quad f(x) = \frac{d\mu_{x_1}}{dx} (x).$$

(4) p > 4 and  $E[\Theta] = 0$ .

On the other hand, we have the following.

**Theorem 4.** (1) If  $\sum_{k} a_{k}^{2} < \infty$  implies  $\mu_{X+a\theta} \sim \mu_{X}$ , then we have  $\limsup_{x\to\infty} |E[\Theta: |\Theta| \le x]| < \infty$ .

(2) If  $\sum_k a_k^4 < \infty$  implies  $\mu_{X+a\theta} \sim \mu_X$ , then we have  $E[\Theta^2] < \infty$  and  $E[\Theta] = 0$ .

If  $\Theta \ge 0$  a.s., then  $\limsup_{x \to \infty} |E[\Theta: |\Theta| \le x]| < \infty$  in Theorem 4 (1) implies  $E[\Theta] < \infty$ , so that we have the following.

**Corollary 2.** Assume  $\Theta \ge 0$  a. s.. If  $\sum_k a_k^2 < \infty$  implies  $\mu_{X+n\Theta} \sim \mu_X$ , then we have  $E[\Theta] < \infty$ .

**Example 1.** Let  $\mathbf{X} = \{X_k\}_k$  be an IID such that  $I_1(\mathbf{X}) < \infty$ , and let  $\mathbf{Y} = \{Y_k\}_k$  be an independent random sequence, independent of  $\mathbf{X}$ , such that each  $Y_k$  is exponentially distributed. Then the following  $(1) \sim (4)$  are equivalent:

- (1)  $\mu_{x+y} \sim \mu_x$ .
- (2)  $\sum_{k} E[Y_{k}]^{2} < \infty$ .
- (3)  $\sum_{k} \mathrm{E}[Y_{k}^{2}] < \infty$ .

(4) 
$$\sum_{k} Y_{k}^{2} < \infty$$
 a. s..

In fact,  $\mu_Y$  is expressed as  $\mu_Y = \mu_{a\theta}$ , where  $\Theta$  is exponentially distributed,  $E[\Theta] = 1$ , and  $a_k = E[Y_k]$ ,  $k \in \mathbb{N}$ . Then by Theorem 3 (1),  $\mu_{X+a\theta} \sim \mu_X$  is equivalent to  $\sum_k a_k^2 < \infty$ . They are also equivalent to  $\sum_k E[Y_k^2] < \infty$  because  $E[Y_k^2] = a_k^2 E[\Theta^2] = 2a_k^2$ . Moreover, we have by Kolmogorov's three series theorem,  $\sum_k Y_k^2 < \infty$  a. s. if and only if  $\sum_k a_k^2 < \infty$ .

**Example 2.** Let  $\mathbf{X} = \{X_k\}_k$  be an IID such that  $I_2(\mathbf{X}) < \infty$ , and let  $\mathbf{Y} = \{Y_k\}_k$  be an independent random sequence, which is independent of  $\mathbf{X}$ , such that each  $Y_k$  is a symmetric  $\alpha$ -stable random variable, where  $0 < \alpha \le 2$ . Let

$$\mathrm{E}[e^{itY_k}] = e^{-c_k + t + \alpha}$$
, where  $c_k \ge 0$ ,  $k \in \mathbb{N}$ .

Then  $\mu_{\mathbf{Y}}$  is expressed as  $\mu_{\mathbf{Y}} = \mu_{\mathbf{a}\mathbf{\theta}}$ , where  $\mathbf{E}[e^{it\theta}] = e^{-|t|^{\alpha}}$  and  $a_k = c_k^{\frac{1}{\alpha}}$ ,  $k \in \mathbb{N}$ . In addition, we have by Blumenthal and Getoor [1, Theorem 2.1],

$$0 < \liminf_{x \to \infty} x^{\alpha} P(\mid \Theta \mid >x) \le \sup_{x \ge 0} x^{\alpha} P(\mid \Theta \mid >x) < \infty.$$

Hence by Proposition 1,  $\sum_k |a_k|^a < \infty$  implies  $\mu_{x+y} \sim \mu_x$ , and by Theorem 1 (3),  $\mu_{x+y} \sim \mu_x$  implies  $\sum_k |a_k|^{2a} < \infty$ .

#### § 2. General Case

In this section we do not assume that **Y** is similar. We first give preliminaries and then prove a variation of Theorem 3 of Sato and Watari [8]. A general characterization of  $\mu_{x+y} \sim \mu_x$  has been given by Kitada and Sato [4, Theorem 2] as follows.

**Lemma 1** (Kitada and Sato [4]). Assume  $\mu_{X_k+Y_k} \sim \mu_{X_k}$  for every  $k \in \mathbb{N}$ , and define

$$Z_k(x) = \frac{d\mu_{X_k + Y_k}}{d\mu_{X_k}}(x) - 1, \qquad k \in \mathbb{N}.$$

Then  $\mu_{x+y} \sim \mu_x$  holds if and only if the following hold:

$$\sum_{k} \mathbb{E}[Z_{k}(X_{k}) : Z_{k}(X_{k}) \geq 1] < \infty,$$
  
$$\sum_{k} \mathbb{E}[Z_{k}(X_{k})^{2} : |Z_{k}(X_{k})| < 1] < \infty.$$

This is a necessary and sufficient condition, but  $Z_k(x)$  depends on the distribution of  $X_1$  and is not always easily estimated. Starting from Lemma 1, Hino [2, Theorem 1.8] proved certain conditions for the relation  $\mu_{X+Y} \sim \mu_X$  as follows. His conditions are described in terms only of the distribution of Y, but they are necessary or sufficient conditions.

**Lemma 2** (Hino [2]). (1) If  $\mu_{X+Y} \sim \mu_X$ , then we have for every  $\varepsilon > 0$ ,  $\sum_k P(|Y_k| > \varepsilon)^2 + \sum_k E[Y_k : |Y_k| \le \varepsilon]^2 + \sum_k E[Y_k^2 : |Y_k| \le \varepsilon]^2 < \infty$ .

(2) Assume  $I_2(\mathbf{X}) < \infty$ . If there exists  $\varepsilon > 0$  such that

$$\sum_{k} P(\mid Y_{k}\mid >_{\mathcal{E}}) + \sum_{k} E[Y_{k}:\mid Y_{k}\mid \leq_{\mathcal{E}}]^{2} + \sum_{k} E[Y_{k}^{4}:\mid Y_{k}\mid \leq_{\mathcal{E}}] < \infty,$$

then we have  $\mu_{x+y} \sim \mu_x$ .

(3) If there exists  $\varepsilon > 0$  such that

$$\sup_{|z| < \varepsilon} (\varepsilon - |z|)^2 \int_{-\infty}^{+\infty} \frac{f''(x+z)^2}{f(x)} dx < \infty$$

and

$$\sum_{k} \mathrm{P}(\mid Y_{k}\mid > \varepsilon) + \sum_{k} \mathrm{E}[Y_{k} : \mid Y_{k}\mid < \varepsilon]^{2} + \sum_{k} \mathrm{E}[Y_{k}^{2} : \mid Y_{k}\mid \leq \varepsilon]^{2} < \infty,$$

then we have  $\mu_{x+y} \sim \mu_x$ .

Applying Lemma 2, we have the following theorem.

**Theorem 5.** If  $I_2(\mathbf{X}) < \infty$ ,  $E[Y_k] = 0$ ,  $\sup_k E[Y_k^2] < \infty$  and  $\sum_k Y_k^4 < \infty$  a. s., then we have  $\mu_{\mathbf{X}+\mathbf{Y}} \sim \mu_{\mathbf{X}}$ .

*Proof.* Since  $\sum_k Y_k^4 < \infty$  a.s., we have by Kolmogorov's three series theorem,

$$\sum_{k} P(\mid Y_{k}\mid >1) < \infty$$
 and  $\sum_{k} E[Y_{k}^{4}:\mid Y_{k}\mid \leq 1] < \infty$ .

Since  $E[Y_k] = 0$ , we have

$$E[Y_k: | Y_k | \leq 1] = -E[Y_k: | Y_k | > 1].$$

and thus, by the Schwarz inequality,

$$\sum_{k} E[Y_{k}: | Y_{k}| \le 1]^{2} = \sum_{k} E[Y_{k}: | Y_{k}| > 1]^{2} \le \sum_{k} E[Y_{k}^{2}] P(| Y_{k}| > 1)$$

$$\le \left(\sup_{k} E[Y_{k}^{2}]\right) \sum_{k} P(| Y_{k}| > 1) < \infty.$$

Hence by Lemma 2 (2), we have  $\mu_{x+y} \sim \mu_x$ .

Corollary 3.  $I_2(\mathbf{X}) < \infty$ ,  $\Sigma_k \mathbb{E}[Y_k^2] < \infty$  and  $\mathbb{E}[Y_k] = 0$  together imply  $\mu_{\mathbf{X}+\mathbf{Y}} \sim \mu_{\mathbf{X}}$ .

Sato and Watari [8, Theorem 3] proved that  $\mu_{x+y} \sim \mu_x$  holds if  $I_2(\mathbf{X}) < \infty$ ,  $\sum_k Y_k^4 < \infty$  a. s. and each  $Y_k$  is symmetric. We assume  $\mathrm{E}[Y_k] = 0$  and  $\sup_k \mathrm{E}[Y_k^2] < \infty$  instead of assuming the symmetry of  $Y_k$ . Then the case p=4 of Theorem 2 (2) is a special case of Theorem 5. In fact,  $\mathrm{E}[\Theta^4] < \infty$  and  $\sum_k a_k^4 < \infty$  together imply  $\sum_k \mathrm{E}[a_k^4 \Theta_k^4] < \infty$ , so that we have  $\sum_k a_k^4 \Theta_k^4 < \infty$  a. s. and  $\sup_k \mathrm{E}[a_k^2 \Theta_k^2] < \infty$ .

## § 3. Proofs

*Proof of Theorem 1.* (1) Since  $P(\Theta \neq 0) > 0$ , there exists K > 0 such that

$$P(0 < |\Theta| \le K) > 0$$
 and  $P(|\Theta| \ge K) > 0$ .

Then by Lemma 2 (1), we have

$$\infty > \sum_{a_{k} \neq 0} P(|\Theta| > \frac{1}{|a_{k}|})^{2} \ge \sum_{|a_{k}|K>1} P(|\Theta| > \frac{1}{|a_{k}|})^{2}$$

$$\ge \sum_{|a_{k}|K>1} P(|\Theta| \ge K)^{2},$$

so that  $|a_k|K>1$  holds for finitely many  $k \in \mathbb{N}$ . Hence there exists  $k_0 \in \mathbb{N}$  such that

$$|a_k|K \leq 1$$
 for  $k \geq k_0$ .

We therefore have, by Lemma 2 (1),

$$\begin{split} & \infty > \sum_{k \geq k_0} \mathrm{E}[Y_k^2 : \mid Y_k \mid \leq 1]^2 \\ & = \sum_{\substack{k \geq k_0 \\ a_k \neq 0}} a_k^4 \, \mathrm{E}\Big[\Theta_k^2 : \mid \Theta_k \mid \leq \frac{1}{a_k}\Big]^2 \geq \sum_{k \geq k_0} a_k^4 \, \mathrm{E}[\Theta^2 : \mid \Theta \mid \leq K]^2, \end{split}$$

so that  $\sum_{k} a_{k}^{4} < \infty$ .

(2) Let  $m = \frac{1}{2} \lim \inf_{x \to \infty} | E[\Theta : | \Theta | \le x] | > 0$ . In the case that  $| \Theta | \le M$  a. s. for some M > 0, we have  $m = \frac{1}{2} | E[\Theta] |$ , so that by Lemma 2 (1), we have

$$\infty > \sum_{k} E[Y_{k}: | Y_{k} | \leq M]^{2} = \sum_{k} E[Y_{k}]^{2} = \sum_{k} a_{k}^{2} E[\Theta]^{2} = 4m^{2} \sum_{k} a_{k}^{2}.$$

In the case that  $P(|\Theta| > x) > 0$  for all x > 0, there exists L > 0 such that

$$| E[\Theta : |\Theta| \le x] | \ge m \text{ for } x \ge L.$$

By Lemma 2 (1), we have

$$\infty > \sum_{k} P(\mid Y_{k} \mid > 1)^{2} = \sum_{a_{k} \neq 0} P\left(\mid \Theta \mid > \frac{1}{\mid a_{k} \mid}\right)^{2}$$

$$\geq \sum_{\mid a_{k} \mid L > 1} P\left(\mid \Theta \mid > \frac{1}{\mid a_{k} \mid}\right)^{2} \geq \sum_{\mid a_{k} \mid L > 1} P(\mid \Theta \mid > L)^{2}.$$

Then since  $P(|\Theta|>L)>0$ ,  $|a_k|L>1$  holds for finitely many  $k\in\mathbb{N}$ , so that there exists  $n_0\in\mathbb{N}$  such that  $|a_k|L\leq 1$  for  $k\geq n_0$ . We thus have

$$\left| E\left[\Theta: \mid \Theta \mid \leq \frac{1}{\mid a_k \mid} \right] \right| \geq m \text{ for } k \geq n_0 \text{ with } a_k \neq 0.$$

Therefore by Lemma 2 (1), we have

$$\infty > \sum_{k} \mathbb{E}[Y_{k}: | Y_{k} | \leq 1]^{2} \geq \sum_{\substack{k \geq n_{0} \\ a_{k} \neq 0}} a_{k}^{2} \mathbb{E}\left[\Theta_{k}: | \Theta_{k} | \leq \frac{1}{|a_{k}|}\right]^{2}$$
$$\geq m^{2} \sum_{k \geq n} a_{k}^{2}.$$

(3) Let  $r = \frac{1}{2} \liminf_{x \to \infty} x^p P(|\Theta| > x) > 0$ . Then there exists L > 0 such that  $x^p P(|\Theta| > x) \ge r$  for  $x \ge L$ .

Since  $P(|\Theta| > L) > 0$ , as in the proof of (1), we know there exists  $n_0 \in \mathbb{N}$  such that

$$\mid a_k \mid L \leq 1 \text{ for } k \geq n_0.$$

Hence by Lemma 2 (1), we have

$$\infty > \sum_{k \geq n_0} P(\mid Y_k \mid > 1)^2 \geq \sum_{\substack{k \geq n_0 \\ a_k \neq 0}} P\left(\mid \Theta \mid > \frac{1}{\mid a_k \mid}\right)^2 \geq r^2 \sum_{k \geq n_0} \mid a_k \mid ^{2p},$$

so that  $\sum_{k} |a_{k}|^{2p} < \infty$ .

Proof of Theorem 2. (1) Since  $\mathbb{E}\left[\sum_{k}a_{k}^{2}\Theta_{k}^{2}\right] = \mathbb{E}\left[\Theta^{2}\right]\sum_{k}a_{k}^{2} < \infty$ , we have  $\sum_{k}a_{k}^{2}\Theta_{k}^{2} < \infty$  a. s., that is,  $\mu_{ae}(l_{2}) = 1$ , and since **X** and  $\mathbf{a}\Theta$  are independent, we therefore have  $\mu_{\mathbf{X}+ae} = \mu_{\mathbf{X}} * \mu_{ae}$ . It follows that

$$\mu_{\mathbf{X}+\mathbf{a}\mathbf{\Theta}}(A) = \int_{l_2} \mu_{\mathbf{X}+\mathbf{y}}(A) d\mu_{\mathbf{a}\mathbf{\Theta}}(\mathbf{y})$$

for every Borel set A in  $\mathbf{R}^{\mathbf{N}}$ . On the other hand, since  $I_1(\mathbf{X}) < \infty$ , we have by Theorem A,  $\mu_{\mathbf{X}+\mathbf{y}} \sim \mu_{\mathbf{X}}$  for every  $\mathbf{y} \in l_2$ . If  $\mu_{\mathbf{X}+\mathbf{a}\mathbf{\theta}}(A) = 0$ , then  $\mu_{\mathbf{X}+\mathbf{y}}(A) = 0$  for some  $\mathbf{y} \in l_2$ , so that  $\mu_{\mathbf{X}}(A) = 0$ . Conversely, if  $\mu_{\mathbf{X}}(A) = 0$ , then  $\mu_{\mathbf{X}+\mathbf{y}}(A) = 0$  for all  $\mathbf{y} \in l_2$ , so that  $\mu_{\mathbf{X}+\mathbf{a}\mathbf{\theta}}(A) = 0$ . We therefore have  $\mu_{\mathbf{X}+\mathbf{a}\mathbf{\theta}} \sim \mu_{\mathbf{X}}$ .

(2) Since  $E[ | \Theta |^p] < \infty$ , we have  $M = \sup_x |x|^p P( | \Theta | >x) < \infty$ . Then for every  $k \in \mathbb{N}$ ,

$$P(\mid Y_k \mid >1) = P(\mid \Theta \mid > \frac{1}{\mid a_k \mid}) \leq M \mid a_k \mid p,$$

$$\begin{split} \mathrm{E}[Y_k:\mid Y_k\mid \leq 1]^2 &= a_k^2 \mathrm{E}\Big[\Theta:\mid \Theta\mid \leq \frac{1}{\mid a_k\mid}\Big]^2 = a_k^2 \mathrm{E}\Big[\Theta:\mid \Theta\mid > \frac{1}{\mid a_k\mid}\Big]^2 \\ &\leq a_k^2 \mathrm{E}[\Theta^2] \mathrm{P}\Big(\mid \Theta\mid > \frac{1}{\mid a_k\mid}\Big) \\ &\leq a_k^2 \mathrm{E}[\Theta^2] a_k^2 \mathrm{E}[\Theta^2] = a_k^4 \mathrm{E}[\Theta^2]^2 \end{split}$$

and 
$$\mathrm{E}[Y_k^4:\mid Y_k\mid \leq 1] \leq \mathrm{E}[\mid Y_k\mid^{p\wedge 4}:\mid Y_k\mid \leq 1]$$

$$= \mid a_k\mid^{p\wedge 4}\mathrm{E}\Big[\mid \Theta\mid^{p\wedge 4}:\mid \Theta\mid \leq \frac{1}{\mid a_k\mid}\Big]$$

$$\leq \mid a_k\mid^{p\wedge 4}\mathrm{E}[\mid \Theta\mid^{p\wedge 4}].$$

Hence by Lemma 2 (2),  $\sum_{k} |a_{k}|^{p \wedge 4} < \infty$  implies  $\mu_{X+ae} \sim \mu_{X}$ .

*Proof of Theorem 3.* (1) By Corollary 1,  $\mu_{X+ae} \sim \mu_X$  implies  $\sum_k a_k^2 < \infty$ , and by Theorem 2 (1),  $\sum_k a_k^2 < \infty$  implies  $\mu_{X+ae} \sim \mu_X$ .

(2) By Theorem 1 (1) and Theorem 2 (2), we have  $\mu_{X+a\theta} \sim \mu_X$  if and only if  $\sum_k a_k^4 < \infty$ .

*Proof of Proposition 1.* Let  $M = \sup_{x \ge 0} x^p P(|\Theta| > x) < \infty$ . Then we have for 0 ,

$$P(\mid Y_{k}\mid >1) = P(\mid \Theta\mid > \frac{1}{\mid a_{k}\mid}) \leq M \mid a_{k}\mid^{p}$$
and  $E[Y_{k}^{4}:\mid Y_{k}\mid \leq 1] = a_{k}^{4}E\Big[\Theta^{4}:\mid \Theta\mid \leq \frac{1}{\mid a_{k}\mid}\Big] = -a_{k}^{4}\int_{0}^{\frac{1}{\mid a_{k}\mid}}x^{4}dP(\mid \Theta\mid > x)$ 

$$\leq 4a_{k}^{4}\int_{0}^{\frac{1}{\mid a_{k}\mid}}x^{3}P(\mid \Theta\mid > x)dx \leq 4Ma_{k}^{4}\int_{0}^{\frac{1}{\mid a_{k}\mid}}x^{3-p}dx$$

$$= \frac{4M}{4-p}\mid a_{k}\mid^{p}.$$

Therefore in (1) and (2), it is sufficient by Lemma 2 (2) to prove that  $\sum_{k} |a_{k}|^{p} < \infty$  implies  $\sum_{k} \mathrm{E}[Y_{k}: |Y_{k}| \leq 1]^{2} < \infty$ .

Assume  $\sum_{k} |a_{k}|^{p \wedge 4} < \infty$ . Then there exists  $k_{0} \in \mathbb{N}$  such that

$$|a_k| \leq 1$$
 for  $k \geq k_0$ .

We thus have for  $k \ge k_0$ ,

$$\begin{split} | & \operatorname{E}[Y_k: | Y_k | \leq 1] | \leq | a_k | \operatorname{E}[ | \Theta | : | \Theta | \leq \frac{1}{|a_k|} ] \\ & \leq | a_k | + | a_k | \operatorname{E}[ | \Theta | : 1 < | \Theta | \leq \frac{1}{|a_k|} ] \\ & \leq | a_k | + | a_k | \int_{1}^{\frac{1}{|a_k|}} \operatorname{P}( | \Theta | > x ) dx \\ & \leq | a_k | + M | a_k | \int_{1}^{\frac{1}{|a_k|}} x^{-p} dx. \end{split}$$

(1) If  $0 , then for <math>k \ge k_0$ ,

$$| E[Y_{k}: | Y_{k} | \leq 1] | \leq | a_{k} | + M | a_{k} | \int_{1}^{\frac{1}{|a_{k}|}} x^{-p} dx$$

$$= \left(1 - \frac{M}{1 - p}\right) | a_{k} | + \frac{M}{1 - p} | a_{k} |^{p}.$$

Therefore  $\sum_{k} |a_{k}|^{p} < \infty$  implies  $\sum_{k} \mathbb{E}[Y_{k}: |Y_{k}| \leq 1]^{2} < \infty$ . If p=1, then for  $k \geq k_{0}$ ,

$$| E[Y_k: | Y_k | \le 1] | \le | a_k | + M | a_k | \int_1^{\frac{1}{|a_k|}} \frac{1}{x} dx$$

$$= | a_k | (1+M | \log | a_k | | | ).$$

Therefore  $\sum_{k} |a_{k}| < \infty$  implies  $\sum_{k} \mathbb{E}[Y_{k}: |Y_{k}| \leq 1]^{2} < \infty$ . If  $1 , then for <math>k \geq k_{0}$ ,

$$| E[Y_k: | Y_k | \le 1] | \le | a_k | + M | a_k | \int_1^{\frac{1}{|a_k|}} x^{-p} dx \le \frac{M+1}{p-1} | a_k |.$$

Therefore  $\sum_{k} |a_{k}|^{p} < \infty$  implies  $\sum_{k} \mathbb{E}[Y_{k}: |Y_{k}| \leq 1]^{2} < \infty$ .

(2) Since

$$\mathbb{E}\Big[\Theta: \mid \Theta \mid \leq \frac{1}{\mid a_k \mid}\Big] = -\mathbb{E}\Big[\Theta: \mid \Theta \mid > \frac{1}{\mid a_k \mid}\Big] \quad \text{and} \quad \mathbb{E}\Big[\Theta^2\Big] < \infty,$$

the following holds:

$$E[Y_k: | Y_k| \le 1]^2 = a_k^2 E\left[\Theta: | \Theta| \le \frac{1}{|a_k|}\right]^2 = a_k^2 E\left[\Theta: | \Theta| > \frac{1}{|a_k|}\right]^2 \\
\le a_k^2 E\left[\Theta^2\right] P\left(| \Theta| > \frac{1}{|a_k|}\right) \\
\le a_k^2 E\left[\Theta^2\right] a_k^2 E\left[\Theta^2\right] = a_k^4 E\left[\Theta^2\right]^2,$$

so that  $\sum_{k} |a_{k}|^{p} < \infty$  implies  $\sum_{k} \mathbb{E}[Y_{k}: |Y_{k}| \leq 1]^{2} < \infty$ .

(3) We have

$$P(\mid Y_{k}\mid >_{\varepsilon}) = P\left(\mid \Theta\mid > \frac{\varepsilon}{\mid a_{k}\mid}\right) \leq \frac{M}{\varepsilon^{4}} a_{k}^{4},$$

and since

$$\mathrm{E}\Big[\Theta: \mid \Theta \mid \leq \frac{\varepsilon}{\mid a_{k}\mid}\Big] = -\mathrm{E}\Big[\Theta: \mid \Theta \mid > \frac{\varepsilon}{\mid a_{k}\mid}\Big] \quad \text{and} \quad \mathrm{E}[\Theta^{2}] < \infty,$$

if follows that

$$\begin{split} \mathrm{E}[Y_{\mathbf{k}}:\mid Y_{\mathbf{k}}\mid \leq & \epsilon]^{2} = a_{\mathbf{k}}^{2}\mathrm{E}\Big[\Theta:\mid \Theta\mid \leq \frac{\varepsilon}{\mid a_{\mathbf{k}}\mid}\Big]^{2} = a_{\mathbf{k}}^{2}\mathrm{E}\Big[\Theta:\mid \Theta\mid > \frac{\varepsilon}{\mid a_{\mathbf{k}}\mid}\Big]^{2} \\ \leq & a_{\mathbf{k}}^{2}\mathrm{E}\left[\Theta^{2}\right]\mathrm{P}\Big(\mid \Theta\mid > \frac{\varepsilon}{\mid a_{\mathbf{k}}\mid}\Big) \\ \leq & a_{\mathbf{k}}^{2}\mathrm{E}\left[\Theta^{2}\right]\frac{a_{\mathbf{k}}^{2}}{\varepsilon^{2}}\;\mathrm{E}\left[\Theta^{2}\right] = a_{\mathbf{k}}^{4}\;\frac{\mathrm{E}\left[\Theta^{2}\right]^{2}}{\varepsilon^{2}}\;\text{,} \end{split}$$

and 
$$\mathbb{E}[Y_k^2: | Y_k | \leq \varepsilon]^2 = a_k^4 \mathbb{E}\Big[\Theta_k^2: | \Theta_k | \leq \frac{\varepsilon}{|a_k|}\Big]^2 \leq a_k^4 \mathbb{E}[\Theta^2]^2$$
.

Hence by Lemma 2 (3),  $\sum_{k} a_{k}^{4} < \infty$  implies  $\mu_{X+a\theta} \sim \mu_{X}$ .

(4) If p > 4, then we have  $E[\Theta^4] < \infty$ . This case is proved in Theorem 2 (2).

Proof of Theorem 4. (1) Assume  $\limsup_{x\to\infty} |\mathbb{E}[\Theta: |\Theta| \le x]| = \infty$ . Then for  $T(x) = \mathbb{E}\Big[\Theta: |\Theta| \le \frac{1}{x}\Big]^2$ , we have  $\limsup_{x\to 0} T(x) = \infty$ , so that there exists, by Shepp [9, Lemma 4], a sequence  $\mathbf{a} = \{a_k\}_k$  such that

$$\sum_{k} a_k^2 < \infty \quad \text{and} \quad \sum_{k} \mathbb{E}[Y_k : |Y_k| \le 1]^2 = \sum_{a_k \ne 0} a_k^2 \mathbb{E}[\Theta : |\Theta| \le \frac{1}{|a_k|}]^2 = \infty.$$

Therefore by Lemma 2 (1), we have  $\mu_{X+s\theta} \nsim \mu_X$ .

(2) We first prove  $E[\Theta^2] < \infty$ . If  $E[\Theta^2] = \infty$ , it follows that  $\lim_{x\to +0} E[\Theta: \Theta: \Theta: \le \frac{1}{\sqrt{x}}]^2 = \infty$ . Hence by Shepp [9, Lemma 4], there exists a sequence  $\mathbf{a} = \{a_k\}_k$  such that

$$\sum_{k} a_k^4 < \infty \quad \text{and} \quad \sum_{k} \mathbb{E}[Y_k^2 : |Y_k| \le 1]^2 = \sum_{a_k \ne 0} a_k^4 \mathbb{E}\Big[\Theta_k^2 : |\Theta_k| \le \frac{1}{\sqrt{a_k^2}}\Big]^2 = \infty.$$

Therefore by Lemma 2 (1), we have  $\mu_{X+a\theta} \not\sim \mu_X$ .

Next we prove  $E[\Theta] = 0$ . If  $E[\Theta] \neq 0$ , then  $\lim_{x\to\infty} |E[\Theta: |\Theta| \leq x]| = |E[\Theta]| > 0$ . Hence by Theorem 1 (2),  $\mu_{X+a\Theta} \sim \mu_X$  implies  $\sum_k a_k^2 < \infty$ . Then  $\sum_k a_k^4 < \infty$  implies  $\mu_{X+a\Theta} \sim \mu_X$ , and  $\mu_{X+a\Theta} \sim \mu_X$  implies  $\sum_k a_k^2 < \infty$ , so that  $\sum_k a_k^4 < \infty$  implies  $\sum_k a_k^2 < \infty$ . This is a contradiction.

It is therefore shown that  $E[\Theta^2] < \infty$  and  $E[\Theta] = 0$ .

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#### References

- [1] R. M. Blumenthal and R. K. Getoor, Some theorems on stable processes, Trans. Amer. Math. Soc., 95 (1960), 263 273.
- [2] M. Hino, On equivalence of product measures by random translation, J. Math. Kyoto Univ., 34-4 (1994), 755-765.
- [3] S. Kakutani, On equivalence of infinite product measures, Ann. Math., 49 (1948), 214 224.
- [4] K. Kitada and H. Sato, On the absolute continuity of infinite product measure and its convolution, Probab. Th. Rel. Fields, 81 (1989), 609 627.
- [5] Y. Okazaki, On equivalence of product measure by symmetric random l<sub>4</sub>-translation, J. Funct. Anal., 115 (1993), 100 103.
- [6] Y. Okazaki and H. Sato, Distinguishing a random sequence from a random translate of itself, Ann. Probab., 22 (1994), 1092 – 1096.
- [7] H. Sato, Infinite products and infinite sum, in Stochastic analysis on infinite dimensional spaces, 289 296, Longman Scientific & Technical, 1994.
- [8] H. Sato and C. Watari, Some integral inequalities and absolutely continuity of a symmetric random translation, J. Funct. Anal., 114-1 (1993), 257-266.
- [9] L. A. Shepp, Distinguishing a sequence of random variables from a translate of itself, Ann. Math. Stat., 36 (1965), 1107 - 1112.