

Some Limit Theorems for Random Fields

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Abstract. We prove a central limit theorem with remainder and an iterated logarithm law for collections of mixing random variables indexed by $Z^d, d \geq 1$. These results are applicable to certain Gibbs random fields.

1. Introduction

In this paper we extend various classical limit theorems to collections of random variables indexed by $Z^d, d \geq 1$. The study of such collections is motivated by Gibbs random fields. For example we may define our probability space (Ω, \mathcal{F}, P) in the following way. Let $\Omega = \{-1, +1\}^{Z^d}$, with \mathcal{F} the σ -algebra generated by finite-dimensional cylinder sets. Then if Φ is a translation-invariant, finite-range, real-valued potential function on the finite subsets of Z^d with $\Phi(\emptyset) = 0$, the Gibbs state for the potential Φ is a probability measure P on (Ω, \mathcal{F}) for which, if $n \in Z^d$,

$$(1.1) \quad P\left(\omega(n) \middle| \mathcal{F}_{Z^d - \{n\}}\right) = \left(1 + \exp\left(2 \sum_{A \ni n} (\Phi(A)) \prod_{m \in A} \omega(m)\right)\right)^{-1}$$

is a regular conditional probability distribution for the “spin” at site n given the configuration on $Z^d - \{n\}$. Here \mathcal{F}_A is the σ -algebra depending on the coordinates in the set $A \subset Z^d$.

In [2] a unique such P is shown to exist for certain choices of Φ . Moreover conditions on Φ which imply that mixing occurs are stated. That is, it is shown that for certain Φ and for $A, B \subset Z^d$ the dependence of sets in \mathcal{F}_A on sets in \mathcal{F}_B decreases as $d(A, B)$, the usual Euclidean distance between A and B , increases, but may increase as $|A|$, the cardinality of A , increases.

In [5] the F. K. G. inequalities are used to determine conditions on Φ which imply that events at sites $n, m \in Z^d$ are positively correlated. For example if Φ is a nearest-neighbor potential

$$\Phi(\{n, m\}) > 0 \text{ for } d(n, m) = 1$$

is sufficient.

Thus we are led to study random variables $(X_n), n \in Z^d$, which are positively

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correlated and which satisfy some sort of mixing condition. In the above example, for instance, we may consider $X_n(\omega) = \omega(n)$, the “spin” at n . We are interested here in asymptotic properties of sums $S_A = \sum_{n \in A} X_n$. In [9] it is shown that a sufficiently strong mixing condition together with positive correlation between sites and certain moment conditions implies a central limit theorem for the sequence S_A when $A \uparrow \mathbb{Z}^d$ sufficiently nicely. In this paper we show that the same properties of P and the random variables X_n may be used to estimate the speed at which S_A converges to a normal random variable. This estimate is then used to obtain results about *a.s.* convergence, namely an iterated logarithm law and its consequences. The techniques used are modifications of those which work for $d = 1$ [see e.g., 14, chapter 5.] The question of when the central limit theorem holds for random fields has, of course, been studied repeatedly (see, e.g., [3]), but questions about *a.s.* convergence seem not to have been looked at much up to now.

The statements of our results are in Section 3 and the proofs are in Section 5. Section 2 contains some necessary notation and some basic estimates. In Section 4 we discuss some of the known results on mixing for random variables indexed by \mathbb{Z}^d . So far it seems that in models where mixing can be shown to occur it occurs at an exponential rate. However some of the results in [10] strongly suggest that slower mixing rates are possible, perhaps for infinite range potentials. Because of this and also because our proofs in the exponential case are easily modified to work for more slowly decreasing functions we have included such functions in the statements of the results.

2. Notation and Some Basic Estimates

Let $d \geq 1$ be fixed. We assume throughout that we are given a probability space (Ω, \mathcal{F}, P) , with $\Omega = \mathbb{R}^{\mathbb{Z}^d}$ and \mathcal{F} the σ -algebra generated by finite-dimensional cylinder sets, and a collection of real-valued random variables $(X_n, n \in \mathbb{Z}^d)$, with $X_n \in \mathcal{F}_{\{n\}}$. We also assume

$$(2.1) \quad E(X_n) = 0 \text{ and } 0 < E(X_n^2) < \sigma^2 \text{ for all } n.$$

For simplicity we assume A_N is a d -dimensional cube of side N or sphere of radius N . We set

$$S_N = \sum_{n \in A_N} X_n, \text{ and}$$

$$\Sigma_N^2 = \text{Var } S_N = E \left(\sum_{n \in A_N} X_n \right)^2,$$

and we assume that there is a $C > 0$ for which, if $A \subset \mathbb{Z}^d$,

$$(2.2) \quad E \left(\sum_{n \in A} X_n \right)^2 \geq C |A|.$$

Condition (2.2) is implied, of course, if the X_n are positively correlated or independent and identically distributed or if events at sites m and $n, m \neq n$, are sufficiently

close to being uncorrelated. If the X_n are independent we are reduced to the case $d = 1$. Here we assume instead an asymptotic independence or “mixing”.

Definition (2.3). We shall say the measure P satisfies mixing condition (M) if there exists a continuous function $\alpha: [1, \infty) \rightarrow (0, \infty)$ such that $\alpha(h) \downarrow 0$ as $h \uparrow \infty$ and, for $A, B \subset Z^d$ with $d(A, B) = h, E_1 \in \mathcal{F}_A$ and $E_2 \in \mathcal{F}_B$,

$$(M) \quad |P(E_1 E_2) - P(E_1)P(E_2)| \leq \alpha(h) |A| |B|.$$

Obviously this definition is nontrivial only if A and B are both finite. Now (M) implies the following estimates, the proofs of which may be found in [1] and [16]:

Lemma (2.4). *Suppose (M) holds, $A, B \subset Z^d$ with $d(A, B) = h$, and $f \in \mathcal{F}_A, g \in \mathcal{F}_B$ with $\|f\|_p < \infty$ and $\|g\|_q < \infty$ for some $p, q \geq 1$, where the norms are with respect to the measure P . Then if $p, q, r \geq 1$ satisfy $\frac{1}{p} + \frac{1}{q} + \frac{1}{r} = 1$,*

$$(2.5) \quad |E(fg) - E(f)E(g)| \leq (4\alpha(h)|A||B|)^{1/r} \|f\|_p \|g\|_q.$$

If $p = q = \infty$

$$(2.6) \quad |E(fg) - E(f)E(g)| \leq 4\alpha(h)|A||B| \|f\|_\infty \|g\|_\infty.$$

3. The Main Theorems

Let $\phi(x)$ denote the distribution function of a normal random variable with mean 0 and variance 1. Then we have the following results:

Theorem (3.1). *Suppose (2.2) holds and (M) holds with $\alpha(h) = e^{-\beta h}$ for some $\beta > 0$. Suppose also that there is a $C > 0$ with $E|X_n|^5 \leq C$ for all $n \in Z^d$. Then for $x \in \mathbb{R}$ and $d > 1$*

$$P(X_N/\Sigma_N < x) = \phi(x) + O(\Sigma_N^{-(3/5d)}).$$

If $d = 1$

$$P(X_N/\Sigma_N < x) = \phi(x) + O(\Sigma_N^{-1/2}).$$

Theorem (3.2). *Suppose (2.2) holds and (M) holds with*

$$\int_1^\infty \alpha^{1/5}(h) h^{2d-1} dh < \infty.$$

Suppose also that there is a $C > 0$ with $E|X_n|^5 \leq C$ for all $n \in Z^d$. Then for $x \in \mathbb{R}$

$$(3.3) \quad P(X_N/\Sigma_N < x) = \phi(x) + O(\Sigma_N^{-(1/4d)}).$$

Theorem (3.4). *Suppose (2.2) holds and (M) holds with*

$$\alpha(h) = O(h^{-2d-\varepsilon} \wedge h^{-3})$$

for some $\varepsilon > 0$. Suppose there is a $C > 0$ with $\|X_n\|_\infty \leq C$ for all $n \in Z^d$. Then (3.3) is true.

These results slightly better the remainder terms obtained in [12] for $d = 1$,

but the error is still considerably larger than in the independent case, where the example of coin-tossing random variables, $X_n = \pm 1$, each with probability $1/2$, $d = 1$, shows that $O(\Sigma_N^{-1})$ is best possible.

These theorems have the following consequences:

Theorem (3.5). *Suppose the X_n satisfy the hypotheses of (3.1) or (3.2). If we set*

$$(3.6) \quad t_N = 2\Sigma_N^2 \log(\log \Sigma_N^2),$$

then

$$(3.7) \quad \overline{\lim}_{N \rightarrow \infty} \frac{S_N}{t_N^{1/2}} = 1 \text{ a.s.}$$

and

$$(3.8) \quad \lim_{N \rightarrow \infty} \frac{S_N}{t_N^{1/2}} = -1 \text{ a.s.}$$

Theorem (3.9). *Suppose the X_n satisfy the hypotheses of (3.4) with $\alpha(h) = O(h^{-8d-6-\epsilon})$ for some $\epsilon > 0$. Then (3.7) and (3.8) are true for t_N given by (3.6).*

Theorem (3.9), of course, improves (3.5) and (3.10) below only for $d \geq 3$.

Corollary (3.10). *If the X_n satisfy the hypotheses of (3.5) or (3.8) and if $(a_N)_{N \geq 1}$ are constants for which*

$$t_N^{1/2} = o(a_N),$$

where t_N is defined by (3.8) then

$$(3.11) \quad \lim_{N \rightarrow \infty} \frac{S_N}{a_N} = 0 \text{ a.s.}$$

Notice that (3.10) implies

$$\lim_{N \rightarrow \infty} \frac{S_N}{\Sigma_N^p} = 0 \text{ a.s.}$$

for any $p > 1$. In [9] it is shown that if $E(X_n^4) \leq C$ for all n , and (M) holds with

$$\int_1^\infty \alpha^{1/2}(h)h^{d-1} < \infty,$$

then (3.11) is true for $a_N = \Sigma_N^{2k}$, $k > 1/2$, if we replace “a.s.” by “in probability”, while (3.11) holds a.s. for $k \geq 1$ if $d = 1$ and $k > \frac{1}{2} \left(1 + \frac{1}{d}\right)$ if $d > 1$. These results do not require (2.2).

Results similar to those stated here are also proved in [13, especially Chapter 7] and [15], while in [17] a functional law of the iterated logarithm is obtained for $d \geq 1$ for processes with independent increments.

4. Examples of Mixing

In general determining when mixing occurs and at what rate seems to be an open problem. Perhaps the simplest method which has been developed to show that mixing occurs is that outlined in [2]. We describe the method and show how it applies to a simple example. Let us define (Ω, \mathcal{F}) as in Section 1. For $\omega \in \Omega, n \in \mathbb{Z}^d$, let ${}_n\omega$ be $\omega|_{\mathbb{Z}^d - \{n\}}$ and let

$$q({}_n\omega, \cdot) = P(\cdot | {}_n\omega)$$

where $P(\cdot | \cdot)$ is given by (1.1). For $\omega_1, \omega_2 \in \Omega$ let $\rho(q({}_n\omega_1, \cdot), q({}_n\omega_2, \cdot))$ be the distance in variation between the measures $q({}_n\omega_1, \cdot)$ and $q({}_n\omega_2, \cdot)$,

$$(4.1) \quad \rho(q({}_n\omega_1, \cdot), q({}_n\omega_2, \cdot)) \\ = \frac{1}{2} \{ |q({}_n\omega_1, 1) - q({}_n\omega_2, 1)| + |q({}_n\omega_1, -1) - q({}_n\omega_2, -1)| \}.$$

If $n, m \in \mathbb{Z}^d, n \neq m$, let

$$\rho_{n,m} = \sup \rho(q({}_n\omega_1, \cdot), q({}_m\omega_2, \cdot)),$$

where the sup is over all pairs ${}_n\omega_1, {}_m\omega_2$ with ${}_n\omega_1(l) = {}_m\omega_2(l)$ for $l \neq m$. Finally if $S \subset \mathbb{Z}^d$ set

$$\alpha(S, (S)_h) = \sum \rho_{n_2, n_1} \cdots \rho_{n_k, n_{k-1}},$$

where the sum is over all paths n_1, \dots, n_k with $n_1 \in S, n_j \notin S$ for $1 < j \leq k, n_{j+1} \neq n_j$ for $j = 1, \dots, k - 1$, and $d(n_k, S) \geq h$. Then the following is true:

Theorem (4.2). (Theorems 2 and 5, [2]). *Suppose there is an $\alpha < 1$ with*

$$(4.3) \quad \sum_{m \in \mathbb{Z}^d - \{n\}} \rho_{n,m} < \alpha$$

for all $n \in \mathbb{Z}^d$. Then there is a unique (translation invariant) probability measure P on (Ω, \mathcal{F}) with regular conditional probability distribution given by (1.1) If in addition

$$(4.4) \quad \sum_{n \in \mathbb{Z}^d - \{m\}} \rho_{n,m} < \alpha$$

for all $m \in \mathbb{Z}^d$ and if $S_1, S_2 \subset \mathbb{Z}^d$ with $d(S_1, S_2) \geq h$ then

$$(4.5) \quad \sup_{A \in \mathcal{F}_{S_1}, B \in \mathcal{F}_{S_2}} |P(AB) - P(A)P(B)| \leq \alpha(S_1, (S_1)_h).$$

It is obvious that (4.5) is actually stronger than (M), since the right hand side of (4.5) does not depend on $|S_2|$. Theorem (4.1) as stated in [2] is actually true with $\{-1, 1\}$ replaced by any finite set and with the definitions above modified accordingly.

As a specific case of (4.1) let us consider the Ising model with

$$\Phi(A) = \left\{ \begin{array}{ll} K/2 & \text{if } |A| = 1 \\ J/2 & \text{if } A = \{n, m\} \text{ and } d(n, m) = 1 \\ 0 & \text{otherwise} \end{array} \right\},$$

where K and J are real numbers. Then if $x, y = \pm 1, d(n, m) = 1$, and ${}_n\omega(m) = y$,

$$q_n(\omega, x) = [1 + \exp(Kx + Jxy + Jx\sum\omega(j))]^{-1},$$

where the sum is over all $j \neq m$ with $d(j, n) = 1$. Thus if we set $\mu = \min |K + J\sum\omega(y)|$ we find that

$$\rho_{n,m} = \begin{cases} \sinh |J|/(\cosh |J| + \cosh(\mu)) & \text{if } d(n, m) = 1 \\ 0 & \text{otherwise} \end{cases}$$

and that $\delta_{n,m}$ does not depend on the sign of K or of J . Since each $n \in \mathbb{Z}^d$ has exactly $2d$ neighbors at distance 1, the sums in (4.3) and (4.4) are identical and thus Theorem (4.2) tells us that in this case there will be (a unique P and) exponential mixing exactly when, for some $\alpha < 1$,

$$\sinh |J|/(\cosh |J| + \cosh \mu) < \alpha/2d.$$

For example if $J = 0$ then there is no interaction and trivially mixing occurs. If $K = 0$ we must have $\tanh |J| < 1/d$, and thus for $d = 1$ there is mixing for all choices of J , while if $d = 2$ we can show mixing only if $\tanh |J| < 1/2$. When $K \neq 0$ and J is fixed and nonzero, K sufficiently large implies mixing while $|K| = |J|$ implies mixing occurs only if $\sinh |J|/(1 + \cosh |J|) < 1/2d$, a more stringent condition than in the case $K = 0$. In fact it is known [7] that for this model, if $K \neq 0$, there is a unique P with the given conditional distributions. However if $d = 1$ and $K = J = 2$ we have $\sinh |J|/(\cosh |J| + 1) > 3/4$ and (4.3) and (4.4) are not satisfied. Thus Theorem (4.2), while easy to apply, is not best possible.

A result similar to Theorem (4.2) is proved in [6, Theorem (4.24)]. In the example above, condition (a) Theorem (4.24) [6] is the same as our condition (4.4). The two conditions, in fact, would be equivalent for (1.1) with any pair potential and possible states ± 1 . Mixing conditions of a slightly different type are obtained in [4].

Now if (4.3) is satisfied for this model the right-hand side of (4.5) is less than or equal to $|S_1| \alpha^h/(1 - \alpha)$ and so (M) certainly holds with $\alpha(h) = \alpha^h/(1 - \alpha)$. Physically-motivated examples where mixing occurs at a slower than exponential rate seem hard to come by. Suppose, however, we set, for $n, m \in \mathbb{Z}^d, \|n, m\| = \max_{1 \leq i \leq d} |n_i - m_i|$ and let $u(\{n, m\}, K)$ be the two point Ursell function for $\Phi(\{n\}) = K$, i.e.,

$$u(\{n, m\}, K) = E(X_n X_m) - EX_n EX_m.$$

Then in [10] it is shown that whenever

$$\lim_{\|n, m\| \rightarrow \infty} \sup \ln |\phi(n, m)|/\|n, m\| = 0$$

there is a K_0 for which

$$\lim_{\|n, m\|} \sup \ln |u\{x, y\}, K|/\|n, m\| = 0$$

for almost all K with $K < K_0$. By Lemma (2.4) if mixing occurs the function

$\alpha(h)$ must satisfy, for $d(n, m) = h$,

$$\alpha(h) \geq A |u(\{n, m\}, K)|$$

for some constant A . Thus physically-motivated examples of slower rates of mixing may exist.

5. Proofs of the Theorems

We first decompose S_N into “almost independent” summands and a “negligible” summand, a standard technique for dealing with mixing random variables.

Definition (5.1). We define a sequence of annuli

$$A_{N,1}, A'_{N,1}, A_{N,2}, A'_{N,2}, \dots, A_{N,l_N}, A'_{N,l_N}, A'_{N,l_N+1},$$

all contained in A_N , and random variables

$$\begin{aligned} Y_{N,j} &= \sum_{n \in A_{N,j}} X_n, \\ Z_{N,j} &= \sum_{n \in A'_{N,j}} X_n, \\ Y_N &= \sum_j Y_{N,j}, \text{ and} \\ Z_N &= \sum_j Z_{N,j}, \end{aligned}$$

where we choose $A_{N,j}$ to be the largest annulus outside $A'_{N,j-1}$ for which

$$E(Y_{N,j}^2) \leq N^\eta$$

and $A'_{N,j}$ to be the annulus outside $A_{N,j}$ with radius k_N , where η and k_N will be defined in the following lemmas.

Lemma (5.2). *Suppose the hypotheses of (3.1) are satisfied. Then if we set*

$$\eta = d - 1 + \varepsilon, 0 < \varepsilon < 1,$$

and

$$k_N = r \log N, (d + 2 - 2\varepsilon)/\beta < r,$$

we have

$$(5.3) \quad E(Y_{N,j}^2) = N^\eta(1 + o(1)), j \leq l_N,$$

$$(5.4) \quad E(Y_{N,l_N+1}^2) \leq N^\eta(1 + o(1)),$$

$$(5.5) \quad \sum_{j \leq l_N+1} E(Y_{N,j}^2) = l_N N^\eta(1 + O(N^{-\eta/2})),$$

$$l_N = \left\{ \begin{array}{l} \Sigma_N^2 N^{-\eta}(1 + O(N^{-\varepsilon} r \log N)) \text{ if } d > 1 \\ \Sigma_N^2 N^{-\eta}(1 + O(N^{-\eta/2})) \text{ if } d = 1 \end{array} \right\},$$

and

$$(5.6) \quad E\left(\sum_{j \leq l_N} Z_{N,j}\right)^2 = N^{d-\varepsilon}(r \log N)(1 + O(1)).$$

Lemma (5.7). *If the hypotheses of (3.2) or (3.4) are satisfied, and if we set*

$$\eta = d - \varepsilon, 0 < \varepsilon < 1,$$

and

$$k_N = N^{\varepsilon_1}, \varepsilon_1 \geq \varepsilon, 1 - \varepsilon_1 > \varepsilon, 1 > \varepsilon + \varepsilon_1,$$

then (5.3), (5.4) and (5.5) are satisfied, while

$$l_N = \Sigma_N^2 N^{-\eta}(1 + O(N^{\varepsilon_1 + \varepsilon - 1})),$$

and

$$(5.8) \quad E\left(\sum_{j \leq l_N} Z_{N,j}\right)^2 = O(N^{d-1+\varepsilon+\varepsilon_1}).$$

Proof of (5.2) and (5.7). The proof is similar to the proof of Lemma 5 in [11]. Note that (2.1), (2.2), (2.4), and (5.3) together imply that $\Sigma_N^2 \sim N^d$ and that the number of terms in a $Y_{N,j}$ is about N^η , while the number of terms in a $Z_{N,j}$ is $O(k_N N^{d-1})$.

Theorem (5.9). *Suppose the hypotheses of (3.1), (3.2), or (3.4) are satisfied. Then the central limit theorem holds for X_N/Σ_N , i.e., the distribution function of X_N/Σ_N converges to the distribution function of ϕ .*

Proof. As usual, we prove that $E(\exp(itX_N/\Sigma_N))$ converges to $e^{-t^2/2}$ for $t \in \mathbb{R}$. In each case $N^{2d-\eta} l_N \alpha(k_N) \rightarrow 0$, so that applying (2.6) l_N times gives

$$E(\exp(itY_N/\Sigma_N)) = \prod_{j \leq l_N+1} E(\exp(itY_{N,j}/\Sigma_N)) + o(1).$$

Also (5.6) or (5.8) implies $E(\exp(itX_N/\Sigma_N)) = E(\exp(itY_N/\Sigma_N)) + o(1)$, so it is enough to show that

$$\prod_{j \leq l_N+1} E(\exp(itY_{N,j}/\Sigma_N))$$

converges to $e^{-t^2/2}$. Now if p_j is the number of summands in $Y_{N,j}$, (2.5), the hypotheses of (3.1), (3.2), or (3.4), and the fact that $\Sigma_n^2 \sim N^d$ can be used to show

$$E(Y_{N,j}^4) = O(p_j^2) = O(N^{2\eta}).$$

Thus

$$\begin{aligned} E|Y_{N,j}|^3 &= O(N^{2\eta})^{3/4} \\ &= O(N^{3\eta/2}) \end{aligned}$$

and thus

$$(5.10) \quad \Sigma_N^{-3} \sum_{j \leq l_N+1} E|Y_{N,j}|^3 = O(N^{-3d/2} l_N N^{3\eta/2}) = o(1),$$

which proves (5.9) by Liapounov's theorem [8, p. 275].

Proof of (3.1), (3.4), and (3.6). We follow the procedure outlined in [12]. Set $g_N^3 = 8\Sigma_N^{-3} \sum_{j \leq l_{N+1}} E|Y_{N,j}|^3$. Now, following the proof of Lemma 3 in [12], we may show that for large N and for $0 \leq t \leq 2/g_N^3$

$$\left| \prod_{j < l_{N+1}} E(\exp(itY_{N,j}/\Sigma_N) - e^{-t^2/2}) \right| \leq e^{-t^2/6} 2t^2(tg_N^3 + R)$$

where

$$(5.11) \quad R = \left\{ \begin{array}{ll} O(N^{-\epsilon} r \log N) & \text{for (3.1), } d > 1 \\ O(N^{-\eta/2}) & \text{for (3.1), } d = 1 \\ O(N^{\epsilon_1 + \epsilon - 1}) & \text{for (3.2) or (3.4)} \end{array} \right\}.$$

Now the same proof as that of Lemma 4 in [12] gives

$$E(\exp(itY_N/\Sigma_N)) - \prod_{j \leq l_{N+1}} E(\exp(itY_{N,j}/\Sigma_N)) = O(t^2)$$

uniformly in $0 \leq t \leq 1$, and then the proof of (14) in [12] yields that for $T = \min(2g_N^{-3}, \Sigma_N^2)$,

$$\int_0^T |E(\exp(itY_N/\Sigma_N)) - e^{-t^2/2}| t^{-1} dt = O(g_N^3 + R).$$

Thus by the basic inequality [8, p. 285]

$$(5.12) \quad P(Y_N/\Sigma_N < x) - \phi(x) = O(g_N^3 + R).$$

Now for $\epsilon_N > 0$

$$\begin{aligned} (5.13) \quad P(X_N/\Sigma_N < x) &= P\left(((Y_N + Z_N)/\Sigma_N < x) \cap \left(\frac{|Z_N|}{\Sigma_N} \geq \epsilon_N \right) \right) \\ &\quad + P\left(((Y_N + Z_N)/\Sigma_N < x) \cap \left(\frac{|Z_N|}{\Sigma_N} < \epsilon_N \right) \right) \\ &= P(Y_N/\Sigma_N < x + \epsilon_N) + O\left(\frac{E(Z_N^2)}{\Sigma_N^2 \epsilon_N^2} \right) \\ &= \phi(x) + O(g_N^3 + R) + O\left(\frac{E(Z_N^2)}{\Sigma_N^2 \epsilon_N^2} \right) + O(\epsilon_N) \end{aligned}$$

by (5.12) and Chebyshev's inequality. Now set $\epsilon = .395$, $\epsilon_N = N^{-.3}$ if the hypotheses of (3.1) are satisfied with $d > 1$, $\epsilon = .5$, $\epsilon_N = N^{-.35}$ if the hypotheses of (3.1) are satisfied with $d = 1$, and $\epsilon = 1/4$, $\epsilon_1 = 3/8$, $\epsilon_N = N^{-1/8}$ if the hypotheses of (3.2) or (3.4) are satisfied. Together with (5.4), (5.11), and (5.6) or (5.8) this allows us to show that (5.13) gives the desired estimate.

Proof of (3.5) and (3.9). Since (3.7) implies (3.8) when we replace S_N by $-S_N$, it is enough to show (3.7) is true. This is most easily accomplished in the manner

outlined in [14, pp. 320 ff.], where results from [15] are proved in detail. First for $|\delta| < 1$ and $\xi > 0$

$$(5.14) \quad (\log \Sigma_N^2)^{-(1+\delta)^2(1+\xi)} < P(S_N > (1 + \delta)(t_N^{1/2})) < (\log \Sigma_N^2)^{-(1+\delta)^2(1-\xi)}$$

follows from (3.1), (3.2), or (3.4). We also need that for $x > 0$ and $N \geq 1$

$$(5.15) \quad P\left(\max_{J \leq N} S_J > x\right) \leq 2P(S_N > x - 2\Sigma_N) + CN^{-\xi}$$

for $0 < \xi < 1/4$ and $C > 0$. This we may prove as follows:

Setting $B_J = (S_J \leq x, I < J, S_J > x), 1 \leq J \leq N$, we have

$$P\left(\max_{J \leq N} S_J > x\right) = \sum' P(B_J) + CN^{-\xi},$$

where \sum' is over all J with $P(B_J) > CN^{-(1+\xi)}$. So it suffices to show

$$\sum' P(B_J) \leq 2P(S_N > x - 2\Sigma_N).$$

Now

$$\begin{aligned} \sum_{J=1}^N P(|S_N - S_J| \leq 2\Sigma_N | B_J) P(B_J) &= \sum_{J=1}^N P((|S_N - S_J| \leq 2\Sigma_N) \cap B_J) \\ &\leq P(S_N > x - 2\Sigma_N) \end{aligned}$$

and thus it suffices to show

$$P(|S_N - S_J| > 2\Sigma_N | B_J) \leq 1/2$$

if $P(B_J) > CN^{-(1+\xi)}$. Now for N large enough and for some $\gamma, 0 < \gamma + \xi < 1/2, \alpha(N^\gamma) \leq 10^{-5}$. Fix $\tau, 0 < \tau < 1/2$, and $J \geq 1$ with $P(B_J) > CN^{-(1+\xi)}$. If $N - J > N^\gamma$, then

$$\begin{aligned} P(|S_N - S_J| > 2\Sigma_N | B_J) &\leq P\left(\left|\sum_{n \in A_{J+N^\gamma} - A_J} X_n\right| > \tau\Sigma_N | B_J\right) \\ &\quad + P\left(\left|\sum_{n \in A_N - A_{J+N^\gamma}} X_n\right| > (2 - \tau)\Sigma_N | B_J\right) \\ &\leq \frac{E\left|\sum_{n \in A_{J+N^\gamma} - A_J} X_n\right|^3}{P(B_J)\tau^3\Sigma_N^3} \\ &\quad + CN^{1+\xi}N^{2d}\alpha(N^\gamma) + P\left(\left|\sum_{n \in A_N - A_{J+N^\gamma}} X_n\right| \geq \frac{3\Sigma_N}{2}\right), \end{aligned}$$

where we have used Chebyshev's inequality and (M) to obtain the first and second terms respectively. But

$$E\left|\sum_{n \in A_{J+N^\gamma} - A_J} X_n\right|^3 = O\left(\frac{N^{1+\xi}(N^\gamma N^{d-1})^{3/2}}{N^{3d/2}}\right)$$

$$= O(N^{\xi+\gamma-1/2}) < \frac{1}{100}$$

for N large enough, and, again by Chebyshev’s inequality and (M) ,

$$CN^{1+\xi}N^{2d}\alpha(N^\gamma) + P\left(\left|\sum_{n \in A_N - A_{J+N^\gamma}} X_n\right| > \frac{3\Sigma_N}{2}\right) = O\left(N^{2d+1+\xi}\alpha(N^\gamma) + \frac{4}{9}\left(\frac{N^{d-1}(N-J-N^\gamma)}{N^d}\right)\right) = \frac{4}{9} + o(1),$$

by the conditions on $\alpha(\cdot)$ specified by (3.5) or (3.9), if we choose $\xi = \frac{1}{100}, \gamma = \frac{4}{10}$. If, on the other hand, $N - J \leq N^\gamma$,

$$P(|S_N - S_J| > 2\Sigma_N | B_J) \leq \frac{E\left|\sum_{n \in A_N - A_J} X_n\right|^3 N^{1+\xi}}{8\Sigma_N^3 C} = O\left(\frac{N^{1+\xi}(N^{d-1}+\gamma)^{3/2}}{N^{3d/2}}\right) = O(N^{-(1/2)+\xi+\gamma}).$$

Thus (5.15) is true. Now, as in the standard proof of the law of the iterated logarithm, we have, for $\delta > 0, \rho > 1$, and $N_j = \rho^j, j \geq 1$, by (5.15),

$$\sum_{j=1}^\infty P\left(\max_{N \leq N_{j+1}} \frac{S_N}{t_N^{1/2}} > 1 + \delta\right) \leq \sum_{j=1}^\infty 2P\left(S_{N_{j+1}} > (1 + \delta)\epsilon_{N_j}^{1/2} - 2\Sigma_{N_{j+1}}\right) + \sum_{j=1}^\infty C(N_{j+1})^{-\xi} < \infty,$$

and thus

$$(5.16) \quad P\left(\max_{N \leq N_{j+1}} \frac{S_N}{t_N^{1/2}} > 1 + \delta \text{ i.o.}\right) = 0.$$

Now for $N > 1$ and $\delta > 0$, if we set

$$C_k = (S_{N^k} - S_{N^{k-1} + N^{k/2}} > (1 - \delta)(2\chi_k^2 \log(\log \chi_k^2))^{1/2})$$

where

$$\chi_k^2 = \text{Var}(S_{N^k} - S_{N^{k-1} + N^{k/2}}) \sim N^{kd} - (N^{k-1} + N^{k/2})^d,$$

then by (5.14), for k large,

$$P(C_k) > (\log \chi_k^2)^{-(1-\delta)^2(1+\xi)}$$

for $|\delta| > 1$ and $\xi > 0$. If ξ is chosen so that $(1 - \delta)^2(1 + \xi) < 1$ then

$$(5.17) \quad \sum_{k=1}^\infty P(C_k) = \infty,$$

since for N and k sufficiently large

$$N^{kd} - (N^{k-1} + N^{k/2})^d > \frac{N^{kd}}{2}.$$

Now, by (M) and (3.5) or (3.9),

$$\left| P\left(\bigcap_{k=n}^{n+m} C_k^c\right) - \prod_{k=n}^{n+m} P(C_k^c) \right| = O\left(\int_{N^{(k+1)/2}}^{\infty} \alpha(h)h^{4d}dh\right) = o(1).$$

This estimate, together with (5.17) gives

$$(5.18) \quad P(C_k \text{ i.o.}) = 1.$$

If

$$B_k = (S_{N^{k+1} + N^{k/2}} > -2(2\tilde{\chi}_k^2 \log(\log \tilde{\chi}_k^2))^{1/2}),$$

where

$$\tilde{\chi}_k^2 = \text{Var}(S_{N^{k+1} + N^{k/2}}) \sim (N^{k+1} + N^{k/2})^d,$$

then by (5.16) and (5.18)

$$(5.19) \quad 1 = P(B_k \cap C_k \text{ i.o.}) < P(S_{N^k} > (1 - \delta')t_N^{1/2} \text{ i.o.})$$

for $\delta' > \delta$. Now (3.5) and (3.9) follow from (5.16) and (5.19).

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