

Asymptotic Expansion of the Distribution of the Generalized Variance in the Non-central Case

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1. Introduction and Summary

The generalized variance (the determinant of the sample variance and covariance matrix) was defined by Wilks [8] as a measure of the spread of observations. In this paper we study asymptotic expansion of the distribution of the generalized variance in the non-central case. In general, if the rows of a $n \times p$ matrix \mathbf{X} are independently normally distributed with common covariance matrix Σ and mean $E[\mathbf{X}] = \mathbf{M}$, then the generalized variance is defined as the determinant of a matrix $\mathbf{S} = (1/n)\mathbf{X}'\mathbf{X}$. Asymptotic expansion of the distribution of $|\mathbf{S}|$ depends on the order of the non-centrality matrix $\Sigma^{-\frac{1}{2}}\mathbf{M}'\mathbf{M}\Sigma^{-\frac{1}{2}} = \Omega$ with respect to n . It is in general true that $\Omega = \mathbf{O}(1)$ or $\Omega = \mathbf{O}(n)$, which means that all elements of Ω are $\mathbf{O}(1)$ or $\mathbf{O}(n)$ as $n \rightarrow \infty$.

In section 2 we derive the limiting distribution of $|\mathbf{S}|$ under the assumption that $\Omega = n\theta_n = \mathbf{O}(n)$ and $\lim_{n \rightarrow \infty} \sqrt{n}(\theta_n - \theta) = \mathbf{0}$. If Ω may be regarded as a constant matrix, asymptotic expansion of the distribution of $|\mathbf{S}|$ is obtained up to the order $n^{-\frac{3}{2}}$ by inverting the characteristic function expressed in terms of hypergeometric function with matrix argument (see section 3).

2. Limiting distribution of $|\mathbf{S}|$ when $\Omega = \mathbf{O}(n)$

In this section we assume that $\Omega = n\theta_n = \mathbf{O}(n)$ and $\lim_{n \rightarrow \infty} \sqrt{n}(\theta_n - \theta) = \mathbf{0}$. At first we shall consider limiting distribution of a function of the non-central Wishart matrix $\mathbf{X}'\mathbf{X}$. Let $C_{\mathbf{X}'\mathbf{X}}(\mathbf{T})$ be the characteristic function of $\mathbf{X}'\mathbf{X}$, where \mathbf{T} is the $p \times p$ symmetric matrix having $\{(1 + \delta_{ij})/2\}t_{ij}$ as its (i, j) element with Kronecker delta δ_{ij} . From the result of Anderson [1] $C_{\mathbf{X}'\mathbf{X}}(\mathbf{T})$ can be expressed by our notation as

$$(2.1) \quad C_{\mathbf{X}'\mathbf{X}}(\mathbf{T}) = |\mathbf{I} - 2i\Sigma^{\frac{1}{2}}\mathbf{T}\Sigma^{\frac{1}{2}}|^{-(n/2)} \operatorname{etr} \left\{ -\frac{1}{2}\Omega + \frac{1}{2}\Omega^{\frac{1}{2}}(\mathbf{I} - 2i\Sigma^{\frac{1}{2}}\mathbf{T}\Sigma^{\frac{1}{2}})^{-1}\Omega^{\frac{1}{2}} \right\}.$$

Put $\mathbf{S}^* = \sqrt{n} \{ \Sigma^{-\frac{1}{2}}\mathbf{S}\Sigma^{-\frac{1}{2}} - (\mathbf{I} + \theta) \}$. Then we can express the characteristic function of \mathbf{S}^* as

$$(2.2) \quad C_S(\mathbf{T}) = \left| \mathbf{I} - \frac{2i}{\sqrt{n}} \mathbf{T} \right|^{-(n/2)} \text{etr} \left\{ -\frac{n}{2} \boldsymbol{\theta}_n + \frac{n}{2} \left(\mathbf{I} - \frac{2i}{\sqrt{n}} \mathbf{T} \right)^{-1} \boldsymbol{\theta}_n - \sqrt{n} i \mathbf{T} (\mathbf{I} + \boldsymbol{\theta}) \right\}.$$

$C_S(\mathbf{T})$ can be expanded by using the well known asymptotic formulas

$$(2.3) \quad \left| \mathbf{I} - \frac{2i}{\sqrt{n}} \mathbf{T} \right|^{-(n/2)} = \exp \left\{ -\frac{n}{2} \log \left| \mathbf{I} - \frac{2i}{\sqrt{n}} \mathbf{T} \right| \right\} \\ = \text{etr} \{ \sqrt{n} i \mathbf{T} - \mathbf{T}^2 \} \{ 1 + o(n^{-\frac{1}{2}}) \}$$

$$(2.4) \quad \left(\mathbf{I} - \frac{2i}{\sqrt{n}} \mathbf{T} \right)^{-1} = \mathbf{I} + \frac{2i}{\sqrt{n}} \mathbf{T} + \left(\frac{2i}{\sqrt{n}} \mathbf{T} \right)^2 + o(n^{-\frac{3}{2}}),$$

which hold for large n such that the maximum of the absolute values of the characteristic roots of $(2i/\sqrt{n})\mathbf{T}$ is less than unity. Applying the formulas (2.3) and (2.4) to the expression of $C_S(\mathbf{T})$ in (2.2), we get

$$(2.5) \quad C_S(\mathbf{T}) = \text{etr} \{ \sqrt{n} i \mathbf{T} (\boldsymbol{\theta}_n - \boldsymbol{\theta}) \} \text{etr} \{ -\mathbf{T}^2 (\mathbf{I} + 2\boldsymbol{\theta}_n) \} \{ 1 + o(n^{-\frac{1}{2}}) \}.$$

Therefore we have

$$(2.6) \quad \lim_{n \rightarrow \infty} C_S(\mathbf{T}) = \text{etr} \{ -\mathbf{T}^2 (\mathbf{I} + 2\boldsymbol{\theta}) \}.$$

(2.6) shows that the limiting distribution of $\mathbf{S}^* = (s_{ij}^*)$ is the multivariate normal distribution with mean zero and covariances $\mathbf{E}[s_{ij}^* s_{kl}^*] = q_{ijkl}$, where q_{ijkl} is defined by

$$(2.7) \quad 2 \text{tr} \mathbf{T}^2 (\mathbf{I} + 2\boldsymbol{\theta}) = \sum_{i \leq j} \sum_{k \leq l} q_{ijkl} t_{ij} t_{kl}.$$

Now we will generalize the well known result for obtaining limiting distributions of statistics (for example, Theorem 4.2.5 in Anderson [2] and Siotani and Hayakawa [6]) to the non-central case.

LEMMA. Let $n\mathbf{S}$ have the non-central Wishart distribution with n degrees of freedom and the non-centrality matrix $\boldsymbol{\Omega}$ such that $\boldsymbol{\Omega} = n\boldsymbol{\theta}_n = \mathbf{O}(n)$ and $\lim_{n \rightarrow \infty} \sqrt{n}(\boldsymbol{\theta}_n - \boldsymbol{\theta}) = \mathbf{0}$. Suppose $f(\mathbf{W})$ is a real valued function of a $p \times p$ symmetric matrix \mathbf{W} with first and second derivatives existing in a neighborhood of $\mathbf{W} = \mathbf{I} + \boldsymbol{\theta}$. Then the statistic $\sqrt{n} \{ f(\boldsymbol{\Sigma}^{-\frac{1}{2}} \mathbf{S} \boldsymbol{\Sigma}^{-\frac{1}{2}}) - f(\mathbf{I} + \boldsymbol{\theta}) \}$ has asymptotically the normal distribution with mean zero and variance $2 \text{tr} \mathbf{F}^2 (\mathbf{I} + 2\boldsymbol{\theta})$, where $\mathbf{F} = (\{ (1 + \delta_{ij})/2 \} f_{ij})$ and $f_{ij} = \partial f(\mathbf{W}) / \partial w_{ij} |_{\mathbf{W} = \mathbf{I} + \boldsymbol{\theta}}$.

PROOF. From (2.6) and Theorem 4.2.5 in Anderson [2] we can see that the asymptotic distribution of $\sqrt{n} \{ f(\boldsymbol{\Sigma}^{-\frac{1}{2}} \mathbf{S} \boldsymbol{\Sigma}^{-\frac{1}{2}}) - f(\mathbf{I} + \boldsymbol{\theta}) \}$ is normal with mean zero. By using (2.7) its asymptotic variance can be expressed as

$$\sum_{i \leq j} \sum_{k \leq l} \frac{\partial f(\mathbf{W})}{\partial w_{ij}} \bigg|_{\mathbf{W}=\mathbf{I}+\boldsymbol{\theta}} \frac{\partial f(\mathbf{W})}{\partial w_{kl}} \bigg|_{\mathbf{W}=\mathbf{I}+\boldsymbol{\theta}} q_{ijkl} = 2tr\mathbf{F}^2(\mathbf{I}+2\boldsymbol{\theta}).$$

Putting $f(\mathbf{W})=|\mathbf{W}|$ in the above lemma and noting that the equality $(\{1+\delta_{ij}\}/2)\partial|\mathbf{W}|/\partial w_{ij}=|\mathbf{W}|\mathbf{W}^{-1}$ holds for any symmetric matrix \mathbf{W} , we have the following theorem.

THEOREM 1. *Let $n\mathbf{S}$ have the non-central Wishart distribution with n degrees of freedom and the non-centrality matrix $\boldsymbol{\Omega}$ such that $\boldsymbol{\Omega}=n\boldsymbol{\theta}_n=\mathbf{O}(n)$ and $\lim_{n \rightarrow \infty} \sqrt{n}(\boldsymbol{\theta}_n-\boldsymbol{\theta})=\mathbf{0}$. Then the distribution of $\sqrt{n}\{(|\mathbf{S}|/|\boldsymbol{\Sigma}|-|\mathbf{I}+\boldsymbol{\theta}|)\}$ is asymptotically normal with mean zero and variance $2|\mathbf{I}+\boldsymbol{\theta}|^2tr(\mathbf{I}+2\boldsymbol{\theta})(\mathbf{I}+\boldsymbol{\theta})^{-2}$.*

3. Asymptotic expansion of the distribution of $|\mathbf{S}|$ when $\boldsymbol{\Omega}=\mathbf{0}(1)$

In this section we shall obtain asymptotic expansion of the distribution of $|\mathbf{S}|$ under the assumption that the non-centrality matrix $\boldsymbol{\Omega}$ is a constant matrix. Constantine [3] showed that the h th moment of $|\mathbf{S}|$ in the non-central case could be expressed by the hypergeometric function of matrix argument. His result can be expressed by our notation as

$$(3.1) \quad \mathbf{E}[|\mathbf{S}|^h] = |\boldsymbol{\Sigma}|^h \left(\frac{2}{n}\right)^{ph} \frac{\Gamma_p\left(\frac{n}{2}+h\right)}{\Gamma_p\left(\frac{n}{2}\right)} {}_1F_1\left(-h; \frac{n}{2}; -\frac{1}{2}\boldsymbol{\Omega}\right),$$

where $\Gamma_p(a)$ and the hypergeometric function ${}_1F_1$ are defined by

$$(3.2) \quad \Gamma_p(a) = \pi^{p(p-1)/4} \prod_{\alpha=1}^p \Gamma(a-(\alpha-1)/2)$$

$${}_1F_1(a; b; Z) = \sum_{k=0}^{\infty} \sum_{(\kappa)} \frac{(a)_{\kappa}}{(b)_{\kappa}} \frac{C_{\kappa}(Z)}{k!}$$

$$(a)_{\kappa} = \prod_{\alpha=1}^p (a-(\alpha-1)/2)(a+1-(\alpha-1)/2)\dots(a+k_{\alpha}-1-(\alpha-1)/2).$$

The function $C_{\kappa}(Z)$ is a zonal polynomial of the $p \times p$ symmetric matrix Z corresponding to the partition $\kappa=(k_1, k_2, \dots, k_p)$, with $k_1+\dots+k_p=k$, $k_1 \geq \dots \geq k_p \geq 0$. The symbol $\sum_{(\kappa)}$ means the sum of all such partitions.

Put $\hat{\lambda}=\sqrt{n} \log(|\mathbf{S}|/|\boldsymbol{\Sigma}|)$. We can express the characteristic function of $\hat{\lambda}$ as

$$(3.3) \quad C(t) = \mathbf{E}[e^{it\hat{\lambda}}] = \mathbf{E}[(|\mathbf{S}|/|\boldsymbol{\Sigma}|)^{it\sqrt{n}}]$$

$$= \left(\frac{2}{n}\right)^{it\sqrt{n}} \frac{\Gamma_p\left(\frac{n}{2} + it\sqrt{n}\right)}{\Gamma_p\left(\frac{n}{2}\right)} {}_1F_1\left(-it\sqrt{n}; \frac{n}{2}; -\frac{1}{2}\Omega\right).$$

Applying the Kummer transformation formula ${}_1F_1(a; b; Z) = (\text{etr } Z) \cdot {}_1F_1(b-a; b; -Z)$ (see Herz [4]) to this last expression, we can write $C(t)$ as

$$(3.4) \quad \left(\frac{2}{n}\right)^{it\sqrt{n}} \frac{\Gamma_p\left(\frac{n}{2} + it\sqrt{n}\right)}{\Gamma_p\left(\frac{n}{2}\right)} \text{etr}\left(-\frac{1}{2}\Omega\right) {}_1F_1\left(\frac{n}{2} + it\sqrt{n}; \frac{n}{2}; \frac{1}{2}\Omega\right) \\ = C_1(t)C_2(t).$$

In the case that the non-centrality matrix Ω is equal to zero, $\text{etr}\left(-\frac{1}{2}\Omega\right) \cdot {}_1F_1\left(\frac{n}{2} + it\sqrt{n}; \frac{n}{2}; \frac{1}{2}\Omega\right)$ which we shall denote by $C_2(t)$ is equal to unity. So $(2/n)^{it\sqrt{n}} \Gamma_p\left(\frac{n}{2} + it\sqrt{n}\right) / \Gamma_p\left(\frac{n}{2}\right)$ gives us the characteristic function of $\hat{\lambda}$ in the central case, which we shall denote by $C_1(t)$. We shall use the following asymptotic formula for the gamma function as in Anderson ([2], p. 204).

$$(3.5) \quad \log \Gamma(x+h) = \log \sqrt{2\pi} + \left(x+h-\frac{1}{2}\right) \log x - x - \sum_{r=1}^m \frac{(-1)^r B_{r+1}(h)}{r(r+1)x^r} \\ + O(|x|^{-m-1})$$

which holds for large $|x|$ and fixed h with the Bernoulli polynomial $B_r(h)$ of degree r , $B_2(h) = h^2 - h + (1/6)$, $B_3(h) = h^3 - (3/2)h^2 + (1/2)h$, etc. Applying the formula (3.5) to each gamma function in $C_1(t)$, we get

$$(3.6) \quad \log C_1(t) = -pt^2 - \frac{1}{\sqrt{n}} \left\{ qit + \frac{2}{3} p(it)^3 \right\} + \frac{1}{n} \left\{ q(it)^2 + \frac{2}{3} p(it)^4 \right\} \\ - \frac{1}{n\sqrt{n}} \left\{ \frac{1}{12} p(2p^2 + 3p - 1)(it) + \frac{4}{3} q(it)^3 + \frac{4}{5} p(it)^5 \right\} + O(n^{-2}),$$

where $q = p(p+1)/2$. This formula implies the asymptotic expansion of $C_1(t)$.

$$(3.7) \quad C_1(t) = e^{-pt^2} \left[1 - \frac{1}{\sqrt{n}} \left\{ qit + \frac{2}{3} p(it)^3 \right\} + \frac{1}{n} \left\{ \frac{1}{2} q(q+2)(it)^2 \right. \right. \\ \left. \left. + \frac{2}{3} p(q+1)(it)^4 + \frac{2}{9} p^2(it)^6 \right\} - \frac{1}{n\sqrt{n}} \left\{ \frac{1}{12} p(2p^2 + 3p - 1)(it) \right. \right. \\ \left. \left. + \frac{1}{6} q(q+2)(q+4)(it)^3 + \frac{1}{15} p(5q^2 + 20q + 12)(it)^5 \right. \right]$$

$$+ \frac{2}{9} p^2(q+2)(it)^7 + \frac{4}{81} p^3(it)^9 \Big\} + 0(n^{-2}) \Big].$$

Now we shall consider the second term $C_2(t)$ of (3.4). From definition (3.2) we have

$$(3.8) \quad C_2(t) = \text{etr} \left(-\frac{1}{2} \boldsymbol{\Omega} \right) \sum_{k=0}^{\infty} \sum_{(\kappa)} \frac{\left(\frac{n}{2} + it\sqrt{n} \right)_{\kappa} C_{\kappa} \left(\frac{1}{2} \boldsymbol{\Omega} \right)}{\left(\frac{n}{2} \right)_{\kappa} k!}.$$

The coefficient of each term can be arranged according to the descending order of powers of n as

$$(3.9) \quad \left(\frac{n}{2} + it\sqrt{n} \right)_{\kappa} = \left(\frac{n}{2} \right)^k \left[1 + \frac{2}{\sqrt{n}} itk + \frac{1}{n} \left\{ \sum_{\alpha=1}^p k_{\alpha}(k_{\alpha} - \alpha) + 2(it)^2 k(k-1) \right\} \right. \\ \left. + \frac{2}{n\sqrt{n}} \left\{ it(k-1) \sum_{\alpha=1}^p k_{\alpha}(k_{\alpha} - \alpha) + \frac{2}{3} (it)^3 k(k-1)(k-2) \right\} + 0(n^{-2}) \right]$$

$$(3.10) \quad \left(\frac{n}{2} \right)_{\kappa} = \left(\frac{n}{2} \right)^k \left\{ 1 + \frac{1}{n} \sum_{\alpha=1}^p k_{\alpha}(k_{\alpha} - \alpha) + 0(n^{-2}) \right\}.$$

Hence we can write $C_2(t)$ as

$$(3.11) \quad C_2(t) = \text{etr} \left(-\frac{1}{2} \boldsymbol{\Omega} \right) \sum_{k=0}^{\infty} \sum_{(\kappa)} \frac{C_{\kappa} \left(\frac{1}{2} \boldsymbol{\Omega} \right)}{k!} \left[1 + \frac{2}{\sqrt{n}} itk + \frac{2}{n} (it)^2 k(k-1) \right. \\ \left. - \frac{2}{n\sqrt{n}} \left\{ it \sum_{\alpha=1}^p k_{\alpha}(k_{\alpha} - \alpha) - \frac{2}{3} (it)^3 k(k-1)(k-2) \right\} + 0(n^{-2}) \right].$$

Now we shall evaluate each term of the above infinite series. Since the identity $(trZ)^k = \sum_{(\kappa)} C_{\kappa}(Z)$ (see James [5]) holds for any symmetric matrix Z ,

We have

$$(3.12) \quad \sum_{k=0}^{\infty} \sum_{(\kappa)} \frac{C_{\kappa}(Z)}{k!} k(k-1)\dots(k-r+1) \\ = \sum_{k=r}^{\infty} \sum_{(k)} \frac{C_{\kappa}(Z)}{(k-r)!} \\ = \sum_{k=r}^{\infty} \frac{(trZ)^k}{(k-r)!} = (trZ)^r \text{etr} Z,$$

which holds for any non-negative integer r . Sugiura [7] proved the following formula.

$$(3.13) \quad \sum_{k=0}^{\infty} \sum_{(\kappa)} \frac{C_{\kappa}(Z)}{k!} \left\{ \sum_{\alpha=1}^p k_{\alpha}(k_{\alpha} - \alpha) \right\} = \text{tr} Z^2 \text{etr} Z.$$

Applying the formula (3.12) and (3.13) to the expression of $C_2(t)$ in (3.11), we can simplify the expression (3.11) for the function $C_2(t)$ as

$$(3.14) \quad C_2(t) = 1 + \frac{it}{\sqrt{n}} \text{tr} \mathbf{\Omega} + \frac{(it)^2}{2n} (\text{tr} \mathbf{\Omega})^2 - \frac{1}{6n\sqrt{n}} \{3(it)\text{tr} \mathbf{\Omega}^2 - (it)^3(\text{tr} \mathbf{\Omega})^3\} + 0(n^{-2}).$$

Combining this result with the expression for $C_1(t)$ in (3.7), we obtain the following asymptotic expansion of the characteristic function $C(t)$.

$$(3.15) \quad C\left(\frac{t}{\sqrt{2p}}\right) = e^{-\frac{t^2}{2}} \left\{ 1 - n^{-\frac{1}{2}} A_1 + n^{-1} A_2 - n^{-\frac{3}{2}} A_3 + 0(n^{-2}) \right\},$$

where the coefficients A_1 , A_2 and A_3 are given by

$$\begin{aligned} A_1 &= \frac{1}{3\sqrt{2p}} \{3it(q - \text{tr} \mathbf{\Omega}) + (it)^3\} \\ A_2 &= \frac{1}{36p} \{9(it)^2[q(q+2) - 2q\text{tr} \mathbf{\Omega} + (\text{tr} \mathbf{\Omega})^2] + 6(it)^4[q+1 - \text{tr} \mathbf{\Omega}] + (it)^6\} \\ A_3 &= \frac{1}{180\sqrt{2pp}} \{15it[p^2(2p^2 + 3p - 1) + 6p\text{tr} \mathbf{\Omega}^2] + 15(it)^3[q(q+2)(q+4) \\ &\quad - 3q(q+2)\text{tr} \mathbf{\Omega} + 3q(\text{tr} \mathbf{\Omega})^2 - (\text{tr} \mathbf{\Omega})^3] + 3(it)^5[5q^2 + 20q + 12 \\ &\quad - 10(q+1)\text{tr} \mathbf{\Omega} + 5(\text{tr} \mathbf{\Omega})^2] + 5(it)^7[q+2 - \text{tr} \mathbf{\Omega}] + (5/9)(it)^9\}. \end{aligned}$$

By inverting this characteristic function, we can finally obtain the following theorem.

THEOREM 2. *Let $n\mathbf{S}$ have the non-central Wishart distribution with n degrees of freedom and the non-centrality matrix $\mathbf{\Omega} = \mathbf{\Sigma}^{-\frac{1}{2}} \mathbf{M} \mathbf{M} \mathbf{\Sigma}^{-\frac{1}{2}}$. Assume that the non-centrality matrix $\mathbf{\Omega}$ may be regarded as a constant matrix with respect to n . Then the asymptotic expansion of the distribution of $|\mathbf{S}|$ can be obtained up to the order $n^{-\frac{3}{2}}$ in the following way. Let $\lambda = (\sqrt{n}/\sqrt{2p}) \log(|\mathbf{S}|/|\mathbf{\Sigma}|)$. Then we have*

$$(3.16) \quad \begin{aligned} P(\lambda \leq x) &= \Phi(x) + \frac{1}{3\sqrt{2pn}} \{3\Phi'(x)(q - \text{tr} \mathbf{\Omega}) + \Phi^{(3)}(x)\} \\ &\quad + \frac{1}{36pn} \{9\Phi^{(2)}(x)[q(q+2) - 2q\text{tr} \mathbf{\Omega} + (\text{tr} \mathbf{\Omega})^2] + 6\Phi^{(4)}(x)[q+1 - \text{tr} \mathbf{\Omega}] + \Phi^{(6)}(x)\} \\ &\quad + \frac{1}{180\sqrt{2pp}\sqrt{nn}} \{15\Phi'(x)[p^2(2p^2 + 3p - 1) + 6p\text{tr} \mathbf{\Omega}^2] + 15\Phi^{(3)}(x)[q(q+2)(q+4) \\ &\quad - 3q(q+2)\text{tr} \mathbf{\Omega} + 3q(\text{tr} \mathbf{\Omega})^2 - (\text{tr} \mathbf{\Omega})^3] + 3\Phi^{(5)}(x)[5q^2 + 20q + 12 - 10(q+1)\text{tr} \mathbf{\Omega} \\ &\quad + 5(\text{tr} \mathbf{\Omega})^2] + 5\Phi^{(7)}(x)[q+2 - \text{tr} \mathbf{\Omega}] + (5/9)\Phi^{(9)}(x)\} + 0(n^{-2}), \end{aligned}$$

where $q = p(p+1)/2$ and $\Phi^{(r)}(x)$ denotes the r th derivative of the standard normal distribution function $\Phi(x)$.

COROLLARY 1. *If the non-centrality matrix Ω is the null matrix, λ can be expanded asymptotically as*

$$(3.17) \quad P(\lambda \leq x) = \Phi(x) + \frac{1}{3\sqrt{2pn}} \{3q\Phi'(x) + \Phi^{(3)}(x)\} + \frac{1}{36pn} \{9\Phi^{(2)}(x)q(q+2) \\ + 6\Phi^{(4)}(x)(q+1) + \Phi^{(6)}(x)\} + \frac{1}{180\sqrt{2p}p\sqrt{n}} \{15\Phi'(x)p^2(2p^2+3p-1) \\ + 15\Phi^{(3)}(x)q(q+2)(q+4) + 3\Phi^{(5)}(x)(5q^2+20q+12) + 5\Phi^{(7)}(x)(q+2) \\ + (5/9)\Phi^{(9)}(x)\} + O(n^{-2}).$$

This corollary will be obtained at once by putting $\Omega = \mathbf{0}$ in (3.16). The asymptotic expansion (3.16) may be useful not only in the case of $\Omega = \mathbf{0}(1)$, but also in the case of $\Omega = \mathbf{0}(n)$. However, we could not succeed in deriving it.

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