

CORRECTION

COMPOUND POISSON APPROXIMATIONS FOR SUMS OF RANDOM VARIABLES

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The proof of Theorem 1.1 contains the erroneous statement that

$$P\left(\sum_{i=1}^n X_i \in B\right) = E[F_1' * \cdots * F_n'(B)].$$

The following modifications to Theorem 1.1 and Lemma 2.2 will correct this error. To simplify the notation, we let $d(X|\mathcal{F}, Y|\mathcal{G})$ denote the random total-variation distance $d(P(X \in \cdot|\mathcal{F}), P(Y \in \cdot|\mathcal{G}))$.

(a) For the statement of Theorem 1.1, let $\zeta_i = I(X_i \neq 0)$ and redefine F_i as

$$F_i(B) = P(X_i \in B|\mathcal{F}_{i-1}, X_i \neq 0, \zeta_k, k \neq i).$$

Then in the bounds (1.3), (1.4), (1.5), replace each d_i by $\zeta_i d_i$. The error was that the original F_i did not contain the additional conditioning on $\zeta_k, k \neq i$, which is needed in the proof.

(b) Replace assertion (iv) of Lemma 2.2 by the following:

(iv) If X_1, \dots, X_n are adapted to the increasing σ -fields $\{\mathcal{F}_i\}_{i=0}^n$ and Y_1, \dots, Y_n are independent with respective distributions G_1, \dots, G_n , then

$$(2.4) \quad D \leq E\left[\sum_{i=1}^n d(X_i|\mathcal{F}_{i-1}, Y_i)\right].$$

(There is no conditioning on the Y_i .)

(c) Replace the last paragraph of the proof of Lemma 2.2 by the following:

We prove statement (iv) by induction. It is true for $n = 1$ since $d(X_1, Y_1) \leq E[d(X_1|\mathcal{F}_0, Y_1)]$ by Lemma 2.1. Now assume it is true for $n - 1$. Let $X^* = (X_1, \dots, X_{n-1}, Y_n)$, where Y_n is independent of X_1, \dots, X_{n-1} . Then by (i) and Lemma 2.1, we have

$$\begin{aligned} D &\leq d(X, Y) \leq d(X, X^*) + d(X^*, Y) \leq E[d(X|\mathcal{F}_{n-1}, X^*|\mathcal{F}_{n-1})] + d(X^*, Y) \\ &= E[d(X_n|\mathcal{F}_{n-1}, Y_n)] + d((X_1, \dots, X_{n-1}), (Y_1, \dots, Y_{n-1})), \end{aligned}$$

which equals the right-hand side of (2.4).

(d) In the proof of Theorem 1.1, replace the five lines after expression (3.1) with the following:

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We can write $Y = \sum_{i=1}^n \zeta_i Y_i$, where Y_1, \dots, Y_n are independent with the common distribution F and are independent of ζ_1, \dots, ζ_n . Let \mathcal{H} denote the σ -field generated by ζ_1, \dots, ζ_n . Then applying (2.2) and Lemma (2.2)(iv) with $X_i, Y_i, P(\cdot)$ equal to $\zeta_i X_i, \zeta_i Y_i, P(\cdot|\mathcal{H})$, respectively, we have

$$\begin{aligned}
 d\left(\sum_{i=1}^n X_i, Y\right) &= d\left(\sum_{i=1}^n \zeta_i X_i, \sum_{i=1}^n \zeta_i Y_i\right) \\
 &\leq E\left[d\left(\sum_{i=1}^n \zeta_i X_i|\mathcal{H}, \sum_{i=1}^n \zeta_i Y_i|\mathcal{H}\right)\right] \\
 (3.2) \quad &\leq E\left\{\sum_{i=1}^n E[d(\zeta_i X_i|(\mathcal{F}_{i-1}, \mathcal{H}), \zeta_i Y_i|\mathcal{H})|\mathcal{H}]\right\} \\
 &= E\left(\sum_{i=1}^n \zeta_i d_i\right).
 \end{aligned}$$

To obtain the last inequality, first note that the standard properties of conditional probabilities allow us to write

$$\begin{aligned}
 P(\zeta_i X_i \in B|\mathcal{F}_{i-1}, \mathcal{H}) \\
 = P(X_i \in B|\mathcal{F}_{i-1}, \zeta_i = 1, \zeta_k = 0, k \neq i)I(\zeta_i = 1) + I(0 \in B)I(\zeta_i = 0).
 \end{aligned}$$

Then using this and a similar expression for $P(\zeta_i Y_i \in B|\mathcal{H})$, it follows from the definition of d that

$$d(\zeta_i X_i|(\mathcal{F}_{i-1}, \mathcal{H}), \zeta_i Y_i|\mathcal{H}) = I(\zeta_i = 1)d_i = \zeta_i d_i.$$

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