

88. On the Asymptotic Distribution of the Sum of Independent Random Variables.

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§1. Let $\{X_i\}$ $i = 1, 2, \dots$ be a sequence of independent random variables defined in a probability space (Ω, \mathcal{F}, P) . The so-called central limit theorem¹⁾ states that when a sequence $\{X_i\}$ satisfies certain conditions then

$$\lim_{n \rightarrow \infty} P\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n X_i(\omega) \leq a\right) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^a e^{-u^2/2} du = G(a), \quad (\text{I})$$

where $\sqrt{\frac{1}{n}} \sum_{i=1}^n X_i$ denotes suitably normalized variable. Concerning this theorem we consider following two generalizations:

1° Replace a constant upper limit a of summation by a measurable function $g(\omega)$ defined in Ω .

2° Replace the number n of random variables of summation by a random function $N_n(\omega)$ defined in Ω .

On these generalizations many theorems have been proved²⁾. Let $\{X_i\}$ be a sequence of independent random variables satisfying the central limit theorem (I). For any real numbers a and b , we define the sets $E_{a,b}^i = [\omega; a \leq X_i(\omega) < b]$ and denote by \bar{F} the smallest Borel field which includes all the sets $E_{a,b}^i$ defined for any a, b and $i = 1, 2, \dots$. We complete \bar{F} with respect to the measure P and denote it by \bar{F} . In §3 we prove the following:

Theorem 1. If $E \in \bar{F}$, then

$$\lim_{n \rightarrow \infty} P\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n X_i(\omega) \leq a, E\right) = P(E)G(a).$$

In order to prove this theorem we show some lemmas in §2, and in §4 we consider the above generalizations by using Theorem I.

To define and to discuss the problems on $\{X_i\}$, it is sufficient to consider the probability space (Ω, \mathcal{F}, P) as (Ω, \bar{F}, P) . So the theorems proved in §4 give the answer of the above generalizations for independent sequence.

§2. First of all we consider a sequence $\{X_i\}$ which satisfies following conditions:

1) H. Cranier, Random variable and its probability distribution. Cambridge (1937).
2) J. C. Smith, On the asymptotic distribution of the sum of Rademacher functions. Bull. Amer. Math. Soc., vol. 51 (1945).

H. Robbins, On the sum of random number of random variables. Bull. Amer. Math. Soc., vol. 54 (1948).

S. Takahashi, On the central limit theorem (under the press).

- 1°. $\{X_i\}$ is an independent sequence.
- 2°. $\{X_i\}$ satisfies the central limit theorem (I).
- 3°. For each i , the set of values which X_i takes is at most enumerable.

Let α_k^i be the values which X_i takes. Put $P_k^i = P[X_i = \alpha_k^i]$, $A_k^i = [\omega; X_i = \alpha_k^i]$ and F° the smallest Borel field which includes all the sets A_k^i $k = 1, 2, \dots, i = 1, 2, \dots$. We assume that P_k^i is a non-increasing sequence of k for each i . Then (Ω, F°, P) is also a probability space. For any sequence (finite or infinite) of integers i_1, i_2, \dots, i_n we define the set

$$A_{i_1, i_2, \dots, i_n} = \bigcap_{l=1}^n A_{i_l}^{i_l}.$$

Then, from the independency of $\{X_i\}$

$$\begin{aligned} P(A_{i_1, i_2, \dots, i_n}) &= P\left(\bigcap_{l=1}^n A_{i_l}^{i_l}\right) = P\left(\bigcap_{l=1}^n X_l = \alpha_{i_l}^{i_l}\right) \\ &= \prod_{l=1}^n P(X_l = \alpha_{i_l}^{i_l}) = \prod_{l=1}^n P_{i_l}^{i_l} \end{aligned} \tag{II}$$

If for some infinite sequence $i_1, i_2, \dots, i_n, \dots$, $P(A_{i_1, i_2, \dots, i_n, \dots}) = P > 0$, then $\lim_{n \rightarrow \infty} P\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n X_i = \sqrt{\frac{1}{n}} \sum_{i=1}^n \alpha_{i_l}^{i_l}\right) = P$. This contradicts with the assumption 2°.

Hence for any infinite sequence $i_1, i_2, \dots, i_m, \dots$

$$P(A_{i_1, i_2, \dots, i_n, \dots}) = 0.$$

(A): For any two sets A_{i_1, i_2, \dots, i_n} and A_{j_1, j_2, \dots, j_m} , it is seen:

- 1°. If there exists at least one k such that $k \leq m, n$ and $i_k \neq j_k$, then $A_{i_1, i_2, \dots, i_n} \cap A_{j_1, j_2, \dots, j_m} = \emptyset$, where \emptyset denotes the empty set.
- 2°. If $m = n$ and $i_k = j_k$ for all $k \leq m = n$, then

$$A_{i_1, i_2, \dots, i_n} = A_{j_1, j_2, \dots, j_m}.$$

- 3°. If $m = n$ and $i_k = j_k$ for all $k \leq m$, then

$$A_{i_1, i_2, \dots, i_n} \subset A_{j_1, j_2, \dots, j_m}.$$

Next we define a probability space (T, B, m) as follows:

- 1°. T is the interval $[0, 1[$.
- 2°. B is the class of B -measurable sets in $[0, 1[$.
- 3°. m is Lebesgue measure.

In T we define the set E_{i_1, i_2, \dots, i_n} for any sequence of integers i_1, i_2, \dots, i_n .

$$1^\circ. E_{i_1} = \left[t : \sum_{K=1}^{i_1-1} P_K^1 \leq t < \sum_{K=1}^{i_1} P_K^1 \right].$$

2°. If $E_{i_1, i_2, \dots, i_{n-1}}$ has been defined and the interval $[d_1, d_2[$ denotes then

$$E_{i_1, i_2, \dots, i_n} = \left[t; d_1 + (d_2 - d_1) \sum_{k=1}^{i_{n-1}} P_k^{i_{n-1}} \leq t < d_1 + (d_2 - d_1) \sum_{k=1}^{i_n} P_k^{i_n} \right].$$

From the above construction and (II)

$$m(E_{i_1, i_2, \dots, i_n}) = \prod_{i=1}^n P_{i_i}^i = P(A_{i_1, i_2, \dots, i_n}) \text{ for all } i_1, i_2, \dots, i_n. \quad (IV)$$

For any infinite sequence $i_1, i_2, \dots, i_n, \dots$

$$m(E_{i_1, i_2, \dots, i_n, \dots}) = 0. \quad (V)$$

By (V), it is seen that the smallest Borel field which includes all the sets E_{i_1, i_2, \dots, i_n} is identical with the class B .

(B): For any two sets E_{i_1, i_2, \dots, i_n} and E_{j_1, j_2, \dots, j_m} , it is seen :

1°. If there exists at least one k such that $i_k \neq j_k, k \leq m, n$, then

$$E_{i_1, i_2, \dots, i_n} \cap E_{j_1, j_2, \dots, j_m} = \emptyset.$$

2°. If $m = n$ and $i_k = j_k$ for all $k \leq m = n$, then

$$E_{i_1, i_2, \dots, i_n} = E_{j_1, j_2, \dots, j_m}.$$

3°. If $m < n$ and $i_k = j_k$ for all $k \leq m$, then

$$E_{i_1, i_2, \dots, i_n} \subset E_{j_1, j_2, \dots, j_m}.$$

By (A) and (B) we can define a transformation φ from B to F° as follows :

1°. $\varphi(\emptyset) = \emptyset$.

2°. $\varphi(E_{i_1, i_2, \dots, i_n}) = A_{i_1, i_2, \dots, i_n}$ for all i_1, i_2, \dots, i_n .

3°. $\varphi\left(\bigcup_{i=1}^{\infty} E_i\right) = \bigcup_{i=1}^{\infty} \varphi(E_i), \quad E_i \in B$.

From 1°-3° it follows that

4°. $\varphi(T) = \Omega$ and $\varphi(E) = \varphi(T) - \varphi(E')$ where E' denotes the complement of E .

5°. $\varphi\left(\bigcap_{i=1}^{\infty} E_i\right) = \bigcap_{i=1}^{\infty} \varphi(E_i), \quad E_i \in B$.

Hence φ is a one to one correspondence between the sets of B and sets of F° .

By (IV) and the properties of φ , it follows that ;

1°. if $E \in B$ then $\varphi(E) \in F^\circ$ and $m(E) = P(\varphi(E))$,

2°. if $E' \in F^\circ$ then $\varphi^{-1}(E') \in B$ and $m(\varphi^{-1}(E')) = P(E')$.

We define a sequence of random variables $\{Y_i(t)\} i = 1, 2, \dots$ as follows, if $t \in E_{k_1, k_2, \dots, k_i}$ then $Y_i(t) = \alpha_{k_i}^i$ for all $k_1, k_2, \dots, k_i, i = 1, 2, \dots$.

Thus defined sequence $\{Y_i(t)\}$ is independent. For, $m\left(\bigcap_{i=1}^n Y_i(t)\right)$

$$= \alpha_{k_i}^i) = m(E_{k_1, k_2, \dots, k_n}) = \prod_{i=1}^n P_{k_i}^i. \text{ On the other hand}$$

$$[t : Y_i(t) = \alpha_{k_i}^i] = \bigcup_{k_1, k_2, \dots, k_{i-1}} E_{k_1, k_2, \dots, k_i},$$

where $\bigcup_{k_1, k_2, \dots, k_{i-1}}$ denotes the summation for all possible combinations of k_1, k_2, \dots, k_{i-1} . According to (B) and (IV)

$$\begin{aligned} m[t : Y_i(t) = \alpha_{k_i}^i] &= \sum_{k_1, k_2, \dots, k_{i-1}} m(E_{k_1, k_2, \dots, k_{i-1}, k_i}) \\ &= \sum_{k_1, k_2, \dots, k_{i-1}} P_{k_1}^1 \prod_{l=1}^{i-1} P_{k_l}^l = P_{k_i}^i. \end{aligned}$$

Hence $m\left(\bigcap_{i=1}^n Y_i(t) = a_{k_i}^i\right) = \prod_{i=1}^n m(Y_i(t) = a_{k_i}^i)$ for all n . By the fact that $Y_i(t)$ takes on E_{k_1, k_2, \dots, k_i} the same value as that X_i takes on A_{k_1, k_2, \dots, k_i} and by the properties of φ it is seen for all n that

$$\varphi\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a\right) = \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n X_i(\omega) \leq a\right).$$

Hence $m\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a\right) = P\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n X_i(\omega) \leq a\right),$

so $\lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a\right) = \lim_{n \rightarrow \infty} P\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n X_i(\omega) \leq a\right) = G(a).$

Lemma 1. If $E \in B$, then

$$\lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, E\right) = m(E) G(a).$$

Proof. For any finite sequence of integers i_1, i_2, \dots, i_l

$$\begin{aligned} & \lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, E_{i_1, i_2, \dots, i_l}\right) \\ &= \lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \left(\sum_{i=1}^l + \sum_{i=l+1}^n\right) Y_i(t) \leq a, E_{i_1, i_2, \dots, i_l}\right) \\ &= \lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=l+1}^n Y_i(t) \leq a, E_{i_1, i_2, \dots, i_l}\right) \\ &= \lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=l+1}^n Y_i(t) \leq a, \bigcap_{k=1}^l X_k = a_{i_k}^{i_k}\right) \\ &= \lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=l+1}^n Y_i(t) \leq a\right) m\left(\bigcap_{k=1}^l X_k = a_{i_k}^{i_k}\right) \\ &= \lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a\right) m(E_{i_1, i_2, \dots, i_l}) = G(a) m(E_{i_1, i_2, \dots, i_l}). \end{aligned}$$

Now let M denote the family of sets E which satisfy the following relation

$$\lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, E\right) = m(E) G(a).$$

Then, about M it is seen that:

1°. M includes E_{i_1, i_2, \dots, i_l} for all i_1, i_2, \dots, i_l ($l = 1, 2, \dots$). For a finite sequence i_1, i_2, \dots, i_l we have proved above, but for infinite sequence it is evident from (V).

2°. If $E \subset E'$ and $E, E' \in M$, then $E' - E \in M$.

$$\begin{aligned} & \text{For, } \lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, E' - E\right) \\ &= \lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, E'\right) - \lim_{n \rightarrow \infty} m\left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, E\right) \\ &= G(a) m(E') - G(a) m(E) = G(a) m(E' - E). \end{aligned}$$

3°. If $E = \bigcup_{j=1}^{\infty} A_j$, $A_j \in M$ and $A_j, A_{j'}$ are non-overlapping ($j \neq j'$), then $E \in M$.

$$\begin{aligned} \text{For, } \lim_{n \rightarrow \infty} m \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, E \right) \\ = \lim_{n \rightarrow \infty} \sum_{j=1}^n m \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, A_j \right). \end{aligned} \tag{VI}$$

For all n , $0 \leq m \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, A_j \right) \leq m(A_j)$

and $\sum_{j=1}^{\infty} m(A_j) = m(E)$.

Hence the convergence of $\sum_{j=1}^{\infty} m \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, A_j \right)$ is uniform with respect to n . So we can exchange the order of \lim and \sum of (VI). We have therefore

$$(VI) = \sum_{j=1}^{\infty} \lim_{n \rightarrow \infty} m \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, A_j \right) = \sum_{j=1}^{\infty} G(a) m(A_j) = G(a) m(E).$$

By 1°-3° and the fact that m denotes Lebesgue measure, it follows that M includes all sets of B .

We complete F^{σ} with respect to the measure P and denote it by \bar{F}^{σ} .

Lemma 2. If $E \in F^{\sigma}$,

then $\lim_{n \rightarrow \infty} P \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n X_i(\omega) \leq a, E \right) = P(E) G(a)$

Proof: It is sufficient to prove this lemma for the case $E \in F^{\sigma}$.

For all n , $\varphi^{-1} \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n X_i(\omega) \leq a, E \right) = \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, \varphi^{-1}(E) \right)$.

$$\begin{aligned} \text{Hence } \lim_{n \rightarrow \infty} P \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n X_i(\omega) \leq a, E \right) \\ = \lim_{n \rightarrow \infty} m \left(\sqrt{\frac{1}{n}} \sum_{i=1}^n Y_i(t) \leq a, \varphi^{-1}(E) \right) \\ = G(a) m(\varphi^{-1}(E)) = G(a) P(E). \end{aligned}$$

§ 3. In this paragraph, we prove the theorem mentioned in § 1. Let $\{X_i\}$ be a sequence of independent random variables satisfying the central limit theorem (I).

Now, let $(h_{i,k})$ ($i = 1, 2, \dots, k = 0, \pm 1, \pm 2, \dots$) be a sequence of real numbers satisfying the following conditions:

1°. $h_{i,k} > 0$ for $k > 0$, $h_{i,k} < 0$ for $k < 0$.

$h_{i,k} = 0$ for $k = 0$.

2°. $(h_{i,k}) \leq h_i$ for all k and $\sum_{i=1}^n h_i = O(\sqrt{n})$ ($n \rightarrow \infty$).

3°. $\sum_{k=0}^{\infty} h_{i,k} = +\infty$ $\sum_{k=0}^{-\infty} h_{i,k} = -\infty$ for each i .

Using this sequence $(h_{i,k})$, we define a sequence of random vari-

ables $\{Z_i(\omega)\}$ ($i = 1, 2, \dots$) as follows :

$$Z_i(\omega) = \sum_{k=0}^n h_{i,k} \quad \text{if } \omega \in E_{i,n} = \left[\omega ; \sum_{k=0}^n h_{i,k} \leq X_i(\omega) < \sum_{k=0}^{n+1} h_{i,k} \right]$$

for $n = 0, \pm 1, \pm 2, \dots, i = 1, 2, \dots$.

Thus defined sequence $\{Z_i(\omega)\}$ $i = 1, 2, \dots$ is independent.

For,
$$P\left(\bigcap_{i=1}^n Z_i(\omega) = \sum_{k=0}^{n_i} h_{i,k}\right) = P\left(\bigcap_{i=1}^n \left[\sum_{k=0}^{n_i} h_{i,k} \leq X_i(\omega) < \sum_{k=0}^{n_i+1} h_{i,k}\right]\right)$$

$$= \prod_{i=1}^n P\left[\sum_{k=0}^{n_i} h_{i,k} \leq X_i(\omega) < \sum_{k=0}^{n_i+1} h_{i,k}\right] = \prod_{i=1}^n P\left(Z_i(\omega) = \sum_{k=0}^{n_i} h_{i,k}\right).$$

$\{Z_i(\omega)\}$ satisfies the central limit theorem (I).

For,
$$P\left(\left|\sum_{i=1}^n (X_i - Z_i) / \sqrt{n}\right| \leq \varepsilon\right) \geq P\left(\sum_{i=1}^n h_i / \sqrt{n} < \varepsilon\right)$$

for all n . On the other hand $\sum_{i=1}^n h_i = o(\sqrt{n})$. Hence

$$\lim_{n \rightarrow \infty} P\left(\left|\sum_{i=1}^n (X_i - X_i) / \sqrt{n}\right| > \varepsilon\right) \geq \lim_{n \rightarrow \infty} P\left(\sum_{i=1}^n h_i / \sqrt{n} < \varepsilon\right) = 1.$$

So
$$\frac{1}{\sqrt{n}} \sum_{i=1}^n (X_i - Z_i) \rightarrow 0 \quad (n \rightarrow \infty) \text{ in probability.} \tag{VII}$$

Therefore
$$\begin{aligned} \lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \sum_{i=1}^n Z_i(\omega) \leq a\right) \\ = \lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \sum_{i=1}^n (Z_i(\omega) - X_i(\omega) + X_i(\omega)) \leq a\right) \\ = \lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \sum_{i=1}^n X_i(\omega) \leq a\right) = G(a). \end{aligned}$$

Now, let F^* be the smallest Borel field which includes $E_{i,n}$ for $n = 0, \pm 1, \pm 2, \dots, i = 1, 2, \dots$. Then by Lemma 2 $E \in F^*$, and

$$\lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \sum_{i=1}^n Z_i(\omega) \leq a, E\right) = P(E) G(a). \tag{VIII}$$

By (VIII) and (VII), if $E \in F^*$,

$$\lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \sum_{i=1}^n X_i(\omega) \leq a, E\right) = P(E) G(a). \tag{IX}$$

Next, we state Theorem 1 by using the definitions of \bar{F} and \tilde{F} mentioned in § 1.

Theorem 1. If $E \in \tilde{F}$, then

$$\lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \sum_{i=1}^n X_i(\omega) \leq a, E\right) = P(E) G(a).$$

Proof. It is sufficient to prove this theorem for the case where $E \in \bar{F}$. To define any set E belonging to \bar{F} , it is sufficient to consider at most enumerable sets of the type of $[\omega ; a \leq X_i(\omega) < b]$ for each i . So we can choose $(h_{i,k})$ such that the set E belongs to F^* determined by $(h_{i,k})$. From (IX), Theorem 1 holds for this set.

§ 4 Theorem 2. Let $g(\omega)$ be a non-negative \bar{F} -measurable function. Then

$$\lim_{n \rightarrow \infty} P \left(\sqrt{\frac{1}{n}} \left| \sum_{i=1}^n X_i(\omega) \right| \leq g(\omega) \right) = \sqrt{\frac{1}{2\pi}} \int_{\Omega} P(d\omega) \int_{-g(\omega)}^{g(\omega)} e^{-u^2/2} du.$$

Proof. It is sufficient to prove this theorem for the case where $g(\omega)$ is an \bar{F} -measurable simple function. Let $g(\omega)$ be a simple function such that $g(\omega) = \{a_i, F_i\}$ ($i = 1, 2, \dots$)

$$\begin{aligned} & \lim_{n \rightarrow \infty} P \left(\sqrt{\frac{1}{n}} \left| \sum_{i=1}^n X_i(\omega) \right| \leq g(\omega) \right) \\ &= \lim_{n \rightarrow \infty} \sum_{j=1}^{\infty} P \left(\sqrt{\frac{1}{n}} \left| \sum_{i=1}^n X_i(\omega) \right| \leq a_j, \quad g(\omega) = a_j \right). \end{aligned} \tag{X}$$

We can exchange the order of \lim and \sum of (X) by the same way as in 3° of Lemma 1. We have

$$\begin{aligned} \text{(X)} &= \sum_{j=1}^{\infty} \lim_{n \rightarrow \infty} P \left(\sqrt{\frac{1}{n}} \left| \sum_{i=1}^n X_i(\omega) \right| \leq a_j, \quad g(\omega) = a_j \right) \\ &= \sqrt{\frac{1}{2\pi}} \sum_{j=1}^{\infty} P(g(\omega) = a_j) \int_{-a_j}^{a_j} e^{-u^2/2} du = \sqrt{\frac{1}{2\pi}} \int_{\Omega} P(d\omega) \int_{-g(\omega)}^{g(\omega)} e^{-u^2/2} du. \end{aligned}$$

Two measurable functions $g_1(\omega)$ and $g_2(\omega)$ have the distribution functions $G_1(u)$ and $G_2(u)$ respectively, and if $G_1(u) = G_2(u)$ holds for the continuous points, then it is said that $g_1(\omega)$ and $g_2(\omega)$ have the same distribution function $G(u)$ (or $G_2(u)$).

Corollary 1. Let $g_1(\omega)$ and $g_2(\omega)$ be non-negative \bar{F} -measurable functions having the same distribution function $\bar{G}(u)$. Then

$$\begin{aligned} & \lim_{n \rightarrow \infty} P \left(\sqrt{\frac{1}{n}} \left| \sum_{i=1}^n X_i(\omega) \right| \leq g_1(\omega) \right) \\ &= \lim_{n \rightarrow \infty} P \left(\sqrt{\frac{1}{n}} \left| \sum_{i=1}^n X_i(\omega) \right| \leq g_2(\omega) \right) = \sqrt{\frac{1}{2\pi}} \int_0^{\infty} d\bar{G}(v) \int_{-v}^v e^{-u^2/2} du. \end{aligned}$$

Proof. From Theorem 2.

$$\begin{aligned} & \lim_{n \rightarrow \infty} P \left(\sqrt{\frac{1}{n}} \left| \sum_{i=1}^n X_i(\omega) \right| \leq g_1(\omega) \right) = \sqrt{\frac{1}{2\pi}} \int_{\Omega} P(d\omega) \int_{-g_1(\omega)}^{g_1(\omega)} e^{-u^2/2} du \\ &= \sqrt{\frac{1}{2\pi}} \int_0^{\infty} d\bar{G}(v) \int_{-v}^v e^{-u^2/2} du = \sqrt{\frac{1}{2\pi}} \int_{\Omega} P(d\omega) \int_{-g_2(\omega)}^{g_2(\omega)} e^{-u^2/2} du \\ &= \lim_{n \rightarrow \infty} P \left(\sqrt{\frac{1}{n}} \left| \sum_{i=1}^n X_i(\omega) \right| \leq g_2(\omega) \right). \end{aligned}$$

Next consider the second generalization.

Theorem 3. Let $N_n(\omega) = nN(\omega) + Q_n(\omega)$, where $N_n(\omega)$ and $N(\omega)$ are P -measurable functions which takes non-negative integers, and $Q_n(\omega) = O(\sqrt{n})$. If $N(\omega)$ is \bar{F} -measurable, then

$$\begin{aligned} & \lim_{n \rightarrow \infty} P \left(\sqrt{\frac{1}{n}} \left| \sum_{i=0}^{N_n(\omega)} X_i(\omega) \right| \leq a \right) \\ &= \sum_{M=0}^{\infty} P(N(\omega) = M) \sqrt{\frac{1}{2\pi}} \int_{-aM^{-1/2}}^{aM^{-1/2}} e^{-u^2/2} du \\ &= \int_{\Omega} P(d\omega) \sqrt{\frac{1}{2\pi}} \int_{-aN(\omega)^{-1/2}}^{aN(\omega)^{-1/2}} e^{-u^2/2} du. \end{aligned}$$

Proof. Let us put $\frac{1}{\sqrt{n}} \sum_{i=0}^{N_n(\omega)} X_i(\omega) = \frac{1}{\sqrt{n}} \sum_{i=0}^{nN(\omega)} X_i(\omega) + S_n(\omega)$.

First, we prove that $S_n(\omega)$ converges in probability to 0. For, $\lim P(|S_n(\omega)| > \varepsilon)$

$$\begin{aligned} &= \lim_{n \rightarrow \infty} P(|S_n(\omega)| > \varepsilon, \bigcup_{M=0}^{\infty} (N(\omega) = M, \bigcup_{k=-\infty}^{\infty} Q_n(\omega) = k)) \\ &= \sum_{M=0}^{\infty} \lim_{n \rightarrow \infty} P(|S_n(\omega)| > \varepsilon, N(\omega) = M, \bigcup_{k=-\infty}^{\infty} Q_n(\omega) = k) \\ &\leq \sum_{M=0}^{\infty} \lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \left| \sum' X_i(\omega) \right| > \varepsilon\right), \end{aligned}$$

where \sum' denotes the summation from nM to $nM + Q_n(\omega)$ or from $nM - |Q_n(\omega)|$ to nM according as $Q_n(\omega) \geq 0$ or $Q_n(\omega) < 0$. On the other hand $\{X_i\} i = 1, 2, \dots$ satisfies the central limit theorem (I) and $Q_n(\omega) = 0$ ($n^{1/2}$) as $n \rightarrow \infty$. Hence

$$\lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \left| \sum' X_i(\omega) \right| > \varepsilon\right) = 0.$$

$$\begin{aligned} \text{So } \lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \left| \sum_{i=0}^{N_n(\omega)} X_i(\omega) \right| \leq a\right) &= \lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \left| \sum_{i=0}^{nN(\omega)} X_i(\omega) \right| \leq a\right) \\ &= \lim_{n \rightarrow \infty} \sum_{M=0}^{\infty} P\left(\frac{1}{\sqrt{n}} \left| \sum_{i=0}^{nM} X_i(\omega) \right| \leq a, N(\omega) = M\right) \\ &= \lim_{n \rightarrow \infty} \sum_{M=0}^{\infty} P\left(\frac{1}{\sqrt{nM}} \left| \sum_{i=0}^{nM} X_i(\omega) \right| \leq aM^{-1/2}, N(\omega) = M\right) \\ &= \sum_{M=0}^{\infty} \lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{nM}} \left| \sum_{i=0}^{nM} X_i(\omega) \right| \leq aM^{-1/2}, N(\omega) = M\right) \\ &= \sum_{M=0}^{\infty} P(N(\omega) = M) \frac{1}{\sqrt{2\pi}} \int_{-aM^{-1/2}}^{aM^{-1/2}} e^{-u^2/2} du \\ &= \frac{1}{\sqrt{2\pi}} \int_{\Omega} P(d\omega) \int_{-aN(\omega)^{-1/2}}^{aN(\omega)^{-1/2}} e^{-u^2/2} du. \end{aligned}$$

In Theorem 3, when $M = N(\omega) = 0$, $\pm aM^{-1/2}$ and $\pm aN(\omega)^{-1/2}$ denote $\pm \infty$.

Corollary 2. If $N'_n(\omega) = nN'(\omega) + Q'_n(\omega)$ and $N''_n(\omega) = nN''(\omega) + Q''_n(\omega)$ satisfy the conditions of $N_n(\omega)$, $N(\omega)$ and $Q_n(\omega)$ in Theorem 3, and $N'(\omega)$, $N''(\omega)$ have the same distribution function

$$\begin{aligned} \bar{G}(u), \text{ then } \lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \left| \sum_{i=0}^{N'_n(\omega)} X_i(\omega) \right| \leq a\right) \\ = \lim_{n \rightarrow \infty} P\left(\frac{1}{\sqrt{n}} \left| \sum_{i=0}^{N''_n(\omega)} X_i(\omega) \right| \leq a\right) = \frac{1}{\sqrt{2n}} \int_0^{\infty} d\bar{G}(v) \int_{-av^{-1/2}}^{av^{-1/2}} e^{-u^2/2} du. \end{aligned}$$

Proof is evident from Theorem 3.