AN INEQUALITY CONCERNING TESTS OF FIT OF THE KOLMOGOROV-SMIRNOV TYPE

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1. Introduction. Let $F_n(x)$ be the sumpolygon (empirical distribution function) of a sample of size n from a continuous distribution function F(x). Let K(x), $G_1(x)$, $G_2(x)$, $H_1(x)$, $H_2(x)$ be functions of x, such that for all x,

$$G_1(x) \ge G_2(x); \qquad H_2(x) \ge H_1(x).$$

The object of the present paper is to prove the following inequalities

(1)
$$P[\inf_{x} (F_{n} - K) \ge 0 \mid \inf_{x} (G_{1} - F_{n}) \ge 0, \inf_{x} (F_{n} - H_{2}) \ge 0]$$

 $\ge P[\inf_{x} (F_{n} - K) \ge 0 \mid \inf_{x} (G_{2} - F_{n}) \ge 0, \inf_{x} (F_{n} - H_{1}) \ge 0],$

(2)
$$P[\inf_x (K - F_n) \ge 0 \mid \inf_x (G_1 - F_n) \ge 0, \inf_x (F_n - H_2) \ge 0]$$

 $\le P[\inf_x (K - F_n) \ge 0 \mid \inf_x (G_2 - F_n) \ge 0, \inf_x (F_n - H_1) \ge 0],$

where all probabilities are supposed to exist. Since these inequalities are symmetrical, it suffices to prove one of them.

These inequalities provide an approximation for the distribution of two-sided statistics of the Kolmogorov-Smirnov type. Such a distribution is written

$$P\{\sup_{x} n^{\frac{1}{2}} | F_n(x) - F(x) | \psi[F(x)] \le \lambda\}$$

or more generally

(3)
$$P_n = P[\inf_x (G_2 - F_n) \ge 0, \inf_x (F_n - H_2) \ge 0].$$

In order to approximate P_n , take $H_1(x)$ and $H_2(x)$ in (1) smaller than zero for all x and replace K(x) in (1) by $H_2(x)$; similarly take $G_1(x)$ and $G_2(x)$ in (2) larger than 1 for all x and replace K(x) in (2) by $G_1(x)$. One then easily obtains the upper bound

(4)
$$P_n \leq P_n' P[\inf_x (G_2 - F_n) \geq 0] \cdot P[\inf_x (F_n - H_2) \geq 0] \cdot \{P[\inf_x (G_1 - F_n) \geq 0] \cdot P[\inf_x (F_n - H_1) \geq 0]\}^{-1}$$

where $P_n' = P[\inf_x (G_1 - F_n) \ge 0, \inf_x (F_n - H_1) \ge 0]$. If now G_1 and H_1 are chosen close to G_2 resp. H_2 , but such that P_n' is more easily calculable than P_n , then (4) provides an interesting approximation of (3). A lower bound can be found in a similar way.

Wald and Wolfowitz [3] and [4] have given the following two bounds for P_n

(5)
$$P_n \leq P[\inf_x (G_2 - F_n) \geq 0] \cdot P[\inf_x (F_n - H_2) \geq 0],$$

$$P_n \geq P[\inf_x (G_2 - F_n) \geq 0] + P[\inf_x (F_n - H_2) \geq 0] - 1.$$

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However they did not prove (5) but only conjectured it. As a matter of fact (5) constitutes the particular form of (4) corresponding to $G_1(x) \ge 1$ and $H_1(x) \le 0$ for all x. These bounds constitute good approximations for large values of P_n , which are the more interesting for purposes of testing. Furthermore they have the advantage of reducing the two-sided case to the one-sided case. On the other hand any other choice of $G_1(x)$ and $H_1(x)$ in (4) provides a closer upper bound than (5).

As Professor W. Hoeffding pointed out to the authors, (5) also is a consequence of a theorem of Lehmann [2]. Moreover he observed that the same theorem implies an analogous inequality for the distribution of the two-sample statistics of the Kolmogorov-Smirnov type. More details will be given in Section 4. Sections 2 and 3 are devoted to the proof of (1).

2. Lemma.

LEMMA. Let X be a many-dimensional random variable and S and T two measurable subsets of the range space of X. Let Y be a one-dimensional random variable. If for all y such that $a \leq y \leq b$

$$P[X \in S \mid (X \in T) \cap (Y = y)]$$

is a non-increasing function of y, then

$$P[X \in S \mid (X \in T) \cap (y_0 \leq Y \leq y_1)]$$

is a non-increasing function of y_0 and y_1 when $a \leq y_0 \leq y_1 \leq b$.

Proof. Let (X', Y') have the conditional distribution of (X, Y) under the condition $X \in T$ and let

$$h(y) = P[X \in S \mid (X \in T) \cap (Y = y)] = P[X' \in S \mid Y' = y].$$

Then

$$P[X \varepsilon S \mid (X \varepsilon T) \cap (y_0 \le Y \le y_1)]$$

$$= P[X' \varepsilon S \mid y_0 \le Y' \le y_1]$$

$$= \int_{y_0 \le y \le y_1} h(y) \cdot dF_{Y'}(y) / \int_{y_0 \le y \le y_1} dF_{Y'}(y)$$

where $F_{Y'}$ denotes the distribution function of Y'. Under the assumption that h(y) is non-increasing in [a, b], it has to be shown that (6) is a non-increasing function of y_0 and y_1 for $a \le y_0 \le y_1 \le b$, which is obviously true.

This proof has been suggested by Professor W. Hoeffding who pointed out that this lemma is closely related to Lemma 4 of Lehmann [2].

3. Proof of the inequality (1). Without loss of generality it will be supposed that $F(x) \equiv x$ for $0 \le x \le 1$. Let $X_1 \le X_2 \le \cdots \le X_n$ be the order statistics of a size n sample from F(x). The requirement that

$$\inf_{x} [F_n(x) - K(x)] \ge 0$$

is equivalent to the requirement that every X_j should not be larger than some well defined number k_j , i.e. $X_j \leq k_j$, $j = 1, \dots, n$, with $0 \leq k_1 \leq k_2 \leq \dots \leq n$

 $k_n \leq 1$. Similarly $\inf_x [F_n(x) - H(x)] \geq 0$ is equivalent to $X_j \leq h_j$, $j = 1, \dots, n$, with $0 \leq h_1 \leq h_2 \leq \dots \leq h_n \leq 1$. On the other hand $\inf_x [G(x) - F_n(x)] \geq 0$ is equivalent to $X_j \geq g_j$, $j = 1, \dots, n$, with $0 \leq g_1 \leq g_2 \leq \dots \leq g_n \leq 1$. Proving the inequality (1) then is equivalent to proving the following theorem.

Theorem. If $g_j \leq h_j$ for all j, then

$$P[\bigcap_{j=1}^{n} (X_j \leq k_j) \mid \bigcap_{j=1}^{n} (g_j \leq X_j \leq k_j)]$$

is a non-increasing function of $g_1, \dots, g_n, h_1, \dots, h_n$.

PROOF. The theorem holds if, for some j, $g_j > k_j$; in the sequel we suppose $g_j \le k_j$ for all j.

The theorem is evidently true for sample size 1. For proving it in general we will proceed by complete induction, supposing it is true for sample sizes $1, 2, \dots, n-1$.

Let i be any integer in [1, n]. We introduce the following vectors;

$$X = (X_1, \dots, X_{i-1}, X_{i+1}, \dots, X_n);$$

 $X' = (X_1, \dots, X_{i-1});$
 $X'' = (X_{i+1}, \dots, X_n);$
 $k = (k_1, \dots, k_{i-1}, k_{i+1}, \dots, k_n);$
 $k' = (k_1, \dots, k_{i-1});$
 $k'' = (k_{i+1}, \dots, k_n);$
 $c' = (c, \dots, c)$ $(i-1 \text{ components});$
 $c'' = (c, \dots, c)$ $(n-i \text{ components}).$

In a similar way we define g, g', g'', h, h' and h''. An inequality of two vectors will mean the simultaneous inequality of all corresponding components, e.g.: $X' \leq h'$ means

$$(X_1 \leq h_1) \cap (X_2 \leq h_2) \cap \cdots \cap (X_{i-1} \leq h_{i-1}).$$

Let $g_{i-1} \leq c \leq h_{i+1}$ (it is understood that $g_0 = 0$ and $h_{n+1} = 1$) and

$$Q = P[X \le k \mid (g \le X \le h) \cap (X_i = c)].$$

If $i \neq 1$ and $i \neq n$,

$$Q = P[(X' \le k') \cap (X'' \le k'') \mid (g' \le X' \le h') \cap (g'' \le X'' \le h'') \cap (X_i = c)]$$

= $Q' \cdot Q''$,

with

$$Q' = P[X' \le k' \mid (g' \le X' \le h') \cap (g'' \le X'' \le h'') \cap (X_i = c)],$$

$$Q'' = P[X'' \le k'' \mid (X' \le k') \cap (g' \le X' \le h') \cap (g'' \le X'' \le h'') \cap (X_i = c)].$$

Since $g_i \leq k_j$ for all j, Q'' is defined. Since X' and X'' are independent under the hypothesis $X_i = c$ (see e.g. Hajos and Rényi [1]), one has

$$Q' = P[X' \le k' \mid (g' \le X' \le h') \cap (X_i = c)],$$

$$Q'' = P[X'' \le k'' \mid (g'' \le X'' \le h'') \cap (X_i = c)].$$

Let $Y' = (Y_1, Y_2, \dots, Y_{i-1})$ be the order statistic of a size (i-1) sample from a rectangular distribution within [0, 1]. Under the hypothesis $Y' \leq c', Y'$ is distributed as the order statistic from a rectangular distribution within [0, c]. The same may be said of X' under the hypothesis $X_i = c$ (see e.g. Hajos and Rényi [1]). We thus have

$$Q' = P(Y' \le k' | (g' \le Y' \le h') \cap (Y' \le c')].$$

Similarly let Y'' be the order statistic of a size (n-i) sample from a rectangular distribution within [0, 1]. We then obtain

$$Q'' = P[Y'' \le k'' \mid (g'' \le Y'' \le h'') \cap (Y'' \ge c'')].$$

According to the present theorem applied to sample sizes smaller than n, Q' and Q'' are non-increasing functions of c when $g_{i-1} \leq c \leq h_{i+1}$. If i = 1 or i = n, we arrive easily at the same conclusion. Thus the same holds for Q.

Let

$$R = P[X \le k \mid (g \le X \le h) \cap (g_i \le X_i \le h_i)].$$

According to the lemma, R is a non-increasing function of g_i and h_i when $g_{i-1} \le g_i \le h_i \le h_{i+1}$. Let now

$$S = P[X_i \le k_i \mid (X \le k) \cap (g \le X \le h) \cap (g_i \le X_i \le h_i)],$$

$$T = P[(X \le k) \cap (X_i \le k_i) \mid (g \le X \le h) \cap (g_i \le X_i \le h_i)].$$

The probability S evidently is a non-increasing function of g_i and h_i . Since $T = R \cdot S$, T has the same property. As i is any integer in [1, n], the theorem is proved.

4. Two corollaries of a theorem of Lehmann (Hoeffding). Two real-valued functions r and s of n arguments are called by Lehmann discordant for the ith coordinate if, considered as functions of the ith coordinate (with all other coordinates held fixed), they are monotone in opposite directions, i.e. either r non-decreasing and s non-increasing or the inverse.

Let $(T_1, U_1), \dots, (T_n, U_n)$ be independent pairs of random variables such that

(7)
$$P(T_i \le t, U_i \le u) \ge P(T_i \le t) \cdot P(U_i \le u)$$

for all u and t, and let

$$T = r(T_1, \dots, T_n), \qquad U = s(U_1, \dots, U_n).$$

Lehmann proves that, if r and s are discordant for all i, then

$$P(T \ge t, U \ge u) \le P(T \ge t) \cdot P(U \ge u)$$

for all u and t.

We now show that this theorem implies inequality (5). Let T_1, \dots, T_n be n independent random variables and let $F_n(x)$ be their sumpolygon. Let $U_i = T_i$. Define

$$T = r(T_1, \dots, T_n) = \inf_x [F_n(x) - H(x)],$$

 $U = s(U_1, \dots, U_n) = \inf_x [G(x) - F_n(x)],$

where G(x) and H(x) are arbitrary functions of x. Then condition (7) is fulfilled and r and s are discordant for all i. This proves inequality (5) even in the case of discontinuous variables.

The same theorem of Lehmann implies an analogous inequality for the two-sample case. Let T_1, \dots, T_k be independent random variables with sumpolygon $F_1(x)$ and let T_{k+1}, \dots, T_n be independent random variables with sumpolygon $F_2(x)$. Let $U_i = T_i$ and define

$$D(x) = F_1(x) - F_2(x),$$

 $T = r(T_1, \dots, T_n) = \inf_x [D(x) - H(x)],$
 $U = s(U_1, \dots, U_n) = \inf_x [G(x) - D(x)],$

where G(x) and H(x) are arbitrary functions of x. Again condition (7) is fulfilled and r and s are discordant for all i. In particular when $G(x) \equiv -H(x) = a > 0$, one has

$$P[\sup_{x} |D(x)| \le a] \le \{P[\sup_{x} D(x) \le a]\}^{2}.$$

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