AN OPTIMAL PROPERTY OF PRINCIPAL COMPONENTS

By J. N. DARROCH

University of Michigan

1. Introduction and summary. Let $x' = (x_1, x_2, \dots, x_p)$ be a random vector and let E[x] = 0, $E[xx'] = \Sigma = (\sigma_{ij})$ where we assume that Σ is non-singular. Further let

$$\Sigma = T\Lambda T' = (t_1, t_2, \cdots, t_p) \begin{bmatrix} \lambda_1 & 0 & \cdot & \cdot & \cdot & 0 \\ 0 & \lambda_2 & \cdot & \cdot & \cdot & 0 \\ \vdots & \vdots & & \vdots & & \vdots \\ 0 & 0 & \cdot & \cdot & \cdot & \lambda_p \end{bmatrix} \begin{bmatrix} t'_1 \\ t'_2 \\ \vdots \\ t'_p \end{bmatrix}$$

where TT'=I and where we suppose that the eigenvalues of Σ are in order of decreasing magnitude, that is $\lambda_1 \geq \lambda_2 \geq \cdots \geq \lambda_p > 0$. The principal components of x, namely $u_1 = t_1'x$, $u_2 = t_2'x$, \cdots , $u_p = t_p'x$ were introduced by Hotelling (1933) who characterised them by certain optimal properties. Since then Girshick (1936), Anderson (1958) and Kullback (1959) have characterised the principal components by slightly different sets of optimal properties. Thus Anderson shows that u_1 is the linear function $\alpha_1'x$ having maximum variance subject to $\alpha_1'\alpha_1 = 1$; u_2 is the linear function $\alpha_2'x$ which is uncorrelated with u_1 and has maximum variance subject to $\alpha_2'\alpha_2 = 1$; and so on.

The above mentioned characterisations have two properties in common; they introduce the principal components one by one and, more importantly, the optimal properties hold only with-in the class of linear functions of x_1 , x_2 , \cdots , x_p .

In the following theorem the first k principal components are characterized by an optimal property within the class of all random variables.

2. The optimal property. Before stating the main result we note that, since

$$TT' = t_1t_1' + t_2t_2' + \cdots + t_nt_n' = I,$$

therefore

$$x = t_1 u_1 + t_2 u_2 + \cdots + t_n u_n.$$

THEOREM. Let A be any $p \times k$ matrix and let $f' = (f_1, f_2, \dots, f_k)$ be any random vector. Then

$$5_1 = \text{trace } E[(x - Af)(x - Af)'] = E[\sum_{i=1}^{p} (x_i - (Af)_i)^2]$$

is minimized with respect to A and f when and only when

$$Af = t_1u_1 + t_2u_2 + \cdots + t_ku_k,$$

and the minimum value of \mathfrak{I}_1 is $\sum_{i=k+1}^{p} \lambda_i$.

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PROOF. Without loss of generality suppose that E[f'] = I and let E[xf'] = B. Then B must satisfy the condition

since the matrix in (1) is the covariance matrix of $(x', f') = (x_1, x_2, \dots, x_p, f_1, f_2, \dots, f_k)$. Now, if Γ_{22} is positive definite, a necessary and sufficient condition for

$$egin{bmatrix} \Gamma_{11} & \Gamma_{12} \ \Gamma_{21} & \Gamma_{22} \end{bmatrix}$$

to be non-negative definite is that $\Gamma_{11} - \Gamma_{12}\Gamma_{22}^{-1}\Gamma_{21}$ is non-negative definite. Therefore condition (1) is equivalent to

(2)
$$\Sigma - BB'$$
 is non-negative definite.

Now

$$E[x - Af)(x - Af)']$$

= $\Sigma - AB' - BA' + AA' = \Sigma - BB' + (A - B)(A - B)'.$

Since tr $(A - B)(A - B)' \ge 0$ with equality if and only if A = B, it follows that the minimum of \mathfrak{I}_1 with respect to A is $\mathfrak{I}_2 = \text{tr }(\Sigma - BB')$. As in Section 1 write $\Sigma = T\Lambda T'$, and define

$$C = \Lambda^{-\frac{1}{2}} T' B.$$

Then

$$egin{aligned} \mathfrak{I}_2 &=& \operatorname{tr} \; (T\Lambda T' - T\Lambda^{rac{1}{2}}CC'\Lambda^{rac{1}{2}}T') \ &=& \operatorname{tr} \; (T'T(\Lambda - \Lambda^{rac{1}{2}}CC'\Lambda^{rac{1}{2}})) \ &=& \operatorname{tr} \; (\Lambda - \Lambda^{rac{1}{2}}CC'\Lambda^{rac{1}{2}}) \ &=& \operatorname{tr} \; \Lambda(I - CC'). \end{aligned}$$

Since C is a $p \times k$ matrix, CC' is at most of rank k. Therefore we can find P such that CC' = PDP', PP' = I, where D is a diagonal matrix of the form

We can now write

$$\mathfrak{I}_{2} = \operatorname{tr} \Lambda (I - PDP')
= \sum_{i=1}^{p} \lambda_{i} - \sum_{i=1}^{p} \lambda_{i} \sum_{j=1}^{k} p_{ij}^{2} dj.$$

Condition (2) has meanwhile become

(4)
$$I - D$$
 is non-negative definite.

Therefore \mathfrak{I}_2 is minimised subject to (4) by choosing $d_1 = \cdots = d_k = 1$. It remains to minimise

$$\mathfrak{I}_3 = \sum_{i=1}^{p} \lambda_i - \sum_{i=1}^{p} \lambda_i w_i$$

with respect to w_1 , w_2 , \cdots , w_p where

$$w_i = \sum_{j=1}^k p_{ij}^2$$
.

Because P is orthogonal,

(5)
$$0 \le w_i \le 1, \qquad \sum_{i=1}^p w_i = \sum_{i=1}^p \sum_{j=1}^k p_{ij}^2 = k.$$

Therefore \mathfrak{I}_3 is minimised with respect to (5) by choosing

(6)
$$w_1 = \cdots = w_k = 1, \quad w_{k+1} = \cdots = w_n = 0$$

and the minimum value of \mathfrak{I}_3 is therefore $\sum_{i=k+1}^p \lambda_i$. Equations (6) are equivalent to

$$P = \begin{bmatrix} P_{11} & 0 \\ 0 & P_{22} \end{bmatrix}$$

where P_{11} is $k \times k$.

Retracing our steps we find that 32 is minimised when

$$C = \begin{bmatrix} Q \\ 0 \end{bmatrix}$$

where Q is any orthogonal $k \times k$ matrix. Therefore \mathfrak{I}_2 is minimised when

(7)
$$B = T\Lambda^{\frac{1}{2}} \begin{bmatrix} Q \\ 0 \end{bmatrix} = G \text{ say.}$$

So far we have shown that, if g is a random vector with the properties that E[xg'] = G, E[gg'] = I, then the minimum value of \mathfrak{I}_1 , namely $\sum_{i=k+1}^{p} \lambda_i$, is attained by taking Af equal to Gg.

Now let v = Hx where H is a $k \times p$ matrix defined by E[xv'] = G. Thus

$$(8) H = G' \Sigma^{-1}$$

and we see that

$$\begin{split} E[gg'] &= H\Sigma H' = G'\Sigma^{-1}\Sigma\Sigma^{-1}G \\ &= (Q'0)\Lambda^{\frac{1}{2}}T'\Sigma^{-1}T\Lambda^{\frac{1}{2}}\begin{bmatrix} Q \\ 0 \end{bmatrix} \\ &= (Q'0)\Lambda^{\frac{1}{2}}\Lambda^{-1}\Lambda^{\frac{1}{2}}\begin{bmatrix} Q \\ 0 \end{bmatrix} \\ &= I. \end{split}$$

Thus v satisfies all the conditions on g. Moreover, neglecting differences which have zero probability measure, it is the *only* vector to do so.

For

$$E[(v - g)(v - g)'] = E[(Hx - g)(Hx - g)']$$

$$= H\Sigma H' - HG - G'H' + I$$

$$= 0$$

by (7) and (8).

Finally, it is easily verified that

$$v = T \begin{bmatrix} I & 0 \\ 0 & 0 \end{bmatrix} T'x$$
$$= t_1(t'_1 x) + \cdots + t_k(t'_k x).$$

Thus 31 is minimised uniquely by taking

$$Af = t_1u_1 + \cdots + t_ku_k$$

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