

WEIGHTED NORM INEQUALITY FOR OPERATOR ON MARTINGALES

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1. Introduction. Let \mathcal{M} be a family of martingales on a probability space (Ω, \mathcal{F}, P) . The norm inequalities for operators of matrix type on \mathcal{M} were obtained by Burkholder, Davis and Gundy [2] [3]. Our purpose in this paper is to prove a weighted norm inequality similar to that of Burkholder-Davis-Gundy. Throughout the paper, we fix a BMO-martingale $M_n = \sum_{k=1}^n m_k$ such that $1 + m_k \geq \varepsilon$ ($k \geq 1$) for some constant ε with $0 < \varepsilon \leq 1$. Then the process Z given by the formula $Z_n = \prod_{k=1}^n (1 + m_k)$ is a positive uniformly integrable martingale and the weighted probability measure $d\hat{P} = (Z_\infty/Z_1)dP$ is equivalent to dP (see [6]).

THEOREM. Let Φ be a non-decreasing continuous convex function on $[0, \infty[$ satisfying $\Phi(0) = 0$ and the growth condition $\Phi(2t) \leq C\Phi(t)$ for all $t \geq 0$. If U and V are two operators of matrix type on \mathcal{M} , then there exists a positive constant $C = C(U, V, \varepsilon, \Phi, M)$ such that the inequality

$$(1) \quad \hat{E}[\Phi(U^{**}(X))] \leq C\hat{E}[\Phi(V^*(X))]$$

holds for all $X \in \mathcal{M}$, where $\hat{E}[\]$ denotes the expectation over Ω with respect to $d\hat{P}$.

The result for the case $Z \equiv 1$ was established by Burkholder, Davis and Gundy [2, Theorem 2.3].

The following inequality was obtained in the continuous parameter case by Bonami and Lepingle [1] and Sekiguchi [9] independently.

COROLLARY. Let us denote the square function operator by $S(X)$ and the maximal operator by X^* . Then the inequality

$$(2) \quad c\hat{E}[\Phi(X^*)] \leq \hat{E}[\Phi(S(X))] \leq C\hat{E}[\Phi(X^*)]$$

is valid for all $X \in \mathcal{M}$.

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2. Preliminaries. The reader is assumed to be familiar with the martingale theory as is given in Meyer [7] and Neveu [8]. Throughout

the paper, let us denote by c or C a positive constant and by $C(p)$ a positive constant depending only on the parameter p . Both letters are not necessarily the same at each occurrence.

1) Notations. Let (Ω, F, P) be a probability space with a non-decreasing sequence $(F_n)_{n \geq 1}$ of sub σ -fields of F such that $\bigvee_{n=1}^{\infty} F_n = F$. Let $X = (X_n; n \geq 1)$ be an (F_n) -adapted process and (x_1, x_2, \dots) be the difference sequence of X so that $X_n = \sum_{k=1}^n x_k$.

A matrix $(u_{jk}; j \geq 1, k \geq 1)$ is said to be of type B-G (B-G stands for Burkholder and Gundy) if it has the following properties:

- (a) Each entry u_{jk} is an F_{k-1} -measurable random variable.
- (b) There is a constant $\alpha > 1$ such that for all $k \geq 1$,

$$(3) \quad 1/\alpha \leq \sum_{j=1}^{\infty} u_{jk}^2 \leq \alpha.$$

We define $U(X)$, $U_n(X)$, $U_n^*(X)$ and $U_n^{**}(X)$ for a matrix (u_{jk}) of type B-G as follows:

$$\begin{aligned} U(X) &= \left(\sum_{j=1}^{\infty} \limsup_{n \rightarrow \infty} \left| \sum_{k=1}^n u_{jk} x_k \right|^2 \right)^{1/2}, \\ U_n(X) &= \left(\sum_{j=1}^{\infty} \left| \sum_{k=1}^n u_{jk} x_k \right|^2 \right)^{1/2}, \\ U_n^*(X) &= \sup_{i \leq n} U_i(X) \end{aligned}$$

and

$$U_n^{**}(X) = \left(\sum_{j=1}^{\infty} \sup_{i \leq n} \left| \sum_{k=1}^i u_{jk} x_k \right|^2 \right)^{1/2}.$$

We write simply $U^*(X)$ and $U^{**}(X)$ instead of $U_{\infty}^*(X)$ and $U_{\infty}^{**}(X)$. $U(X)$ is called an operator of matrix type which was introduced by Burkholder and Gundy [3]. In the same way, for another matrix (v_{jk}) of type B-G, we can define $V(X)$, $V^*(X)$ and $V^{**}(X)$ by using v_{jk} instead of u_{jk} . Typical examples, corresponding to the identity matrix or a single-row matrix, are $S_n(X) = (\sum_{k=1}^n x_k^2)^{1/2}$, $S(X) = (\sum_{k=1}^{\infty} x_k^2)^{1/2}$, $X_n^* = \sup_{k \leq n} |X_k|$ and $X^* = \sup_k |X_k|$. Let us set $X_0 = U_0(X) = U_0^*(X) = 0$, $Z_0 = 1$ and $F_0 = F_1$ for convenience. Now we define \hat{X}_n and \hat{X} as follows:

$$\begin{aligned} \hat{x}_n &= -x_n(Z_{n-1}/Z_n) = -x_n/(1 + m_n), \\ \hat{X}_n &= \sum_{k=1}^n \hat{x}_k, \quad \hat{X} = (\hat{X}_n)_{n \geq 1}. \end{aligned}$$

In particular $\hat{m}_n = -m_n/(1 + m_n)$, $m_n = -\hat{m}_n/(1 + \hat{m}_n)$ and $(1 + m_n)(1 + \hat{m}_n) = 1$. So we obtain

$$(4) \quad \hat{x}_n = -x_n(1 + \hat{m}_n) = -x_n - x_n \hat{m}_n.$$

Let us denote by $\|X\|_{\text{BMO}}$ the smallest positive constant c such that c^2 dominates $E[S(X)^2 - S_{n-1}(X)^2 | F_n]$ P -a.s. for all $n \geq 1$. BMO is the class of those martingales X which satisfy $\|X\|_{\text{BMO}} < \infty$. We choose and fix a constant ε with $1/\varepsilon \geq 1 + \|M\|_{\text{BMO}} \geq 1 + m_n \geq \varepsilon > 0$. Then we get $\hat{M} \in \text{BMO}(\hat{P})$ with $1/\varepsilon \geq 1 + \hat{m}_n \geq \varepsilon$, where $\text{BMO}(\hat{P})$ is the BMO-class with respect to \hat{P} (see [4]). Since the equality

$$(5) \quad \hat{E}[Y | F_n] = E[Y(Z_\infty/Z_n) | F_n] \text{ a.s. under } P \text{ and } \hat{P}$$

holds for all \hat{P} -integrable random variable Y , it is easy to see that \hat{X} is a \hat{P} -martingale for each martingale X . By (4) we have

$$(6) \quad \varepsilon S(X) \leq S(\hat{X}) \leq (1/\varepsilon)S(X) \text{ a.s. .}$$

In this paper, unless otherwise stated, “a martingale” means “a martingale with respect to P ”.

2) Preliminary lemmas. To show our theorem, we need several lemmas.

LEMMA 1. Let (a_{jk}) be a matrix of type B-G. Then there is a positive constant $C(\alpha)$ such that the inequality

$$(7) \quad \hat{E}\left[\left(\sum_{j=1}^{\infty} \sup_n \left|\sum_{k=1}^n a_{jk} x_k \hat{y}_k\right|^2\right)^{1/2}\right] \leq C(\alpha) \|\hat{Y}\|_{\text{BMO}(\hat{P})} \hat{E}[S(X)]$$

is valid for all $\hat{Y} \in \text{BMO}(\hat{P})$, where (\hat{y}_k) is the difference sequence of \hat{Y} .

PROOF. Let us fix a positive integer N . For any $\delta > 0$,

$$\begin{aligned} & \hat{E}\left[\left(\sum_{j=1}^{\infty} \sup_{i \leq N} \left|\sum_{k=1}^i a_{jk} x_k \hat{y}_k\right|^2\right)^{1/2}\right] \\ &= \hat{E}\left[\left\{\sum_{j=1}^{\infty} \sup_{i \leq N} \left|\sum_{k=1}^i x_k (S_k(X) + \delta)^{-1/2} (S_k(X) + \delta)^{1/2} a_{jk} \hat{y}_k\right|^2\right\}^{1/2}\right] \\ &\leq \hat{E}\left[\left\{\sum_{j=1}^{\infty} \sup_{i \leq N} \left(\sum_{k=1}^i x_k^2 / (S_k(X) + \delta)\right) \left(\sum_{k=1}^i a_{jk}^2 (S_k(X) + \delta) \hat{y}_k^2\right)\right\}^{1/2}\right] \end{aligned}$$

by Cauchy-Schwarz's inequality. Moreover,

$$x_k^2 / (S_k(X) + \delta) = (S_k(X)^2 - S_{k-1}(X)^2) / (S_k(X) + \delta) \leq 2(S_k(X) - S_{k-1}(X)).$$

Therefore the left hand side of (7) is dominated by

$$\begin{aligned} & \sqrt{2} \hat{E}\left[\left\{\sum_{j=1}^{\infty} S_N(X) \left(\sum_{k=1}^N a_{jk}^2 (S_k(X) + \delta) \hat{y}_k^2\right)\right\}^{1/2}\right] \\ &= \sqrt{2} \hat{E}\left[S_N(X)^{1/2} \left\{\sum_{k=1}^N (S_k(X) + \delta) \hat{y}_k^2 \left(\sum_{j=1}^{\infty} a_{jk}^2\right)\right\}^{1/2}\right] \end{aligned}$$

$$\begin{aligned}
&\leq C(\alpha) \hat{E} \left[S_N(X)^{1/2} \left\{ \sum_{k=1}^N (S_k(X) + \delta) \hat{y}_k^2 \right\}^{1/2} \right] \quad (\text{by (3)}) \\
&\leq C(\alpha) \hat{E}[S_N(X)]^{1/2} \hat{E} \left[\sum_{k=1}^N (S_k(X) + \delta) \hat{y}_k^2 \right]^{1/2}
\end{aligned}$$

by Schwarz's inequality. The last factor of the above expression is equal to

$$\begin{aligned}
&\hat{E} \left[\sum_{k=1}^N (S_k(X) + \delta) (S_k(\hat{Y})^2 - S_{k-1}(\hat{Y})^2) \right]^{1/2} \\
&= \hat{E} \left[\sum_{k=1}^N (S_k(X) - S_{k-1}(X)) (S_N(\hat{Y})^2 - S_{k-1}(\hat{Y})^2) + \delta S_N(\hat{Y}) \right]^{1/2} \\
&= \hat{E} \left[\sum_{k=1}^N (S_k(X) - S_{k-1}(X)) \hat{E}[S_N(\hat{Y})^2 - S_{k-1}(\hat{Y})^2 | F_k] \right. \\
&\quad \left. + \delta \hat{E}[S_N(\hat{Y}) | F_0] \right]^{1/2} \leq \|\hat{Y}\|_{\text{BMO}(\hat{P})} \hat{E}[S_N(X) + \delta]^{1/2}.
\end{aligned}$$

Letting $\delta \rightarrow 0$ and then $N \rightarrow \infty$, we obtain (7).

LEMMA 2. *The inequality*

$$(8) \quad \hat{E} \left[\sup_n \left| \sum_{k=1}^n x_k \hat{y}_k \right| \right] \leq \sqrt{2} \|\hat{Y}\|_{\text{BMO}(\hat{P})} \hat{E}[S(X)]$$

holds for all $\hat{Y} \in \text{BMO}(\hat{P})$.

The inequality (8) is of Fefferman's type. The proof of Meyer [7, V-9, p. 337] is still valid in our case, where X is a semimartingale with respect to \hat{P} .

We are going to verify the inequality (1) in the case $\Phi(\lambda) = \lambda$.

LEMMA 3. *Let X be a martingale. Then*

$$(9) \quad \hat{E}[U^{**}(X)] \leq C(\alpha, \varepsilon, M) \hat{E}[S(X)].$$

PROOF. By (4) and Lemma 1, we have

$$\begin{aligned}
\hat{E}[U^{**}(X)] &\leq \hat{E}[U^{**}(\hat{X})] + \hat{E} \left[\left(\sum_{j=1}^{\infty} \sup_n \left| \sum_{k=1}^n u_{jk} x_k \hat{m}_k \right|^2 \right)^{1/2} \right] \\
&\leq \hat{E}[U^{**}(\hat{X})] + C(\alpha) \|\hat{M}\|_{\text{BMO}(\hat{P})} \hat{E}[S(X)].
\end{aligned}$$

Applying Burkholder-Davis-Gundy's theorem and (6), we have

$$\hat{E}[U^{**}(\hat{X})] \leq C(\alpha) \hat{E}[S(\hat{X})] \leq C(\alpha, \varepsilon) \hat{E}[S(X)].$$

Thus (9) holds.

Here we define A_k , d_k and $D(X) = (D_n)_{n \geq 1}$ as follows:

$$A_1 = 0, \quad A_k = \sum_{j=1}^{\infty} \left(\sum_{i=1}^{k-1} v_{ji} x_i \right) v_{jk}, \quad d_k = 2A_k x_k, \quad D_n = \sum_{k=1}^n d_k.$$

LEMMA 4. *There exists a positive constant $C(\alpha, \varepsilon, M)$ such that*

$$(10) \quad \hat{E}[S(X)^2/(V^*(X) + \delta)] \leq C(\alpha, \varepsilon, M)\{\hat{E}[V^*(X) + \delta] + \hat{E}[S(X)]\}$$

for every $\delta > 0$ and for every martingale X with $\hat{E}[D(X)^] < \infty$.*

PROOF. We define \hat{H} and \hat{G} as follows:

$$\hat{H}_n = \hat{E}[1/(V^*(X) + \delta) | F_n], \quad \hat{h}_n = \hat{H}_n - \hat{H}_{n-1}, \quad \hat{H} = (\hat{H}_n)_{n \geq 1},$$

$$\hat{G}_n = \sum_{k=1}^n A_k \hat{h}_k \quad \text{and} \quad \hat{G} = (\hat{G}_n)_{n \geq 1}.$$

Note that \hat{H} is a \hat{P} -martingale dominated by $1/\delta$. First we show $\hat{G} \in \text{BMO}(\hat{P})$. By Cauchy-Schwarz's inequality and (3),

$$|A_n| \leq \left(\sum_{j=1}^{\infty} \left| \sum_{k=1}^{n-1} v_{jk} x_k \right|^2 \right)^{1/2} \left(\sum_{j=1}^{\infty} v_{jn}^2 \right)^{1/2} \leq C(\alpha) V_{n-1}^*(X).$$

So we have

$$(11) \quad |A_n \hat{H}_{n-1}| \leq C(\alpha) \hat{E}[V_{n-1}^*(X)/(V^*(X) + \delta) | F_{n-1}] \leq C(\alpha)$$

and $|A_n \hat{H}_n| \leq C(\alpha)$. By using the above inequalities, we obtain

$$\begin{aligned} \hat{E}[S(\hat{G})^2 - S_{n-1}(\hat{G})^2 | F_n] &= \hat{E} \left[\sum_{k=n+1}^{\infty} A_k^2 \hat{h}_k^2 \middle| F_n \right] + A_n^2 (\hat{H}_n - \hat{H}_{n-1})^2 \\ &\leq C(\alpha) \hat{E} \left[\sum_{k=n+1}^{\infty} V_{k-1}^*(X)^2 \hat{h}_k^2 \middle| F_n \right] + 2(|A_n \hat{H}_n|^2 + |A_n \hat{H}_{n-1}|^2) \\ &\leq C(\alpha) \hat{E} \left[\sum_{k=n+1}^{\infty} V_{k-1}^*(X)^2 (\hat{H}_k^2 - \hat{H}_{k-1}^2) \middle| F_n \right] + C(\alpha) \\ &\leq C(\alpha) \hat{E}[V^*(X)^2 \hat{H}_{\infty}^2 | F_n] + C(\alpha) \leq C(\alpha) \end{aligned}$$

from which

$$(12) \quad \|\hat{G}\|_{\text{BMO}(\hat{P})} \leq C(\alpha)$$

follows. Secondly we modify $D_n \hat{H}_n$ as follows:

$$\begin{aligned} D_n \hat{H}_n &= \left(\sum_{k=1}^n d_k \right) \left(\sum_{k=1}^n \hat{h}_k \right) = \sum_{k=1}^n D_{k-1} \hat{h}_k + \sum_{k=1}^n \hat{H}_{k-1} d_k + \sum_{k=1}^n \hat{h}_k d_k \\ &= \sum_{k=1}^n D_{k-1} \hat{h}_k + 2 \sum_{k=1}^n \hat{H}_{k-1} A_k x_k + 2 \sum_{k=1}^n \hat{h}_k A_k x_k \\ &= \sum_{k=1}^n D_{k-1} \hat{h}_k - 2 \sum_{k=1}^n \hat{H}_{k-1} A_k \hat{x}_k - 2 \sum_{k=1}^n \hat{H}_{k-1} A_k x_k \hat{m}_k + 2 \sum_{k=1}^n x_k A_k \hat{h}_k. \end{aligned}$$

Here $\sum D_{k-1} \hat{h}_k$ and $\sum \hat{H}_{k-1} A_k \hat{x}_k$ are \hat{P} -local martingales with value 0 at time 1, so there is a non-decreasing sequence $\{R_n\}$ of stopping times such that $\lim_{n \rightarrow \infty} R_n = \infty$ a.s. and $\hat{E}[\sum_{k=1}^{R_n} D_{k-1} \hat{h}_k] = \hat{E}[\sum_{k=1}^{R_n} \hat{H}_{k-1} A_k \hat{x}_k] = 0$. Therefore we obtain

$$\begin{aligned}
|\hat{E}[D_{R_n}\hat{H}_{R_n}]| &\leq 2\hat{E}\left[\sum_{k=1}^{R_n}|\hat{H}_{k-1}A_k||x_k\hat{m}_k|\right] + 2\hat{E}\left[\sum_{k=1}^{R_n}|x_kA_k\hat{h}_k|\right] \\
&\leq C(\alpha)\hat{E}\left[\sum_{k=1}^{\infty}|x_k\hat{m}_k|\right] + 2\hat{E}\left[\sum_{k=1}^{\infty}|x_kA_k\hat{h}_k|\right] \quad (\text{by (11)}) \\
&\leq C(\alpha)\|\hat{M}\|_{\text{BMO}(\hat{P})}\hat{E}[S(X)] + 2\|\hat{G}\|_{\text{BMO}(\hat{P})}\hat{E}[S(X)] \quad (\text{by Lemma 2}) \\
&\leq C(\alpha, M)\hat{E}[S(X)] \quad (\text{by (12)}) .
\end{aligned}$$

Since $|D_n\hat{H}_n|$ is dominated by $(1/\delta)D(X)^*$, we get $|\hat{E}[D_\infty H_\infty]| \leq C\hat{E}[S(X)]$ by the dominated convergence theorem. On the other hand,

$$V_n(X)^2 - V_{n-1}(X)^2 - \sum_{j=1}^{\infty} v_{jn}^2 x_n^2 = 2 \sum_{j=1}^{\infty} \left(\sum_{k=1}^{n-1} v_{jk} x_k \right) v_{jn} x_n = d_n ,$$

so we find $\sum_{k=1}^{\infty} \sum_{j=1}^{\infty} v_{jk}^2 x_k^2 = V_\infty(X)^2 - D_\infty$. Then

$$\begin{aligned}
\hat{E}[S(X)^2/(V^*(X) + \delta)] &= \hat{E}\left[\left(\sum_{k=1}^{\infty} x_k^2\right)/(V^*(X) + \delta)\right] \\
&\leq C(\alpha)\hat{E}\left[\left(\sum_{k=1}^{\infty} \sum_{j=1}^{\infty} v_{jk}^2 x_k^2\right)/(V^*(X) + \delta)\right] \quad (\text{by (3)}) \\
&\leq C(\alpha)\hat{E}[(V_\infty(X)^2 - D_\infty)/(V^*(X) + \delta)] \\
&\leq C(\alpha)\{\hat{E}[V_\infty(X)^2/(V^*(X) + \delta)] + |\hat{E}[D_\infty \hat{H}_\infty]|\} \\
&\leq C(\alpha, \varepsilon, M)\{\hat{E}[V^*(X) + \delta] + \hat{E}[S(X)]\} .
\end{aligned}$$

Therefore we obtain (10).

LEMMA 5. *The inequality*

$$(13) \quad \hat{E}[U^{**}(X)] \leq C(\alpha, \varepsilon, M)\hat{E}[V^*(X)]$$

holds for all martingale X .

PROOF. Since $S(\hat{X})$ is locally \hat{P} -integrable for a \hat{P} -local martingale \hat{X} , $S(X)$ is also locally \hat{P} -integrable by (6). By using the stopping argument, we may assume $\hat{E}[S(X)] < \infty$ and $\hat{E}[S(D(X))] < \infty$. Applying Lemma 3 to the case where (u_{jk}) is a single-row matrix, we get $\hat{E}[D(X)^*] \leq C\hat{E}[S(D(X))] < \infty$. By Schwarz's inequality and by Lemma 4, we have

$$\begin{aligned}
\hat{E}[S(X)] &= \hat{E}[S(X)(V^*(X) + \delta)^{-1/2}(V^*(X) + \delta)^{1/2}] \\
&\leq \hat{E}[S(X)^2/(V^*(X) + \delta)]^{1/2} \hat{E}[V^*(X) + \delta]^{1/2} \\
&\leq c(\alpha, \varepsilon, M)\{\hat{E}[V^*(X) + \delta] + \hat{E}[S(X)]\}^{1/2} \hat{E}[V^*(X) + \delta]^{1/2} .
\end{aligned}$$

Put $A = \hat{E}[S(X)]$ and $B = \hat{E}[V^*(X) + \delta]$. Then the above inequality is equal to $A \leq c\{(B + A)B\}^{1/2}$, where $A < \infty$. Therefore there exists some constant c' , depending only on c , such that $A \leq c'B$, that is, $\hat{E}[S(X)] \leq c'\hat{E}[V^*(X) + \delta]$. Letting $\delta \rightarrow 0$ and combining this inequality with (9), we obtain (13).

3. Proof of Theorem. By virtue of Neveu-Garsia's lemma (see [8, IX-3-5]), it is sufficient to prove that there is a positive constant c such that

$$(14) \quad \hat{E}[U^{**}(X) - U_{n-1}^{**}(X) | F_n] \leq c \hat{E}[V^*(X) | F_n]$$

for every martingale X and $n \geq 1$. By (5), the inequality (14) coincides with the following inequality:

$$(15) \quad E[(U^{**}(X) - U_{n-1}^{**}(X))(Z_\infty/Z_n) | F_n] \leq c E[V^*(X)(Z_\infty/Z_n) | F_n]$$

for every martingale X over (F_n) . Let A be an element of F_n . Set $d\bar{P} = (Z_\infty/Z_n)dP$ and $X'_k = \{X_{k+n-1} - X_{n-1}\}I_A$ for each martingale X over (F_n) . Then $X' = (X'_k)_{k \geq 1}$ is a martingale over $(F_{k+n-1})_{k \geq 1}$ such that

$$(16) \quad U^{**}(X) - U_{n-1}^{**}(X) \leq U^{**}(X') \quad \text{and} \quad V^*(X') \leq 2V^*(X)$$

on A . Furthermore, it is easy to see that $M' \in \text{BMO}$ with $1/\varepsilon \geq 1 + m'_k \geq \varepsilon$ for the BMO-martingale M with $1/\varepsilon \geq 1 + m_k \geq \varepsilon$. Therefore the inequality (13) is still valid for a martingale X' over $(F_{k+n-1})_{k \geq 1}$ and for the weighted probability measure $d\bar{P}$ instead of dP . Thus we get $\bar{E}[U^{**}(X')] \leq C\bar{E}[V^*(X')]$, that is, $E[U^{**}(X')(Z_\infty/Z_n); A] \leq CE[V^*(X')(Z_\infty/Z_n); A]$ for every X' , where $\bar{E}[\]$ denotes the expectation over Ω with respect to $d\bar{P}$. By (16), we obtain

$$\begin{aligned} E[(U^{**}(X) - U_{n-1}^{**}(X))(Z_\infty/Z_n); A] &\leq E[U^{**}(X')(Z_\infty/Z_n); A] \\ &\leq CE[V^*(X')(Z_\infty/Z_n); A] \leq CE[V^*(X)(Z_\infty/Z_n); A] \end{aligned}$$

for every martingale X . This holds for any $A \in F_n$, so that we get (15). Hence the theorem is established.

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