A THEOREM ON THE GALTON-WATSON PROCESS1

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In this note we will prove a theorem concerning a limiting distribution associated with the Galton-Watson process. Specifically, we consider a stochastic process, $\{Z_n : n = 0, 1, \dots\}$, with the following properties:

- (1) $Z_0 = 1$;
- (2) if P denotes the probability measure associated with the process, then $P(Z_1 = i) = p_i$, $i = 0, 1, \cdots$. Moreover the process is a Markoff process with transition probabilities,

$$P_{ij} = P(Z_{n+1} = j \mid Z_n = i) = \sum_{k_1 + k_2 + \dots + k_i = j} p_{k_1} \cdot p_{k_2} \cdot \dots \cdot p_{k_i},$$

$$i = 1, 2, \dots, j = 0, 1, \dots, P_{0j} = 0, j = 1, 2, \dots, \text{ and } P_{00} = 1;$$

- (3) $p_i \neq 1$ for all i; and
- (4) $E(Z_1) = m > 1$.

We will show that the random variables, (Z_n/m^n) , $n=0,1,\cdots$, converge a.e. to a random variable, W, whose probability distribution has a jump at the origin and a continuous density function on the set of positive real numbers. Levinson and Harris have proved similar theorems but under more restrictive assumptions and by using quite different arguments. Specifically, Levinson [4] by assuming that $E(Z_1 \log Z_1) < \infty$ and Harris [3] by assuming that $E(Z_1^2) < \infty$ have established our result. Harris has also proved that his assumptions imply convergence in the mean of the (Z_n/m^n) 's, and both Harris and Levinson have proved that their assumptions imply that E(W) = 1 and that P(W = 0) = q < 1, where q is a number to be defined later. In contrast under our assumptions we can only prove that $E(W) \le 1$ and that P(W = 0) = q or 1. However we will show in a forthcoming paper with Harry Kesten that if Assumptions 1 through 4 hold and if P(W = 0) = q, then E(W) = 1. Moreover E(W) = 1 only if $E(Z_1 \log Z_1) < \infty$.

The probability generating function of Z_1 will be denoted f(s) and is defined by the equation, $f(s) = \sum_{k=0}^{\infty} p_k s^k$, on the set of all complex numbers s such that $|s| \leq 1$. The probability generating function of Z_n will be denoted by $f_n(\cdot)$. We will make repeated use of a few facts about the $f_n(\cdot)$'s that are stated briefly below.

- (5) $f_{n+k}(s) = f_n(f_k(s)) = f_k(f_n(s)).$
- (6) There exists a unique real number q such that $0 \le q < 1, f(q) = q$.
- (7) For all $s \in [q, 1)$ we have $1 > s \ge f(s) \ge f_2(s) \ge \cdots \ge f_n(s) \ge q$ with

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 $\lim_n f_n(s) = q$. Similarly for all $s \in [0, q]$ we have $0 \le s \le f(s) \le f_2(s) \le \cdots \le f_n(s) \le q$ with $\lim_n f_n(s) = q$.

(8) The $f_n(\cdot)$'s are all differentiable and convex on the unit interval. Moreover for all $s \in [q, 1), f_n'(s) \leq \{f'(s)\}^n$, and for all $s \in [0, q], f_n'(s) \geq \{f'(s)\}^n$.

Doob has pointed out that the random variables, (Z_n/m^n) , $n=0,1,\cdots$ constitute a martingale that converges a.e. to a random variable W with mean less than or equal to one (Harris [2], p. 13). We will denote the characteristic function of W by $\varphi(it)$ where t varies over the interval $(-\infty, \infty)$. It is easy to show that $\varphi(\cdot)$ must satisfy the functional equation, $\varphi(mit) = f(\varphi(it))$. In particular, $\varphi(m^nit) = f_n(\varphi(it))$ for all n. We will next show that either $\varphi(it) = 1$ identically in t or $|\varphi(it)| < 1$ for all $t \neq 0$.

LEMMA 1. Either $\varphi(it) = 1$ for all t or $|\varphi(it)| < 1$ for all $t \neq 0$.

Proof. Throughout this proof we assume that $\varphi(it)$ is not identically equal to one. We will first show that the equality $|\varphi(it)| = 1$ for all t is impossible. Suppose the contrary to be true. Then $\varphi(it) = e^{+ita}$ for some a (Loève [5], p. 202). Moreover, when using the functional equation, $\varphi(mit) = f(\varphi(it))$, we find that $e^{\pm imta} = \sum_{k=0}^{\infty} p_k e^{\pm ikta}$. Hence $\sum_{k=0}^{\infty} p_k (1 - \cos(k - m)ta = 0$ for all t. This implies that m is an integer and that $p_m = 1$ which is contrary to Assumption 3 above.

We will next show that there exists a $\delta > 0$ such that for all $t \varepsilon (-\delta, \delta) - 0$, $|\varphi(it)| < 1$. Suppose not. Then there is a sequence of numbers t_n that converge to zero with the property that $|\varphi(it_n)| = 1$ for all n. We may assume without loss in generality that the t_n 's are all positive. If $|\varphi(it_n)| = 1$, then there exists a number a_n in the interval $[0, 2\pi)$ such that $\varphi(it_n) = e^{+ia_n}$. Moreover, if x is a point of increase of the distribution associated with W, then there exists an integer k_n such that $t_n x = a_n + k_n 2\pi$. Since the t_n 's converge to zero, we conclude that there is a large number N such that $x = (a_n/t_n)$ for all $n \ge N$. This implies that $\varphi(\cdot)$ is degenerate which contradicts our previous result. (The argument used above is due to Levinson [4].)

Finally we will show that $|\varphi(it)| < 1$ for all $t \neq 0$. Let $\delta > 0$ be so small that for all $t \in (-\delta, \delta) - 0$, $|\varphi(it)| < 1$, and let M be an integer so large that $m^{-M} < \delta$. Moreover let $A = [m^{-(M+1)}, m^{-M})$. If t_0 is any positive number greater than δ , then there exists a number $u \in A$ and an integer n such that $t_0 = m^n u$. Hence $|\varphi(it_0)| = |\varphi(im^n u)| = |f_n(\varphi(iu))| \le f_n(|\varphi(iu)|) < 1$. This proves our assertion for all positive t's. Our arguments with only obvious modifications work equally well for negative numbers. Q.E.D.

Lemma 2. If $\varphi(it)$ is not identically equal to one and q=0, then $\lim_{t\to\pm\infty}|\varphi(it)|=0$.

PROOF. To prove this lemma we let δ be a positive number and let A be defined as in the proof of the preceding lemma. Moreover we let $\alpha = \max_{t \in \overline{A}} |\varphi(it)|$, and let N be an integer so large that $f_N(\alpha) < \epsilon$, where ϵ is an arbitrarily chosen positive real number. If t_0 is any number greater than m^{N-M-1} , then there is a number u in A such that $t_0 = m^n u$ for some n. Hence $|\varphi(it_0)| = |\varphi(im^n u)| \leq f_N(|\varphi(iu)|) \leq f_N(|\alpha|) < \epsilon$. The proof that $\lim_{t \to -\infty} |\varphi(it)| = 0$ can be obtained by essentially the same argument. Q.E.D.

We will next show that if q = 0, then $\varphi'(\cdot)$ is absolutely integrable.

LEMMA 3. If $\varphi(it)$ is not identically equal to one, then $\varphi(it)$ is differentiable; and if in addition q = 0, then $\varphi'(it)$ is absolutely integrable.

PROOF. Since $E\left|W\right| \leq 1$, $\varphi(\cdot)$ is obviously differentiable (Doob [1], p. 38). Hence the only thing to prove is that $\varphi'(\cdot)$ is absolutely integrable whenever q=0. To do this we proceed as follows: Since $\lim_{t\to\pm\infty}|\varphi(it)|=0$, we can find a large constant K such that for all t with absolute value greater than or equal to K, $f'(|\varphi(it)|) < 1$. Let $\beta = \max_{t \mid t \mid \in [K, mK]} f'(|\varphi(it)|) < 1$. Then $\int_{m^nK}^{m^n+1K} |\varphi'(it)| \, dt = m^n \int_{K}^{mK} |\varphi'(im^nt)| \, dt \leq \int_{K}^{mK} f_n'(|\varphi(it)|) |\varphi'(it)| \, dt \leq \int_{K}^{mK} f'(|\varphi(it)|)^n |\varphi'(it)| \, dt \leq \beta^n \cdot \int_{mK}^{mK} |\varphi'(it)| \, dt$. Hence for all $T \geq K$ there exists a constant Q independent of T such that $\int_{-T}^{T} |\varphi'(it)| \, dt \leq \int_{-K}^{K} |\varphi'(it)| \, dt + 2Q(m-1)(1/(1-\beta))K < \infty$. Q.E.D.

When using the preceding lemmas we can give an exceedingly simple proof of the following theorem:

THEOREM 1. If $\varphi(it)$ is not identically equal to one, then the distribution of W has a jump of magnitude equal to q at the origin and a continuous density function on the set of positive real numbers.

Proof. We will first prove that the distribution of W is differentiable on the set, $(0, \infty)$ under the additional assumption that q=0. Let $g_T(x)=(1/2\pi)\int_{-T}^T e^{-itx}\varphi(it)\,dt$ for $T=1,\,2,\,\cdots$, and x>0. Clearly, $g_T(\cdot)$ is a continuous function on the set, $(0,\,\infty)$. Moreover by integrating by parts we find that $g_T(x)=(-1/2\pi ix)\{e^{-iTx}\varphi(iT)-e^{iTx}\varphi(-iT)\}+(1/2\pi ix)\int_{-T}^T e^{-itx}\cdot(d\varphi(it)/dt)\,dt$. Hence if $0< x_1< x_2<\infty$, we can use Lemmas 2 and 3 to deduce that as T tends to infinity $g_T(\cdot)$ converges uniformly on $[x_1,\,x_2]$ to a continuous function g(x). Finally if $K(x)=P(W\leqq x)$, then by using the uniform boundedness of $g_T(\cdot)$ on closed bounded intervals we find that $K(x_2)-K(x_1)=\lim_{T\to\infty}\int_{-T}^T ((e^{-itx_2}-e^{-itx_1})/-2\pi it)\varphi(it)\,dt=\lim_{T\to\infty}\int_{x_1}^{x_2} g_T(x)\,dx=\int_{x_1}^{x_2} g(x)\,dx$, which establishes the required differentiability of $K(\cdot)$ for the case, q=0 (the last argument is the same as that used by Harris [3]). To show that K(0+)=0 we need only observe that

$$\lim_{t \to \pm \infty} |\varphi(it)| = \lim_{t \to \pm \infty} |K(0+)| + \int_{0+}^{\infty} e^{+itx} K'(x) dx$$
$$= |K(0+)| + \lim_{t \to \pm \infty} \int_{0+}^{\infty} e^{+itx} K'(x) dx = K(0+).$$

To prove the theorem in the case q>0 we proceed as follows: Let $\Psi(it)=(\varphi((1-q)it)-q)/(1-q)$ and let h(s)=(f((1-q)s+q)-q)/(1-q). Then it is easily shown that on the set, $\{0\leq s\leq 1\}$, $h(\cdot)$ is a convex, differentiable probability generating function such that h(0)=0, $h(s)\neq s^m$, and h'(1)=m. Moreover, if for all n we let $h_n(s)=h(h_{n-1}(s))$, then for all $s\in [0,1)$ we have $\lim_n h_n(s)=\lim_n (f_n((1-q)s+q)-q)/(1-q)=0$. Finally, $h_n'(s)=f_n'((1-q)s+q)\triangleq f'((1-q)s+q)^n=h'(s)^n$. Since $P(W=0)\geq \lim_n f_n(0)=q$, it is also easy to show that $\Psi(it)$ is a characteristic function and that it satisfies the condition, $\Psi(int)=h(\Psi(it))$. From these observations it follows that Lemmas 1, 2, and 3 apply to $\Psi(it)$. Hence $|\Psi(it)|<1$ for all $t\neq 0$, $\lim_{t\to\pm\infty} |\Psi(it)|=0$, and $\Psi'(it)$ is absolutely integrable. When using the in-

equality, $(1-q)|\Psi(it)| = |\varphi((1-q)it) - q| \ge ||\varphi((1-q)it)| - q|$, these results imply that $\lim_{t\to\pm\infty}|\varphi(it)| = q$ and that $\varphi'(it)$ is absolutely integrable. Similarly if $G(\cdot)$ is the probability distribution associated with $\Psi(\cdot)$, then the result obtained in the preceding paragraph applies to $G(\cdot)$. Hence $G(\cdot)$ is differentiable on the set $(0, \infty)$ and G(0+)=0. Finally, when using the definition of $\Psi(\cdot)$ and the unique correspondence between characteristic functions and right-continuous probability distributions we find that

$$G(x) = (K(x/(1-q)) - qE^*(x))/(1-q)$$

for all $x \ge 0$, where $E^*(x) \equiv 1$ for $x \ge 0$ and 0 for x < 0. This implies that $K(\cdot)$ is differentiable on $(0, \infty)$ and that K(0+) = q. Q.E.D.

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