## THE BROWNIAN ESCAPE PROCESS1

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Let X be the Brownian motion process in  $\mathbb{R}^d$ , d > 3 with X(0) = 0. Let  $L_r$  be the last exit time of X from the ball of radius r centered at the origin. Then  $(L_r)$  has independent increments and we compute the distribution of  $L_r$ . When d = 3 this yields a simple proof of a recent result of Pitman.

Let  $X = (X_t)$  be the standard Brownian motion process in  $\mathbb{R}^d$ . If t > 0, define

$$(1) L_r = \sup\{t: |X_t| \le r\}.$$

We shall assume that  $d \ge 3$ . Then almost surely  $|X_t| \to \infty$  as  $t \to \infty$ , and so  $L_r$  is finite almost surely. If  $P = P^0$  is the law of the Brownian motion starting from the origin we shall call the process  $(L_r, r \ge 0; P)$  the Brownian escape process. Clearly  $r \to L_r$  is strictly increasing and  $L_0 = 0$ .

Recently Pitman [5] has shown that when d = 3,  $(L_r, P)$  is a stable subordinator of index  $\frac{1}{2}$  and rate  $2^{\frac{1}{2}}$ ; that is,  $(L_r, P)$  has stationary independent increments and the Laplace transform of its distribution is given by

$$(2) E(e^{-\alpha L_r}) = e^{-r(2\alpha)^{\frac{1}{2}}}.$$

Pitman derives this as a corollary of a deep result connecting one-dimensional Brownian motion and the three-dimensional Bessel process. The purpose of this note is to give a simple direct proof of Pitman's result and to study the escape process when d > 3.

Here is our main result.

THEOREM. Let  $d \ge 3$ . Then the process  $(L_r, P)$  is continuous in probability and has independent increments. If r > 0,  $L_r$  has a continuous density given by (t > 0)

(3) 
$$P[L_r \in dt] = r^{d-2} \left[ 2^{(d-2)/2} \Gamma\left(\frac{d-2}{2}\right) t^{d/2} \right]^{-1} e^{-r^2/2t} dt.$$

The escape process has stationary increments if and only if d = 3.

REMARK. If d = 3, (3) is the familiar density of the stable subordinator whose Laplace transform is given by (2). Although  $(L_r)$  does not have stationary increments when d > 3, it is trying in some average sense to "look like" a stable

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subordinator of index  $\frac{1}{2}$ . More precisely we shall see that

(4) 
$$\int_0^\infty E(e^{-\alpha L_r}) dr = \left[c(d)(2\alpha)^{\frac{1}{2}}\right]^{-1}$$

(5) 
$$c(d) = \frac{\Gamma\left(\frac{d-2}{2}\right)}{\pi^{\frac{1}{2}}\Gamma\left(\frac{d-1}{2}\right)}.$$

Of course, (4) is exactly what one would obtain if  $(L_r)$  were a stable subordinator of index  $\frac{1}{2}$  and rate  $c(d) \cdot 2^{\frac{1}{2}}$ .

PROOF. Using the scaling property of the Brownian motion (i.e., if c > 0,  $c^{-\frac{1}{2}}X(ct)$  is again a Brownian motion), it is immediate that  $L_r$  has the same distribution as  $r^2L_1$  under  $P = P^0$ . Since  $L_r$  is increasing this implies that  $r \to L_r$  is continuous at each fixed r almost surely, and hence it is continuous in probability.

To see that  $(L_r)$  has independent increments we define  $R_t = |X_t|$ . Then  $(R_t; t \ge 0)$  is a continuous strong Markov process on  $\mathbb{R}^+ = [0, \infty)$  (it is the *d*-dimensional Bessel process) and

$$(6) L_{r} = \sup\{t: R_{t} \leqslant r\}$$

is the "last exit" time of  $(R_t)$  from the interval [0, r]. Let  $(\mathcal{G}_t)_{t\geqslant 0}$  denote the usual (completed) fields of the diffusion  $(R_t)$ . A by now standard result (see, e.g., Theorem 37 of [1]) states that under  $P=P^0$  the post  $L_r$  process  $(R_{L_r+t})_{t>0}$  is conditionally independent of  $\mathcal{G}_{L_r}$  given  $R_{L_r}$ . But by the continuity of the paths  $R_{L_r}=r$  almost surely P and so in the present case  $(R_{L_r+t})_{t>0}$  is independent of  $\mathcal{G}_{L_r}$ . Here  $\mathcal{G}_{L_r}$  is the usual field associated with a last exit time. If r < s, then

(7) 
$$L_{s} - L_{r} = \sup\{t > 0: R_{L_{r}+t} \leq s\},$$

and so  $L_s - L_r$  is independent of  $\mathcal{G}_{L_r}$ . Also  $R_{t \wedge L_r}$  is  $\mathcal{G}_{t \wedge L_r}$  measurable. But  $t \wedge L_r \leq L_r$  and is  $\mathcal{G}_{L_r}$  measurable, and consequently  $\mathcal{G}_{t \wedge L_r} \subset \mathcal{G}_{L_r}$ . If a < r, then

$$L_a = \sup\{t: R_{t \wedge L} \leq a\},\,$$

and so each  $L_a$  with  $a \le r$  is  $\mathcal{G}_{L_r}$  measurable. As a result the process  $(L_r, P)$  has independent increments.

Before coming to the proof of (3) we introduce the process

(8) 
$$F_{t} = \inf\{|X_{s}|: s > t\}, \quad t \geq 0.$$

Clearly  $t \to F_t$  is increasing and continuous and  $\lim_{t \to \infty} F_t = \infty$  almost surely. Note that  $F_t = F_0 \circ \theta_t$  and  $F_t = \inf\{|X_s|: s \ge t\}$ . If 0 < a < b, let  $T_{ab} = \inf\{t: |X_t| \notin (a, b)\}$ . It is well known and easily checked that if a < |x| < b, then

(9) 
$$P^{x}[|X(T_{ab})| = b] = \left[ \left( \frac{b}{a} \right)^{d-2} - 1 \right]^{-1} \left[ \left( \frac{b}{a} \right)^{d-2} - \left( \frac{b}{|x|} \right)^{d-2} \right].$$

Letting  $b \to \infty$  in (9) we obtain

(10) 
$$P^{x}[F_{0} > a] = 1 - \left(\frac{a}{|x|}\right)^{d-2} \qquad 0 < a < |x|.$$

We are now prepared to prove (3). Clearly

$$P^{0}[L_{r} < t] = P^{0}[F_{t} > r] = E^{0}\{P^{X(t)}[F_{0} > r]\},$$

and using (10) we have

$$P^{0}[L_{r} < t] = (2\pi t)^{-d/2} \int_{|x| > r} e^{-|x|^{2}/2t} \left[ 1 - \left(\frac{r}{|x|}\right)^{d-2} \right] dx$$

$$= \frac{2}{\Gamma(d/2)(2t)^{d/2}} \int_{r}^{\infty} u \left[ u^{d-2} - r^{d-2} \right] e^{-u^{2}/2t} du$$

$$= \left[ \Gamma\left(\frac{d}{2}\right) \right]^{-1} \left\{ \int_{r^{2}/2t}^{\infty} v^{(d-2)/2} e^{-v} dv - r^{d-2} (2t)^{1-d/2} e^{-r^{2}/2t} \right\}.$$

Differentiating with respect to t we obtain (3).

Let us denote the density in (3) by  $g_d(r, t)$ . Consulting a table of Laplace transforms one finds

(11) 
$$\int_0^\infty e^{-\alpha t} g_3(r, t) dt = e^{-r(2\alpha)^{\frac{1}{2}}}$$

$$(12) \int_0^\infty e^{-\alpha t} g_d(r, t) dt = \left[ 2^{(d-4)/2} \Gamma\left(\frac{d-2}{2}\right) \right]^{-1} \left( r(2\alpha)^{\frac{1}{2}} \right)^{(d-2)/2} K_{(d-2)/2} \left( r(2\alpha)^{\frac{1}{2}} \right)$$

where  $K_r$  is the usual modified Bessel function of the third kind. Of course, when d=3, (12) reduces to (11). Also, if d=5, (12) reduces to  $(1+r(2\alpha)^{\frac{1}{2}})e^{-r(2\alpha)^{\frac{1}{2}}}$ . Either calculating from scratch using (10), or by using (27) on page 51 of [2], we obtain (4) and (5) by integrating (12) in r over  $(0, \infty)$ . Therefore if  $(L_r)$  has stationary independent increments, then one must have

$$\int_{0}^{\infty} e^{-\alpha t} g_{d}(r, t) dt = e^{-rc(d)(2\alpha)^{\frac{1}{2}}}$$

and one easily checks (compare derivatives at  $\alpha = 0$ ) that this contradicts (12) if d > 3. This completes the proof of the theorem.

REMARK. Since any increment of a process with independent increments that is continuous in probability has an infinitely divisible distribution, the density  $g_d(r, t)$  is infinitely divisible. Also if  $0 \le a < b$ , the increment  $L_b - L_a$  is a positive random variable having an infinitely divisible distribution whose Laplace transform, in view of (12), is given by

(13) 
$$\frac{b^{\nu}K_{\nu}(b(2\alpha)^{\frac{1}{2}})}{a^{\nu}K_{\nu}(a(2\alpha)^{\frac{1}{2}})}, \qquad \nu = \frac{d-2}{2}, d \geqslant 3.$$

One can invert this explicitly when d = 5, but the resulting density does not seem to be particularly interesting. On the other hand, starting with the Bessel diffusion

Y with generator  $\frac{1}{2}(D^2 + ((2\nu + 1)/x)D)$  on  $x \ge 0$  where D = d/dx, the same argument shows that if

(14) 
$$f_{\nu}(t) = \left[2^{\nu}\Gamma(\nu)t^{\nu+1}\right]^{-1}e^{-1/2t}, \quad t > 0,$$

then  $a^{-2}f_{\nu}(t/a^2)$  is the density of the escape process  $L_a = \sup\{t: Y_t \le a\}$ . In particular  $f_{\nu}$  is infinitely divisible for  $\nu > 0$ , and (13) is the Laplace transform of an infinitely divisible distribution for  $\nu > 0$ . The fact that  $f_{\nu}$  is an infinitely divisible density may also be obtained as a limiting case of a recently announced result of Ismail and Kelker on the infinite divisibility of the F-distribution. See [4].

It is interesting to compare these results to the familiar first passage results. Let  $T_r = \inf\{t: |X_t| \ge r\}$ . Then using the strong Markov property  $(T_r, P)$  is an increasing process with independent increments for  $d \ge 1$ . Using a standard martingale argument, or the generator if we deal with the Bessel process (see, e.g., [3]), one has

(15) 
$$E^{0}(e^{-\alpha T(r)}) = \left(r(2\alpha)^{\frac{1}{2}}\right)^{\nu} \left[2^{\nu}\Gamma(\nu+1)I_{\nu}\left(r(2\alpha)^{\frac{1}{2}}\right)\right]^{-1}$$

where  $\nu = (d-2)/2$ ,  $d \ge 1$  or  $\nu \ge -\frac{1}{2}$ . Here  $I_{\nu}$  is the modified Bessel function of the first kind. If a < b the increment  $T_b - T_a$  has an infinitely divisible distribution whose Laplace transform is given by

(16) 
$$\frac{b^{\nu}I_{\nu}\left(a(2\alpha)^{\frac{1}{2}}\right)}{a^{\nu}I_{\nu}\left(b(2\alpha)^{\frac{1}{2}}\right)}.$$

If d = 1, (15) reduces to  $(\cosh r(2\alpha)^{\frac{1}{2}})^{-1}$ . It is interesting to note that the densities for the last exit (escape) process  $L_r$  have simple expressions while those of the first passage process  $T_r$  seem to be more complicated. For example, when d = 1, the density of  $T_r$  may be expressed in terms of theta functions.

The expression (3) for the density of  $L_r$  is also an easy consequence of formula (13) in [6].

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