ON THE A PRIORI DISTRIBUTION OF THE COVARIANCE MATRIX¹

By C. VILLEGAS

Universidad de la República, Montevideo, and The University of Rochester

The purpose of this note is to give a fiducial argument in support of the well known non-informative a priori distribution of a covariance matrix. Although the fiducial argument was introduced and promoted energetically by R. A. Fisher as an inference method which presumably would lead to necessary (and therefore unique) a posteriori distributions [4], it was discovered soon that different fiducial arguments could lead to entirely different a posteriori distributions [2], and, in the particular case of fiducial inferences about a covariance matrix, interesting problems arise which require further research [3].

Let X be a random $p \times p$ matrix, which has the multivariate normal distribution with mean value equal to 0 (the null matrix) and covariance matrix Σ . That is, the density of X is proportional to

(1)
$$|\Sigma|^{-p/2} \exp{-\frac{1}{2} \operatorname{tr} \Sigma^{-1} X X'}$$
.

Consider the change of variable

$$(2) Y = \Sigma^{-\frac{1}{2}} X.$$

As is well known, the random matrix Y has the multivariate normal distribution with mean value equal to 0 and covariance matrix I (the identity matrix). Obviously, the transpose Y' has the same distribution, and

$$(3) W = Y'Y = X'\Sigma^{-1}X$$

has the Wishart distribution with a mean value equal to pI, covariance matrix I, and p degrees of freedom, the density of which is proportional to

$$|W|^{-\frac{1}{2}} \exp - \frac{1}{2} \operatorname{tr} W.$$

Suppose now that X is an observed matrix and that Σ is unknown. Since the distribution of W does not depend on Σ , it is natural to assume that the *a posteriori* distribution of W is the same Wishart distribution. A posteriori, X and X' are constant matrices, and (3) is a transformation between the symmetric matrices W and Σ^{-1} , the Jacobian of which is, according to Theorem 3.7 of [1],

$$(5) dW/d\Sigma^{-1} = |X|^{p+1}.$$

Hence the a posteriori density of Σ^{-1} is, up to a constant factor,

(6)
$$|\Sigma|^{\frac{1}{2}} \exp - \frac{1}{2} \operatorname{tr} \Sigma^{-1} X X'.$$

Received 8 July 1968; revised 9 December 1968.

¹ This work has been done with special support from the Universidad de la República, Montevideo, and it was revised while the author was visiting at the University of Rochester.

A simple comparison between (1) and (6) shows that the *a priori* density of Σ^{-1} is

(7)
$$\pi(\Sigma^{-1}) = |\Sigma|^{(p+1)/2}.$$

This a priori distribution was derived from an invariance argument by Jeffreys (1961) for the case p=2; it was considered, for arbitrary values of p, by Geisser and Cornfield (1963) and it was used by Tiao and Zellner (1964) and by Geisser (1965) to develop a Bayesian multivariate theory. The same a posteriori distribution has been obtained recently by Fraser and Haq (see Fraser (1968)) using an interesting new approach.

REFERENCES

- [1] DEEMER, WALTER L. and OLKIN, INGRAM (1951). The Jacobians of certain matrix transformations useful in multivariate analysis (based on lectures of P. L. Hsu). Biometrika 38 345-367.
- [2] Dempster, A. P. (1963). Further examples of inconsistencies in the fiducial argument. Ann. Math. Statist. 34 884-891.
- [3] Dempster, A. P. (1963). On a paradox concerning inference about a covariance matrix. Ann. Math. Statist. 34 1414-1418.
- [4] Fisher, R. A. (1958). Statistical Methods and Scientific Inference (second edition). Hafner, New York.
- [5] Fraser, D. A. S. (1968). The Structure of Inference. Wiley, New York.
- [6] Jeffreys, Harold (1961). Theory of Probability. Clarendon Press, Oxford.
- [7] Geisser, Seymour (1965). Bayesian estimation in multivariate analysis. Ann. Math. Statist. 36 150-159.
- [8] Geisser, Seymour and Cornfield, Jerome (1963). Posterior distributions for multivariate normal parameters. J. Roy. Statist. Soc. Ser. B 25 368-376.
- [9] Tiao, G. C. and Zellner, A. (1964). On the Bayesian estimation of multivariate regression. J. Roy. Statist. Soc. Ser. B 26 277-285.