# TREES, NOT CUBES: <br> HYPERCONTRACTIVITY, COSINESS, AND NOISE STABILITY 

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## Abstract:

Noise sensitivity of functions on the leaves of a binary tree is studied, and a hypercontractive inequality is obtained. We deduce that the spider walk is not noise stable.

## Introduction

For the simplest random walk (Fig. 1a), the set $\Omega_{n}^{\text {simp }}$ of all $n$-step trajectories may be thought of either as (the set of leaves of) a binary tree, or (the vertices of) a binary cube $\{-1,+1\}^{n}$. However, consider another random walk (Fig. 1b); call it the simplest spider walk, since it is a discrete counterpart of a spider martingale, see [2]. The corresponding $\Omega_{n}^{\text {spider }}$ is the set of leaves of a binary tree. It is not quite appropriate to think of such $n$-step "spider walks" as the vertices of a binary cube, since for different $i$ and $j$ in $\{1,2, \ldots, n\}$ it is not necessary that the $j$ 'th step has the same or opposite direction from the $i$ 'th step. Of course, one may choose to ignore this point, and use the $n$ bits given by a point in $\{-1,1\}^{n}$ to describe a spider walk, in such a way that for each $j=1,2, \ldots, n$, the first $j$ bits determine the first $j$ steps of the walk. Such a correspondence would not be unique. To put it differently, the vertices of the cube have a natural associated partial order. When you consider two walks on $\mathbb{Z}$, the associated partial ordering has a natural interpretation, one trajectory is (weakly) larger than the other if whenever the latter moved to the right, the former also moved to the right. However, this interpretation does not make sense for the spider walk in Fig. 1b, and even less to more complicated "spider webs" with several "roundabouts", such as that of Fig. 1c.
Noise sensitivity and stability are introduced and studied in [3] for functions on cubes. Different cube structures on a binary tree are non-equivalent in that respect. It is shown here that a


Figure 1: (a) simple walk; (b) spider walk; (c) a spider web. At each point, there are two equiprobable moves.
natural function on $\Omega_{n}^{\text {spider }}$ is non-stable under every cube structure. One of the tools used is a new hypercontractive inequality, which hopefully may find uses elsewhere.

## 1 Stability and sensitivity on cubes, revisited

A function $f:\{-1,+1\}^{n} \rightarrow \mathbb{C}$ has its Fourier-Walsh expansion,

$$
\begin{aligned}
& f\left(\tau_{1}, \ldots, \tau_{n}\right)= \\
& \quad=\hat{f}_{0}+\sum_{k} \hat{f}_{1}(k) \tau_{k}+\sum_{k<l} \hat{f}_{2}(k, l) \tau_{k} \tau_{l}+\cdots+\hat{f}_{n}(1, \ldots, n) \tau_{1} \ldots \tau_{n}
\end{aligned}
$$

Set

$$
\tilde{f}_{j}\left(\tau_{1}, \ldots, \tau_{n}\right)=\sum_{i_{1}<i_{2}<\cdots<i_{j}} \hat{f}_{j}\left(i_{1}, \ldots, i_{j}\right) \tau_{i_{1}} \tau_{i_{2}} \ldots \tau_{i_{j}}
$$

Since the transform $f \mapsto \hat{f}$ is isometric, we have $\|f\|^{2}=\sum_{0}^{n}\left\|\tilde{f}_{m}\right\|^{2}$, where

$$
\begin{equation*}
\|f\|^{2}=2^{-n} \sum_{\tau_{1}, \ldots, \tau_{n}}\left|f\left(\tau_{1}, \ldots, \tau_{n}\right)\right|^{2} \tag{1.1}
\end{equation*}
$$

The quantities

$$
S_{1}^{m}(f)=\sum_{i=1}^{m}\left\|\tilde{f}_{i}\right\|^{2}, \quad S_{m}^{\infty}(f)=\sum_{i=m}^{n}\left\|\tilde{f}_{i}\right\|^{2}
$$

are used for describing low-frequency and high-frequency parts of the spectrum of $f$.
Given a sequence of functions $F=\left(f_{n}\right)_{n=1}^{\infty}, f_{n}:\{-1,+1\}^{n} \rightarrow \mathbb{C}$, satisfying $0<\liminf _{n \rightarrow \infty}\left\|f_{n}\right\| \leq$ $\limsup _{n \rightarrow \infty}\left\|f_{n}\right\|<\infty$, we consider numbers

$$
\begin{aligned}
& S_{1}^{m}(F)=\limsup _{n \rightarrow \infty} S_{1}^{m}\left(f_{n}\right), \\
& S_{m}^{\infty}(F)=\limsup _{n \rightarrow \infty} S_{m}^{\infty}\left(f_{n}\right)
\end{aligned}
$$

Here is one of equivalent definitions of stability and sensitivity for such $F$, according to [3, Th. 1.8] (indicator functions are considered there):

$$
\begin{array}{rll}
F \text { is stable } & \text { iff } & S_{m}^{\infty}(F) \rightarrow 0 \text { for } m \rightarrow \infty \\
F \text { is sensitive } & \text { iff } & S_{1}^{m}(F)=0 \text { for all } m
\end{array}
$$

Constant components are irrelevant; that is, if $g_{n}=f_{n}+c_{n}, c_{n} \in \mathbb{C}$, then $S_{1}^{m}\left(f_{n}\right)=S_{1}^{m}\left(g_{m}\right)$ and $S_{m+1}^{\infty}\left(f_{n}\right)=S_{m+1}^{\infty}\left(g_{n}\right)$, therefore stability of $\left(f_{n}\right)_{n=1}^{\infty}$ is equivalent to stability of $\left(g_{n}\right)_{n=1}^{\infty}$; the same for sensitivity. If $\left(f_{n}\right)_{n=1}^{\infty}$ is both stable and sensitive, then (and only then) $S_{1}^{\infty}\left(f_{n}\right) \rightarrow$ 0 , that is, $\left\|f_{n}-c_{n}\right\| \rightarrow 0$ for some $c_{n} \in \mathbb{C}$.
A random variable $\tau$ will be called a random sign, if $\mathbb{P}(\tau=-1)=1 / 2$ and $\mathbb{P}(\tau=+1)=1 / 2$. A joint distribution for two random signs $\tau^{\prime}, \tau^{\prime \prime}$ is determined by their correlation coefficient $\rho=\mathbb{E}\left(\tau^{\prime} \tau^{\prime \prime}\right)=1-2 \mathbb{P}\left(\tau^{\prime} \neq \tau^{\prime \prime}\right)$. Given $n$ independent pairs $\left(\tau_{1}^{\prime}, \tau_{1}^{\prime \prime}\right), \ldots,\left(\tau_{n}^{\prime}, \tau_{n}^{\prime \prime}\right)$ of random signs with the same correlation $\rho$ for each pair, we call $\left(\tau_{1}^{\prime}, \ldots, \tau_{n}^{\prime}\right)$ and $\left(\tau_{1}^{\prime \prime}, \ldots, \tau_{n}^{\prime \prime}\right)$ a $\rho$ correlated pair of random points of the cube $\{-1,+1\}^{n}$. (In terms of [3] it is $\left(x, N_{\varepsilon}(x)\right)$ with $\varepsilon=(1-\rho) / 2$.) It is easy to see that

$$
\mathbb{E}\left(\overline{f\left(\tau^{\prime}\right)} f\left(\tau^{\prime \prime}\right)\right)=\sum_{m=0}^{n} \rho^{m}\left\|\tilde{f}_{m}\right\|^{2}
$$

for a $\rho$-correlated pair $\left(\tau^{\prime}, \tau^{\prime \prime}\right)$. We may write it as a scalar product in the space $L_{2}\left(\{-1,+1\}^{n}\right)$ with the norm (1.1),

$$
\begin{equation*}
\mathbb{E}\left(\overline{f\left(\tau^{\prime}\right)} f\left(\tau^{\prime \prime}\right)\right)=\left(\rho^{\mathbf{N}} f, f\right) \tag{1.2}
\end{equation*}
$$

here $\rho^{\mathbf{N}}$ is the operator $\rho^{\mathbf{N}} f=\sum_{n} \rho^{n} \tilde{f}_{n}$. Similarly, $\mathbb{E}\left(\overline{g\left(\tau^{\prime}\right)} f\left(\tau^{\prime \prime}\right)\right)=\left(\rho^{\mathbf{N}} f, g\right)$. On the other hand,

$$
\mathbb{E}\left(\overline{g\left(\tau^{\prime}\right)} f\left(\tau^{\prime \prime}\right)\right)=\mathbb{E}\left(\overline{g\left(\tau^{\prime}\right)} \cdot \mathbb{E}\left(f\left(\tau^{\prime \prime}\right) \mid \tau^{\prime}\right)\right)=\left(\tau^{\prime} \mapsto \mathbb{E}\left(f\left(\tau^{\prime \prime}\right) \mid \tau^{\prime}\right), g\right)
$$

thus,

$$
\begin{equation*}
\mathbb{E}\left(f\left(\tau^{\prime \prime}\right) \mid \tau^{\prime}\right)=\left(\rho^{\mathbf{N}} f\right)\left(\tau^{\prime}\right) \tag{1.3}
\end{equation*}
$$

(Our $\rho^{\mathbf{N}}$ is $T_{\eta}=Q_{\varepsilon}$ of [3] with $\eta=\rho, \varepsilon=(1-\rho) / 2$.) (In fact, let $\mathbf{N} f=\sum_{n} n \tilde{f}_{n}$, then $-\mathbf{N}$ is the generator of a Markov process on $\{-1,+1\}^{n} ; \exp (-t \mathbf{N})$ is its semigroup; note that $\rho^{\mathbf{N}}$ is of the form $\exp (-t \mathbf{N})$. The Markov process is quite simple: during $d t$, each coordinate flips with the probability $\frac{1}{2} d t+o(d t)$. However, we do not need it.) Note also that $\mathbb{E}\left(\left|f_{n}\left(\tau^{\prime \prime}\right)-\left(\rho^{\mathbf{N}} f_{n}\right)\left(\tau^{\prime}\right)\right|^{2} \mid \tau^{\prime}\right)$ is the conditional variance $\operatorname{Var}\left(f_{n}\left(\tau^{\prime \prime}\right) \mid \tau^{\prime}\right)$, and its mean value (over all $\tau^{\prime}$ ) is

$$
\begin{equation*}
\mathbb{E} \operatorname{Var}\left(f_{n}\left(\tau^{\prime \prime}\right) \mid \tau^{\prime}\right)=\left\|f_{n}\right\|^{2}-\left\|\rho^{\mathbf{N}} f_{n}\right\|^{2}=\left(\left(\mathbf{1}-\rho^{2 \mathbf{N}}\right) f_{n}, f_{n}\right) \tag{1.4}
\end{equation*}
$$

Note also that the operator $0^{\mathbf{N}}=\lim _{\rho \rightarrow 0} \rho^{\mathbf{N}}$ is the projection onto the one-dimensional space of constants, $f \mapsto(\mathbb{E} f) \cdot \mathbf{1}$.
Stability of $F=\left(f_{n}\right)_{n=1}^{\infty}$ is equivalent to:

- $\left\|\rho^{\mathbf{N}} f_{n}-f_{n}\right\| \xrightarrow[\rho \rightarrow 1]{ } 0$ uniformly in $n$;
- $\left(\rho^{\mathbf{N}} f_{n}, f_{n}\right) \xrightarrow[\rho \rightarrow 1]{\longrightarrow}\|f\|^{2}$ uniformly in $n$;
- $\left\|f_{n}\right\|^{2}-\left\|\rho^{\mathbf{N}} f_{n}\right\|^{2} \xrightarrow[\rho \rightarrow 1]{\longrightarrow} 0$ uniformly in $n$.

Sensitivity of $F$ is equivalent to:

- $\left\|\left(\rho^{\mathbf{N}}-0^{\mathbf{N}}\right) f_{n}\right\| \xrightarrow[n \rightarrow \infty]{ } 0$ for some (or every) $\rho \in(0,1)$;
- $\left(\left(\rho^{\mathbf{N}}-0^{\mathbf{N}}\right) f_{n}, f_{n}\right) \xrightarrow[n \rightarrow \infty]{ } 0$ for some (or every) $\rho \in(0,1)$.

Combining these facts with the probabilistic interpretation (1.2), (1.3), (1.4) of $\rho^{\mathbf{N}}$ we see that

- $F$ is stable iff $\mathbb{E}\left(\overline{f_{n}\left(\tau^{\prime}\right)} f_{n}\left(\tau^{\prime \prime}\right)\right) \underset{\rho \rightarrow 1}{\longrightarrow} \mathbb{E}\left|f_{n}(\tau)\right|^{2}$ uniformly in $n$ or, equivalently, $\mathbb{E}\left(\operatorname{Var}\left(f_{n}\left(\tau^{\prime \prime}\right) \mid \tau^{\prime}\right)\right) \xrightarrow[\rho \rightarrow 1]{\longrightarrow} 0$ uniformly in $n$;
- $F$ is sensitive iff $\mathbb{E}\left(\overline{f_{n}\left(\tau^{\prime}\right)} f_{n}\left(\tau^{\prime \prime}\right)\right)-\left|\mathbb{E} f_{n}(\tau)\right|^{2} \underset{n \rightarrow \infty}{\longrightarrow} 0$ for some (or every) $\rho \in(0,1)$ or, equivalently, $\mathbb{E}\left|\mathbb{E}\left(f\left(\tau^{\prime \prime}\right) \mid \tau^{\prime}\right)-\mathbb{E} f\right|^{2} \underset{n \rightarrow \infty}{\longrightarrow} 0$ for some (or every) $\rho \in(0,1)$.

These are versions of definitions introduced in $[3$, Sect. 1.1, 1.4].

## 2 Stability and sensitivity on trees

A branch of the $n$-level binary tree can be written as a sequence of sequences ()$,\left(\tau_{1}\right),\left(\tau_{1}, \tau_{2}\right)$, $\left(\tau_{1}, \tau_{2}, \tau_{3}\right), \ldots,\left(\tau_{1}, \ldots, \tau_{n}\right)$. Branches correspond to leaves $\left(\tau_{1}, \ldots, \tau_{n}\right) \in\{-1,+1\}^{n}$. Automorphisms of the tree can be described as maps $A:\{-1,+1\}^{n} \rightarrow\{-1,+1\}^{n}$ of the form

$$
A\left(\tau_{1}, \ldots, \tau_{n}\right)=\left(a() \tau_{1}, a\left(\tau_{1}\right) \tau_{2}, a\left(\tau_{1}, \tau_{2}\right) \tau_{3}, \ldots, a\left(\tau_{1}, \ldots, \tau_{n-1}\right) \tau_{n}\right)
$$

for arbitrary functions $a: \cup_{m=1}^{n}\{-1,+1\}^{m-1} \rightarrow\{-1,+1\}$. (Thus, the tree has $2^{1} \cdot 2^{2} \cdot 2^{4} \cdot \ldots$. $2^{2^{n-1}}=2^{2^{n}-1}$ automorphisms, while the cube $\{-1,+1\}^{n}$ has only $2^{n} n$ ! automorphisms.) Here is an example of a tree automorphism (far from being a cube automorphism):

$$
\left(\tau_{1}, \ldots, \tau_{n}\right) \mapsto\left(\tau_{1}, \tau_{1} \tau_{2}, \ldots, \tau_{1} \ldots \tau_{n}\right)
$$

The function $f_{n}\left(\tau_{1}, \ldots, \tau_{n}\right)=\frac{1}{\sqrt{n}}\left(\tau_{1}+\cdots+\tau_{n}\right)$ satisfies $S_{1}^{1}\left(f_{n}\right)=1, S_{2}^{\infty}\left(f_{n}\right)=0$. However, the function $g_{n}\left(\tau_{1}, \ldots, \tau_{n}\right)=\frac{1}{\sqrt{n}}\left(\tau_{1}+\tau_{1} \tau_{2}+\cdots+\tau_{1} \ldots \tau_{n}\right)$ satisfies $S_{1}^{m}\left(g_{n}\right)=\min \left(\frac{m}{n}, 1\right)$, $S_{m}^{\infty}\left(g_{n}\right)=\max \left(\frac{n-m+1}{n}, 0\right)$. According to the definitions of Sect. $1,\left(f_{n}\right)_{n=1}^{\infty}$ is stable, but $\left(g_{n}\right)_{n=1}^{\infty}$ is sensitive. We see that the definitions are not tree-invariant. A straightforward way to tree-invariance is used in the following definition of "tree stability" and "tree sensitivity". From now on, stability and sensitivity of Sect. 1 will be called "cube stability" and "cube sensitivity".
2.1 Definition (a) A sequence $\left(f_{n}\right)_{n=1}^{\infty}$ of functions $f_{n}:\{-1,+1\}^{n} \rightarrow \mathbb{C}$ is tree stable, if there exists a sequence of tree automorphisms $A_{n}:\{-1,+1\}^{n} \rightarrow\{-1,+1\}^{n}$ such that the sequence $\left(f_{n} \circ A_{n}\right)_{n=1}^{\infty}$ is cube stable.
(b) The sequence $\left(f_{n}\right)_{n=1}^{\infty}$ is tree sensitive, if $\left(f_{n} \circ A_{n}\right)_{n=1}^{\infty}$ is cube sensitive for every sequence $\left(A_{n}\right)$ of tree automorphisms.

The definition can be formulated in terms of $f_{n}\left(A_{n}\left(\tau^{\prime}\right)\right)$ and $f_{n}\left(A_{n}\left(\tau^{\prime \prime}\right)\right)$ where $\left(\tau^{\prime}, \tau^{\prime \prime}\right)$ is a $\rho$-correlated pair of random points of the cube $\{-1,+1\}^{n}$. Equivalently, we may consider $f_{n}\left(\tau^{\prime}\right)$ and $f_{n}\left(\tau^{\prime \prime}\right)$ where $\tau^{\prime}, \tau^{\prime \prime}$ are such that for some $A_{n},\left(A_{n} \tau^{\prime}, A_{n} \tau^{\prime \prime}\right)$ is a $\rho$-correlated pair. That is,

$$
\begin{align*}
& \mathbb{E}\left(\tau_{m}^{\prime} \mid \tau_{1}^{\prime}, \tau_{1}^{\prime \prime}, \ldots, \tau_{m-1}^{\prime}, \tau_{m-1}^{\prime \prime}\right)=\mathbb{E}\left(\tau_{m}^{\prime \prime} \mid \tau_{1}^{\prime}, \tau_{1}^{\prime \prime}, \ldots, \tau_{m-1}^{\prime}, \tau_{m-1}^{\prime \prime}\right)=0,  \tag{2.2}\\
& \mathbb{E}\left(\tau_{m}^{\prime} \tau_{m}^{\prime \prime} \mid \tau_{1}^{\prime}, \tau_{1}^{\prime \prime}, \ldots, \tau_{m-1}^{\prime}, \tau_{m-1}^{\prime \prime}\right)=a\left(\tau_{1}^{\prime}, \ldots, \tau_{m-1}^{\prime}\right) a\left(\tau_{1}^{\prime \prime}, \ldots, \tau_{m-1}^{\prime \prime}\right) \rho, \tag{2.3}
\end{align*}
$$

where $a: \cup_{m=1}^{n}\{-1,+1\}^{m-1} \rightarrow\{-1,+1\}$. On the other hand, consider an arbitrary $\{-1,+1\}^{n} \times$ $\{-1,+1\}^{n}$-valued random variable $\left(\tau^{\prime}, \tau^{\prime \prime}\right)$ satisfying (2.2) (which implies that each one of $\tau^{\prime}, \tau^{\prime \prime}$ is uniform on $\{-1,+1\}^{n}$ ), but maybe not (2.3), and define

$$
\begin{equation*}
\rho_{\max }\left(\tau^{\prime}, \tau^{\prime \prime}\right)=\max _{m=1, \ldots, n} \max \left|\mathbb{E}\left(\tau_{m}^{\prime} \tau_{m}^{\prime \prime} \mid \tau_{1}^{\prime}, \tau_{1}^{\prime \prime}, \ldots, \tau_{m-1}^{\prime}, \tau_{m-1}^{\prime \prime}\right)\right|, \tag{2.4}
\end{equation*}
$$

where the internal maximum is taken over all possible values of $\left(\tau_{1}^{\prime}, \tau_{1}^{\prime \prime}, \ldots, \tau_{m-1}^{\prime}, \tau_{m-1}^{\prime \prime}\right)$. The joint distribution of $\tau^{\prime}$ and $\tau^{\prime \prime}$ is a probability measure $\mu$ on $\{-1,+1\}^{n} \times\{-1,+1\}^{n}$, and we denote $\rho_{\max }\left(\tau^{\prime}, \tau^{\prime \prime}\right)$ by $\rho_{\max }(\mu)$. Given $f, g:\{-1,+1\}^{n} \rightarrow \mathbb{C}$, we denote $\mathbb{E} \overline{f\left(\tau^{\prime}\right)} g\left(\tau^{\prime \prime}\right)$ by $\langle f| \mu|g\rangle$.
2.5 Definition A sequence $\left(f_{n}\right)_{n=1}^{\infty}$ of functions $f_{n}:\{-1,+1\}^{n} \rightarrow \mathbb{C}$, satisfying $0<\lim \inf _{n \rightarrow \infty}\left\|f_{n}\right\| \leq \lim \sup _{n \rightarrow \infty}\left\|f_{n}\right\|<\infty$, is cosy, if for any $\varepsilon>0$ there is a sequence $\left(\mu_{n}\right)_{n=1}^{\infty}, \mu_{n}$ being a probability measure on $\{-1,+1\}^{n} \times\{-1,+1\}^{n}$, such that $\lim \sup _{n \rightarrow \infty} \rho_{\max }\left(\mu_{n}\right)<1$ and $\lim \sup _{n \rightarrow \infty}\left(\left\|f_{n}\right\|^{2}-\left\langle f_{n}\right| \mu_{n}\left|f_{n}\right\rangle\right)<\varepsilon$.
2.6 Lemma Every tree stable sequence is cosy.

Proof. Let $\left(f_{n}\right)$ be tree stable. Take tree automorphisms $A_{n}$ such that $\left(f_{n} \circ A_{n}\right)$ is cube stable. We have $\mathbb{E}\left(\overline{f_{n}\left(A_{n}\left(\tau^{\prime}\right)\right)} f_{n}\left(A_{n}\left(\tau^{\prime \prime}\right)\right)\right) \xrightarrow[\rho \rightarrow 1]{\longrightarrow} \mathbb{E}\left|f_{n}(\tau)\right|^{2}$ uniformly in $n$. Here $\tau^{\prime}, \tau^{\prime \prime}$ are $\rho$-correlated. The joint distribution $\mu_{n}(\rho)$ of $A_{n}\left(\tau^{\prime}\right)$ and $A_{n}\left(\tau^{\prime \prime}\right)$ satisfies $\rho_{\max }\left(\mu_{n}(\rho)\right) \leq \rho$ due to (2.3). Also, $\left\langle f_{n}\right| \mu_{n}\left|f_{n}\right\rangle \xrightarrow[\rho \rightarrow 1]{ }\left\|f_{n}\right\|^{2}$ uniformly in $n$, which means that $\sup _{n}\left(\left\|f_{n}\right\|^{2}-\left\langle f_{n}\right| \mu_{n}\left|f_{n}\right\rangle\right) \rightarrow 0$ for $\rho \rightarrow 1$.
Is there a cosy but not tree stable sequence? We do not know. The conditional correlation given by (2.3) is not only $\pm \rho$, it is also factorizable (a function of $\tau^{\prime}$ times the same function of $\left.\tau^{\prime \prime}\right)$, which seems to be much stronger than just $\rho_{\max }(\mu) \leq \rho$.

## 3 Hypercontractivity

Let $\left(\tau^{\prime}, \tau^{\prime \prime}\right)$ be a $\rho$-correlated pair of random points of the cube $\{-1,+1\}^{n}$. Then for every $f, g:\{-1,+1\}^{n} \rightarrow \mathbb{R}$

$$
\begin{equation*}
\left|\mathbb{E} f\left(\tau^{\prime}\right) g\left(\tau^{\prime \prime}\right)\right|^{1+\rho} \leq\left(\mathbb{E}\left|f\left(\tau^{\prime}\right)\right|^{1+\rho}\right)\left(\mathbb{E}\left|g\left(\tau^{\prime \prime}\right)\right|^{1+\rho}\right), \tag{3.1}
\end{equation*}
$$

which is a discrete version of the celebrated hypercontractivity theorem pioneered by Nelson (see [7, Sect. 3]). For a proof, see [1]; there, following Gross [6], the inequality is proved for $n=1$ (just two points, $\{-1,+1\}$ ) [1, Prop. 1.5], which is enough due to tensorization [1,

Lemma 1.3]. (See also [3, Lemma 2.4].) The case of $f, g$ taking on two values 0 and 1 only is especially important:

$$
\mathbb{P}^{1+\rho}\left(\tau^{\prime} \in S^{\prime} \& \tau^{\prime \prime} \in S^{\prime \prime}\right) \leq \mathbb{P}\left(\tau^{\prime} \in S^{\prime}\right) \mathbb{P}\left(\tau^{\prime \prime} \in S^{\prime \prime}\right)=\frac{\left|S^{\prime}\right|}{2^{n}} \cdot \frac{\left|S^{\prime \prime}\right|}{2^{n}}
$$

for any $S^{\prime}, S^{\prime \prime} \subset\{-1,+1\}^{n}$. Note that $\rho=0$ means independence, ${ }^{1}$ while $\rho=1$ is trivial: $\mathbb{P}^{2}(\ldots) \leq\left(\min \left(\mathbb{P}\left(S^{\prime}\right), \mathbb{P}\left(S^{\prime \prime}\right)\right)\right)^{2} \leq \mathbb{P}\left(S^{\prime}\right) \mathbb{P}\left(S^{\prime \prime}\right)$.
For a probability measure $\mu$ on $\{-1,+1\}^{n} \times\{-1,+1\}^{n}$ we denote by $\langle g| \mu|f\rangle$ the value $\mathbb{E}\left(f\left(\tau^{\prime}\right) g\left(\tau^{\prime \prime}\right)\right)$, where $\left(\tau^{\prime}, \tau^{\prime \prime}\right) \sim \mu$. The hypercontractivity (3.1) may be written as $\left.|\langle g| \mu| f\right\rangle \mid \leq$ $\|f\|_{1+\rho}\|g\|_{1+\rho}$, where $\mu=\mu(\rho)$ is the distribution of a $\rho$-correlated pair. The class of $\mu$ that satisfy the inequality (for all $f, g$ ) is invariant under transformations of the form $A \times B$, where $A, B:\{-1,+1\}^{n} \rightarrow\{-1,+1\}^{n}$ are arbitrary invertible maps (since such maps preserve $\|\cdot\|_{1+\rho}$ ). In particular, all measures of the form (2.2-2.3) fit.
Can we generalize the statement for all $\mu$ such that $\rho_{\max }(\mu) \leq \rho$ ? The approach of Gross, based on tensorization, works on cubes (and other products), not trees. Fortunately, we have another approach, found by Neveu [8], that works also on trees.
3.2 Lemma For every $r \in\left[\frac{1}{2}, 1\right], x, y \in[0,1]$, and $\rho \in\left[-\frac{1-r}{r}, \frac{1-r}{r}\right]$,

$$
\begin{aligned}
& (1+\rho)(1-x)^{r}(1-y)^{r}+(1-\rho)(1-x)^{r}(1+y)^{r}+ \\
& \quad+(1-\rho)(1+x)^{r}(1-y)^{r}+(1+\rho)(1+x)^{r}(1+y)^{r} \leq 4 .
\end{aligned}
$$

Proof. The left hand side is linear in $\rho$ with the coefficient $\left((1+x)^{r}-(1-x)^{r}\right)\left((1+y)^{r}-\right.$ $\left.(1-y)^{r}\right) \geq 0$. Therefore, it suffices to prove the inequality for $\rho=\frac{1-r}{r}, r \in\left(\frac{1}{2}, 1\right)$ (the cases $r=\frac{1}{2}$ and $r=1$ follow by continuity). Assume the contrary, then the continuous function $f_{r}$ on $[0,1] \times[0,1]$, defined by

$$
\begin{aligned}
& f_{r}(x, y)=\frac{1}{r}(1-x)^{r}(1-y)^{r}+\frac{2 r-1}{r}(1-x)^{r}(1+y)^{r}+ \\
&+\frac{2 r-1}{r}(1+x)^{r}(1-y)^{r}+\frac{1}{r}(1+x)^{r}(1+y)^{r}
\end{aligned}
$$

has a global maximum $f_{r}\left(x_{0}, y_{0}\right)>4$ for some $r \in\left(\frac{1}{2}, 1\right)$. The case $x_{0}=y_{0}=0$ is excluded (since $\left.f_{r}(0,0)=4\right)$. Also, $x_{0} \neq 1$ (since $\left.\left.\frac{\partial}{\partial x}\right|_{x=1-} f_{r}(x, y)=-\infty\right)$ and $y_{0} \neq 1$. The new variables

$$
u=\frac{1+x}{1-x} \in[1, \infty), \quad v=\frac{1+y}{1-y} \in[1, \infty)
$$

will be useful. We have

$$
\begin{equation*}
\frac{1+x}{(1-x)^{r}(1-y)^{r}} \frac{\partial}{\partial x} f_{r}(x, y)=u^{r} v^{r}-u-(2 r-1)\left(u v^{r}-u^{r}\right) . \tag{3.3}
\end{equation*}
$$

For $u=1, v>1$ the right hand side is $2(1-r)\left(v^{r}-1\right)>0$; therefore $x_{0} \neq 0$ (since $\left.\left(x_{0}, y_{0}\right) \neq(0,0)\right)$, and similarly $y_{0} \neq 0$. So, $\left(x_{0}, y_{0}\right)$ is an interior point of $[0,1] \times[0,1]$. The corresponding $u_{0}, v_{0} \in(1, \infty)$ satisfy $u_{0}^{r} v_{0}^{r}-u_{0}-(2 r-1)\left(u_{0} v_{0}^{r}-u_{0}^{r}\right)=0$. By subtracting the same expression with $v_{0}$ switched with $u_{0}$, which also vanishes, we get

$$
v_{0}-u_{0}+(2 r-1)\left(u_{0}^{r} v_{0}-u_{0} v_{0}^{r}+u_{0}^{r}-v_{0}^{r}\right)=0
$$

[^0]Aiming to conclude that $u_{0}=v_{0}$, consider the function $u \mapsto v_{0}-u+(2 r-1)\left(u^{r} v_{0}-u v_{0}^{r}+u^{r}-v_{0}^{r}\right)$ on $[1, \infty)$. It is concave, and positive when $u=1$, since $v_{0}-1+(2 r-1)\left(v_{0}-2 v_{0}^{r}+1\right) \geq$ $v_{0}-1+(2 r-1)\left(v_{0}-2 v_{0}+1\right)=\left(v_{0}-1\right)(2-2 r)$. Therefore, the function cannot vanish more than once, and $u=v_{0}$ is its unique root. So, $u_{0}=v_{0}$.
It follows from (3.3) that

$$
\frac{1+x}{(1-x)^{2 r}} \cdot \frac{1}{2} \frac{\partial}{\partial x} f(x, x)=u^{2 r}-u-(2 r-1)\left(u^{r+1}-u^{r}\right)
$$

therefore $u_{0}$ is a root of the equation $u^{2 r-1}-1-(2 r-1)\left(u^{r}-u^{r-1}\right)=0$, different from the evident root $u=1$. However, the function $u \mapsto u^{2 r-1}-1-(2 r-1)\left(u^{r}-u^{r-1}\right)$ is strictly monotone, since

$$
\frac{1}{2 r-1} \frac{\partial}{\partial u}(\ldots)=u^{2 r-2}-r u^{r-1}+(r-1) u^{r-2}=u^{r-2}\left(u^{r}-r u+r-1\right)<0
$$

due to the inequality $u^{r} \leq 1+r(u-1)$ (which follows from concavity of $u^{r}$ ). The contradiction completes the proof.
3.4 Theorem Let $\rho \in[0,1]$, and $\mu$ be a probability measure on $\{-1,+1\}^{n} \times\{-1,+1\}^{n}$ such that $^{2} \rho_{\max }(\mu) \leq \rho$. Then for every $f, g:\{-1,+1\}^{n} \rightarrow \mathbb{C}$

$$
|\langle g| \mu| f\rangle \mid \leq\|f\|_{1+\rho}\|g\|_{1+\rho} .
$$

Proof. Consider random points $\tau^{\prime}, \tau^{\prime \prime}$ of $\{-1,+1\}^{n}$ such that $\left(\tau^{\prime}, \tau^{\prime \prime}\right) \sim \mu$. We have two (correlated) random processes $\tau_{1}^{\prime}, \ldots, \tau_{n}^{\prime}$ and $\tau_{1}^{\prime \prime}, \ldots, \tau_{n}^{\prime \prime}$. Consider the random variables

$$
M_{n}^{\prime}=\left|f\left(\tau_{1}^{\prime}, \ldots, \tau_{n}^{\prime}\right)\right|^{1 / r}, \quad M_{n}^{\prime \prime}=\left|g\left(\tau_{1}^{\prime \prime}, \ldots, \tau_{n}^{\prime \prime}\right)\right|^{1 / r}
$$

and the corresponding martingales

$$
\begin{aligned}
& M_{m}^{\prime}=\mathbb{E}\left(M_{n}^{\prime} \mid \tau_{1}^{\prime}, \tau_{1}^{\prime \prime}, \ldots, \tau_{m}^{\prime}, \tau_{m}^{\prime \prime}\right)=\mathbb{E}\left(M_{n}^{\prime} \mid \tau_{1}^{\prime}, \ldots, \tau_{m}^{\prime}\right), \\
& M_{m}^{\prime \prime}=\mathbb{E}\left(M_{n}^{\prime \prime} \mid \tau_{1}^{\prime}, \tau_{1}^{\prime \prime}, \ldots, \tau_{m}^{\prime}, \tau_{m}^{\prime \prime}\right)=\mathbb{E}\left(M_{n}^{\prime \prime} \mid \tau_{1}^{\prime \prime}, \ldots, \tau_{m}^{\prime \prime}\right)
\end{aligned}
$$

for $m=0,1, \ldots, n$; the equalities for conditional expectations follow from (2.2). For any $m=1, \ldots, n$ and any values of $\tau_{1}^{\prime}, \tau_{1}^{\prime \prime}, \ldots, \tau_{m-1}^{\prime}, \tau_{m-1}^{\prime \prime}$ consider the conditional distribution of the pair $\left(M_{m}^{\prime}, M_{m}^{\prime \prime}\right)$. It is concentrated at four points that can be written as ${ }^{3}\left((1 \pm x) M_{m-1}^{\prime},(1 \pm\right.$ y) $M_{m-1}^{\prime \prime}$ ). The first " $\pm$ " depends only on $\tau_{m}^{\prime}$, the second on $\tau_{m}^{\prime \prime}$ (given the past); each of them is "-" or "+" equiprobably. They have some correlation coefficient lying between $(-\rho)$ and $\rho$. Lemma 3.2 gives

$$
4 \mathbb{E}\left(\left.\left(\frac{M_{m}^{\prime}}{M_{m-1}^{\prime}} \frac{M_{m}^{\prime \prime}}{M_{m-1}^{\prime \prime}}\right)^{r} \right\rvert\, \ldots\right) \leq 4
$$

where $r=\frac{1}{1+\rho}$. Thus, $\mathbb{E}\left(\left(M_{m}^{\prime} M_{m}^{\prime \prime}\right)^{r} \mid \ldots\right) \leq\left(M_{m-1}^{\prime} M_{m-1}^{\prime \prime}\right)^{r}$, which means that the process $\left(M_{m}^{\prime} M_{m}^{\prime \prime}\right)^{r}$ is a supermartingale. Therefore, $\mathbb{E}\left(M_{n}^{\prime} M_{n}^{\prime \prime}\right)^{r} \leq\left(M_{0}^{\prime} M_{0}^{\prime \prime}\right)^{r}$, that is,

$$
\begin{aligned}
\mathbb{E}\left|f\left(\tau_{1}^{\prime}, \ldots, \tau_{n}^{\prime}\right) g\left(\tau_{1}^{\prime \prime}, \ldots, \tau_{n}^{\prime \prime}\right)\right| & \leq\left(\mathbb{E}\left|f\left(\tau_{1}^{\prime}, \ldots, \tau_{n}^{\prime}\right)\right|^{1 / r}\right)^{r} \cdot\left(\mathbb{E}\left|g\left(\tau_{1}^{\prime \prime}, \ldots, \tau_{n}^{\prime \prime}\right)\right|^{1 / r}\right)^{r} \\
& =\|f\|_{1+\rho}\|g\|_{1+\rho}
\end{aligned}
$$

[^1]
## 4 The main result

Return to the spider walk (Fig. 1b). It may be treated as a complex-valued martingale $Z$ (Fig. 2a), starting at the origin. Take each step to have length 1 . The set $\Omega_{n}^{\text {spider }}$ of all $n$ step trajectories of $Z$ can be identified with the set of leaves of a binary tree. The endpoint $Z_{n}=Z_{n}(\omega)$ of a trajectory $\omega \in \Omega_{n}^{\text {spider }}$ is a complex-valued function on $\Omega_{n}^{\text {spider }}$. Taking into account that $\mathbb{E}\left|Z_{n}\right|^{2}=n$, we ask about tree stability of the sequence $\left(Z_{n} / \sqrt{n}\right)_{n=1}^{\infty}$.


Figure 2: (a) the spider walk as a complex-valued martingale; (b) combinatorial distance.
4.1 Theorem The sequence $\left(Z_{n} / \sqrt{n}\right)_{n=1}^{\infty}$ is non-cosy.

By Lemma 2.6 it follows that the sequence $\left(Z_{n} / \sqrt{n}\right)_{n=1}^{\infty}$ is not tree stable. Recently, M. Emery and J. Warren found that some tree sensitive sequences result naturally from their constructions.
In contrast to the spider walk, the simple walk (Fig. 1a) produces a sequence $\left(\left(\tau_{1}+\cdots+\right.\right.$ $\left.\left.\tau_{n}\right) / \sqrt{n}\right)_{n=1}^{\infty}$ that evidently is cube stable, therefore tree stable, therefore cosy.
4.2 Lemma (a) $\lim \sup _{n \rightarrow \infty} \sqrt{n} \mathbb{P}\left(Z_{n}=0\right)<\infty$.
(b) $\liminf _{n \rightarrow \infty}\left(n^{-1 / 2} \sum_{k=1}^{n} \mathbb{P}\left(Z_{k}=0\right)\right)>0$.

The proof is left to the reader. Both (a) and (b) hold for each node of our graph, not just 0. In fact, the limit exists,

$$
\lim _{n \rightarrow \infty}\left(n^{1 / 2} \mathbb{P}\left(Z_{n}=0\right)\right)=\frac{1}{2} \lim _{n \rightarrow \infty}\left(n^{-1 / 2} \sum_{k=1}^{n} \mathbb{P}\left(Z_{k}=0\right)\right) \in(0, \infty)
$$

but we do not need it.
Proof of the theorem. Let $\mu_{n}$ be a probability measure on $\Omega_{n}^{\text {spider }} \times \Omega_{n}^{\text {spider }}$ such that ${ }^{4}$ $\rho_{\max }(\mu) \leq \rho, \rho \in(0,1)$; we'll estimate $\left\langle Z_{n}\right| \mu_{n}\left|Z_{n}\right\rangle$ from above in terms of $\rho$. We have two (correlated) copies $\left(Z_{k}^{\prime}\right)_{k=1}^{n},\left(Z_{k}^{\prime \prime}\right)_{k=1}^{n}$ of the martingale $\left(Z_{k}\right)_{k=1}^{n}$. Consider the combinatorial distance (see Fig. 2b)

$$
D_{k}=\operatorname{dist}\left(Z_{k}^{\prime}, Z_{k}^{\prime \prime}\right)
$$

Conditionally, given the past $\left(Z_{1}^{\prime}, Z_{1}^{\prime \prime}, \ldots, Z_{m-1}^{\prime}, Z_{m-1}^{\prime \prime}\right)$, we have two equiprobable values for $Z_{m}^{\prime}$, and two equiprobable values for $Z_{m}^{\prime \prime}$; the two binary choices are correlated, their correlation

[^2]lying in $[-\rho, \rho]$. The four possible values for $\left(Z_{m}^{\prime}, Z_{m}^{\prime \prime}\right)$ lead usually to three possible values $D_{m-1}-2, D_{m-1}, D_{m-1}+2$ for $D_{m}$, see Fig. 3a; their probabilities depend on the correlation, but the (conditional) expectation of $D_{m}$ is equal to $D_{m-1}$ irrespective of the correlation. Sometimes, however, a different situation appears, see Fig. 3b; here the conditional expectation of $D_{m}$ is equal to $D_{m-1}+1 / 2$ rather than $D_{m-1}$. That happens when $Z_{m-1}^{\prime \prime}$ is situated at the beginning of a ray (any one of our three rays) and $Z_{m-1}^{\prime}$ is on the same ray, outside the central triangle $\Delta\left(A B C\right.$ on Fig. 2b). In that case ${ }^{5}$ we set $L_{m-1}=1$, otherwise $L_{m-1}=0$. We do not care about the case when $Z_{m-1}^{\prime}, Z_{m-1}^{\prime \prime}$ are both on $\Delta$; this case may be neglected due to hypercontractivity, as we'll see soon. Also, the situation where $Z_{m-1}^{\prime}=Z_{m-1}^{\prime \prime}$ may occur, and then $\mathbb{E}\left(D_{m} \mid D_{m-1}\right) \geq D_{m-1}$.

(a)


Figure 3: (a) the usual case, $L=0$ : in the mean, $D$ remains the same; (b) the case of $L=1$ : in the mean, $D$ increase by $1 / 2$. More cases exist, but $D$ never decreases in the mean.

Theorem 3.4, applied to appropriate indicators, gives $\mathbb{P}^{1+\rho}\left(Z_{k}^{\prime} \in \Delta \& Z_{k}^{\prime \prime} \in \Delta\right) \leq \mathbb{P}\left(Z_{k}^{\prime} \in\right.$ $\Delta) \cdot \mathbb{P}\left(Z_{k}^{\prime \prime} \in \Delta\right)$, that is,

$$
\mathbb{P}\left(Z_{k}^{\prime} \in \Delta \& Z_{k}^{\prime \prime} \in \Delta\right) \leq\left(\mathbb{P}\left(Z_{k} \in \Delta\right)\right)^{\frac{2}{1+\rho}}
$$

for all $k=0, \ldots, n$. Combining it with Lemma 4.2 (a) we get

$$
\begin{equation*}
\sum_{k=0}^{n} \mathbb{P}\left(Z_{k}^{\prime} \in \Delta \& Z_{k}^{\prime \prime} \in \Delta\right) \leq \varepsilon_{n}(\rho) \cdot \sqrt{n} \tag{4.3}
\end{equation*}
$$

for some $\varepsilon_{n}(\rho)$ such that $\varepsilon_{n}(\rho) \xrightarrow[n \rightarrow \infty]{ } 0$ for every $\rho \in(0,1)$, and $\varepsilon_{n}(\rho)$ does not depend on $\mu$ as long as $\rho_{\max }(\mu) \leq \rho$.
Now we are in position to show that

$$
\begin{equation*}
\sum_{k=0}^{n} \mathbb{P}\left(L_{k}=1\right) \geq c_{0} \sqrt{n} \tag{4.4}
\end{equation*}
$$

for $n \geq n_{0}(\rho)$; here $n_{0}(\rho)$ and $c_{0}>0$ do not depend on $\mu$. First, Lemma 4.2 (b) shows that $\mathbb{P}\left(Z_{k}^{\prime \prime}=0\right)$ is large enough. Second, (4.3) shows that $\mathbb{P}\left(Z_{k}^{\prime \prime}=0 \& Z_{k}^{\prime} \notin \Delta\right)$ is still large enough. The same holds for $\mathbb{P}\left(Z_{k}^{\prime \prime}=0 \& Z_{k}^{\prime} \notin \Delta_{+2}\right)$, where $\Delta_{+2}$ is the (combinatorial) 2-neighborhood of $\Delta$. Last, given that $Z_{k}^{\prime \prime}=0$ and $Z_{k}^{\prime} \notin \Delta_{+2}$, we have a not-so-small (in fact, $\geq 1 / 4$ ) conditional probability that $L_{k}+L_{k+1}+L_{k+2}>0$. This proves (4.4).

[^3]The process $\left(D_{m}-\frac{1}{2} \sum_{k=0}^{m-1} L_{k}\right)_{m=0}^{n}$ is a submartingale (that is, increases in the mean). Therefore, using (4.4),

$$
\mathbb{E} D_{n} \geq \frac{1}{2} \sum_{k=0}^{n-1} \mathbb{P}\left(L_{k}=1\right) \geq \frac{1}{2} c_{0} \sqrt{n}
$$

for $n \geq n_{0}(\rho)$. Note that $D_{n}=\operatorname{dist}\left(Z_{n}^{\prime}, Z_{n}^{\prime \prime}\right) \leq C_{1}\left|Z_{n}^{\prime}-Z_{n}^{\prime \prime}\right|$ for some absolute constant $C_{1}$. We have

$$
\left(\mathbb{E}\left|Z_{n}^{\prime}-Z_{n}^{\prime \prime}\right|^{2}\right)^{1 / 2} \geq \mathbb{E}\left|Z_{n}^{\prime}-Z_{n}^{\prime \prime}\right| \geq C_{1}^{-1} \mathbb{E} D_{n} \geq \frac{1}{2} C_{1}^{-1} c_{0} \sqrt{n}
$$

and

$$
\left\|Z_{n}\right\|^{2}-\left\langle Z_{n}\right| \mu_{n}\left|Z_{n}\right\rangle=\frac{1}{2} \mathbb{E}\left|Z_{n}^{\prime}-Z_{n}^{\prime \prime}\right|^{2} \geq \frac{1}{4} C_{1}^{-2} c_{0}^{2} n
$$

for $n \geq n_{0}(\rho)$; so,

$$
\limsup _{n \rightarrow \infty}\left(\left\|\frac{Z_{n}}{\sqrt{n}}\right\|^{2}-\left\langle\frac{Z_{n}}{\sqrt{n}}\right| \mu_{n}\left|\frac{Z_{n}}{\sqrt{n}}\right\rangle\right) \geq \frac{c_{0}^{2}}{4 C_{1}^{2}}
$$

irrespective of $\rho$, which means non-cosiness.

## 5 Connections to continuous models

Theorem 4.1 (non-cosiness) is a discrete counterpart of [9, Th. 4.13]. A continuous complexvalued martingale $Z(t)$ considered there, so-called Walsh's Brownian motion, is the limit of our $\left(Z_{n t} / \sqrt{n}\right)$ when $n \rightarrow \infty$. The constants $c_{0}$ and $C_{1}$ used in the proof of Theorem 4.1 can be improved (in fact, made optimal) by using explicit calculations for Walsh's Brownian motion. Cosiness for the simple walk is a discrete counterpart of [9, Lemma 2.5].
Theorem 3.3 (hypercontractivity on trees) is a discrete counterpart of [9, Lemma 6.5]. However, our use of hypercontractivity when proving non-cosiness follows [2, pp. 278-280]. It is possible to estimate $\mathbb{P}\left(Z_{k}^{\prime} \in \Delta \& Z_{k}^{\prime \prime} \in \Delta\right)$ without hypercontractivity, following [5] or [9, Sect. 4].
Cosiness, defined in Def. 2.5, is a discrete counterpart of the notion of cosiness introduced in [9, Def. 2.4]. Different variants of cosiness (called I-cosiness and D-cosiness) are investigated by Emery, Schachermayer, and Beghdadi-Sakrani, see [4] and references therein. See also Warren [13].
Noise stability and noise sensitivity, introduced in [3], have their continuous counterparts, see [10, 11]. Stability corresponds to white noises, sensitivity to black noises. Mixed cases (neither stable nor sensitive, see [3, end of Sect. 1.4]) correspond to noises that are neither white nor black (as in [14]).

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[^0]:    ${ }^{1}$ Equality results from the inequality applied to complementary sets.

[^1]:    ${ }^{2}$ It is assumed that $\mu$ satisfies $(2.2) ; \rho_{\text {max }}$ was defined only for such measures.
    ${ }^{3}$ Of course, $x$ and $y$ depend on $\tau_{1}^{\prime}, \tau_{1}^{\prime \prime}, \ldots, \tau_{m-1}^{\prime}, \tau_{m-1}^{\prime \prime}$.

[^2]:    ${ }^{4}$ It is assumed that $\mu$ satisfies (2.2); $\rho_{\text {max }}$ was defined only for such measures.

[^3]:    ${ }^{5}$ There is a symmetric case $\left(Z_{m-1}^{\prime}\right.$ at the beginning $\left.\ldots\right)$, but we do not use it.

